## Estimating Demand with Recentered Instruments

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### Introduction

Demand for differentiated products is key to many economic analyses

- IO: welfare effects of mergers or new products; conduct testing
- Trade: welfare effects of tariffs; gains from trade or internal migration

Modern demand systems tractably model rich substitution patterns:

- IO: mixed logit (Berry, Levinsohn, Pakes '95), nested logit
- Trade: nested CES (Hottman, Redding, Weinstein '18), mixed CES (Adao, Costinot, Donaldson '17)

But finding credible IVs to estimate these models' parameters can be hard

• We propose a new IV construction under weaker assumptions

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## Mixed/Nested Logit Overview

Markets m with products  $j \in \mathcal{J}_m$  and an outside good j = 0

- Researcher observes prices  $p_{im}$  and product characteristics  $x_{im}$
- + market shares  $s_{jm}$  arising from random utility maximization:

$$s_{jm} = S_j(\boldsymbol{\delta}_m; \sigma, \boldsymbol{x}_m^{(1)}, \boldsymbol{p}_m)$$
 for  $\delta_{jm} = \alpha p_{jm} + \beta' x_{jm} + \xi_{jm}$ 

 $\xi_{jm}$  captures unobserved product characteristics and consumer tastes, lpha governs own-price elasticities,  $\sigma$  governs substitution patterns

Berry '94, Berry et al. '95 famously show the model can be inverted:

$$\mathfrak{D}_{j}(\boldsymbol{s}_{m};\sigma,\boldsymbol{x}_{m}^{(1)},\boldsymbol{p}_{m})=\delta_{jm}$$

Yields moment conditions given suitable instruments...

## **Endogeneity and Conventional IVs**

E.g. in nested logit:

$$\log(s_{jm}/s_{0m}) - \sigma\log(s_{jm}/\sum_{k \in \mathsf{nest}(j)} s_{km}) = \delta_{jm} = \alpha p_{jm} + \xi_{jm}$$

- ullet Endogeneity #1:  $p_{jm}$  is likely correlated with  $\xi_{jm}$ 
  - Natural IV: exogenous cost shocks  $g_{jm}$  (assume available)
- ullet Endogeneity #2:  $oldsymbol{s}_m$  is correlated with  $oldsymbol{\xi}_{jm}$  even if prices are random

Standard IVs for  $\sigma$ : fn's of rivals' characteristics (BLP '95, Gandhi-Houde '20)

- ullet E.g. # of rivals w/ characteristics close to j (nested logit: in j's nest)
- Usually relevant: e.g. larger nests have smaller within-nest shares
- Excluded from utility
- But validity requires characteristics to be econometrically exogenous
  - Violated e.g. if firms introduce more products in nests for which consumers have a higher preference  $\xi$

### Our Instruments

Fn's of cost shocks and characteristics of rival products that are "recentered": by construction uncorrelated with any fn's of characteristics

- E.g. average cost shock of products with characteristics similar to  $x_{jm}$  (nested logit: products in j's nest)
- If cost shocks are exogenous, valid even with endogenous characteristics, thanks to recentering (Borusyak-Hull 2023)
- Relevant: e.g. favorable cost shocks of rivals reduce s<sub>jm</sub>
  - Choose the fn optimally, a la Chamberlain (BLP '99, Borusyak-Hull '25)
  - Bonus: varies across markets even if the choice set does not (Nevo '01)
- Intuition: Think of product substitution as spillover effects
  - Nested logit  $\equiv$  peer effects model:  $\log \frac{s_{jm}}{s_{0m}} = \alpha p_{jm} + \sigma \log \frac{s_{jm}}{s_{\text{nest}(j)m}} + \xi_{jm}$
  - Peer groups can be endogenous if shocks to peers are exogenous

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### Related Literatures

#### Demand estimation without exogenous characteristics:

- Restricting unobservables: Sweeting '13, Moon et al. '18
- Modeling product entry: Fan '13, Petrin et al. '22
  - ▶ We leave unobservables and entry unrestricted
- Using IVs orthogonal to characteristics: Ackerberg and Crawford '09
  - ▶ We propose a way to construct such IVs by reusing cost shocks

