

Measuring the Cost of Living in Mexico and the US

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Price Measurement Across Countries

- Long-standing problems of measurement
 - Sampling: collected from stores instead of consumers
 - Quality: brain surgery in Nairobi vs Tokyo
 - Variety: product availability

This Paper

① Nielsen US and Mexico

- Representative panel of households with purchases of consumer goods.
- Products matched at the barcode level across countries.

② Quantify potential biases behind the ICP using non-homothetic price index

- A new decomposition framework to quantify sampling bias, quality bias, and variety bias independently

Preview of Findings

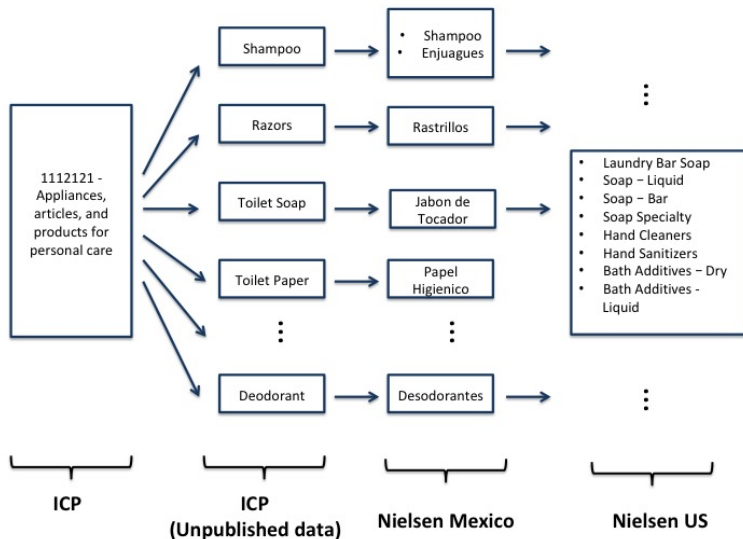
- Price level ratio between Mexico and the US:

$$\begin{aligned} \text{NH}^M &= \Theta^M \times \text{ICP}^M \\ 0.72 &= 0.90 \times 0.80 \end{aligned}$$

- $\Theta^M = S^M \times Q^M \times V^M$
 - Sampling bias: $S^M = 0.82$
 - Quality bias: $Q^M = 1.45$
 - Engel-curve Variety bias: $V^M = 0.75$

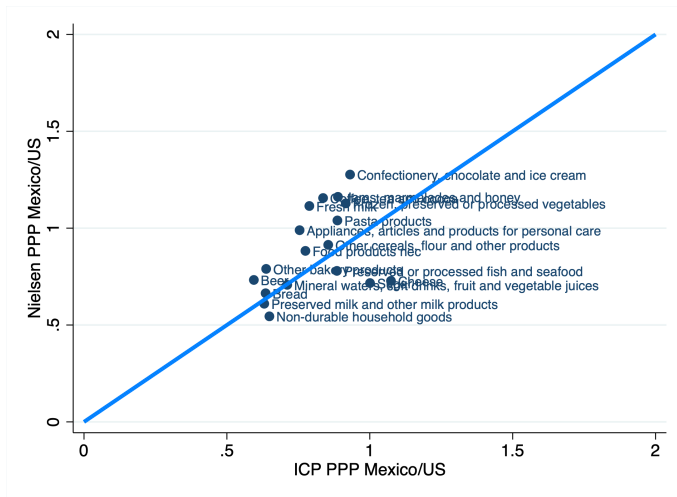
- ICP 2011
 - 155 basic headings
 - Thousands of comparable items
- Nielsen Mexico
 - Representative sample of 5,000 households for 2012-2013.
 - Households visited biweekly report consumption diary information.
- Nielsen US
 - Representative sample of 60,000 households.
 - Panelists use in-home scanners to record their purchases.

Matched sample: Nielsen data, ICP data



- 1 Among multiple barcodes for specific item in a store, pick one barcode to represent the item (\bar{p}_{sib}).
- 2 Aggregate across stores with store size weights (\bar{p}_{ib}).
- 3 Jevons index across items (ICP_b).

Nielsen vs ICP PPP by Basic Heading



- Nielsen data **mimics well** the prices constructed by the ICP.

