Intangible capital can be used to create new goods and services (product intangibles) or to improve the efficiency of the firm (process intangibles). We report and study a new empirical fact: Executive and skilled labor pay is increasing in firm process intensity (the fraction of intangibles corresponding to process intangibles). We rationalize this fact in a dynamic principal-agent model, with the optimal contract uncovering process intensity’s direct and indirect effect on compensation. The direct effect is a level effect: Higher process intensity increases the returns to shirking. The indirect effect is a slope effect: Higher complementarity between process intangibles and physical capital investment increases the agent’s hold-up power over the firm for any level of process intensity. We verify these effects in the data. Importantly, we show that these effects are present in executive compensation and in the wages of highly skilled innovative employees, which we can measure using proprietary granular vacancy posting data from a labor-market data firm. In our baseline specification, a one standard deviation increase in process intensity is associated with an 8% increase in executive pay and a 3% increase in skilled labor wages relative to industry peers.

**Keywords:** Process Intangibles, Intangible capital, Dynamic contracting, Compensation.

**JEL Classification:** D21, E22, G31, G32, L22, O31, O34.

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1 Introduction

Intangible and innovative capital can play multiple roles within the firm. For example, intangible capital can be used with other inputs to create new output. We call these intangibles “product intangibles”. Other intangibles can be used to make the firm more efficient, by, for example, reducing costs or better organizing resources. We refer to these intangibles as “process intangibles”.\(^1\) Do these different uses of intangibles affect the compensation of executives and skilled employees? The data show that the answer is yes: Higher process intensity (process intangibles relative to total intangibles) is associated with higher pay.

![Figure 1: Average Executive Compensation by Process Intensity Bin](image)

This figure shows the mean executive compensation per unit of physical capital (times 100). The bins on the x-axis are created by sorting firms based on their process intensity each year. The bins are equally spaced. The y-axis is created by taking total executive compensation in the year and dividing it by the physical capital stock.

Figure 1 plots average executive compensation as a function of the firm process intensity.\(^2\) As the level of process intensity increases, the average executive compensa-

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\(^1\)The OECD’s Oslo Manual defines product innovation as *the introduction of a good or service that is new or significantly improved with respect to its characteristics or intended uses* and simultaneously defines process and organization innovation as *implementation of a new or significantly improved production or a new organizational method in the firm’s business practices, workplace organization or external relations* (*OECD* (2005)).

\(^2\)See Section 5 for exact details on construction.
In this paper, we rationalize this new empirical fact using a dynamic principal-agent model in which process intangibles are exposed to agency frictions and verify the model’s predictions in the data.

Our model is motivated by a novel empirical stylized fact: physical investment and process intangibles are complements. This complementarity creates an agency problem. The agent’s shirking effort in process innovation can reduce the efficacy of physical investment and reduce the firm’s value. This fact motivates us to model the physical capital accumulation by a CES aggregation function between physical investment and process intangibles.

In our model, if the agent exerts full effort, process intangibles are used to increase the efficiency of physical capital investment (i.e., the firm gets more “bang for the buck” per dollar of physical investment). This modelling choice is consistent with the existing literature. Parisi, Schiantarelli and Sembenelli (2006) and Bena, Ortiz-Molina and Simintzi (2022) both find that process improvements are associated with more capital investment. If the agent shirks, they enjoy private benefits proportional to the difference in physical capital growth with and without process intangibles.

We solve for the optimal contract that induces the agent to provide effort and find that there are two channels through which process intensity and compensation are linked. We call these the direct and indirect effects. The direct effect can be considered a level effect: Holding other variables and parameters fixed, as the process intensity of the firm increases, so does the promised utility to the agent. This effect arises because the agent’s benefit from shirking increases as process intensity increases, ceteris paribus. Therefore, the owners of the firm must promise the agent more utility to induce them to provide effort.

The indirect effect is akin to a slope effect: The process intensity-compensation association becomes stronger as the complementarity between process intangibles and physical investment becomes larger. This effect is a hold-up problem. As process intangibles become more important to the physical capital growth process, the agent can extract

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3This graph is unconditional on the level of intangibles. We show that this results holds conditionally on intangibles and other covariates as well, see Section 2.

4Pan and Li (2016) model process innovations as cost reducing. They also state, “[Process innovations] may involve investment in new technology embodied in machinery and equipment...” In unreported results, we do not find any negative relationship between costs of goods sold (COGS) and process intangibles.

5We could also motivate our modelling choice following arguments by Jensen and Meckling (1976). Intangibles are generally harder for outsiders (owners) to monitor (Lev (2000)), and this monitoring issue is exacerbated by the internal-to-the-firm nature of process intangibles.
more rents from the firm. This is most easily in the extreme cases. When physical investment and process intangibles are perfect substitutes, any rent extraction by the agent can be perfectly offset by an equivalent increase in physical capital investment. The level of physical capital growth is affected, but the marginal product of investment is not. In the other extreme case, process intangibles and physical investment are perfect complements. In this case, the agent must be induced to provide effort; otherwise, all physical investment is wasted: The agent can block the firm from growing until he is compensated enough. In reality, most firms are somewhere between these two extremes.

We measure process intensity in the data using information contained in patent claims (Bena and Simintzi (2019)). They scrape the text of filed patents, looking for phrases like a process for... or a product for... to determine the type of patent. There are other methods of measuring product versus process intensity. We focus on the Bena and Simintzi (2019) data and method because it is straightforward, publicly-available, and has already been used successfully in the previous paper.

We calibrate other model parameters, in particular, the CES parameter in the physical capital accumulation, using the empirical mean of the physical investment rate together with the total compensation and physical capital ratio in different quintiles of intangible-physical ratio. The calibrated model produces the flat physical investment rate in the lower four quintiles, which is also rationalized by Ward (2022) using a model with agency friction in the intangible accumulation. Our model also generates the elevated physical investment rate in the top quintile, but still under estimates the compensation in the top quintile. After splitting the data into low and high process intensity subsamples, the model also matches well the conditional moments of physical investment rate and compensation in both subsamples.

In our empirical analysis, we measure our main outcome variable, compensation, in two different ways. First, we use executive compensation, both total and deferred, from Compustat. This is the standard data used in the literature to test dynamic principal-agent models (Ward (2022)). The argument for using this data is that executives are the most powerful people in a firm and are best positioned to extract rents. However, it is not clear that executive effort actually matters for process intangibles to be effective. Our second measure overcomes that issue. We gather wage data on vacancy postings.

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6We cite these papers below.

7For example, when Nissan had a break-through in its car production methods (Link), its CEO was embroiled in a serious legal scandal. Exerting effort over process innovations was surely the last thing on his mind.
from Burning Glass Technologies (BGT). BGT is a firm whose competitive advantage is its unique vacancy posting data. The main benefit of this data is that BGT provides a large and standardized set of skills associated with each vacancy posting. Therefore, we can look at the posted wages for workers with skills specific to innovation, process improvement, and research and development (R&D).

Empirically, we verify the direct and indirect effects identified in the model, conditional on several covariates and fixed effects. We find that a one standard deviation increase in the process intensity of the firm is associated with an 8.4% increase in total compensation, a 7.6% increase in deferred compensation, and a 1.7% increase in the fraction of compensation deferred for executives. These are all measured relative to firm physical capital, which is consistent with the normalization in the model. The wage data from BGT is not a total flow, so it must be normalized differently. We normalize with respect to the wage of job postings requiring similar skills at other firms within the same industry-year (i.e., the leave-one-out industry-date mean). A one standard deviation increase in process intensity is associated with a 3.1% increase in this relative skilled wage. We first measure the complementarity of physical capital investment and process intensity to test the indirect effect. We do this by examining how the marginal product of physical investment on actual physical capital growth varies with the level of process intensity. We then sort firms based on our measure of complementarity. High complementarity firms have uniformly stronger associations between process intensity and compensation. A one standard deviation increase in process intensity is associated with a 16.5% increase in total compensation in the high complementarity firms, compared with a 6.7% increase in the low complementarity firms. For deferred compensation and the BGT relative skilled wage, these numbers are 17.7% versus 7% and 4.9% versus 2.4%, respectively.

These empirical results are robust to several different specifications. First, we repeat our main regressions at the executive level instead of the firm level. Results are qualitatively similar, even when we restrict our sample to executives who changed firms at least once. Second, we exploit the granularity of the BGT data even further and split our high skilled workers into those with product-focused innovative skills and those with

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8BGT has since merged with Lightcast, and the merged entity uses the Lightcast name.
9Eisfeldt, Falato and Xiaolan (2021) show that skilled labor is increasingly being paid via equity compensation. We focus on wages due to data availability, though our theory speaks to deferred compensation of skilled labor.
10Due to data limitations, we estimate complementarity at the four-digit NAICS level and assign all firms in that industry to have the same complementarity.
process-focused innovative skills. We find our effects are present in the second group, but not the first. Third, we test another channel connecting process intensity and compensation that is specific to the model. In the model, agency friction becomes stronger as uncertainty about capital growth increases. This implies that we should see another indirect effect in the data: Higher uncertainty in capital growth should lead to a stronger process intensity-compensation connection. We show that this is also empirically true.

We make three main contributions. First, we present a new finding that heterogeneity in the uses of intangibles is associated with heterogeneity in pay. In particular, higher process intensity is associated with higher pay. Second, we develop a dynamic principal-agent model with heterogeneity in intangibles that can rationalize the empirical phenomenon. Importantly, the model shows that there is a direct and indirect effect of process intensity on compensation: The level effect comes from variation in the shirking benefit, and the slope effect comes from variation in the complementarity between process intangibles and physical capital investment. Third, we show that both the direct and indirect effects exist in the data, not only for executives but also for the skilled works whose effort determines the efficacy of process intangibles.

This paper sits at the intersection of three different literature. First, we contribute to the literature on dynamic agency theory (e.g., DeMarzo and Sannikov (2006), Biais et al. (2007), DeMarzo and Fishman (2007a, b), Sannikov (2008), DeMarzo et al. (2012)). Our model extends this framework to include heterogeneous forms of intangible capital. This adds new testable predictions (the core of our paper) and new state variables, adding computational complexity. More closely related, Ward (2022) studies the role of agency frictions on intangibles, but does not distinguish between different types of intangibles. We consider the papers complement, as we essentially take Ward (2022)’s result as a starting point (there are agency frictions on intangibles) and take the next natural steps to document the heterogeneity of intangible capital and its implications. To the best of our knowledge, there are no other papers relating intangibles and agency frictions.11

Second, we contribute to the literature connecting intangible capital and finance (e.g., Lev and Radhakrishnan (2005), Eisfeldt and Papanikolaou (2013), Kung and Schmid (2015), Peters and Taylor (2017), Crouzet and Eberly (2018), Ewens, Peters and Wang (2019), Crouzet et al. (2022)). These papers do not study agency conflict, nor do they seek to quantify heterogeneous intangible capital.12 We use the methods of Ewens, Pe-

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11 Grabner (2014) studies the empirical relationship between “creativity-dependent” firms and incentive pay.
12 In fact, we combine patent data and intangible capital measures to get our heterogeneous measures.
ters and Wang (2019) and Peters and Taylor (2017) to create our firm-level measure of intangible capital and investment. We fit within the subset of this literature that looks at the relationship between pay and innovation/intangibles (Lerner and Wulf (2007), Lustig, Syverson and Van Nieuwerburgh (2011), Kline et al. (2019), Song et al. (2019), Sun and Xiaolan (2019), Kogan et al. (2020), Bhandari and McGrattan (2021)). These papers also do not look at agency conflicts or the heterogeneous nature of intangibles.

Third, we contribute to the small literature on process versus product innovation and finance. Our measure of process innovation intensity comes from Bena and Simintzi (2019). Ganglmair, Robinson and Seeligson (2022) provide a survey of the empirical evidence on process claims over time and provide their measure of process intensity. None of these papers are concerned with compensation or agency. To the best of our knowledge, we are the first to explicitly tie a formal model of the firm (with or without agency) to the empirical data on process versus product innovation.

2 Stylized Facts

This section presents three key empirical stylized facts that motivate our model in the next section. The model will then provide further implications for us to test. These facts also serve as a summary of our main results. We rely on simple double sorts or regressions in this section and defer the more detailed empirical work to Section 6. We leave formal data description and variable construction to Section 5.

The facts we present are the following: First, firms with higher process intensity provide higher compensation for their executives. Second, the association between process intensity and compensation increases in the amount of physical capital investment. Third, firms with higher process intensity have lower contemporaneous sales. All three facts are conditional on the level of the intangible capital stock.

Figure 2 displays the first stylized fact. To construct this figure, we independently sorted firms into three bins based on their process intensity and three bins based on their intangible capital to physical capital ratios. We see two effects here. First, the average

13 Angenendt (2018) also estimates process intensity.
14 Mohnen and Hall (2013) provide an overview of the empirical evidence linking firm outcomes to process and product innovation.
15 This section focuses on total executive compensation, one of our three compensation and salary measures. We leave the other two, deferred compensation and skilled labor salaries, to the main Results section.
16 The bins are rebalanced every year. Sorting based on intangible capital is done conditional on industry,
This figure shows the mean executive compensation per unit of physical capital. The bins on the x-axis are created by sorting firms based on their process intensity each year. Each sub-graph and color is created by annually sorting firms based on their intangible capital to physical capital stocks. Bins are rebalanced each year for both variables, and the intangible capital bin assignments are conditional on industry. The measure of process intensity is scaled by the industry average process intensity. The compensation variable as been multiplied by 100 to remove small decimal numbers.

level of executive compensation is increasing as the intangible capital bin increases (intangible capital level increases). This is implied by Ward (2022). Second, within each intangible capital bin (conditioning on the intangible capital level), executive compensation increases with the level of process intensity. This novel fact suggests that not only does the level of intangibles matter, but that the type of intangibles matter, too. This empirical fact has not been documented in the data, nor has it been explained by the existing models.

To illustrate the economic importance of our channel, we use Figure 3 to compare high and low process intense firms. In particular, we focus on the subset of firms that only filed process focused patents (full process) versus those that filed only product as Eisfeldt, Kim and Papanikolaou (2020) suggest. Process intensity is already normalized by the industry average process intensity.
This figure shows the average executive compensation per unit capital for full versus full product intense firms within intangibility quintiles. Full process intense firms issue patents in which 100% of the claims are related to process improvements, while full product firms issue patents in which 100% of the claims are related to product improvements. The intangibility quintiles are rebalanced every year. The average compensation per unit capital is computed by first taking the cross-sectional median within each quintile-date (for full product and full process firms separately) and then taking the time-series average of these medians.

Figure 3 shows that the process intensity effect is economically meaningful. For example, in the highest intangibility quintile, the full process executives are paid approximately five times as much as their peers at the full product firms, per unit physical capital. Using a back-of-the-envelope calculation, in dollar terms, this is a $450,000 difference per year.

Figure 4 displays our second stylized fact. This figure shows the sensitivity of executive compensation to a one-standard deviation increase in process intensity by physical

focused patents (full product). We split the sample in intangible over physical capital (intangibility) quintiles and look at the average compensation per unit physical capital for executives at these full process or product firms.
capital investment bin. The 95% confidence intervals are also displayed.

The key takeaway here is that the sensitivity is increasing as physical investment increases. This captures our idea of the hold-up problem inherent in process intangibles. Firms undertaking more physical investment are more “dependent” on the efforts of the agents to fully realize the benefits of the investment. The agent can thus extract rents from the firm. This effect is increasing with physical investment. This positive relationship is predicated by the assumption that physical investment and process intangibles are complements. This assumption is verified in the Appendix E. Consider the extreme case where physical investment and process intangibles are perfect substitutes. Then, there should be no interaction between process intangibles and physical investment in executive compensation.

The third stylized fact is that sales are decreasing with process intensity, conditional on the level of intangible capital. To show this fact, we estimate:

\[
Sales_{ft} = y_t + y_j + \beta_1 ProcIn_{ft} + \beta_2 IK_{ft} + \beta_3 iB/M_{ft} + \beta_4 Size_{ft} + \varepsilon_{ft}
\]  

(2.1)

where sales are divided by physical capital, ProcIn is our main process intensity measure as explained in Section 5, iB/M is the book-to-market ratio with intangibles added to the book value, and Size is market capitalization. The two fixed effects, \(y_t\) and \(y_j\), control for date and industry effects, respectively. Standard errors are clustered at the firm level. All variables are logged, except for process intensity (since 0 is meaningfully frequent), which is expressed in standard deviation units.

Table 1 displays the results. The top row shows the estimates for the effects of process intensity on sales. The effects are negative and significant across specifications. A one standard deviation increase in process intensity is associated with about a 10% decrease in sales. Column (1) displays the sparsest specification using only intangible capital and process intensity without any controls or fixed effects. Column (2) adds controls. Column (3) adds fixed effects, without controls, and Column (4) includes all controls and fixed effects.