### Identifying models with shift-share and other "formula" IVs:

- Linear models: Borusyak, Hull, Jaravel '22, '25a,b; Borusyak, Hull '23, '25
- Linearized models: Adao, Arkolakis, Esposito '25, Borusyak, Dix-Carneiro, Kovak '23
  - ► We work with a nonlinear model directly
- Recentered IVs yield robustness to misspecification: Andrews et al. '25
  - ▶ We show a different advantage and propose a specific IV construction

## Outline

- 1. Introduction ✓
- 2. Theoretical Results
  - Identification assumption
  - IV construction
  - Special cases
  - Estimation and asymptotics
  - Extensions
- 3. Monte Carlo Simulations

## **Exogenous Cost Shocks**

Observe supply-side shocks  $g_{jm}$ : to input costs, productivity, taxes/subsidies, firm ownership

**Assumption 1:**  $\mathbb{E}\left[\xi_{jm} \mid \boldsymbol{g}_{m}, \boldsymbol{x}_{m}, \boldsymbol{q}_{m}\right] = \mathbb{E}\left[\xi_{jm} \mid \boldsymbol{x}_{m}, \boldsymbol{q}_{m}\right]$  where optional  $q_{m}$  collects other data, e.g. lagged prices and shares

- $g_m$  should not affect product entry or consumer preferences, or correlate with any variables affecting these
- E.g. exchange rate fluctuations when estimating the demand for cars
- Relative to standard assumption  $\mathbb{E}\left[\xi_{jm} \mid \boldsymbol{g}_{m}, \boldsymbol{x}_{m}\right] = 0$ , allows observed characteristics to correlate with own & rivals'  $\boldsymbol{\xi}$

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## Optimal IV (extending Borusyak and Hull '25)

**Lemma:**  $\mathbb{E}[z_{jm}\xi_{jm}] = 0$  follows from Assumption 1 if and only if  $z_{jm}$  consists of recentered formula IVs:

$$z_{jm} = f_{jm}(\boldsymbol{g}_m, \boldsymbol{x}_m, \boldsymbol{q}_m)$$
 such that  $\mathbb{E}\left[f_{jm}(\boldsymbol{g}_m, \boldsymbol{x}_m, \boldsymbol{q}_m) \mid \boldsymbol{x}_m, \boldsymbol{q}_m\right] = 0$ 

**Proposition**: For  $\nabla_{jm} = \frac{\partial \xi_{jm}}{\partial \theta}$ , the asymptotically optimal IV is  $(z_{jm}^*)_{j,m} = \mathbb{E}\left[\xi \xi' \mid \mathbf{x}, \mathbf{q}\right]^{-1} (\mathbb{E}\left[\nabla \mid \mathbf{g}, \mathbf{x}, \mathbf{q}\right] - \mathbb{E}\left[\nabla \mid \mathbf{x}, \mathbf{q}\right])$ 

Our proposed IV approximates the numerator of  $z_{im}^*$ :

- 1) Approximate  $\mathbb{E}[\nabla_{jm} | \boldsymbol{g}, \boldsymbol{x}, \boldsymbol{q}]$  by  $\hat{\nabla}_{jm}(\boldsymbol{g}_m, \boldsymbol{x}_m, \boldsymbol{q}_m)$
- 2) Recenter it:  $z_{jm} = \hat{\nabla}_{jm} \mu_{jm}$  for  $\mu_{jm} = \mathbb{E}\left[\hat{\nabla}_{jm}(\boldsymbol{g}_m, \boldsymbol{x}_m, \boldsymbol{q}_m) \mid \boldsymbol{x}_m, \boldsymbol{q}_m\right]$
- $\hookrightarrow z_{jm} = \text{unexpected part of } \nabla_{jm} \text{ due to } \boldsymbol{g}_m$

# Our IV Construction: For lpha and eta

Recall 
$$\xi_{jm} = \mathfrak{D}_j\left(\boldsymbol{s}_m; \sigma, \boldsymbol{x}_m^{(1)}, \boldsymbol{p}_m\right) - \alpha p_{jm} - \beta' x_{jm}$$

For  $\alpha$ : predicted response of  $\frac{\partial}{\partial \alpha} \xi_{jm} = -p_{jm}$  to shocks

- ullet Just use  $g_{jm}$ : implicit pass-through model  $p_{jm}=\pi g_{jm}+\omega_{jm}$
- Can improve power by adding a supply side

Can't identify eta using cost shocks: they don't affect  $rac{\partial}{\partial eta} \xi_{jm} = -x_{jm}$ 

- ullet But cross-price elasticities do not involve  $oldsymbol{eta}$  (Ackerberg-Crawford '09)
- ullet Same for counterfactuals that hold  $x_{jm}$  fixed, even new products
- ullet Can control for  $x_{jm}$  for efficiency, like pre-period covariates in an RCT

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## Our IV Construction: For $\sigma$



- 1) Pick preliminary parameter estimates  $\check{\alpha}, \check{\sigma}$  (needn't be consistent)
  - E.g. using characteristic-based IVs
- 2) Construct a "no-shock scenario"  $(\check{p}_m, \check{s}_m)$  as a fn of  $(x_m, q_m)$ 
  - E.g. pre-period prices and shares, if available
- 3) Construct eq'm changes due to cost shocks as a fn of  $(g_m, x_m, q_m)$ :
  - Prices:  $\hat{p}_{jm} \check{p}_{jm} = \check{\pi}g_{jm}$  for pass-through  $\check{\pi}$
  - Mean utilities:  $\hat{\delta}_{jm} \check{\delta}_{jm} = \check{\alpha} \check{\pi}_{jm} g_{jm}$
  - Shares:  $\hat{s}_{jm} \check{s}_{jm} = S_j(\hat{\boldsymbol{\delta}}_m; \check{\boldsymbol{\sigma}}, \boldsymbol{x}_m^{(1)}, \hat{\boldsymbol{p}}_m) S_j(\check{\boldsymbol{\delta}}_m; \check{\boldsymbol{\sigma}}, \boldsymbol{x}_m^{(1)}, \check{\boldsymbol{p}}_m)$
  - $\hat{\nabla}_{jm} \check{\nabla}_{jm} = \frac{\partial}{\partial \sigma} \mathcal{D}_j(\hat{\mathbf{s}}_m; \check{\sigma}, \mathbf{x}_m^{(1)}, \hat{\mathbf{p}}_m) \frac{\partial}{\partial \sigma} \mathcal{D}_j(\check{\mathbf{s}}_m; \check{\sigma}, \mathbf{x}_m^{(1)}, \check{\mathbf{p}}_m)$
  - Or 1st-order approximation:  $\hat{\nabla}_{jm} \check{\nabla}_{jm} \approx \sum_{k \in \mathcal{J}_m} w_{jk}(\mathbf{x}_m, \mathbf{q}_m) \cdot \mathbf{g}_{km}$
- 4) Recenter  $\hat{\nabla}_{jm}$ : set  $z_{jm} = \hat{\nabla}_{jm} \mathbb{E}\left[\hat{\nabla}_{jm} \mid \mathbf{x}_m, \mathbf{q}_m\right]$ . How?

# How to Recenter $\hat{\nabla}_{jm}$ ? (see Borusyak and Hull '23)

Compute  $\mu_{jm} = \mathbb{E}\left[\hat{\nabla}_{jm} \mid \pmb{x}_m, \pmb{q}_m\right]$  without nonparametric regression:

• Use formula for  $\hat{\nabla}_{jm}$  + view cost shocks as a natural experiment, i.e. a draw from some "serendipitous randomized trial" (DiNardo '04)

Example: leveraging exchange rates of the country of car production

- Assume a random walk for each currency
- ullet Define  $g_{jm}$  as exchange rate increments,  $\mathbb{E}\left[g_{jm} \mid \pmb{x}_m, \pmb{q}_m\right] = 0$
- Shift-share IVs don't need recentering:

$$\mu_{jm} = \mathbb{E}\left[\sum_{k \in \mathcal{J}_m} w_{jk}(\boldsymbol{x}_m, \boldsymbol{q}_m) \cdot g_{km} \mid \boldsymbol{x}_m, \boldsymbol{q}_m\right] = 0$$

• For nonlinear IVs, draw many counterfactuals  $\mathbf{g}_{m}^{(c)}$  as permutations of increments over time for each currency and compute

$$\mu_{jm} = \frac{1}{C} \sum_{c} \hat{\nabla}_{jm}(\boldsymbol{g}_{m}^{(c)}, \boldsymbol{x}_{m}, \boldsymbol{q}_{m})$$