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17The sensitivity is measured as the simple regression coefficient on process intensity when log compensation is regressed on process intensity, intangible capital, and firm size. Physical capital investment bins are created analogously to the process intensity bins above.

18Moreover, we also document there that the complementarity between physical investment and process intangibles is much stronger than the complementarity between physical investment and product intangibles.
Figure 4: Sensitivity of Executive Compensation to Process Intensity by Physical Investment Bin

This figure shows the sensitivity of log executive compensation (over physical capital) to a one-standard deviation increase in process intensity by physical investment bin. The sensitivity is measured as the simple regression coefficient on process intensity when log compensation is regressed on process intensity. The investment bins are based on the physical investment to physical capital ratio of firms. They are rebalanced every year. 95% confidence intervals are displayed using firm level clustering.
Table 1 Here

We next turn to the model, which will capture these stylized facts and provide further implications for the connection between compensation and process intensity.

3 Model setting

3.1 Capital, investment, and agency

The firm produces output using both physical and intangible capital, whose stock values are \( K \) and \( O \), respectively. The firm determines its investment, \( I \), in the physical capital and its investment, \( S \), in the intangible capital. Both \( I \) and \( S \) are assumed to be non-negative, so that investment is irreversible. Physical and intangible investments are subject to convex adjustment costs \( C_K(I) \) and \( C_O(S) \), respectively. The instantaneous cash flow produced by the firm is

\[
Y = \mu \left[ (1 - \phi) K^\psi + \phi (\theta O)^\psi \right]^{1/\psi} - I - S - C_K(I) - C_O(S),
\]

(3.1)
after netting investments and adjustment costs. The production requires a combination of physical capital and a fraction, \( \theta \), of intangibles that are used towards product innovation. They are aggregated by a CES production function \( \mu \left[ (1 - \phi) K^\psi + \phi (\theta O)^\psi \right]^{1/\psi} \) with the productivity rate \( \mu \). Following Eisfeldt and Papanikolaou (2013), we assume the adjustment costs for physical and intangible investment as

\[
C_K(I) = \frac{Q_K}{c_K} \left( \frac{I}{K} \right)^{c_K} K \quad \text{and} \quad C_O(I) = \frac{Q_O}{c_O} \left( \frac{O}{O} \right)^{c_O} O,
\]

(3.2)
respectively, with constants \( c_K, c_O, Q_K, \) and \( Q_O \).

The intangible capital \( O \) evolves according to

\[
dO_t = (S_t - \delta_O O_t) dt,
\]

(3.3)
where \( \delta_O \) is the depreciation rate of the intangible capital. The evolution of \( K \) follows

\[
dK_t = (D(e_t, I_t, O_t) - \delta_K K_t) \ dt + \sigma_K dZ^e_t,
\]

(3.4)
where \( \delta_K \) is the depreciation rate of the physical capital and \( Z^e \) is a Brownian motion.
describing shocks to the physical capital. Accumulation of physical capital depends on both investment $I$ and a fraction, $1 - \theta$, of intangibles. The production function $D$ takes a CES form

$$D(e, I, O) = \frac{A}{a^{1/\rho}} \left[ a I^\rho + e(1-a) \left((1-\theta)O\right)^\rho \right]^{1/\rho}. \quad (3.5)$$

The presence of intangibles in the production function $D$ models the process innovation, which makes physical investment more efficient: a larger value of $(1 - \theta)O$ increases the physical capital $K$ more for the given physical investment $I$. The CES parameter $\rho$ measures the complementarity between the physical investment and the intangibles in the physical capital accumulation. The lower $\rho$ is, the more complementarity between the two components. The factor $a^{-1/\rho}$ in front of the CES function is a normalization factor so that without intangible the production function takes the standard form $D(e, I, 0) = AI$. See Lin (2012) for a more detailed discussion on the production function $D$. This production function is also consistent with empirical evidences. Parisi, Schiantarelli and Sembenelli (2006) find that process innovations are more strongly associated with capital investment than product innovations. Bena, Ortiz-Molina and Simintzi (2022) find that firms with high innovation ability use process improvements to adjust their investment rates.

In the production function $D$, the agent’s effort $e$ is either 0 or 1. When $e = 1$, the agent exerts full effort and works efficiently, the physical capital increases by $D(1, I, O)$ and the efficacy of physical investment is improved by process intangibles. When $e = 0$, the agent shirks his effort, and the physical capital only increases by $D(0, I, O) = AI$, which is independent of process intangibles. The dependence of $D$ on the agent’s effort models the agency friction on process innovations. Define

$$\Lambda(I, O) = D(1, I, O) - D(0, I, O). \quad (3.6)$$

The function $\Lambda$ measures the increment of the firm’s physical capital accumulation due to the agent’s effort, conditional on the physical capital investment and intangible capital. Therefore, $\Lambda$ measures how much stake the agent controls in the process innovation.

Four properties of $\Lambda$ are important for our results: (i) $\Lambda(I, O)$ increases with $O$, indicating a more important role agent’s effort plays in the physical capital accumulation, hence a higher shirking benefit, when a firm possesses more intangibles; (ii) $\Lambda(I, O)$ increases in $1 - \theta$, implying that more process innovation increases the agency friction; 

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19He (2009) also makes the benefit function dependent on the investment rate.
(iii) When $\rho < 1$, $\frac{\partial \Lambda}{\partial (1 - \theta)}$ increases in $I$.\(^{20}\)

This indicates a hold up problem for the physical investment and the problem is more severe when the firm invests more. (iv) When $\rho_1 < \rho_2$,

$$\frac{\partial \Lambda_{\rho_1}}{\partial (1 - \theta)} > \frac{\partial \Lambda_{\rho_2}}{\partial (1 - \theta)} > 0,$$

for fixed $I$ and $O$. Therefore, the agency friction is more severe when $\rho$ is smaller and process intensity is higher.

When the agent shirks, he enjoys a flow of private benefits, which is assumed to be $\lambda \Lambda$. Conditional on the agent’s impact on the physical capital accumulation, $\lambda$ measures the magnitude of agent’s private benefit from shirking. We will show later that $\lambda$ is also the ratio between compensation volatility and physical capital volatility, and this ratio increases in the firm’s level of intangible capital in our data. Therefore, we will assume later that $\lambda$ increases with firm’s intangibles.

We assume that the firm’s owner (principal) only observes the dynamics of $O$ and $K$, but cannot observe the agent’s effort $e$ due to the random shocks in $Z^e$. This introduces agency friction in process innovation.\(^{21}\)

Our model setting mirrors the stylized facts documented in the previous section. Among the intangible capital $O$, $\theta O$ is used in the product innovation to generate output, $(1 - \theta)O$ is utilized in the process innovation to improve the efficacy of physical investment. Therefore, we call $1 - \theta$ firm’s process intensity. As $\theta$ decreases (equivalently, $1 - \theta$ increases), the CES production function $\mu \left[ (1 - \phi) K^{\psi} + \phi (\theta O)^\psi \right]^{1/\psi}$ decreases, which maps to the third stylized fact that sales decrease in process intensity. Complementarity between physical investment and process intangibles in the second stylized fact motivates the CES aggregation between physical investment and process intangibles in (3.5). Agent’s effort $e$ in the CES function represents the hold up problem: shirking reduces the efficacy of physical investment. The property (iii) of the shirking benefit $\Lambda$ also echoes the empirical pattern in Figure 4.

\(^{20}\)See Acemoglu (2008) for a discussion of complementarity/substitutability in the CES production function.

\(^{21}\)We can also consider the case where the dynamics of $O$ is subject to random shocks, for example, $dO = (S_t - \delta_t O_t)dt + \sigma_t O_t dW_t$ for another Brownian motion $W$ independent of $Z^e$. However, contracting on $O$ does not provide an incentive to the agent in our model and makes the agent’s continuation utility more volatile. We will show later that the principal’s value function is concave in the agent’s continuation utility. Therefore, the principal is implicitly risk averse in the agent’s continuation utility, hence does not load on the intangibles in the optimal contract.
Next, we present a contracting problem between the owner of our model firm (principal) and an executive or a skilled employee (agent) who has expertise in process innovation.

3.2 Contracting problem

The principal offers a contract with a cumulative compensation of $C$ to the agent. The agent does not subsidize the firm by accepting negative compensation. Therefore, $C$ is a non-decreasing process. For a given compensation plan $C$, the agent’s continuation utility $U$ is

$$U_t = \max_{e \in \{0, 1\}} \mathbb{E}^e_t \left[ \int_t^\tau e^{-\gamma(s-t)} [dC_s + (1-e_s)\lambda_s \Lambda_s ds] \right]. \quad (3.8)$$

The expectation is taken with respect to a probability $\mathbb{P}^e$, which is induced by the agent’s effort $e$. The Brownian motion $Z^e$ in (3.4) is under the measure $\mathbb{P}^e$. When $e = 0$, we suppress the superscript. The agent is assumed to be risk neutral, discounting future compensation and potential private shirking benefit using a subjective discounting rate of $\gamma$. When the agent’s continuation utility decreases to the outside value, normalized to be zero, at an endogenously determined stopping time $\tau$, the contract is terminated. Then, the agent leaves the firm, production continues, but the physical capital accumulation is less efficient with its expected value reduced to $D(0, I, O) = AI$.

The principal of the firm chooses a contract to maximize the expected future cash flow net compensation discounted by the interest rate $r$, which is assumed to be strictly less than $\gamma$. Principal chooses among the contracts which incentivize the agent’s full effort $e = 1$. Therefore, the principal’s optimization problem at time zero is

$$\max_{I, S, C} \mathbb{E}^{e^*} \left[ \int_0^\tau e^{-rs} [Y_s ds - dC_s] + e^{-r\tau} V_\tau \right], \quad (3.9)$$

subject to agent’s incentive compatibility constraint that the agent chooses the full effort optimally, i.e., $e^* = 1$ and agent’s participation constraint $U_0 \geq 0$. In (3.9), the contract termination time is

$$\tau = \inf \{ t \geq 0 : U_t = 0 \},$$

when the agent’s continuation value from the contract reaches his outside value (normalized to be zero). The contract is terminated at $\tau$ to protect the agent’s limited liability with respect to his outside value. After $\tau$, the efficacy of firm’s physical capital accumu-
lation is reduced:
\[ dK_t = AI_t \, dt + \sigma K_t \, dZ_t. \] (3.10)

Nevertheless, production continues without process intangibles. Hence, the firm’s value at the contract termination is
\[ V_\tau = \max_{I, S} E_\tau \left[ \int_\tau^\infty e^{-r(s-\tau)} Y_s \, ds \right], \] (3.11)
subject to (3.3) and (3.10).

4 Optimal contract and implications

4.1 Optimal contract

In order to incentivize the agent’s full effort, the principal exposes the agent’s continuation utility to variations in \( K \). Introducing a pay-performance sensitivity \( \phi \) to \( dK \) yields the benefit of working \( \phi \Lambda \) for the agent. Comparing to the cost of working (losing the shirking benefits) \( \lambda \Lambda \), the principal needs to choose \( \phi \geq \lambda \) to incentivize the agent’s full effort. The following result summarizes the agent’s optimal effort choice and dynamics of the continuation utility.

Lemma 4.1 For a given cumulative compensation \( C \), there exists a process \( \phi \) such that the agent’s continuation utility follows
\[ dU_t = \gamma U_t \, dt + \phi_t K_t \sigma dZ_t^* - dC_t, \] (4.1)
where the agent’s optimal effort is
\[ e_t^* = \begin{cases} 
1, & \phi_t \geq \lambda_t, \\
0, & \text{otherwise}.
\end{cases} \] (4.2)

Therefore, in order to incentivize full effort, agent’s incentive compatibility constraint is
\[ \phi_t \geq \lambda_t. \] (4.3)

We now turn to the principal’s problem (3.9). Introduce the principal’s value function
as
\[
V(K_t, U_t, O_t) = \max_{I,S,C} \mathbb{E}_t^\pi \left[ \int_t^\tau e^{-r(s-t)} (Y_s ds - dC_s) + e^{-r(\tau-t)} V_\tau \right].
\] (4.4)

The homogeneity in \( K \) allows us to introduce a function \( v \) via
\[
V(K, U, O) = K v(u, o),
\] (4.5)

where
\[
u = U/K \quad \text{and} \quad o = O/K
\]
are the continuation utility to physical capital ratio and the intangible to physical capital ratio, respectively.\(^{22}\) Physical capital accumulation takes the form
\[
D(e, I, O) = K d(e, i, o), \quad \text{where} \quad d(e, i, o) = \frac{A a^{1/\rho}}{\varphi^{1/\rho}} \left[ a^\rho + e(1-a) ((1-\theta) o)^\rho \right]^{1/\rho}. \tag{4.6}
\]

When \( e = 1 \), we denote \( d(1, i, o) \) by \( d(i, o) \) to simplify notation. Using \( u \) and \( o \) as two state variables for the principal’s problem, the optimal contract, and the optimal investment strategies are characterized by the following result.

**Proposition 4.1** The function \( v \), the optimal contract, and the optimal investment are described as follows:

(i) The function \( v \) satisfies the HJB equation
\[
0 = \max \left\{ -(r + \delta_K) v + \max_{i \geq 0, s \geq 0, \varphi \geq \lambda} \left\{ (v - o \partial_o v - u \partial_u v) d(i, o) \right. \right. \\
+ (s - (\delta_O - \delta_K) o) \partial_o v + (\gamma + \delta_K) u \partial_u v \\
+ \frac{1}{2} o^2 \sigma^2 \partial^2_{oo} v + \frac{1}{2} (\varphi - u)^2 \sigma^2 \partial^2_{uu} v - o (\varphi - u) \sigma^2 \partial^2_{ov} v \\
+ \mu \left[ 1 - \varphi + \varphi (\theta o)^\psi \right]^{1/\psi} - i - s - C_K(i) - C_O(s/o) o \left. \right\}, -\partial_u v - 1 \right\}. \tag{4.7}
\]

(ii) Define \( \bar{\pi}(o) = \inf \{ u : \partial_u v(u, o) = -1 \} \). The optimal compensation is a reflection type. Whenever \( u_t < \bar{\pi}(o_t) \), no compensation is paid, i.e., \( dC^*_t = 0 \). Only when \( u_t = \bar{\pi}(o_t) \), compensation is paid to keep the state process \((u, o)\) below \( \bar{u} \).

\(^{22}\)We choose to normalize by physical capital to better map to the existing literature, see eg. Eisfeldt and Papanikolaou (2013).
(iii) When
\[ \partial_{uu}^2 v < 0 \quad \text{and} \quad \lambda > u + \frac{\partial_{uu}^2 v}{\partial_{uv}^2} \]
the optimal contract sensitivity \( \varphi^* \) is \( \lambda \).

(iv) When \( v - o \partial_v - u \partial_u v > 0 \), the optimal physical investment and physical capital ratio, \( i^* \), satisfies the first order condition
\[ (v - o \partial_v - u \partial_u v) \partial_i (i^*, o) = 1 + Q_K(i^*)^{c_k-1}; \]
on otherwise, \( i^* = 0 \). If \( \partial_o v > 1 \), the optimal intangible investment and physical capital ratio, \( s^* \), is
\[ s^* = o (\partial_o v - 1) \frac{1}{Q_o} \]
on otherwise \( s^* = 0 \).

To understand the HJB equation (4.7), we first use (3.3), (3.4), and (4.1) to derive the dynamics of \( u = U/K \) and \( o = O/K \):
\[ \begin{align*}
\mathrm{d}o_t &= [s_t - (\delta_O - \delta_K) o_t - o_t d(i_t, o_t) + o_t \sigma^2] \, \mathrm{d}t - o_t \sigma \mathrm{d}Z_t, \\
\mathrm{d}u_t &= [(\gamma + \delta_K) u_t - u_t d(i, o) + \sigma^2 (u_t - \varphi_t)] \, \mathrm{d}t + \sigma (\varphi_t - u_t) \, \mathrm{d}Z_t - \frac{1}{K} \, \mathrm{d}C_t,
\end{align*} \]
where \( d(i, o) = A[a^\rho + (1 - a)(1 - \theta)^\rho \theta^\rho]^{1/\rho} \), \( i = I/K \), \( o = O/K \), and the superscript 1 is suppressed on \( Z^1 \) to simplify notation. Equation (4.7) divides the state space into two regions: (i) continuation region where
\[ r v K = \mathbb{E}[d(Kv)] + \mathbb{E}[dK] + \mathbb{E}[dK dv] + Y, \]
where \( \mathcal{L}_{u,o} \) is the infinitesimal generator of \( (u, o) \) in (4.11) and (4.12); (ii) compensation region, where the marginal benefit of compensation \( -\partial_u v \) equals the unit marginal cost. The right-hand side of (4.7) compares two groups of terms corresponding to continuation and compensation, respectively. Only one group equals zero for each point in the state space. The boundary between the continuation and the compensation region is \( \pi \). The optimal compensation satisfies \( dC_t^* = 0 \) when \( u_t < \bar{u}(o_t) \) and \( dC_t^* > 0 \) when \( u_t = \)}
\( \bar{u}(o_t) \). This compensation maintains the state process to be lower than the compensation boundary and reflects the state process whenever the compensation boundary is reached.

The optimal pay-performance sensitivity is determined by the constrained optimization problem

\[
\max_{\phi \geq \lambda} \left\{ \frac{1}{2} \left( \phi - u \right)^2 \sigma^2 \frac{\partial^2 v}{\partial u^2} \sigma \left( \phi - u \right) \sigma^2 \frac{\partial^2 v}{\partial u \partial o} \right\},
\]

where the pay-performance sensitivity \( \phi \) is subject to the incentive compatibility constraint \( \phi \geq \lambda \). When the conditions (4.8) are satisfied, the incentive compatibility constraint is binding, i.e., \( \phi^* = \lambda \). Conditions (4.8) will be verified numerically in our experiments later.

The optimal investments are determined jointly by their first-order conditions and the non-negativity constraint. When \( i^* > 0 \), it satisfies the first order condition (4.9), where the right-hand side is the marginal cost of physical investment. The left-hand side of (4.9) consists of two components. First, the marginal impact of physical investment on the growth rate of the physical capital is \( \partial_d(i^*, o) \). Therefore, the marginal benefit on the value function, due to the change of physical capital accumulation, is \( v \partial_d(i^*, o) \). Second, the growth in physical capital reduces the intangible and physical capital ratio, at the rate of \( o \partial_d(i^*, o) \), and also reduces the continuation utility and physical capital ratio, at the rate of \( u \partial_d(i^*, o) \). Both reductions introduce the marginal cost \( o \partial_o v + u \partial_u v \partial_d(i^*, o) \). The optimal investment in the physical capital balances the net marginal benefit on the left-hand side of (4.9) and the marginal cost on the right-hand side. The optimal investment in the intangible capital satisfies the following first-order condition, when \( \partial_o v > 0 \),

\[
\partial_o v = 1 + Q_0 (s^*/o)^{\sigma o - 1},
\]

where the marginal cost on the right-hand side matches the marginal benefit \( \partial_o v \) on the left. This first-order condition yields the optimal choice of \( s^* \) in (4.10).

The HJB equation (4.7) is combined with several boundary conditions. When \( U \) reaches 0, the contract terminates, and the firm continues production without process intangibles. Therefore, the boundary condition at \( u = 0 \) is

\[
v(0, o) = v^T(o), \quad (4.13)
\]
where the contract termination value $v^T$ satisfies the HJB equation

\[
(r + \delta_K)v^T = \max_{i \geq 0, s \geq 0} \left\{ (v^T - o\partial_o v^T)A_i + (s - (\delta_o - \delta_K)o)\partial_o v^T + \frac{1}{2}o^2\sigma^2\partial^2_{oo} v^T + \mu[1 - \phi + \phi(\theta o)^{1/\psi} - i - s - C_K(i) - C_O(s/o)o] \right\}.
\] (4.14)

After the equation (4.7) is solved, the compensation boundary $\bar{u}$ is determined endogenously via $\bar{u}(o) = \inf\{u : \partial_u v(u, o) = -1\}$. Several other technical boundary conditions are discussed in Appendix C.

### 4.2 Stationary distribution

After the optimal contract and investment strategies are characterized for an individual firm in the previous section, we examine in this section the stationary distribution of the state variables.\(^{23}\) This helps us to better match model predictions with empirical observations.

Because the volatility of $u$ in (4.12) is non-degenerate at $u = 0$, firm liquidation happens with positive probability under the optimal contract. In order to maintain a stationary mass of firms, we introduce firm entry. The stationary density $g$ of the state variable $(u, o)$ satisfies the stationary Fokker-Planck-Kolmogorov equation:

\[
L^{*}_{u,o} g(u, o) + m \psi(u, o) = 0,
\] (4.15)

where $L^{*}_{u,o}$ is the adjoint operator of the infinitesimal generator $L_{u,o}$, $\psi(u, o)$ represents an entry density integrating to one, and $m$ is an entry rate. To ensure that the stationary density $g$ integrates into one, the entry rate $m$ is chosen to match the existing mass:

\[
m = - \int_0^\infty \int_0^{\bar{u}(o)} L^{*}_{u,o} g(u, o) \, du \, do.
\]

Coefficients in the infinitesimal generator $L_{u,o}$ depend on the optimal investment strategies and the agent’s optimal effort under the optimal contract. Therefore, the stationary density $g$ describes the behavior of the equilibrium state variables.

4.3 Quantitative model implications

We examine the quantitative implications of our model in this section. We calibrate several model parameters to the data.

Table 2

The data panel in Table 2 presents the mean physical investment rate and the ratio between total compensation and physical capital in different intangibility quantiles in our sample. Intangibility is measured by the intangible-physical capital ratio. We calibrate our model parameters, in particular the CES parameter $\rho$ in the physical capital accumulation, using these empirical moments. Our calibrated value for $\rho$ is 0.6 and the empirical measurement of $\theta$ using patent data is 0.7, both of which are consistent with their values obtained by Lin (2012).

It follows from (4.1) that the sensitivity of changes in $U$ with respect to changes in $K$ is $\lambda_t$ when the incentive compatibility constraint is satisfied. Table 3 shows that this sensitivity increases with the intangible-physical capital ratio. Motivated by this empirical fact, we assume

$$\lambda_t = \bar{\lambda}_0 t,$$  \hspace{1cm} (4.16)

for a constant parameter $\bar{\lambda}$.

Table 3

We estimate the volatility parameter $\sigma$ using the standard deviation of annual changes in the log physical capital stock. The agent’s impatience parameter $\gamma$ is set so that the level of compensation is consistent with the data. For firm entry, we assume that the principal has all bargaining power so that a new firm starts at $u^e(o)$ which maximizes the principal’s value $v(\cdot, o)$ for a given $o$. The entry density $\psi(u, o)$ is assumed to have the decomposition $\psi(u, o) = \zeta(o)\xi(u|o)$, where $\zeta$ is the density of a log normal distribution with parameters $\mu_\psi$ and $\sigma_\psi$, and $\xi(\cdot|o)$ has a unit mass at $u^e(o)$. \hspace{1cm} (4.24) All other model parameters are summarized in Table 4. They are all consistent with the parameter choice in the literature.

Table 4

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We now present the model implications of our calibrated model. The left panel of Figure 5 presents the compensation boundary \( \bar{u} \) (black solid line) and the mean of agent’s continuation utility conditioning on \( o \) (red dotted line) in the stationary distribution. The compensation boundary \( \bar{u} \) increases with \( o \). This is due to two effects. First, the production function \( d \) of the physical capital investment increases with \( o \). A higher intangible-physical capital ratio improves the efficiency of physical capital investment via process innovation. However, it also elevates the importance of the agent’s effort in the physical capital accumulation. Second, the private benefit rate \( \lambda \) increases with \( o \) in (4.16). Both effects imply that the agency friction worsens with more intangibles. In order to mitigate inefficient liquidation, the principal increases the compensation boundary \( \bar{u} \) to build up the agent’s continuation by deferring more compensations into the future. The conditional mean of \( u \) (red dotted line in the left panel) also increases with \( o \), following the same pattern of the compensation boundary and indicating a positive relationship between the average deferred compensation and intangible capital.

The middle panel of Figure 5 presents the physical investment rate \( i^* \) at the payment boundary \( \bar{u} \) (black solid line) and the mean physical investment rate conditioning on \( o \)

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\[ u = U/K \] and physical investment is \( i^* = I^*/K \). The black solid lines in the left and middle panels present the value of compensation and physical investment at the payment boundary \( \bar{u} \). The red dotted lines plot the mean conditioning on the intangible-physical capital ratio \( o \). The right panel presented the stationary density of \( o \). All parameters are listed in Table 4.

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\[ \mu_\psi = 0.5 \text{ and } \sigma_\psi = 0.5 \text{ so that the median value of } o \text{ in the stationary distribution is consistent with the empirical value in Table 2.} \]
in the stationary distribution. Both investment rates increase with \( o \) and are convex for large value of \( o \). As we will see later, this convexity pattern helps our model to match the elevated empirical physical investment rate in the top quintile of intangibility. Finally, the stationary distribution of the intangible-physical capital ratio, presented in the right panel of Figure 5, displays right skewness.

The model panel in Table 2 presents the model generated average physical investment rate and compensation in different quintiles of the intangible-physical capital ratio. The model generates the elevated physical investment rate in the top quintile, but still underestimates the compensation in the top quintile. Even though the model is not calibrated using the moments conditioning on the process intensity, it performs well conditionally. We split the data into a low process intensity and a high process intensity subsample and measure the conditional moments of the physical investment rate and compensation. We solve the model under low process intensity (0.2) and high process intensity (.0.4) parameters. Our parameter choices are within one standard deviation of the empirical value in the low and high process intensity subsample. The remaining parameters are unchanged from the main calibration. The models generate physical investment rates and compensations that match well with their empirical counterparts.

The impact of \( \theta \) on the compensation is presented in Figure 6. When \( \theta \) increases, more proportion of the intangible capital is used in the product innovation, and less proportion
is used in the process innovation. Given $i$ and $o$, the physical capital accumulation function $d$ decreases in $\theta$ when the agent exerts full effort. As a result, the physical capital accumulation depends less on the agent’s effort and the agency friction is less severe when $\theta$ increases. The left panel of Figure 6 shows that the conditional mean of $u$ decreases with $\theta$, implying that the principal defers less compensation into the future when less proportion of the intangible capital is used for process innovation. The right panel of Figure 6 shows the same pattern as physical capital accumulation becoming less efficient with lower process intensity.

The impact of $\rho$ is presented in Figure 7. When $\rho$ increases, the complementarity between physical capital investment and the intangible capital used in process innovation weakens. Substitution between physical capital investment and process intangibles become easier. Hence, the physical investment rate decreases with $\rho$ in the right panel of Figure 7. Meanwhile, the agent’s effort becomes less important in the physical capital accumulation and the agency friction subsides. The left panel of Figure 7 shows that the impact of $\rho$ on the conditional mean of $u$ is marginal in the calibrated model. But the conditional mean of $u$ decreases with $\rho$ for large values of $o$.

These model predictions on the intensity of process innovation and the complementarity between the physical capital investment and the intangible capital will be tested in our empirical analysis next.
4.4 Testable Implications

This subsection details our testable hypotheses. The model is stylized and contains variables that do not have immediate empirical counterparts, or have many plausible counterparts, so explaining what we will try to verify in data is important. We will explicitly state the claims and hypotheses, and in the end, we have five main claims.

1. Manager compensation and process intangibles are positively related. In the model, the shirking benefit function, \( \Lambda \), is increasing, ceteris paribus, in process intangibles. Therefore, the agent can extract a higher average compensation from the firm (see Figure 6). This, in a broad sense, constitutes the main idea of this paper. Empirically, we equate “manager” with “executive,” allowing us to leverage available data. We expect to find a positive association between process intangibles and executive compensation.

2. Process intangibles and “promised utility” are positive related. In the model, the agent is usually compensated through promised utility, not a continuous wage payment. Once the promised utility reaches a certain threshold, the agent is paid a lump sum. This process then repeats until termination.\(^{25}\) To that end, the most model-relevant measure of compensation is a variable that captures the idea of “promised” utility. We view deferred compensation (awarded stocks and options) as such a variable. In particular, we study the fraction of total compensation that is deferred compensation (henceforth, fraction of compensation deferred). We expect to find a positive association between process intangibles and the fraction of compensation deferred.

3. Salaries of high-skill employees and process intangibles are positively related. In the model, and within the agency theory field in general, the agent is the person whose effort affects output, production, etc. While the decisions of executives do indeed affect firm policy, it is more realistic to say that the efforts of lower ranking employees directly affect output, for example.\(^{26}\) The efficiency of process intangibles is determined by how well skilled employees, particularly those with skills in technology and logistics, employ the intangibles. The efforts of these employees matter, and if they shirk, the process

\(^{25}\)Of course, in reality, compensation is more “continuous.” Adding other forms of payment complicates the model, and we choose to focus on the part of the contract due to agency frictions.

\(^{26}\)There is a tension here. A production line worker at a large firm is close to infinitesimal in the scheme of firm output. The CEO is too far removed to have short term impact on output. We are considering a middle-management, knowledge worker class. These are relatively highly paid and highly educated workers. For example, a senior software engineer or a portfolio manager would fall under this umbrella.
intangibles will not be put to best use. We collect data on the posted wages of these high-skill workers and study the association with process intangibles. **We expect to find a positive association between process intangibles and the wages of high-skill workers.**

4. **Process intangibles and physical investment are complements in physical capital production.** In the model, we have assumed that $\rho < 1$, where $\rho$ affects the physical capital production function $D$ in (3.5). This parameter is important because its sign and magnitude determine the complementarity/substitutability of process intangibles and physical investment. In particular, in this claim, we focus on the sign: Is $\rho < 1$, and, therefore, are process intangibles and physical investment complements? **We expect to find evidence of complementarity.**

5. The positive association between process intangibles and compensation is higher when physical investment is higher. In the model, due to the positive complementarity of physical investment and process intangibles we have:

$$\frac{\partial \Lambda}{\partial I \partial (1 - \theta)} > 0.$$  

That is, the effect of process intangibles of physical capital production is increasing in physical investment. This is a slope effect. Comparing this claim to claim 1. above we can rephrase the current hypothesis as: The effect in claim 1. is increasing in the physical capital investment. The intuition is that, with positive complementarity, these process intangibles are more “important” at these firms and the hold power of the agent is stronger. **We expect the effects of claim 1. to be stronger in high physical investment firms compared to low physical investment firms.**

5 **Data**

This section describes our data sources and how we construct our final data set. We also provide a set of stylized facts. These facts will constitute the main empirical phenomena that we are trying to understand.
5.1 CRSP and Compustat

We begin by describing our data preparation procedure for CRSP/Compustat. These data sets give information on the firm balance sheet and income statement variables. The key variables we will construct from CRSP/Compustat are investment rates (intangible and physical) and capital stocks. We will also construct a number of variables commonly used in the finance literature as controls in our regressions.

We employ a number of standard filters on our data. First, we only retain firms traded on AMEX, NASDAQ, or NYSE stock exchanges. Second, following Fama and French (2015), we drop the first two years a firm appears in the data. Third, we drop firms in the Transportation, Finance, and Public industries. Fourth, we drop micro-cap firms as defined by Fama and French (2015).

We describe the construction of our intangible capital stock and investment variables in the next subsection. We will describe the other CRSP/Compustat variables as we use them, since they are more standard.

5.2 Definition of Intangible Capital

Internally generated intangible capital stocks and their associated investment rates are not reported on firm balance sheets, so we must construct these variables ourselves. To do so, we follow Peters and Taylor (2017). First, if any of the following Compustat variables are NAs, we set the values to 0: xrd (R&D), xsga (Selling, General, and Administrative), rdip (R&D in progress), cogs (Costs of Goods Sold). Second, we construct a variable called SGA.

SGA is defined as follows. If R&D is greater than Selling, General, and Administrative expenses and R&D is less than Costs of Goods Sold, then we set SGA equal to Selling, General, and Administrative expenses. Otherwise, we set SGA equal to Selling, General, and Administrative expenses minus the sum of R&D and R&D in progress.

The third and final part of the Peters and Taylor (2017) method uses the perpetual inventory method to construct the “Knowledge Capital” (\( K_{\text{Know}} \)) and “Organization Cap-

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27 We drop entirely firms that do not have beyond two full years of data.
28 Micro-caps are defined as firms whose market capitalization is less than the market capitalization of the 20th percentile NYSE firm’s size.
29 Our results are similar using the Eisfeldt, Kim and Papanikolaou (2020) method of construction. Results are available upon request.
ital” ($K_{Org}$) stocks.

$$K_{Know,ft} = (1 - \delta_{Know})K_{Know,ft-1} + \frac{R&D_{ft}}{CPI_t}$$  \hspace{1cm} (5.1)

$$K_{Org,ft} = (1 - \delta_{Org})K_{Org,ft-1} + (0.3)\frac{SGA_{ft}}{CPI_t}$$  \hspace{1cm} (5.2)

where $CPI_t$ is the consumer price index.\textsuperscript{30} We follow Ewens, Peters and Wang (2019) when we select $\delta_{Know}$ and $\delta_{Org}$. Ewens, Peters and Wang (2019) show that there is heterogeneity in these parameters across industries.\textsuperscript{31} We use their estimates from their pooled estimation, leading to $\delta_{Know} = 0.28$ and $\delta_{Org} = 0.3$.

We define intangible capital as the sum of Knowledge Capital and Organization Capital, $K_{Int} = K_{Know} + K_{Org}$.\textsuperscript{32} It follows from our definition of intangible capital that we construct intangible investment as $R&D_{ft} + SGA_{ft}$.