## Special Cases

• In nested logit,  $\partial \xi_{jm}/\partial \sigma = \log(s_{jm}/s_{\text{nest}(j)m}) \Longrightarrow$ 





$$z_{jm} \propto g_{jm} - \sum_{k \in \text{nest}(j)} \frac{\check{s}_{km}}{\check{s}_{\text{nest}(j)m}} g_{km}$$

- Increased prices of rivals in the nest raise  $\log(s_{jm}/s_{0m}) \Longrightarrow \sigma > 0$
- In the "local to logit" approximation to mixed logit, i.e.  $\check{\sigma} \approx 0$  (Salanie-Wolak '22, here for a non-price random coefficient)

$$z_{jm} \approx x_{jm} \cdot \sum_{k \in \mathcal{J}_m} \check{s}_{km} (x_{km} - \bar{x}_m) g_{km}, \qquad \bar{x}_m = \sum_{k \in \mathcal{J}_m} \check{s}_{km} x_{km}$$

• Increased prices of rivals with higher-than-average  $x_{km}$  raise market shares of products with high  $x_{jm} \Longrightarrow \sigma > 0$ 

### Extensions

Improve IV power by leveraging an auxiliary supply model



- Incorporate observed consumer characteristics
- Use input shocks with estimated mapping from inputs to products
- ullet Identify eta using exogenous shocks to non-price characteristics
- Alternative demand models: nested/mixed CES, Hotelling model (Houde 2012), "principles of differentiation" (Bresnahan et al. 1997), etc.
- Nonparametric identification: straightforward without a random coefficient on price

  Details Conjecture

## Outline

- 1. Introduction ✓
- 2. Theoretical Results ✓
- 3. Monte Carlo Simulations
  - Power comparison with exogenous characteristics
  - Bias comparison with endogenous characteristics

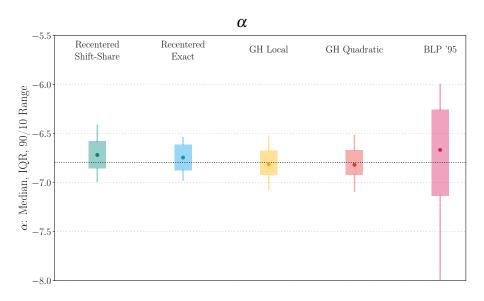
## Monte Carlo I: Exogenous Characteristics

## DGP based on Gandhi and Houde (2020), 100 simulations:



- 100 regional markets r over time periods t = 1,2; m = (r,t)
- 15 products per market with 2 observed time-invariant characteristics and random coefficients on them
- Shocks  $g_{jr2}$  affect costs in period 2
- Estimate  $(\alpha, \sigma_1, \sigma_2)$  with our exact and shift-share IVs (continuously updating, in differences)
- Compare with Gandhi-Houde local and quadratic IVs and BLP '95 IV (in period 2)

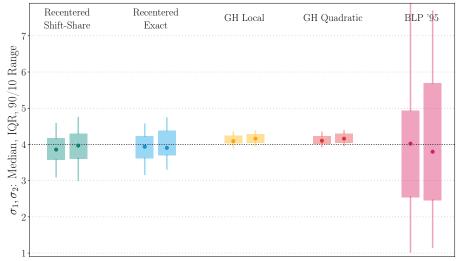
# Price Coefficient (Exogenous Characteristics)



# Nonlinear Parameters (Exogenous Characteristics)





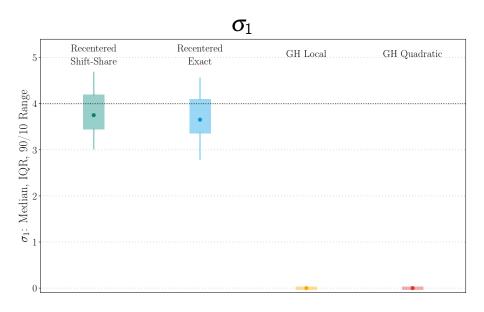


## Monte Carlo II: Endogenous Characteristics

Assume each region has a "bliss point"  $B_r$  for the first characteristic

- ullet Products near  $B_r$  are more popular (subtract  $3(x_{jr1}-B_r)^2$  from  $\xi_{jrt}$ )
- More entry near the bliss point  $(x_{jr1} \sim N(B_r, 1))$
- ⇒ GH instruments are invalid: popular products are in the dense part of the local distribution of characteristics