\section{5.3 Execucomp}

We use Execucomp to calculate the compensation to top executives at a firm.\textsuperscript{33} Our main measure of compensation from Execucomp is total compensation (data item: TDC1). This total compensation measure includes salary, bonus, long-term incentive plans, option awards, and stock awards.

In order to capture a more direct measure of continuation utility (the variable $U$ in the model), we also look at deferred compensation. FASB Statement NO. 123 (revised 2004), “... requires a public entity to measure the cost of employee services received in exchange for an award of equity instruments based on the grant-date fair value of the award.”\textsuperscript{34} We use this fair value of equity based compensation (e.g., stocks and options) as a measure of future promise utility. We also use the fraction of deferred compensation in the total compensation as another measure of promised utility.

\textsuperscript{30}The CPI is gathered from the Bureau of Economic Analysis.
\textsuperscript{31}For example, their estimates of $\delta_{Know}$ range from 0.18 to 0.31.
\textsuperscript{32}If either $K_{Know}$ or $K_{Org}$ is less than 0, we set $K_{int}$ to zero.
\textsuperscript{33}Execucomp usually includes the compensation for the top five executives at the firm. Sometimes the compensation for the top nine is included.
\textsuperscript{34}Link to statement.
5.4 Burning Glass Technologies

We are interested in, not only, the payments to top executives, but also the payments to specialists/skilled labor. Though executives are unlikely to be directly involved in innovation activities, they are arguably the best positioned to extract rents from the firm. Indeed, most papers studying agency conflicts such as the one we study use data from Execucomp to test their predictions. On the other hand, it is plausible that the skilled labor directly involved with innovation has the most information about the technology in question. Therefore, these workers are also well positioned to extract knowledge based rents.\(^{35}\)

Burning Glass Technologies (BGT) is a labor market data firm that collects vacancy and resume data from the Internet using machine learning techniques. The data set we use is collected by an “electronic spider” that scrapes job posting sites like Indeed.com and Monster.com for information about the vacancies posted there.

BGT collects the unstructured data on the websites and arranges them in a database with standardized variables. This allows cross-firm and intertemporal comparisons. Most importantly for us, BGT standardizes the set of skilled jobs firms looking for. For example, one firm may want to hire someone "proficient at Microsoft Word." Another firm might simply state that "the job will require a good deal of writing, so facility with word processors like Microsoft Word is a must". BGT would assign "Microsoft Word" as a skill for both firms. For each job posting, BGT assigns a number for employee skills.\(^{36}\) These skills are drawn from a list created by BGT, which means that the subjective nature of this data is somewhat reduced.

There are three lists of skills, and the difference between these lists is the level of granularity. For example, the least granular list has 29 different levels, such as Administration, Design, Business, and Health Care. We use the middle list (in terms of granularity) that has 677 levels. Examples of skills here include Litigation, Water Testing and Treatment, and Technical Support.

We classify certain skills as being innovation intensive (II) versus not. We call a job posting an innovation intensive job (II job) posting if it has one of these skills assigned to it.\(^{37}\) Our selection of skills for this categorization is subjective. We ask ourselves “What

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\(^{35}\)For example, if computer programmers are trying to solve a complex and highly specialized problem, it is feasible that even their direct managers cannot tell the difference between slow progress and shirking.

\(^{36}\)7% of jobs have no assigned skills or skills are assigned to a non-existent job posting. We drop these cases from the data set.

\(^{37}\)We require only one skill to be II because it is not true that more non-II skills reduce the innovativeness
skills are associated with the creation of new ideas and products?” Note that this related to, but different than, “high skill.” For example, medical doctors are highly skilled and educated, but we do not consider them to typically be involved in the creation of new products or processes. Consequently, medical doctors are not “innovation” job holders.

Around 5% of all BGT job postings have an associated salary. These are the salaries the employer is offering for the position. For each firm-year, we compute the average II job salary. Similarly, we compute the average II job salary within an industry year.

There are two drawbacks to the BGT data. First, BGT data only go back to 2010. As we will see, this reduces our sample size to less than 2000 observations for regressions using BGT data. Second, as alluded to above, even though the II workers in BGT are the ones actually undertaking the innovative work, it is not clear how much power these workers have to extract rents from shirking. This second drawback is not so problematic, since we have Execucomp data, as well. Our results are consistent with either the executives or the II workers, or both, being subject to agency conflicts.

5.5 Process Claims Data

Our data for process claims comes from the data set compiled by Bena and Simintzi (2019). The authors collect data from the U.S. Patent and Trademark Office (USPTO) up to 2021. They parse the structured-text of each patent to identify the claims section of the patent. Patent claims delineate the scope of the patent in the eyes of the law. To that end, they are important and precisely written. For example, the outcomes of patent infringement lawsuits frequently depend on these claims. Within the claims section of the patent, the authors then classify each claim as being either process or product oriented.

Though definitions are subjective, the existing literature (Bena and Simintzi (2019), Ganglmair, Robinson and Seeligson (2022)) generally defines process innovations as those that improve firm productivity/production methods or reduce costs, meanwhile product innovations introduce new products.

Within each-firm year, we compute the total number of process claims across all patents and divide that sum by the total number of claims, processes, and products. of the job, so to speak. For example, one company may want someone who understands artificial intelligence, while another company wants this same role to also manage people and write reports. The second company’s posting would have a smaller fraction of skills classified as II, but that role just described is no less innovative.

We refer the reader to that paper’s Internet Appendix for further details not discussed here.
This measure aggregates information from all the patents filed by the firm that year. This measure is similar to that used by Bena and Simintzi (2019). Note that in the model process intensity, $1 - \theta$, is a parameter. Our measure of process intensity in the data is allowed to vary by firm-year. However, most of the variation in process intensity can be captured by a firm-level fixed effect.\(^{39}\) Thus, our measures do a good job of sorting firms into different, relatively invariant, groups, which is in line with our theory.

By using patent data to construct the process intensity of the firm, we are assuming that this patent-level measure is a good proxy for the overall-firm level measure. We use the patent data because no firm-level measure of process intensity exists. If, for example, firm-level process intensity, $p^f$, is:

$$p^f = \beta p^p + e$$

where $\beta > 1$, $p^p$ is the patent level measure, and $e$ is noise, then we have classic errors-in-variables on the right-hand side. This will not lead to problems in inference, since we are interested in cross-firm comparisons.

Throughout, we drop firms with no claims of any kind (i.e., no patents). This is implicit in our measures of process intensity that are defined as the number of process claims over total claims.

### 5.6 Summary Statistics

Table 5 displays summary statistics. We allocate firms to different portfolios based on their process intensity measure, $1 - \theta$, and the averages of select variables are computed for each portfolio. The firms are assigned to a portfolio each year.

The first column lists the portfolio, where a higher portfolio number indicates a larger average process intensity. The second column lists what we call the “iB/M” ratio.\(^{40}\) The iB/M ratio is constructed similarly to the classic book-to-market ratio.\(^{41}\) Instead of simply taking the ratio of book equity to market capitalization, we add the intangible capital stock to book equity before computing the ratio. The standard measures of book equity fail to account for internally generated intangibles, which are becoming an increasingly

\(^{39}\)50% of the variation in process intensity is captured by firm-fixed effects. Adding a full set of controls, including industry fixed effects, increases the $R^2$ of the regression by only 8%.

\(^{40}\)This terminology follows Park (2019) and Kazemi (2022).

\(^{41}\)We construct firm book equity following the standard method outlined in, e.g., Bali, Engle and Murray (2016).
important part of the firm’s capital stock.\textsuperscript{42} The iB/M ratio is almost monotonically decreasing in the firm’s process intensity. Though we do not explore the iB/M ratio in the model, this result is in line with the production-based asset pricing literature. According to Lin (2012), as the process intensity increases, the marginal product of physical investment increases. This increase in marginal product increases what Kogan and Papanikolaou (2014) call the “present value of growth opportunities.” Kogan and Papanikolaou (2014) show that, everything else equal, a larger present value of growth opportunities leads to a lower book-to-market ratio.\textsuperscript{43} Thus, this empirical result is consistent with our interpretation of process intensity.

Table 5 Here

The next three columns show the intangible investment rate, physical investment rate, and intangible capital stock, all scaled by physical capital. None of these variables have a monotonic relationship with our process intensity portfolio ranking, but the difference between the averages in the fifth and first portfolios are all large and positive.

The final three columns show our compensation and salary measures.\textsuperscript{44}

6 Empirical Results

This section displays our main empirical results. First, we explore our claims 1-3 in subsection 4.4. Second, we study claims 4-5. This division can be thought of as studying the “direct effect” and the “indirect effect,” respectively, of process intangibles on compensation.

The direct effect says that an increase in process intangibles increases the benefit from shirking, and, therefore, increases optimal compensation. The indirect effect says that the strength of that relationship varies depending on the amount of physical investment and complementarity between investment and process intangibles.

We end with a robustness subsection. In that subsection, we provide three further tests. First, we exploit the granularity of the BGT data and show that II job salaries with a process focus are more affected by firm-level process intensity than II job salaries with a product focus. Second, we re-estimate our Execucomp tests, this time restricting our

\textsuperscript{42}See Peters and Taylor (2017), Kazemi (2022), Belo et al. (2022).
\textsuperscript{43}Kazemi (2022) shows the same result for the iB/M ratio.
\textsuperscript{44}Compensation is scaled by physical capital, and the II salary is scaled by the industry average.
sample to executives who worked in at least two firms. This alleviates concerns about higher pay being a firm characteristic. Third, we study the relationship between uncertainty, compensation, and process intensity. If the compensation and process intensity connection is due to agency frictions, as we propose in this paper, we should see the association between process intensity and compensation strengthen when there are larger agency frictions. We find that this is the case.

6.1 Process Intensity, Compensation, and Salaries

6.1.1 The Direct Effect

We show that higher process intensity is associated with higher total and deferred compensation, as well as higher salaries for II job employees relative to their industry peers. These results correspond to what we called the direct effect of agency frictions on the process intensity-compensation association.

We estimate specifications of the form:

\[
\text{Compensation Measure}_{ft} = y_t + y_j + \beta_1 \text{ProcIn}_{ft} + \beta_2 \text{IK}_{ft} + \beta_3 \text{Size}_{ft} + \beta_4 \frac{iB}{M_{ft}} + \beta_5 \text{Sales}_{ft} + \epsilon_{ft}
\]  

(6.1)

where the variables are the same as before. One difference here is in the final listed control. For example, when use deferred compensation as our dependent variable, we will control for total compensation. Similarly, when the relative II job salary is the dependent variable, we will control for total compensation. The dependent variables are either total compensation divided by physical capital, deferred compensation (stocks and option awards) divided by total compensation, or the posted II job salary relative to industry peers in the same year.

We begin by testing Claim 1. from subsection 4.4, which states that we expect to find a positive relationship between process intangibles and executive compensation. We measure the latter as total compensation or as deferred compensation.

Table 6 displays the results when the dependent variables are either total or deferred compensation, both from Execucomp. Looking at row one, we see that the increases in process intensity are associated with increases in both types of executive compensation. A one standard deviation increase in process intensity is associated with a 5.5% increase in total executive compensation, for a given quantity of intangibles. The effects on de-
ferred compensation are similar. A one standard deviation increase in process intensity is associated with an 6% increase in deferred compensation. The sign on intangible capital (row 2) is positive throughout. This is consistent with the model: We can either think of fixing the intangible capital level and increasing process intensity to increase agency frictions, or we can fix the process intensity but increase the amount of intangibles subject to this emphasis on process innovation to increase agency frictions. Firm size is associated with decreases in total compensation and deferred compensation. This is consistent with results in Hall and Liebman (1998) and Murphy (1999) which document that pay-performance sensitivity decreases with firm size. Higher iB/M ratios are also associated with lower compensation. Book-to-market ratios can be used as measures of performance. For example, a low iB/M ratio implies market values the firm much more than its balance sheet shows. This higher “bang for the buck” could be associated with better management and, therefore, higher pay for executives. Higher sales are associated with higher total compensation and deferred compensation.

Table 6 Here

Now, we test Claim 2. from subsection 4.4. The previous results show that both forms (total and deferred) of compensation are increasing in process intensity, but they do not tell us how the composition of payment changes. According to the model, the payment boundary and average promised utility increase in process intensity (Figure 6). The empirical analog of this result is that deferred compensation should become a larger fraction of total compensation as process intensity increases. We re-estimate equation (6.1) using the ratio of deferred to total compensation as our dependent variable.

Table 7 displays the results. The top row shows us that process intensity increases are associated with a larger fraction of compensation deferred. The same is true of intangible capital (row two). At first blush, it may seem that the two coefficients (on process intensity and intangible capital) should be equal, since it is the quantity \((1 - \theta)\sigma\) that affects agency frictions and investment. However, \(\sigma\) is a state variable, so increasing \(\sigma\) can affect other firm decision variables, which can then affect the coefficient estimates. The key is that both coefficients have the same sign. The effect sizes for process intensity are smaller than in the specifications that look solely at total or deferred compensation. A one standard deviation increase in process intensity is associated with a 1-2% increase in the fraction of total compensation deferred. The fact that this effect is smaller than
the previous ones tells us that process intensity is affecting both the current and future components of compensation.

Table 7 Here

Up to this point, we have examined executive pay, but now we turn to Claim 3 from subsection 4.4, which concerns itself with the pay of high-skilled II job workers. Executives are unlikely to be directly involved in the innovation or investment process. At the same time, executives probably have the most scope for extracting rents from their firms. II job employees, though less powerful than c-suite executives, are directly involved in implementing and developing new processes. It is their efforts that determine success or failure. Because internal process innovations and improvements are inherently opaque (especially to outsiders), it is difficult to assess the efficacy of the hours worked by II employees even if their managers can see that the quantity of working hours is high. In the next table, we will show that II employee wages and salaries are also increasing in process intensity, lending credence to the hypothesis that non-executives can extract rents, too. To the best of our knowledge, we are one of the first papers to test the consequences of dynamic agency theory in compensation of non-executives.

Table 8 shows our estimation results when we use the relative II job salary as the dependent variable in equation (6.1). As explained in the Data section, these are posted salaries, not total wage bills. Therefore, we cannot scale by firm size or capital. Instead, we scale by the leave-one-out industry-year mean of the II job salary. The coefficients can be interpreted as “how much more, relative to similar peers, does a firm pay for a given set of skills?” The first row shows that a one standard deviation increase in process intensity is associated with a 3% increase in relative II job salary. The coefficient on process intensity is similar across specifications. Note that our sample size is much smaller, given that Burning Glass Data only starts in 2010 and not all vacancy postings have wage data.

Table 8 Here

6.1.2 The Indirect Effect

This section verifies what we have called the indirect effect of agency frictions on the process intensity-compensation association. We focus on \( \rho \), which measures the comple-
mentarity between physical investment and process intangibles (see equation 3.6). This section tests Claims 4-6 from subsection 4.4.

We start with Claim 4, which asks about the relationship between physical capital and process intensity. This relationship hinges on the model parameter, \( \rho \). In the model, physical capital growth follows:

\[
\frac{dK_t}{K_t} = \left( -\delta_K + A \left[ a \rho_t + (1 - a) ((1 - \theta) o_t)^{\rho} \right] \right) dt + \sigma dZ_t.
\]

Thus, the parameter \( \rho \) reflects the substitution elasticity. A priori, \( \rho \) could imply that physical investment and process intangibles are substitutes or complements in the physical capital production process. We have assumed they are complements in the model. Here we test that assumption. If physical investment and process intangibles are substitutes, then conditional on the stock of intangibles, a higher process intensity should reduce physical capital investment. The reverse is true if they are complements.

We estimate the following panel regression:

\[
\text{PhysInv}_{f,t} = y_t + y_j + \beta_1 \text{ProcIn}_{f,t} + \beta_2 \text{IK}_{f,t} + \beta_3 iB / M_{f,t} + \beta_4 \text{Size}_{f,t} + \beta_5 \text{Sales}_{f,t} + \varepsilon_{f,t} \tag{6.2}
\]

where \( \text{PhysInv}_{f,t} \) is physical capital investment (Compustat: capex) divided by the physical capital stock.

Table 9 displays the results. Looking at the top row, we see that the sign of the process intensity coefficient is positive across specifications. Taking the final three columns together, we conclude that, on average, physical investment and process intangibles are complementary in the creation of physical capital. That is, \( \rho < 1 \). The effect is statistically and economically significant: a one standard deviation increase in process intensity is associated with around a 3% increase in physical capital investment.

Table 9 Here

This result and the results of Table 1 are (almost) mirror images of each other. We can imagine freezing the amount of intangible capital and physical capital in the firm and simply varying the process intensity of the firm. According to the model, when process intensity is lowered, current sales should go up, and on the flip side, as we increase process intensity, more of the firm’s intangible capital stock is devoted to physical capital production. While this second effect depends on the parameter \( \rho \), under our specification, we have shown that increases in process intensity are associated with increases in
physical capital investment. Thus, we see a clear pendulum: as we vary the process intensity of the firm, we are shifting resources from current sales to “future sales” in the form of investment.

Now, we turn to Claim 5 from subsection 4.4. This claim hinged on the assumption of positive complementarity between process intangibles and physical investment. Estimating equation (6.2) revealed that physical capital investment and process intensity are complements.\(^{45}\) When these two variables are complements, we have:

$$\frac{\partial \Lambda}{\partial i (1 - \theta)} > 0.$$  

This is also Claim 5.

What that equation states is that the benefits of shirking are more sensitive to process intensity for agents employed at firms undertaking more physical capital investment, everything else equal.

We test this implication in two steps. First, we divide firms into high and low physical investment portfolios. Each year, we assign a firm to a portfolio depending on whether that firm is above or below the median physical capital investment rate that year. Second, we re-estimate our compensation equations, allowing the coefficients to vary across the two investment portfolios:

$$\text{Compensation Measure}_{ft} = \alpha_b + \beta_{b1} \text{Process Intensity}_{ft} + X_{ft} \beta_b + \epsilon_{ft}. \quad (6.3)$$

The subscript \(b\) indicates the bin/portfolio. The set of regressors is the same as in the previous compensation equations.

Table 10 displays the results of our estimation. For brevity and space, we have suppressed coefficients on all variables aside from process intensity.\(^{46}\) The first row shows the estimated process intensity-compensation sensitivity for low physical investment firms. The second row shows the sensitivity for high investment firms. Only the first column, total compensation, has a significant association in the first bin. Even then, the point estimate is larger in the second bin. For fraction of compensation deferred and II job salaries, the high physical investment bin is where all the positive, significant association between process intensity and compensation lies. The coefficients in the high

\(^{45}\) Namely, ceteris paribus, an increase in process intensity is associated with an increase in physical capital investment. 

\(^{46}\) The full table is available from the authors upon request.
physical investment bins are larger than the corresponding ones from the earlier compensation regressions (Table 7). This is to be expected, since those earlier regressions are estimated on a pooled sample that includes low complementarity firms.\textsuperscript{47}

Table 10 Here

Claim 6 from subsection 4.4 tells us that we should expect the results of the previous table to vary across levels of complementarity. For example, firms with high complementarity and high physical capital investment should have the largest process intensity-compensation sensitivity.

We estimate:

$$
\frac{K_{f,t+1}}{K_{ft}} = \alpha_f + \beta_{1f} \text{Physical Investment}_{ft} + \beta_{2f} \text{Process Intangibles}
+ \beta_{3f} \text{Physical Investment}_{ft} \times \text{Process Intangibles}_{ft} + \varepsilon_{ft}
$$

(6.4)

Note that we use process intangibles, $o(1 - \theta)$, as a regressor instead of splitting out the intangible capital and process intensity separately. This is because we wish to remain as close as possible to the CES specification in the model.\textsuperscript{48}

Notice that the coefficients have $f$ subscripts. This is because we estimate the previous equation for each firm. \textsuperscript{49} $\beta_{3f}$ is our measure of $1/\rho$. This coefficient measures how much process intensity increases the efficiency of physical investment.\textsuperscript{50}

With these estimates in hand, we then group firms based on $1/\rho_f$. We form two bins based on $\beta_{3f}$.\textsuperscript{51} Firms in the first bin have smaller estimates of $\beta_{3f}$, or, larger estimates of $\rho_f$.

We re-estimate our physical investment bin-interacted regressions within each complementarity bin. To be precise, we estimate two versions of equation (6.3): Once using only low complementarity firms and once using only high complementarity firms. We expect the effects of process intensity on compensation as one moves from low to high

\textsuperscript{47}In fact, some firms may find process intensity and investment to be substitutes. The specificity of firm-level production and capital is not easily captured by any theoretical framework. We choose to focus on the average firm.

\textsuperscript{48}We have estimated the previous equation using intangible capital and process intensity separately, as well. These results are available from the authors upon request.

\textsuperscript{49}We require a minimum of ten years of data for each firm to estimate the equation.

\textsuperscript{50}Note that $1/\rho_f$ and $\beta_{3f}$ are not equivalent. The latter is a proxy of the former. Since we are interested in relative ranks across firms, this estimate is sufficient.

\textsuperscript{51}This is analogous to our physical investment sorting procedure above.
physical investment to be larger in the second regression (within high complementarity firms). Due to data limitations (BGT data only extends to 2010) we only estimate compensation regressions using Execucomp data.

Tables 11 and 12 display the results of the previous procedure. Table 11 shows results for the low complementarity firms. The first column shows that sensitivity to process intensity is actually decreasing in low complementarity firms. While this goes against our theoretical framework, we emphasize that we are not claiming to explain all factors driving executive compensation. We interpret this reversed result as roughly supporting our key mechanism: When complementarity is low, the agency frictions are relaxed and other determinants of pay start to matter more. The second column shows that deferred over total compensation increases as expected with physical investment, even amongst low complementarity firms. However, looking at column two in Table 12, we see that the increase across investment bins is smaller amongst low complementarity firms versus high. This difference is the channel we are emphasizing here. The effects of going from low investment to high investment are greater in the high complementarity firms than the low complementarity firms. Indeed, looking at column one, the results for total compensation are back in line with our theory.

This section has established our main results: Compensation and process intensity are tightly linked. This statement applies to total and deferred executive pay. It also applies to the salaries of highly skilled, innovation-based workers. Finally, these results interact with process innovation and physical capital investment complementarity predictably. The more important process intensity is to the firm’s capital growth process, the more rents the agent can extract, ceteris paribus.

### 6.2 Robustness and Further Results

This subsection presents three robustness and placebo tests meant to rule out alternative hypotheses.

In the first test, we exploit the granularity of the Burning Glass data further and look at the difference in salaries between II jobs with a process focus and II jobs with a more product focus. For example, a chemist working at a drug company is likely working on product innovations (new drugs) versus process innovations. On the other hand, an organizational specialist working on supply chain management is more likely to be

\[52\text{Of course, much of this is subjective. Is discovering a new chemical compound that leads to multiple new drugs a process or product innovation?}\]
working on process innovations. We exploit the richness of the Burning Glass Data to distinguish between these types of skills. We define “R&D skills” to be those focused on new scientific discoveries or similar advances.\textsuperscript{53} Examples include Medical Research, Quantum Mechanics, Neuroscience, and Clinical Research. The other category we define is “process skills”. These could also be called organizational capital skills.\textsuperscript{54} Examples here include Logistics, Process Improvement, Operations Analysis, and Supplier Relationship Management. With these definitions in hand, we re-estimate our II job salary regressions using each of these subcategories as the dependent variable.\textsuperscript{55} Our hypothesis is that Process skill job salaries should be more affected by the agency frictions in our model, and, thus, their salaries should be increasing more in process intensity.\textsuperscript{56}

Table \textbf{13} displays the results of estimating equation (6.1) restricted to either R&D skills jobs salaries or Process skills jobs salaries. The set of controls is the same as in the pooled regression. The top row shows that there is no significant relationship between process intensity and the salaries of R&D skills jobs. However, process intensity is associated with an increase in process skills jobs’ salaries. Notice that the coefficient is similar in magnitude to the analogous coefficients in Table \textbf{8}. Within the set of skilled, II job employees, some are more likely to be involved with process innovations relative to product innovations. We have shown that, even with such a fine distinction, there are differences in the sensitivities of salaries to process intensity. This is in line with the empirical interpretation of the model: Those employees best able to extract rents from the process innovations are the ones whose compensation should increase with process intensity.

\begin{center}
\textbf{Table 13 Here}
\end{center}

In the second test, we ask if better executives simply self-select into process intense firms and hence receive higher compensation. To test this, we estimate executive-level regressions on the subset of executives who change firms in our sample. That is, we

\textsuperscript{53}While this is not the same as product innovation, as our previous example showed, they are more likely to be tilted in that direction.

\textsuperscript{54}Eisfeldt and Papanikolaou (2013).

\textsuperscript{55}When we compute the leave-one-out industry average, we do so using the subcategory.

\textsuperscript{56}Our classification system for jobs can assign a job to both categories simultaneously. This makes sense since high skill employees are frequently asked to have an assortment of skills.
This equation looks similar to (6.1). The key difference is in the dependent variable, which is measured at the executive-firm-date (ift) level. In (6.1) we looked at firm-date level regressions.

Table 14 displays the results. The coefficient estimates on process intensity are similar to the full sample, firm-level estimates: A one standard deviation increase in process intensity is associated with a 9% increase in executive compensation.

Table 14 Here

In our third test, we look at firm-level measures of agency frictions to test if the association between process intensity and compensation is higher when frictions are stronger. According to the model, firms facing more uncertainty in the physical capital growth process should be exposed to greater agency frictions.

We compute and industry-year level of uncertainty and assign all firms within that industry-year to have the same level of uncertainty. Within each industry-year, we calculate the cross-sectional variance of physical capital growth at the firm level. This is our measure of $\sigma$ in the model and proxies for the severity of agency frictions.

We split the firm-level distribution of uncertainty into a high and low bin each year. Then, like our complementarity regressions, we estimate the following panel regressions:

\[
\text{Compensation Measure}_{ft} = \alpha_b + \beta_{b1} \text{Process Intensity}_{ft} + \mathbf{X}_{ft} \beta_b + \epsilon_{ft} \quad (6.6)
\]

where now $b$ refers to the uncertainty bin.

Table 15 displays the results. Once again, we see that firms in the high bin have significant and positive coefficients on process intensity. On top of that, the high bin coefficients are always larger than the low bin ones. For example, when moving from a low uncertainty to high uncertainty firm, the association between process intensity and total executive compensation increases by 4%.

Table 15 Here
7 Conclusion

We presented and studied a new empirical fact: Higher process intensity is associated with higher pay for executives and skilled employees. To rationalize this fact, we developed a dynamic principal-agent model in which agent effort determined the efficacy of process intangibles on the physical capital growth process. The model delivered two key channels: A direct effect and indirect effect of process intensity on compensation. The direct effect states that higher process intensity increases the benefits of shirking, so the agent must be further compensated to ensure full effort. The indirect effect states that for a given level of process intensity, higher process intangible-physical capital investment complementarity increases the hold up power the agent has over the firm. This leads to a larger effect of process intensity on compensation for all levels of process intensity.

We verified these two main effects in the data using measures of executive and skilled labor pay. Our baseline specifications showed that a one standard deviation increase in process intensity is associated with an 8% increase in executive pay and a 3% increase in skilled labor pay. When process intangible-physical investment complementarity is high (i.e., the hold up problem is serious), these numbers increase to 16% and 5%, respectively.

We have taken the level of process intensity as given. However, even if changing this ratio is costly, over the medium to long-term we expect it to be endogenous. Studying this choice is left for future work. We have also not considered the asset pricing implications of process intensity and agency. Adding a stochastic discount factor as in Kogan and Papanikolaou (2014) to the model would provide further interesting testable implications.
References


Bena, Jan, and Elena Simintzi. 2019. “Machines could not compete with Chinese labor: Evidence from US firms innovation.” *Available at SSRN 2613248.*


Table 1: Sales and Process Intensity

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<thead>
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<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
<th>(4)</th>
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<tbody>
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<td><strong>Dependent variable:</strong></td>
<td>Sales / Physical Capital</td>
<td>Sales / Physical Capital</td>
<td>Sales / Physical Capital</td>
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<tr>
<td>Process Intensity</td>
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<td>-0.110***</td>
<td>-0.126***</td>
<td>-0.114***</td>
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<td>(0.012)</td>
<td>(0.011)</td>
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<td></td>
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<td>(0.014)</td>
<td>(0.013)</td>
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<tr>
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<td>-0.034*</td>
<td>-0.042**</td>
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<tr>
<td></td>
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</table>

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between firm-level sales (divided by physical capital) and process intensity (in standard deviation units). The results come from estimating panel regressions (2.1). Data definitions are found in the Data section of the paper, and Compensation refers to total executive compensation from Execucomp. All variables aside from process intensity are logged. Industry fixed effects use one broad one digit SIC codes (after 2002) or one digit NAICS (before 2002), and date fixed effects correspond to years. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
Table 2: Quintile Statistics

<table>
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<tbody>
<tr>
<td></td>
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</tr>
<tr>
<td>Intangible $O/K$</td>
<td>0.378</td>
</tr>
<tr>
<td>Physical Investment $I/K$</td>
<td>0.080</td>
</tr>
<tr>
<td>Compensation $U/K$</td>
<td>0.005</td>
</tr>
</tbody>
</table>

| Data | Low Process Intensity | Physical Investment $I/K$ | 0.083   | 0.087   | 0.096   | 0.120   | 0.148   |
|      | Mean: 0.18, Std: 0.095 | Compensation $U/K$ | 0.006   | 0.009   | 0.016   | 0.032   | 0.070   |
|      | High Process Intensity | Physical Investment $I/K$ | 0.084   | 0.091   | 0.113   | 0.138   | 0.164   |
|      | Mean: 0.45, Std: 0.215 | Compensation $U/K$ | 0.005   | 0.009   | 0.018   | 0.045   | 0.076   |

| Model | Low Process Intensity | Physical Investment $I/K$ | 0.057   | 0.075   | 0.084   | 0.103   | 0.125   |
|       | $1 - \theta = 0.2$ | Compensation $U/K$ | 0.024   | 0.034   | 0.041   | 0.052   | 0.057   |
|       | High Process Intensity | Physical Investment $I/K$ | 0.090   | 0.107   | 0.117   | 0.132   | 0.165   |
|       | $1 - \theta = 0.4$ | Compensation $U/K$ | 0.030   | 0.038   | 0.045   | 0.055   | 0.068   |

This table presents the physical investment and compensation for intangible quintiles in both data and the calibrated model. Intangible quantiles are formed by $O/K$. Physical investment is defined as capital expenditures (Compustat item CAPX) divided by lagged physical capital (PPEGT). Compensation is total compensation (Execucomp item TDC1) divided by lagged physical capital. In the calibrated model, quintiles are formed using the stationary distribution of $O/K$. The physical investment is the conditional mean of $I/K$ and the compensation is the conditional mean of $U/K$. 

\[ 1 - \theta = 0.2 \]
Table 3: Ratio between the compensation volatility and the physical capital volatility

This table presents the ratio between the compensation volatility and the physical capital volatility conditional on the intangible capital. This ratio increases with intangibles, but is insensitive to process intensity in each intangible quintile. Intangible quintiles are constructed using ratio between intangible capital and physical capital. The volatilities are measured over five year rolling windows at the firm level. Within each intangibility quintile-year, we compute the median level of the firm-level volatilities. Then, we compute the time-series average of these medians.