## Characteristic-Based IVs Have Strong Bias



## Conclusion

A new IV construction for nested/mixed logit and similar demand models

- Leverages product characteristics but does not require their exogeneity
- Inspired by thinking of product substitution as spillovers

### Open questions:

- Usefulness with microdata (e.g., individual-level choice data)?
  - For MLE or "micro-BLP" GMM estimation of mixed logit
- Usefulness beyond demand, e.g. for estimation of games?
  - Key: IV-GMM estimation

## In-progress applications:

- Demand for automobiles, with exchange rate shocks
- Demand for contraceptives, with price cap shocks





# Random Utility Model (Back)

#### Consumers max:

$$u_{ijm} = \delta_{jm} + \eta_{i0} \rho_{jm} + \eta_i' x_{jm}^{(1)} + \varepsilon_{ijm}, \quad \delta_{jm} = \alpha \rho_{jm} + x_{jm}' \beta + \xi_{jm}$$

for observed characteristics  $x_{jm} = (x_{jm}^{(1)}, x_{jm}^{(2)})$  and random coefficients  $\eta_{i\ell}$ 

- Assume  $(\eta_{i\ell})_{\ell} \stackrel{\textit{iid}}{\sim} \mathscr{P}(\cdot; \sigma)$  for known  $\mathscr{P}$  (e.g. Gaussian)
- Nested logit corresponds to  $x_{jm}^{(1)} =$  nest dummies and a special distribution of  $\eta_i$

# Estimator and Asymptotics (Back)

Given initial  $\check{\theta}=(\check{\alpha},\check{\sigma})$  and  $\check{\pi}$ , we estimate  $\theta$  using moment conditions:

$$\mathbb{E}\left[z_{jm}(\check{\boldsymbol{\theta}},\check{\boldsymbol{\pi}})\cdot\left(\mathfrak{D}_{j}(\boldsymbol{s}_{m};\boldsymbol{\sigma},\boldsymbol{x}_{m}^{(1)},\boldsymbol{p}_{m})-\alpha\rho_{jm}-\mathfrak{B}_{j}(\boldsymbol{x}_{m},\boldsymbol{q}_{m};\boldsymbol{\gamma},\check{\boldsymbol{\theta}})\right)\right]=0$$

where  $\mathscr{B}_j$  generalizes non-causal controls, e.g.  $\mathscr{B}_j=\xi_{jm}^{\rm pre}$  (estimation in differences) or  $\mathscr{B}_j=\gamma'x_{jm}$ , with  $\hat{\gamma}(\theta)$  minimizing residual sum of squares

Initial estimates  $\check{ heta}$  can come from conventional characteristic-based IVs

• Better: "continuously updating"  $z_{jm}(\theta, \check{\pi})$ 

Consistency, asymptotic normality, and inference:

- When markets are iid, follow from standard GMM results
- With non-*iid* markets but if  $z_{jm}$  is a shift-share, extend shock-level asymptotics of Adao, Kolesar, Morales (2019), Borusyak, Hull, Jaravel (2022)
  - E.g. if all regions are affected by the same exchange rate shocks

# Nested Logit Shares (Back)

Let  $\delta_{jm} = \alpha p_{jm} + \xi_{jm}$  and  $D_{nm} = \sum_{j \in \mathcal{J}_m} d_{jn} \exp(\delta_{jm}/(1-\sigma))$ . Then nested logit shares satisfy:

$$egin{aligned} rac{s_{jm}}{s_{n(j)m}} &= rac{\exp\left(\delta_{jm}/(1-\sigma)
ight)}{D_{n(j)m}}, \ s_{nm} &= rac{D_{nm}^{1-\sigma}}{1+\sum_{n'}D_{nm}^{1-\sigma}}, \ s_{0m} &= rac{1}{1+\sum_{n'}D_{nm}^{1-\sigma}} \end{aligned}$$

Manipulating these terms yields share inversion:

$$\log(s_{jm}/s_{0m}) = \alpha p_{jm} + \sigma \log(s_{jm}/s_{n(j)m}) + \xi_{jm}$$