<table>
<thead>
<tr>
<th>Intangible quintiles</th>
<th>1</th>
<th>2</th>
<th>3</th>
<th>4</th>
<th>5</th>
</tr>
</thead>
<tbody>
<tr>
<td>Volatility ratio (%)</td>
<td>0.15</td>
<td>0.30</td>
<td>0.46</td>
<td>1.07</td>
<td>1.85</td>
</tr>
</tbody>
</table>

Table 3: Ratio between the compensation volatility and the physical capital volatility
Table 4: Model Parameters

<table>
<thead>
<tr>
<th>Symbol</th>
<th>Variable Description</th>
<th>Value</th>
<th>Reference</th>
</tr>
</thead>
<tbody>
<tr>
<td>$a$</td>
<td>Weight of the physical investment in the physical capital accumulation</td>
<td>0.65</td>
<td>Calibrated from data</td>
</tr>
<tr>
<td>$A$</td>
<td>Scale parameter in the physical capital accumulation</td>
<td>0.45</td>
<td>Calibrated from data</td>
</tr>
<tr>
<td>$\theta$</td>
<td>Percentage of intangibles used in the product innovation</td>
<td>0.7</td>
<td>Calibrated from data</td>
</tr>
<tr>
<td>$\rho$</td>
<td>CES parameter in the physical capital accumulation</td>
<td>0.6</td>
<td>Calibrated from data</td>
</tr>
<tr>
<td>$\lambda$</td>
<td>Agent shirking benefit parameter</td>
<td>0.03</td>
<td>Calibrated from data</td>
</tr>
<tr>
<td>$\phi$</td>
<td>Weight of the product intangibles used in the production</td>
<td>0.2</td>
<td>Calibrated from data</td>
</tr>
<tr>
<td>$\psi$</td>
<td>CES parameter in the production</td>
<td>0.3</td>
<td>Calibrated from data</td>
</tr>
<tr>
<td>$\sigma$</td>
<td>Volatility of $\log K$</td>
<td>0.29</td>
<td>Calibrated from data</td>
</tr>
<tr>
<td>$\delta_K$</td>
<td>Physical capital deprecation rate</td>
<td>0.12</td>
<td>Lin (2012)</td>
</tr>
<tr>
<td>$\delta_O$</td>
<td>Intangible capital deprecation rate</td>
<td>0.2</td>
<td>Lin (2012)</td>
</tr>
<tr>
<td>$Q_K$</td>
<td>Scale parameter of the physical investment adjustment cost</td>
<td>56.55 (monthly)</td>
<td>Eisfeldt and Papanikolaou (2013)</td>
</tr>
<tr>
<td>$c_K$</td>
<td>Convexity parameter of the physical investment adjustment cost</td>
<td>2 (monthly)</td>
<td>Eisfeldt and Papanikolaou (2013)</td>
</tr>
<tr>
<td>$Q_O$</td>
<td>Scale parameter of the intangible investment adjustment cost</td>
<td>625 (monthly)</td>
<td>Eisfeldt and Papanikolaou (2013)</td>
</tr>
<tr>
<td>$c_O$</td>
<td>Convexity parameter of the physical investment adjustment cost</td>
<td>3.2 (monthly)</td>
<td>Eisfeldt and Papanikolaou (2013)</td>
</tr>
<tr>
<td>$\mu$</td>
<td>Productivity rate</td>
<td>0.45</td>
<td>Ward (2022)</td>
</tr>
<tr>
<td>$\gamma$</td>
<td>Agent impatient parameter</td>
<td>0.08</td>
<td>Calibration from data</td>
</tr>
<tr>
<td>$r$</td>
<td>Interest rate</td>
<td>0.06</td>
<td></td>
</tr>
</tbody>
</table>

This table shows the parameters used in our simulations. Citations are given for the parameters which are consistent with the existing literature. The parameters $Q_K$ and $Q_O$ are in monthly unit. To annualize them, we use the following argument from the first best case. Let $b$ be the marginal benefit of investment $i$. Then the first-best optimization in $i$ is $\max_{i \geq 0} \{ b \cdot i - i - \frac{Q_K}{c_K} \bar{r}^k \}$. Therefore, the optimal monthly investment is $i^* = \left( \frac{b}{Q_K} \right)^{\frac{1}{c_K}}$. Scaling monthly $Q_K$ to $Q_K/12^{c_K-1}$, we obtain the annualized optimal investment $12i = \left( \frac{b-1}{Q_K/12^{c_K-1}} \right)^{\frac{1}{c_K-1}}$. 

50
Table 5: Summary Statistics by Process Intensity

<table>
<thead>
<tr>
<th></th>
<th></th>
<th></th>
<th></th>
<th></th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>0.637</td>
<td>0.724</td>
<td>0.117</td>
<td>2.796</td>
<td>49.690</td>
<td>0.945</td>
</tr>
<tr>
<td>2</td>
<td>0.672</td>
<td>0.514</td>
<td>0.131</td>
<td>1.997</td>
<td>51.442</td>
<td>1.025</td>
</tr>
<tr>
<td>3</td>
<td>0.532</td>
<td>0.720</td>
<td>0.139</td>
<td>3.327</td>
<td>65.725</td>
<td>1.094</td>
</tr>
<tr>
<td>4</td>
<td>0.520</td>
<td>0.943</td>
<td>0.149</td>
<td>3.834</td>
<td>136.783</td>
<td>1.143</td>
</tr>
<tr>
<td>5</td>
<td>0.529</td>
<td>1.495</td>
<td>0.146</td>
<td>6.599</td>
<td>155.727</td>
<td>1.122</td>
</tr>
</tbody>
</table>

This table shows means of select variables by process intensity level. Each year firms are sorted into five equally spaced portfolios (bins) based on their values of $1 – \theta$. Bins are re-balanced each year. Averages of the variables displayed as column titles are computed for each bin. The iB/M ratio is the book-to-market ratio when book equity has intangible capital added in. All variables aside from the iB/M ratio and II Job Wage are relative to the firm’s physical capital stock. Intangible investment is computed as described in the Data section. Physical capital investment is capital expenditures in Compustat. The intangible capital stock is as described in the Data section. Compensation is total compensation from Execucomp. II Job Wage is the average posted II job wage from BGT divided by the industry average II job wage. Deferred compensation is stock and option awards from Execucomp.
Table 6: Executive Compensation and Process Intensity

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>Total Compensation / Physical Capital</th>
<th>Deferred Compensation / Physical Capital</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
</tr>
<tr>
<td>Process Intensity</td>
<td>0.037*</td>
<td>0.068***</td>
</tr>
<tr>
<td></td>
<td>(0.020)</td>
<td>(0.011)</td>
</tr>
<tr>
<td>Intangible Capital</td>
<td>0.879***</td>
<td>0.702***</td>
</tr>
<tr>
<td></td>
<td>(0.029)</td>
<td>(0.020)</td>
</tr>
<tr>
<td>Size</td>
<td>−0.982***</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.023)</td>
<td></td>
</tr>
<tr>
<td>iB/M Ratio</td>
<td>−0.569***</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.013)</td>
<td></td>
</tr>
<tr>
<td>Sales</td>
<td>0.227***</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.024)</td>
<td></td>
</tr>
<tr>
<td>Fixed effects</td>
<td>Industry + Date</td>
<td>Industry + Date</td>
</tr>
<tr>
<td>Observations</td>
<td>11,590</td>
<td>11,398</td>
</tr>
</tbody>
</table>

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between firm-level executive compensation (divided by physical capital) and process intensity (in standard deviation units). Executive compensation is either total compensation or deferred compensation. The results come from estimating panel regressions (6.1). Data definitions are found in the Data section of the paper, and Compensation refers to total executive compensation from Execucomp. All variables aside from process intensity are logged. Industry fixed effects use one broad one digit SIC codes (after 2002) or one digit NAICS (before 2002), and date fixed effects correspond to years. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
Table 7: Fraction of Compensation Deferred and Process Intensity

<table>
<thead>
<tr>
<th></th>
<th>Dependent variable:</th>
<th>Deferred Compensation / Total Compensation</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>(1)</td>
</tr>
<tr>
<td>Process Intensity</td>
<td>0.034***</td>
<td>0.022**</td>
</tr>
<tr>
<td></td>
<td>(0.012)</td>
<td>(0.010)</td>
</tr>
<tr>
<td>Intangible Capital</td>
<td>0.054***</td>
<td>0.114***</td>
</tr>
<tr>
<td></td>
<td>(0.014)</td>
<td>(0.016)</td>
</tr>
<tr>
<td>Size</td>
<td></td>
<td>0.137***</td>
</tr>
<tr>
<td></td>
<td>(0.010)</td>
<td>(0.010)</td>
</tr>
<tr>
<td>iB/M Ratio</td>
<td></td>
<td>−0.035</td>
</tr>
<tr>
<td></td>
<td>(0.023)</td>
<td>(0.023)</td>
</tr>
<tr>
<td>Sales</td>
<td></td>
<td>−0.086***</td>
</tr>
<tr>
<td></td>
<td>(0.021)</td>
<td>(0.021)</td>
</tr>
</tbody>
</table>

Fixed effects: Industry + Date

Observations 5,408 5,306

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between the firm-level fraction of executive compensation deferred and process intensity (in standard deviation units). The dependent variable is the ratio of deferred compensation in the form of stock and option grants to total executive compensation. The results come from estimating panel regressions (6.1). Data definitions are found in the Data section of the paper. All variables aside from process intensity are logged. Industry fixed effects use one broad one digit SIC codes (after 2002) or one digit NAICS (before 2002), and date fixed effects correspond to years. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
Table 8: Innovation Intensive Salaries and Process Intensity

<table>
<thead>
<tr>
<th></th>
<th>(1)</th>
<th>(2)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Process Intensity</td>
<td>0.040***</td>
<td>0.032*</td>
</tr>
<tr>
<td></td>
<td>(0.018)</td>
<td>(0.018)</td>
</tr>
<tr>
<td>Intangible Capital</td>
<td>0.042***</td>
<td>0.074***</td>
</tr>
<tr>
<td></td>
<td>(0.015)</td>
<td>(0.021)</td>
</tr>
<tr>
<td>Size</td>
<td></td>
<td>0.058*</td>
</tr>
<tr>
<td></td>
<td></td>
<td>(0.030)</td>
</tr>
<tr>
<td>iB/M Ratio</td>
<td></td>
<td>0.064***</td>
</tr>
<tr>
<td></td>
<td></td>
<td>(0.011)</td>
</tr>
<tr>
<td>Sales</td>
<td>−0.078**</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.035)</td>
<td></td>
</tr>
<tr>
<td>Fixed effects</td>
<td>Industry + Date</td>
<td>Industry + Date</td>
</tr>
<tr>
<td>Observations</td>
<td>1,562</td>
<td>1,525</td>
</tr>
</tbody>
</table>

Note: *p < 0.1; **p < 0.05; ***p < 0.01

This table shows the relationship between the firm-level average innovation intensive job salary (relative to industry average) and process intensity (in standard deviation units). The dependent variable is defined as the average posted wage for innovation intensive jobs in a given firm-year from Burning Glass Technologies. The industry average used to normalize is a leave-one-out mean of the firm-year innovation intensive salaries. The results come from estimating panel regressions (6.1). Data definitions are found in the Data section of the paper. All variables aside from process intensity are logged. Industry fixed effects use one broad one digit SIC codes (after 2002) or one digit NAICS (before 2002), and date fixed effects correspond to years. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
Table 9: Physical Investment and Process Intensity

<table>
<thead>
<tr>
<th></th>
<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
<th>(4)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Physical Capital Investment</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Process Intensity</td>
<td>−0.001</td>
<td>0.021***</td>
<td>0.031***</td>
<td>0.035***</td>
</tr>
<tr>
<td></td>
<td>(0.009)</td>
<td>(0.008)</td>
<td>(0.009)</td>
<td>(0.008)</td>
</tr>
<tr>
<td>Intangible Capital</td>
<td>0.159***</td>
<td>0.031***</td>
<td>0.209***</td>
<td>0.118***</td>
</tr>
<tr>
<td></td>
<td>(0.010)</td>
<td>(0.011)</td>
<td>(0.011)</td>
<td>(0.012)</td>
</tr>
<tr>
<td>Size</td>
<td>−0.339***</td>
<td></td>
<td>−0.320***</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.013)</td>
<td></td>
<td>(0.014)</td>
<td></td>
</tr>
<tr>
<td>iB/M Ratio</td>
<td>−0.047***</td>
<td></td>
<td>0.024***</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.006)</td>
<td></td>
<td>(0.007)</td>
<td></td>
</tr>
<tr>
<td>Sales</td>
<td>0.341***</td>
<td></td>
<td>0.289***</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.018)</td>
<td></td>
<td>(0.018)</td>
<td></td>
</tr>
</tbody>
</table>

Fixed effects: No Industry + Date  No Industry + Date
Observations: 21,675 21,189 21,675 21,189

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between firm-level physical capital investment (divided by physical capital) and process intensity (in standard deviation units). The results come from estimating panel regressions (6.2). Data definitions are found in the Data section of the paper, and Compensation refers to total executive compensation from Execucomp. All variables aside from process intensity are logged. Industry fixed effects use one broad one digit SIC codes (after 2002) or one digit NAICS (before 2002), and date fixed effects correspond to years. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
Table 10: Process Intensity, Compensation, and Investment Level

<table>
<thead>
<tr>
<th></th>
<th>Total Compensation</th>
<th>Fraction Compensation</th>
<th>Deferred</th>
<th>Skilled Wage Relative to Industry Average</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Bin 1 × Process Intensity</strong></td>
<td>0.067***</td>
<td>−0.004</td>
<td>0.027</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.014)</td>
<td>(0.015)</td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Bin 2 × Process Intensity</strong></td>
<td>0.097***</td>
<td>0.039***</td>
<td>0.059**</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.017)</td>
<td>(0.013)</td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Fixed effects</strong></td>
<td>Bin+Date</td>
<td>Bin+Date</td>
<td>Bin+Date</td>
<td></td>
</tr>
<tr>
<td><strong>Controls?</strong></td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td></td>
</tr>
<tr>
<td><strong>Observations</strong></td>
<td>11,337</td>
<td>5,305</td>
<td>1,524</td>
<td></td>
</tr>
</tbody>
</table>

**Note:** *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between different firm-level compensation measures and process intensity (in standard deviation units) interacted with an indicator for the level of physical investment. Firms in Bin 1 have low physical capital investment, and vice versa for firms in Bin 2. In column one, the dependent variable is total executive compensation divided by physical capital. In column two it is deferred compensation divided by physical capital. In column three it is the average posted wage for innovation intensive jobs in a given firm-year from Burning Glass Technologies. The industry average used to normalize the Burning Glass salaries is a leave-one-out mean of the firm-year innovation intensive salaries. The results come from estimating panel regressions (6.3). Data definitions are found in the Data section of the paper. All variables aside from process intensity are logged. Fixed effects are at the bin and date level. Clustering is performed at the firm level.
Table 11: Process Intensity, Compensation, and Investment Level: Low Complementarity

<table>
<thead>
<tr>
<th></th>
<th>Total Compensation</th>
<th>Deferred Compensation</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
</tr>
<tr>
<td>Bin 1 × Process Intensity</td>
<td>0.169***</td>
<td>−0.002</td>
</tr>
<tr>
<td></td>
<td>(0.031)</td>
<td>(0.029)</td>
</tr>
<tr>
<td>Bin 2 × Process Intensity</td>
<td>0.138***</td>
<td>0.052***</td>
</tr>
<tr>
<td></td>
<td>(0.032)</td>
<td>(0.019)</td>
</tr>
<tr>
<td>Fixed effects</td>
<td>Bin+Date</td>
<td>Bin+Date</td>
</tr>
<tr>
<td>Controls?</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Observations</td>
<td>4,494</td>
<td>2,099</td>
</tr>
</tbody>
</table>

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between different firm-level compensation measures and process intensity (in standard deviation units) interacted with an indicator for the level of physical investment. This sample restricts firms to have low physical investment-process intensity complementarity. Complementarity is estimated by equation (6.4). Firms in Bin 1 have low physical capital investment, and vice versa for firms in Bin 2. In column one, the dependent variable is total executive compensation divided by physical capital. In column two it is deferred compensation divided by physical capital. In column three it is the average posted wage for innovation intensive jobs in a given firm-year from Burning Glass Technologies. The industry average used to normalize the Burning Glass salaries is a leave-one-out mean of the firm-year innovation intensive salaries. The results come from estimating panel regressions (6.3). Data definitions are found in the Data section of the paper. All variables aside from process intensity are logged. Fixed effects are at the bin and date level. Clustering is performed at the firm level.
Table 12: Process Intensity, Compensation, and Investment Level: High Complementarity

<table>
<thead>
<tr>
<th></th>
<th>Total Compensation</th>
<th>Deferred Compensation</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bin 1 × Process Intensity</td>
<td>0.039 (0.030)</td>
<td>-0.018 (0.022)</td>
</tr>
<tr>
<td>Bin 2 × Process Intensity</td>
<td>0.165*** (0.036)</td>
<td>0.092*** (0.029)</td>
</tr>
</tbody>
</table>

Fixed effects Bin+Date        Bin+Date
Controls? Yes                 Yes
Observations 4,674            2,102

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between different firm-level compensation measures and process intensity (in standard deviation units) interacted with an indicator for the level of physical investment. This sample restricts firms to have high physical investment-process intensity complementarity. Complementarity is estimated by equation (6.4). Firms in Bin 1 have low physical capital investment, and vice versa for firms in Bin 2. In column one, the dependent variable is total executive compensation divided by physical capital. In column two it is deferred compensation divided by physical capital. In column three it is the average posted wage for innovation intensive jobs in a given firm-year from Burning Glass Technologies. The industry average used to normalize the Burning Glass salaries is a leave-one-out mean of the firm-year innovation intensive salaries. The results come from estimating panel regressions (6.3). Data definitions are found in the Data section of the paper. All variables aside from process intensity are logged. Fixed effects are at the bin and date level. Clustering is performed at the firm level.
Table 13: Process versus R&D Jobs

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>Process Skill Salaries</th>
<th>R&amp;D Skill Salaries</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
</tr>
<tr>
<td>Process Intensity</td>
<td>0.034**</td>
<td>0.024</td>
</tr>
<tr>
<td></td>
<td>(0.017)</td>
<td>(0.022)</td>
</tr>
<tr>
<td>Intangible Capital</td>
<td>0.080***</td>
<td>0.086***</td>
</tr>
<tr>
<td></td>
<td>(0.025)</td>
<td>(0.030)</td>
</tr>
<tr>
<td>Size</td>
<td>0.018</td>
<td>−0.014</td>
</tr>
<tr>
<td></td>
<td>(0.020)</td>
<td>(0.025)</td>
</tr>
<tr>
<td>iB/M Ratio</td>
<td>0.042***</td>
<td>0.029**</td>
</tr>
<tr>
<td></td>
<td>(0.009)</td>
<td>(0.012)</td>
</tr>
<tr>
<td>Sales</td>
<td>−0.083**</td>
<td>−0.067</td>
</tr>
<tr>
<td></td>
<td>(0.040)</td>
<td>(0.043)</td>
</tr>
<tr>
<td>Fixed effects</td>
<td>Industry + Date</td>
<td>Industry + Date</td>
</tr>
<tr>
<td>Observations</td>
<td>1,507</td>
<td>1,010</td>
</tr>
</tbody>
</table>

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between different II job salaries and process intensity (in standard deviation units). In column one, the dependent variable is the average firm-date salary for process focused jobs divided by the leave-one-out mean of firm-date salaries. In column two it is the average firm-date salary for R&D or product focused jobs divided by the leave-one-out mean of firm-date salaries. The results come from estimating panel regressions (6.1). Data definitions are found in the Data section of the paper. All variables aside from process intensity are logged. Fixed effects are at the bin-date level. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
Table 14: Compensation of Executives Who Switch Firms

<table>
<thead>
<tr>
<th></th>
<th>Total Compensation</th>
<th>Deferred Compensation</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
</tr>
<tr>
<td>Process Intensity</td>
<td>0.086***</td>
<td>0.092***</td>
</tr>
<tr>
<td></td>
<td>(0.018)</td>
<td>(0.021)</td>
</tr>
<tr>
<td>Intangible Capital</td>
<td>0.481***</td>
<td>0.554***</td>
</tr>
<tr>
<td></td>
<td>(0.022)</td>
<td>(0.025)</td>
</tr>
<tr>
<td>Size</td>
<td>−0.524***</td>
<td>−0.430***</td>
</tr>
<tr>
<td></td>
<td>(0.011)</td>
<td>(0.014)</td>
</tr>
<tr>
<td>iB/M Ratio</td>
<td>−0.488***</td>
<td>−0.503***</td>
</tr>
<tr>
<td></td>
<td>(0.032)</td>
<td>(0.035)</td>
</tr>
<tr>
<td>Sales</td>
<td>0.583***</td>
<td>0.530***</td>
</tr>
<tr>
<td></td>
<td>(0.028)</td>
<td>(0.035)</td>
</tr>
</tbody>
</table>

Fixed effects Industry + Date Industry + Date
Observations 5,639 5,802

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between executive compensation (divided by the capital stock) and process intensity (in standard deviation units) amongst the set of executives who switch firms at least once in our sample. In column one, the dependent variable is the total executive compensation. In column two it is the deferred executive compensation. Data definitions are found in the Data section of the paper. All variables aside from process intensity are logged. Fixed effects are at the bin-date level. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
Table 15: Process Intensity, Compensation, and Uncertainty

<table>
<thead>
<tr>
<th></th>
<th>Dependent variable:</th>
<th></th>
<th></th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Total Compensation</td>
<td>Deferred Compensation</td>
<td>Skilled Wage Relative to Industry Average</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Bin 1 × Process Intensity</td>
<td>0.083***</td>
<td>0.080***</td>
<td>0.004</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.011)</td>
<td>(0.024)</td>
<td>(0.022)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Bin 2 × Process Intensity</td>
<td>0.128***</td>
<td>0.143***</td>
<td>0.053**</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.014)</td>
<td>(0.022)</td>
<td>(0.026)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Fixed effects</td>
<td>Bin-Date</td>
<td>Bin-Date</td>
<td>Bin-Date</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Controls?</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Observations</td>
<td>11,633</td>
<td>5,334</td>
<td>1,316</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between different firm-level compensation measures and process intensity (in standard deviation units) interacted with an indicator for the level of capital growth uncertainty. Firms in Bin 1 have low physical capital growth uncertainty and vice versa for firms in Bin 2. In column one, the dependent variable is total executive compensation divided by physical capital. In column two it is deferred compensation divided by physical capital. In column three it is the average posted wage for innovation intensive jobs in a given firm-year from Burning Glass Technologies. The industry average used to normalize the Burning Glass salaries is a leave-one-out mean of the firm-year innovation intensive salaries. The results come from estimating panel regressions (6.6). Data definitions are found in the Data section of the paper. All variables aside from process intensity are logged. Fixed effects are at the bin-date level. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
<table>
<thead>
<tr>
<th>Specification</th>
<th>Ratio Regression ($\tilde{\beta}_5$)</th>
<th>Levels Regressions ($\beta_4$)</th>
<th>Levels Regressions ($\beta_5$)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Mean Industry Coef.</td>
<td>0.105</td>
<td>1.152</td>
<td>0.055</td>
</tr>
<tr>
<td>Full Sample (S.E.)</td>
<td>0.001 (0.005)</td>
<td>0.012*** (0.004)</td>
<td>0.010** (0.004)</td>
</tr>
<tr>
<td>Pos. Proc. In. (S.E.)</td>
<td>0.014** (0.005)</td>
<td>0.016** (0.006)</td>
<td>0.006 (0.004)</td>
</tr>
</tbody>
</table>

This table shows the difference in complementarity (with physical capital investment) between process intangibles and product intangibles. Each row displays our measure of complementarity using a different specification or sample. Each column displays a coefficient related to complementarity. The first row shows the average coefficient across industries when regressions (E.1) and (E.2) are estimated within industry. The second row shows estimated coefficients from those specifications when using a pooled sample panel regression. The third row shows estimates from those specifications using only the sample of firms with non-zero process intensity, again using a panel regression. The first column is the estimate of $\tilde{\beta}_5$ from equation (E.2). The second and third columns are estimates of $\beta_4$ and $\beta_5$ from regression (E.1). Standard errors are in parentheses (clustered at the industry-date level, where industry is 4 digit NAICS after 2002 and 3 digit SIC before 2002).
Table 17: Physical Investment and Process Intensity (EP 2013 Method)

<table>
<thead>
<tr>
<th></th>
<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
<th>(4)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Process Intensity</td>
<td>0.016 *</td>
<td>0.039 ***</td>
<td>0.040 ***</td>
<td>0.039 ***</td>
</tr>
<tr>
<td></td>
<td>(0.009)</td>
<td>(0.008)</td>
<td>(0.009)</td>
<td>(0.008)</td>
</tr>
<tr>
<td>Intangible Capital</td>
<td>0.066 ***</td>
<td>-0.062 ***</td>
<td>0.173 ***</td>
<td>0.044 ***</td>
</tr>
<tr>
<td></td>
<td>(0.007)</td>
<td>(0.008)</td>
<td>(0.009)</td>
<td>(0.011)</td>
</tr>
<tr>
<td>Size</td>
<td>-0.165 ***</td>
<td></td>
<td>-0.128 ***</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.014)</td>
<td></td>
<td>(0.014)</td>
<td></td>
</tr>
<tr>
<td>iB/M Ratio</td>
<td>0.002</td>
<td></td>
<td>0.057 ***</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.006)</td>
<td></td>
<td>(0.007)</td>
<td></td>
</tr>
<tr>
<td>Sales</td>
<td>0.446 ***</td>
<td></td>
<td>0.369 ***</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.020)</td>
<td></td>
<td>(0.021)</td>
<td></td>
</tr>
<tr>
<td>Fixed effects</td>
<td>No</td>
<td>No</td>
<td>Industry + Date</td>
<td>Industry + Date</td>
</tr>
<tr>
<td>Observations</td>
<td>21,081</td>
<td>20,636</td>
<td>21,081</td>
<td>20,636</td>
</tr>
</tbody>
</table>

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between firm-level physical capital investment (divided by physical capital) and process intensity (in standard deviation units). The results come from estimating panel regressions (6.2). Data definitions are found in the Data section of the paper, and Compensation refers to total executive compensation from Execucomp. All variables aside from process intensity are logged. Industry fixed effects use one broad one digit SIC codes (after 2002) or one digit NAICS (before 2002), and date fixed effects correspond to years. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
Table 18: Executive Compensation and Process Intensity (EP 2013 Method)

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>Total Compensation / Physical Capital</th>
<th>Deferred Compensation / Physical Capital</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
</tr>
<tr>
<td>Process Intensity</td>
<td>0.084^{***}</td>
<td>0.079^{***}</td>
</tr>
<tr>
<td></td>
<td>(0.024)</td>
<td>(0.015)</td>
</tr>
<tr>
<td>Intangible Capital</td>
<td>0.407^{***}</td>
<td>0.204^{***}</td>
</tr>
<tr>
<td></td>
<td>(0.026)</td>
<td>(0.018)</td>
</tr>
<tr>
<td>Size</td>
<td>−0.755^{***}</td>
<td>−0.693^{***}</td>
</tr>
<tr>
<td></td>
<td>(0.031)</td>
<td>(0.049)</td>
</tr>
<tr>
<td>iB/M Ratio</td>
<td>−0.548^{***}</td>
<td>−0.435^{***}</td>
</tr>
<tr>
<td></td>
<td>(0.016)</td>
<td>(0.021)</td>
</tr>
<tr>
<td>Sales</td>
<td>0.670^{***}</td>
<td>0.640^{***}</td>
</tr>
<tr>
<td></td>
<td>(0.041)</td>
<td>(0.066)</td>
</tr>
</tbody>
</table>

Fixed effects Industry + Date Industry + Date Industry + Date Industry + Date
Observations 11,416 11,234 5,372 5,272

Note: \* \* \* p<0.1; ** p<0.05; *** p<0.01

This table shows the relationship between firm-level executive compensation (divided by physical capital) and process intensity (in standard deviation units). Executive compensation is either total compensation or deferred compensation. The results come from estimating panel regressions (6.1). Data definitions are found in the Data section of the paper, and Compensation refers to total executive compensation from Execucomp. All variables aside from process intensity are logged. Industry fixed effects use one broad one digit SIC codes (after 2002) or one digit NAICS (before 2002), and date fixed effects correspond to years. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
Table 19: Fraction of Compensation Deferred and Process Intensity (EP 2013 Method)

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>Deferred Compensation / Total Compensation</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
</tr>
<tr>
<td>Process Intensity</td>
<td>0.024** (0.012)</td>
</tr>
<tr>
<td>Intangible Capital</td>
<td>0.047*** (0.010)</td>
</tr>
<tr>
<td>Size</td>
<td></td>
</tr>
<tr>
<td>iB/M Ratio</td>
<td></td>
</tr>
<tr>
<td>Sales</td>
<td></td>
</tr>
</tbody>
</table>

Fixed effects       | No Industry + Date |
Observations         | 5,368  | 5,270  |

*Note:* *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between the firm-level fraction of executive compensation deferred and process intensity (in standard deviation units). The dependent variable is the ratio of deferred compensation in the form of stock and option grants to total executive compensation. The results come from estimating panel regressions (6.1). Data definitions are found in the Data section of the paper. All variables aside from process intensity are logged. Industry fixed effects use one broad one digit SIC codes (after 2002) or one digit NAICS (before 2002), and date fixed effects correspond to years. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
Table 20: Innovation Intensive Salaries and Process Intensity (EP 2013 Method)

<table>
<thead>
<tr>
<th>Dependent variable: Innovation Intensive Wage Relative to Industry Average</th>
<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
<th>(4)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Process Intensity</td>
<td>0.039**</td>
<td>0.033*</td>
<td>0.036**</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.017)</td>
<td>(0.018)</td>
<td>(0.017)</td>
<td></td>
</tr>
<tr>
<td>Intangible Capital</td>
<td></td>
<td></td>
<td>0.032*</td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td>(0.018)</td>
<td></td>
</tr>
<tr>
<td>Size</td>
<td>0.0002</td>
<td>0.002</td>
<td>−0.005</td>
<td>−0.005</td>
</tr>
<tr>
<td></td>
<td>(0.011)</td>
<td>(0.012)</td>
<td>(0.014)</td>
<td>(0.015)</td>
</tr>
<tr>
<td>iB/M Ratio</td>
<td>0.019</td>
<td>0.011</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.023)</td>
<td>(0.025)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Sales</td>
<td>0.048***</td>
<td>0.050***</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.011)</td>
<td>(0.011)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>log(SaK)</td>
<td>0.002</td>
<td>−0.006</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.028)</td>
<td>(0.035)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Fixed effects</td>
<td>No</td>
<td>Industry + Date</td>
<td>No</td>
<td>Industry + Date</td>
</tr>
<tr>
<td>Observations</td>
<td>1,562</td>
<td>1,562</td>
<td>1,524</td>
<td>1,524</td>
</tr>
</tbody>
</table>

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between the firm-level average innovation intensive job salary (relative to industry average) and process intensity (in standard deviation units). The dependent variable is defined as the average posted wage for innovation intensive jobs in a given firm-year from Burning Glass Technologies. The industry average used to normalize is a leave-one-out mean of the firm-year innovation intensive salaries. The results come from estimating panel regressions (6.1). Data definitions are found in the Data section of the paper. All variables aside from process intensity are logged. Industry fixed effects use one broad one digit SIC codes (after 2002) or one digit NAICS (before 2002), and date fixed effects correspond to years. Clustering is performed at the industry-date level, where industry is either three digit SIC or four digit NAICS (delineated by 2002, again), and date is year.
Table 21: Process Intensity, Compensation, and Investment Level (EP 2013 Method)

<table>
<thead>
<tr>
<th></th>
<th>Total Compensation</th>
<th>Deferred Compensation</th>
<th>Skilled Wage Relative to Industry Average</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bin 1 × Process Intensity</td>
<td>0.090*** (0.019)</td>
<td>−0.006 (0.018)</td>
<td>0.016 (0.024)</td>
</tr>
<tr>
<td>Bin 2 × Process Intensity</td>
<td>0.144*** (0.020)</td>
<td>0.068*** (0.017)</td>
<td>0.066*** (0.025)</td>
</tr>
</tbody>
</table>

Fixed effects | Bin+Date | Bin+Date | Bin+Date |
Controls? | Yes | Yes | Yes |
Observations | 11,175 | 4,146 | 1,523 |

Note: *p<0.1; **p<0.05; ***p<0.01

This table shows the relationship between different firm-level compensation measures and process intensity (in standard deviation units) interacted with an indicator for the level of physical investment. Firms in Bin 1 have low physical capital investment, and vice versa for firms in Bin 2. In column one, the dependent variable is total executive compensation divided by physical capital. In column two it is deferred compensation divided by physical capital. In column three it is the average posted wage for innovation intensive jobs in a given firm-year from Burning Glass Technologies. The industry average used to normalize the Burning Glass salaries is a leave-one-out mean of the firm-year innovation intensive salaries. The results come from estimating panel regressions (6.3). Data definitions are found in the Data section of the paper. All variables aside from process intensity are logged. Fixed effects are at the bin and date level. Clustering is performed at the firm level.
**B Proofs for Lemma 4.1 and Proposition 4.1**

**Proof of Lemma 4.1**

Consider a probability measure \( \mathbb{P}^0 \) under which

\[
dK_t = \sigma K_t dZ_t^0
\]

with a \( \mathbb{P}^0 \)-Brownian motion \( Z^0 \). Introduce an equivalent probability measure \( \mathbb{P}^e \) such that \( Z^e \), defined via

\[
dZ_t^e = dZ_t - \frac{D(e_t, I_t, O_t) - \delta_K K_t}{\sigma K_t}\]

is a Brownian motion under \( \mathbb{P}^e \). Then \( K \) follows the dynamics \((3.4)\).

Under \( \mathbb{P}^0 \), the agent’s continuation value \( U \) in \((3.8)\) has the semimartingale decomposition

\[
dU_t = dH_t + \varphi_t dK_t,
\]

where \( \varphi \) arises from the martingale representation theorem. We will use dynamic programming to determine the finite variation process \( H \). To this end, it follows from \((3.8)\) and the dynamic programming that \( \tilde{U}_t = e^{-\gamma t} U_t + \int_0^t e^{-\gamma s} (\lambda \Lambda_s (1 - e_s) ds + dC_s) \) is a super-martingale under \( \mathbb{P}^e \) for arbitrary effort \( e \) and a martingale for the optimal effort \( e^* \). We obtain from Itô’s formula that

\[
d\tilde{U}_t = e^{-\gamma t} \left\{ - \gamma U_t dt + dH_t + \lambda \Lambda_t (1 - e_t) dt + dC_t + \varphi_t K_t d\frac{d(e_t, i_t, o_t)}{i} - \delta_K dK_t \right\},
\]

where \( d(e, i, o) = A[a_i^\rho + e(1 - a)(1 - \theta)^{\rho_0}]^{1/\rho} \). The drift of \( \tilde{U} \) is nonpositive for an arbitrary effort \( e \) and is zero for the optimal effort \( e^* \). Therefore,

\[
dH_t = (\gamma U_t + \varphi_t \delta_K K_t) dt - dC_t - \max_{e \in \{0,1\}} \left\{ \lambda \Lambda_t (1 - e) + \varphi_t K d(e, i_t, o_t) \right\} dt.
\]

The optimal effort \( e^*_t = 1 \) if and only if

\[
\varphi_t K d(1, i, o) \geq \lambda \Lambda_t + \varphi_t K d(0, i, o).
\]

Recall the definition of \( \Lambda \) from \((3.6)\), the previous incentive compatibility condition is
equivalent to

$$\phi_t \geq \lambda.$$  

When the previous condition holds, $e^* = 1$ and we obtain from (B.1) and (B.2) that $U$ follows (4.1).