# IV Construction for Nested Logit (Back)

**Exact prediction**: For  $\hat{p}_{jm} = \check{\pi}_0 + \check{\pi}g_{jm}$  and  $\delta_{jm} = \check{\alpha}\hat{p}_{jm}$ :

$$\begin{split} \widehat{\log} \frac{s_{jm}}{s_{n(j)m}} &= \frac{\check{\alpha}}{1 - \check{\sigma}} \left( \check{\pi}_0 + \check{\pi} g_{jm} \right) - \log \sum_{k \in \mathcal{J}_m} d_{kn} \exp \left( \frac{\check{\alpha}}{1 - \check{\sigma}} \left( \check{\pi}_0 + \check{\pi} g_{jm} \right) \right) \\ &= \frac{\check{\alpha}}{1 - \check{\sigma}} \check{\pi} g_{jm} - \log \sum_{k \in \mathcal{J}_m} d_{kn} \exp \left( \frac{\check{\alpha}}{1 - \check{\sigma}} \check{\pi} g_{jm} \right). \end{split}$$

**First-order approximation** around  $g_{km} = \mu_g$ :

$$\begin{split} \widehat{\log} \frac{s_{jm}}{s_{n(j)m}} &\approx \frac{\check{\alpha}}{1-\check{\sigma}} \check{\pi} \mu_g - \log \sum_{k \in \mathcal{J}_m} d_{kn} \exp\left(\frac{\check{\alpha}}{1-\check{\sigma}} \check{\pi} \mu_g\right) \\ &+ \frac{\check{\alpha}}{1-\check{\sigma}} \check{\pi} (g_{jm} - \mu_g) - \frac{\sum_{k \in \mathcal{J}_m} d_{kn} \exp\left(\frac{\check{\alpha}}{1-\check{\sigma}} \check{\pi} \mu_g\right) \frac{\check{\alpha}}{1-\check{\sigma}} \check{\pi} (g_{jm} - \mu_g)}{\sum_{k \in \mathcal{J}_m} d_{kn} \exp\left(\frac{\check{\alpha}}{1-\check{\sigma}} \check{\pi} \mu_g\right)} \\ &= -\log N_{n(j)m} + \frac{\check{\alpha}}{1-\check{\sigma}} \check{\pi} \left(g_{jm} - \frac{1}{N_{n(j)m}} \sum_{k \in \mathcal{J}_m} d_{kn} g_{jm}\right). \end{split}$$

# Incorporating a Supply Model Back

Assume constant MC  $c_{jm}$  + Nash-Bertrand:

$$oldsymbol{p}_m^* = oldsymbol{c}_m - \left(oldsymbol{H}_m \odot rac{d}{doldsymbol{p}_m'} oldsymbol{s}_m(oldsymbol{p}_m^*)
ight)^{-1} oldsymbol{s}_m(oldsymbol{p}_m^*)$$

- **1** Given  $(\check{\alpha}, \check{\sigma})$  solve for  $c_{jm}$ . Regress  $c_{jm} = \hat{\pi} \tilde{g}_{jm} + \text{error}_{jm}$
- ② Solve for costs  $\check{c}_{jm}$  corresponding to  $(\check{p}_m, \check{s}_m)$
- **3** Predict costs  $\hat{c}_{jm} = \check{c}_{jm} + \hat{\pi} \tilde{g}_{jm}$
- Predict prices  $\hat{\pmb{\rho}}_m = \pmb{p}_m^*(\hat{\pmb{c}}_m)$  (or via 1st-order approximation in  $g_{km}$ )
- **1** Use  $\hat{\boldsymbol{p}}$  to construct recentered IVs

## Nonparametric Identification (\*Back)

Demand model w/index restriction  $p_{jm} + \xi_{jm} = \mathcal{D}_j(s_m, x_m)$  for unknown  $\mathcal{D}_j$  (generalizes mixed logit with no random coefficient on price)

- Assume (i) shock exogeneity:  $\mathbb{E}[\xi_{jm} \mid \mathbf{g_m}, \mathbf{x_m}] = \mathbb{E}[\xi_{jm} \mid \mathbf{x_m}],$  (ii) completeness:  $\mathbb{E}[h(\mathbf{s_m}, \mathbf{x_m}) \mid \mathbf{g_m}, \mathbf{x_m}] \stackrel{a.s.}{=} 0 \implies h(\mathbf{s_m}, \mathbf{x_m}) \stackrel{a.s.}{=} 0$
- Then  $\mathcal{D}_j$  and  $\xi_{jm}$  are identified up to an additive function  $\beta(\mathbf{x}_m)$ ; enough for counterfactuals that hold  $x_{jm}$  fixed