**Proof of Proposition 4.1**

We drive the HJB equation (4.7) from the dynamic programming principle. To this end, it follows from the dynamic programming principle that $\bar{V}_t = e^{-rt}Kv(o_t,u_t) + \int_0^t e^{-rs}(Y_sds - dC_s)$ is a supermartingale under arbitrary strategy $(i,s,C)$ and a martingale under the optimal strategy. Using Itô's formula, together with (4.11) and (4.12), we calculate

$$d(Kv(o,u)) = \left\{Kv d(i,o) - \delta_K Kv + \partial_o v\left[s - (\delta_O - \delta_K)o - o d(i,o) + \sigma^2 o\right] + \partial_u v\left[(\gamma + \delta_K)u - u d(i,o) + \sigma^2(u - \phi)\right] + \frac{1}{2}Ko^2 \sigma^2 \partial_{oo}^2 v + \frac{1}{2}Ko^2(\phi - u)^2 \partial_{uu}^2 v - Ko \sigma^2(\phi - u) \partial_{vu}^2 v + Ko^2\left[-o \partial_o v + (\phi - u) \partial_u v\right]\right\}dt$$

$$+ Ko\left[v - o \partial_o v + (\phi - u) \partial_u v\right]dZe^* - \partial_u v dC_t.$$  

The drift of $\bar{V}$, divided throughout by $e^{-rt}K$, is

$$- rv + v d(i,o) - \delta_K v + \partial_o v\left[s - (\delta_O - \delta_K)o - o d(i,o) + \sigma^2 o\right] + \partial_u v\left[(\gamma + \delta_K)u - u d(i,o) + \sigma^2(u - \phi)\right] + \frac{1}{2}o^2 \sigma^2 \partial_{oo}^2 v + \frac{1}{2}(\phi - u)^2 \sigma^2 \partial_{uu}^2 v - o(\phi - u) \sigma^2 \partial_{vu}^2 v + \sigma^2\left[-o \partial_o v + (\phi - u) \partial_u v\right] + \mu[1 - \phi + \phi(\theta o)^\psi]^{1/\psi} - i - s - \frac{Q_K}{c_K} \bar{v}_K - \frac{Q_O}{c_O}(i/o)^{io} o + (\partial_u v + 1)\left(-\frac{1}{K} \frac{dC_t}{dt}\right).$$
Therefore the dynamic programming principle implies that the HJB equation satisfied by $v$ is

$$(r + \delta K)v = \max_{i,s,\phi,C} \left\{ (v - o \partial_o v - u \partial_u v)d(i,o) + (s - (\delta_O - \delta_K)o)\partial_o v + (\gamma + \delta_K)u \partial_u v \ight.$$  

$$+ \frac{1}{2}o^2\sigma^2 \partial^2_{oo}v + \frac{1}{2}(\varphi - u)^2\sigma^2 \partial^2_{uu}v - o(\varphi - u)\sigma^2 \partial^2_{ou}v$$  

$$+ \mu \left[1 - \phi + \phi(\theta_o)^\psi\right]^{\phi/\psi} - i - s - \frac{Q_K}{c_K}i - \frac{Q_O}{c_O}(i/o) + (\partial_u v + 1)\left(- \frac{1}{K} \frac{dC_t}{dt}\right) \right\}.$$  

(B.3)

Because $dC_t/dt$ is can be infinite, if $\partial_u v + 1 < 0$, the right-hand side of the previous equation can be infinite by choosing infinite $dC_t/dt$. Therefore, the wellposedness of the HJB equation requires that $\partial_u v + 1 \geq 0$. As a result, the equation (B.3) is transformed to (4.7). In order to incentivize the full effort $e^* = 1$, the incentive compatibility condition restricts $\varphi \geq \lambda$.

C Numeric algorithm

In this section, we describe the numeric procedure to solve the HJB equation (4.7) and the Fokker-Planck-Kolmogorov equation (4.15).

For the HJB equation (4.7), we employ the penalty approach to transform (4.7) into

$$0 = - (r + \delta K)v + \max_{i \geq 0, s \geq 0, \varphi \geq \lambda} \left\{ (v - o \partial_o v - u \partial_u v)d(i,o) \ight.$$  

$$+ (s - (\delta_O - \delta_K)o)\partial_o v + (\gamma + \delta_K)u \partial_u v$$  

$$+ \frac{1}{2}o^2\sigma^2 \partial^2_{oo}v + \frac{1}{2}(\varphi - u)^2\sigma^2 \partial^2_{uu}v - o(\varphi - u)\sigma^2 \partial^2_{ou}v$$  

$$+ \mu \left[1 - \phi + \phi(\theta_o)^\psi\right]^{\phi/\psi} - i - s - C_K(i) - C_O(s/o)$$  

$$- \min_{P \in [0, P_{max}]} P[\partial_u v + 1],$$  

(C.1)

where $P_{max}$ is a large positive constant ($10^6$ in the implementation). In the previous equation, when $\partial_u v + 1 \geq 0$, the optimal $P$ is zero and the penalty term $P[\partial_u v + 1] = 0$, hence the first four lines in (C.1) sum to zero; when $\partial_u v + 1 < 0$, the optimal $P = P_{max}$, leading to a negative penalty term $P_{max}[\partial_u v + 1]$, hence the first four lines of (C.1) sum to be negative, consistent with the requirement in (4.7) that the first group of term on
the right-hand side is always nonpositive. In (C.1), we also set \( \varphi = \lambda \) in (4.16) to obtain

\[
0 = -(r + \delta_K)v + \max_{i \geq 0, s \geq 0} \left\{ (v - o \partial_o v - u \partial_u v) d(i, o) + (s - (\delta_O - \delta_K)o) \partial_o v + (\gamma + \delta_K)u \partial_u v + \frac{1}{2}\sigma^2 \partial^2_{oo} v + \frac{1}{2}(\lambda - u)^2 \sigma^2 \partial^2_{uu} v - o(\lambda - u)\sigma^2 \partial^2_{ou} v + \mu \left[ 1 - \phi + \phi(\theta o)^{\psi \psi - i - s - C_K(i) - C_O(s/o)o} \right] \right\} - \min_{P \in [0, P_{\max}]} P[\partial_u v + 1].
\]

(C.2)

After the numeric solution for the previous equation is obtained, we verify whether the condition (4.8) holds. This condition is satisfied in all our numeric experiments.

We choose a domain \([0, o_{\max}] \times [0, u_{\max}]\) with sufficiently large \(o_{\max}\) and \(u_{\max}\). Equation (C.2) is coupled with the following boundary conditions. When \(u = 0\), \(v\) satisfies (4.14), which is solved before (C.2). When \(o = 0\), (4.11) shows that the drift of do is non-negative and the volatility vanishes. Therefore, the boundary condition at \(o = 0\) is not needed in an upwind numeric scheme. When \(u = u_{\max}\), we impose the Neumann boundary condition

\[
\partial_u v(u_{\max}, o) = -1.
\]

When \(o = o_{\max}\), we impose a technical Neumann boundary condition

\[
\partial_o v(u, o_{\max}) = 0.
\]

With aforementioned boundary conditions, our numeric experiments show that the function \(v\) in a fixed bounded domain is not sensitive to the choice of \(u_{\max}\) and \(o_{\max}\) when they are sufficiently large.

We apply policy iteration methods to solve (C.2) (see (Kushner and Dupuis, 2001, Chapter 5 and 6)).

(i) Start with initial guess \(i, s, P = 0\).

(ii) For given \(i, s,\) and \(P\), solve (C.2) with the fixed \(i, s,\) and \(P\) using the upwind finite difference scheme in the domain \([0, o_{\max}] \times [0, u_{\max}]\).

(iii) Using the obtained \(v\) and their derivatives \(\partial_o v\) and \(\partial_u v\) (approximated by finite difference) to solve the maximization problem in (C.2). Update \(i, s,\) and \(P\) using
the corresponding maximizer.

(iv) Go back to Step (ii), until the difference between \( v \) and its value in the previous iteration is less than some small error tolerance \( \epsilon \).

To solve the Fokker-Planck-Kolmogorov equation (4.15), we use the finite difference scheme in (Achdou et al., 2022, Section 5.2). To pin down the entry rate \( m \), we use bisection search: when the integral of the density \( g \) on \([0, u_{\text{max}}] \times [0, o_{\text{max}}]\) less than 1, \( m \) is increased; when the integral of \( g \) is larger than 1, \( m \) is decreased. Iterate a bisection search for \( m \) until the integral of \( g \) is sufficiently close to 1.

## D First best benchmark

To compare with the main model, we study in this section the first best benchmark, where the investment does not subject to agency friction. The firm’s problem is

\[
V(K, O) = \max_{I,S} \mathbb{E} \left[ \int_0^\infty e^{-rs}Y ds \middle| K_0 = K, O_0 = O \right],
\]

subject to (3.1), (3.3), and (3.4) with \( e = 1 \).

The homothetic property in \( K \) allows us to introduce the following decomposition of the value function:

\[
V(K, O) = Kv(o),
\]

where \( o = O/K \).

**Proposition D.1** The function \( v \) in (D.2) satisfies the following HJB equation

\[
(r + \delta_K)v = \max_{i,s \geq 0} \left\{ (v - o\partial_o v)d(i, o) + (s - (\delta_O - \delta_K)o)\partial_o v + \frac{1}{2}o^2\sigma^2\partial^2_{oo}v + \mu [1 - \phi + \phi(\theta o)^\psi]^{1/\psi} - i - s - \frac{Q_K}{c_K}i - \frac{Q_O}{c_O} \left( \frac{s}{o} \right) c_o \right\},
\]

where \( d(i, o) = A[i + (1-a)(1-\theta)^t o]^{1/\rho} \). When \( (v - o\partial_o v)d(0, o) > 1 \), the optimal investment in the physical capital satisfies the first order condition

\[
(v - o\partial_o v)\partial_i d(i^*, o) = 1 + Q_K(i^*)^{c_K-1};
\]

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First best: value function \( v \), optimal investment in the physical capital. The parameters are listed in Table 4.

otherwise, \( i^* = 0 \). If \( \partial_o v > 1 \), then the optimal investment in the intangible capital is

\[
s^* = o \left( \frac{\partial_o v - 1}{Q_o} \right)^{\frac{1}{1-o}} ;
\]

otherwise \( s^* = 0 \).

Figure 8 provides the first best solution with the parameters in Table 4. As the intangible-physical capital ratio increases, investment in the physical capital becomes more efficient, hence both the principal’s value and the physical capital investment over physical capital stock ratio increase. Comparing with the physical investment rate in the middle panel of Figure 5, the physical investment rate is higher in the first best case without agency friction. Because the principal can dictate agent’s effort in the first best case, the agent only receives his outside value as the compensation.

**Proof of Proposition D.1**

Recall the value function \( V \) in (4.4). It follows from the dynamic programming principle that \( \tilde{V}_t = e^{-rt}V(K_t, O_t) + \int_0^t e^{-rs}(dY_s - dC_s) \) is a supermartingale for an arbitrage strategy \( (i, s) \) and is a martingale under the optimal strategy. Using (D.2) and (4.11), we obtain
from Itô’s formula that

$$d(Kv(o)) = \left\{ Kv(d(i, o) - \delta_K Kv + K\partial_o v[s - (\delta_O - \delta_K)o - o d(i, o)] + \frac{1}{2} Kv^2 \sigma^2 \partial^2_{oo} v \right\} dt + K\sigma(v - o \partial_o v) dZ_t.$$ 

The drift of $\tilde{V}$ (divided throughout by $e^{-rt}K$) is

$$-rv + vd(i, o) - \delta_K v + \partial_o v[s - (\delta_O - \delta_K)o - o d(i, o)] + \frac{1}{2} \sigma^2 \partial^2_{oo} v$$

$$+ \mu [1 - \phi + \phi(\theta o)^{\psi}]^{1/\psi} - i - s - \frac{Q_K}{C_K} i - \frac{Q_O}{C_O} \left( \frac{s}{o} \right) C_o + \epsilon_t.$$ 

Therefore the HJB equation (D.3) follows from the fact that the drift of $\tilde{V}$ is nonpositive for any $i, s$ and is zero for optimal $i^*$ and $s^*$. The first order conditions in $i^*$ and $s^*$ follow from the same argument as in Proposition 4.1.

## E Complementarity

This appendix summarizes our results on the complementarity of different types of intangibles with respect to physical investment. This is important because a claim in the model is that only process intangibles are complementary with physical investment. This assumption leads to the hold-up channel we emphasize. We check this claim in the data by estimating panel regressions with physical investment interacted with different types of intangibles.

We estimate two different regressions. First, we estimate:

$$\frac{K_{f,t+2}}{K_{f,t-1}} = \alpha + \beta' X_{ft} + \beta_1 \text{Phys. Inv.}_{ft}$$

$$+ \beta_2 \text{Proc. Int.}_{ft} + \beta_3 \text{Prod. Int.}_{ft}$$

$$+ \beta_4 \text{Proc. Int.}_{ft} \times \text{Phy. Inv.}_{ft} + \beta_5 \text{Prod. Int.}_{ft} \times \text{Phy. Inv.}_{ft} + \epsilon_{ft}.$$ 

(E.1)

The dependent variable is three-year physical capital growth, $X_{ft}$ is a vector of controls (size, iB/M ratio, and sales to capital ratio), Phys. Int. is physical capital investment (divided by physical capital), Proc. Int. is the ratio of process intangibles to physical
capital, and Prod. Int. is the ratio of product intangibles to physical capital. Second, we estimate:

\[
\frac{\text{Three Year Capital}}{K_{f,t-1}} = \alpha + \beta' X_{f,t} + \beta_1 \text{Phys. Inv.}_{f,t} + \beta_2 \text{Int. Cap.}_{f,t} + \beta_3 \text{ProcIn}_{f,t} + \beta_4 \text{Int. Cap.}_{f,t} \times \text{Phys. Inv.}_{f,t} + \beta_5 \text{ProcIn}_{f,t} \times \text{Phys. Inv.}_{f,t} + \epsilon_{f,t}.
\]

(E.2)

These two regressions capture the same idea: If \( \beta_5 > \beta_4 \) in equation (E.1), then process intangibles are more complementary with physical investment than product intangibles, and if \( \bar{\beta}_5 > 0 \) in equation (E.2) then the same thing is true. This is the source of the hold-up friction and motivation for our modeling assumptions. The first regression looks at effects of changing the level of one type of intangible while fixing the level of the other. The second regression looks at changing the composition of intangibles while fixing the total level of intangibles.

As in the body of the paper, we estimate the previous two regressions within industry, which means all the coefficients are industry dependent. We summarize the results of the estimates by looking at the mean value of the coefficients of interest across industries.

We also estimate the above two regressions using pooled panel regressions with industry and date fixed effects and clustering at the industry-date level. We estimate each panel regression on two different subsamples: The full sample and the sample restricted to firms with non-zero process intensity.

Table 16 displays the results. The first row shows the average coefficient value across industries when we estimate complementarity within industry (as in the body of the paper). The first column shows the average value of \( \bar{\beta}_5 \) in equation (E.2). The second column shows the average value of \( \beta_4 \) in equation (E.1). The final column shows the average value \( \beta_5 \) in equation (E.1). The first row is consistent with our hypothesis: \( \bar{\beta}_5 > 0 \) and \( \beta_4 > \beta_5 \). The second row shows the results from the pooled panel regression when using the full sample. Once again, the results are consistent with the hypothesis.\(^{57}\) The final row repeats the panel regression but restricts the sample to firms with non-zero process intangibles. The anticipated effects are stronger than in the full sample results. This is encouraging, since the model is meant to capture firms that engage in both types

\(^{57}\) \( \bar{\beta}_5 \) is statistically insignificant, however.
of intangible use.

Table 16 Here

F Results using Eisfeldt and Papanikolaou (2013) Method for Constructing Intangible Capital

In the main body of the paper, we used the methods from Peters and Taylor (2017) and Ewens, Peters and Wang (2019) to construct intangible capital from Compustat data. We have also used the method of Eisfeldt and Papanikolaou (2013), which we present here. The latter method differs in two ways. First, they do not distinguish between SG&A and R&D. Instead, they simply capitalize SG&A fully. The authors claim that this is more robust to variations in reporting standards across industries. Second, Eisfeldt and Papanikolaou (2013) recommend conditioning on industry when forming portfolio ranks. Because we do not use a portfolio approach, our use of industry fixed effects is important here.

Tables 17-21 display the results of repeating our main empirical exercises using the Eisfeldt and Papanikolaou (2013) to construct intangibles. The results are almost entirely unchanged.