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- Nonparametric demand with an index restriction on a non-price characteristic:  $x_{im} + \xi_{im} = \mathcal{D}_i(\boldsymbol{p}_m, \boldsymbol{s}_m)$  (fixing other characteristics)
- Berry and Haile (2014): if  $\mathbb{E}[x_{jm} + \xi_{jm} \mid \mathbf{x}_m, \mathbf{g}_m] = x_{jm}$ , NPIV identifies  $\mathcal{D}_j(\cdot)$  under completeness of  $(\mathbf{p}_m, \mathbf{s}_m) \mid (\mathbf{g}_m, \mathbf{x}_m)$ 
  - Conclusion: not enough to have exogenous  $g_m$ ; also need exogenous  $x_m$
  - But interactions of endogenous  $x_m$  & cost shocks are valid IVs!
- Conjecture (Borusyak, Chen, Hull in progress):
   If g<sub>m</sub> ⊥ ξ<sub>m</sub> | x<sub>m</sub>, then 𝒩(·) is identified up to invertible transformations r(·) ⇒ "conditional demand" is point-identified
  - So far proved assuming  $p_m \perp x_m \mid (\delta_m, g_m)$  (index restriction on prices, includes the case of exogenous prices)
  - In general,  $\mathbb{E}[\mathfrak{D}(\boldsymbol{p}_m, \boldsymbol{s}_m) \mid \boldsymbol{g}_m, \boldsymbol{x}_m] = \boldsymbol{h}(\boldsymbol{x}_m)$  for some  $\boldsymbol{h}(\cdot)$
  - Given  $h(\cdot)$ , completeness implies at most one  $f(\cdot)$  via NPIV
  - This equation is satisified by  $\mathfrak{D}(\cdot)$  but also  $r(\mathfrak{D}(\cdot))$  for any  $r(\cdot)$
  - For any  $h(\cdot)$  such that  $\exists r \colon \mathbb{E}[r(x_m + \xi_m) \mid x_m] = h(x_m)$  a.s., there is no other solution. But we don't know about other  $h(\cdot)$  functions yet

## Monte Carlo I: Details (Back)

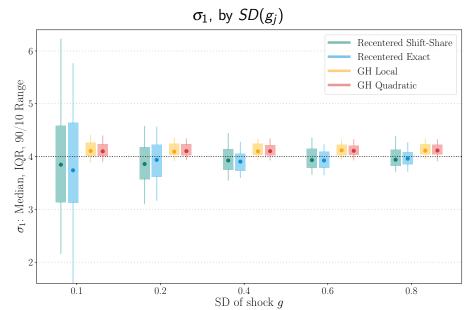
### DGP:

- $x_{jr\ell} \sim N(0,1)$ ,  $\xi_{jr1} \sim N(0,1)$ ,  $\xi_{jr2} = 0.5\xi_{jr1} + \sqrt{1 0.5^2} \cdot N(0,1)$
- Costs:  $c_{jrt} = \gamma' x_{jrt} + g_{jrt} + \omega_{jrt}$ ,  $\omega_{jr1} \sim N(0,1)$ ,  $g_{jr1} = 0$ ,  $\omega_{jr2} = 0.9 c_{jr1} + \sqrt{1 0.9^2} \cdot N(0,1)$ ,  $g_{jr2} \sim N(0,0.04)$
- Prices set by simultaneous Nash-Bertrand

Estimate  $(\alpha, \sigma)$  via alternative moment conditions:

- $\mathbb{E}[\Delta \xi_{jr} \cdot (g_{jr}, z_{jr})] = 0$  where  $z_{jm}$  is continuously-updating recentered IV from first-order approximation or exact prediction
- $\mathbb{E}\left[\xi_{jr2}\cdot(g_{jr},x_{jr},z_{jr}^C)\right]=0$  where  $z_{jr}^C$  is BLP or Differentiation IVs

# Recentered IVs Benefit from Cost Shock Variation



## GH IVs Benefit from Characteristic Variation