Facts on Sampling, Quality and Variety

- 1 Mexican households shop more frequently and visit more stores. Therefore, Mexicans buy a larger share of items at stores where they are cheaper. [▶ Details](#)
- 2 The distribution of prices in the US has a higher mean and a longer right tail, but these patterns are attenuated when we compare common goods. [▶ Details](#)
- 3 A significant presence of US brands in the Mexican market gives more variety to Mexican consumers. [▶ Details](#)

- basic headings b , items i , barcodes k , stores s
- CES aggregation across **basic headings**
- CES aggregation across **items**
- Non-homothetic CES aggregation across **barcodes**

$$1 = \sum_{k \in \Omega_{ib}^M} \left(\frac{\varphi_{kib}^M C_{kib}^M}{(C_{ib}^M)^{(\varepsilon_{kib} - \sigma_{ib}) / (1 - \sigma_{ib})}} \right)^{\frac{\sigma_{ib} - 1}{\sigma_{ib}}}$$

where ε_{kib} is the elasticity of a barcode k with respect to item-level consumption C_{ib}^M

- Cobb-Douglas Aggregation across **stores**.

Non-homothetic Price Index

$$\text{NIH}_{ib}^M \equiv \frac{P_{ib}^M}{P_{ib}^U} = \prod_{k \in \Omega_{ib}} \left(\frac{p_{kib}^M}{p_{kib}^U} \right)^{\omega_{kib} \frac{1}{1-\theta_{ib}}} \times \left(\frac{\lambda_{ib}^M}{\lambda_{ib}^U} \right)^{\frac{1}{\sigma_{ib}-1} \frac{1}{1-\theta_{ib}}} \times \left(\frac{E_{ib}^M}{E_{ib}^U} \right)^{\frac{\theta_{ib}}{\theta_{ib}-1}}$$

where

$$\theta_{ib} \equiv \sum_{k \in \Omega_{ib}} \omega_{kib} \frac{\varepsilon_{kib} - 1}{\sigma_{ib} - 1}$$

- Sato-Vartia index across common barcodes
- Variety correction
- Engel-curve adjustment

Decomposition of Non-homothetic Price Index

$$\text{NH}_b^M = \Theta_b^M \times \text{ICP}_b^M$$
$$\Theta_b^M \equiv S_b^M \times Q_b^M \times V_b^M$$

- S_b^M : Sampling Bias
- Q_b^M : Quality Bias
- V_b^M : Engel-curve Variety Bias

S_b^M : Sampling Bias

$$S_b^M \equiv \left(\left(\prod_{i \in \Omega_b} \frac{\bar{p}_{ib}^M}{\bar{p}_{ib}^U} \right)^{\frac{-1}{N_b}} \times \prod_{i \in \Omega_b} \left(\frac{\bar{p}_{ib}^M}{\bar{p}_{ib}^U} \right)^{\omega_{ib}} \right) \times \left(\prod_{i \in \Omega_b} \left(\frac{\hat{p}_{ib}^M / \bar{p}_{ib}^M}{\hat{p}_{ib}^U / \bar{p}_{ib}^U} \right)^{\omega_{ib}} \right)$$

where

$$\bar{p}_{ib}^M = \prod_{s \in \Psi^M} \left(\bar{p}_{sib}^M \right)^{\phi_s^M} \quad \text{and} \quad \hat{p}_{ib}^M \equiv \prod_{s \in \Psi^M} \left(\bar{p}_{sib}^M \right)^{\phi_{sib}^M}$$

- Bias comes from missing expenditures for each item.
- Bias depends on covariance between expenditures and prices across items.
 - No significant difference between two countries [▶ Details](#)

S_b^M : Sampling Bias

$$S_b^M \equiv \left(\left(\prod_{i \in \Omega_b} \frac{\bar{p}_{ib}^M}{\bar{p}_{ib}^U} \right)^{\frac{-1}{N_b}} \times \prod_{i \in \Omega_b} \left(\frac{\bar{p}_{ib}^M}{\bar{p}_{ib}^U} \right)^{\omega_{ib}} \right) \times \left(\prod_{i \in \Omega_b} \left(\frac{\hat{p}_{ib}^M / \bar{p}_{ib}^M}{\hat{p}_{ib}^U / \bar{p}_{ib}^U} \right)^{\omega_{ib}} \right)$$

where

$$\bar{p}_{ib}^M = \prod_{s \in \Psi^M} \left(\bar{p}_{sib}^M \right)^{\phi_s^M} \quad \text{and} \quad \hat{p}_{ib}^M \equiv \prod_{s \in \Psi^M} \left(\bar{p}_{sib}^M \right)^{\phi_{sib}^M}$$

- Bias comes from missing expenditures for each item at each store.
- Bias depends on covariance between expenditures and prices across stores.
 - Significant difference between two countries [▶ Details](#)

Q_b^M : Quality Bias, V_b^M : Engel-curve Variety Bias

$$Q_b^M \equiv \prod_{i \in \Omega_b} \left(\left(\frac{\hat{p}_{ib}^M}{\hat{p}_{ib}^U} \right)^{-1} \times \prod_{k \in \Omega_{ih}} \left(\frac{p_{kib}^M}{p_{kib}^U} \right)^{\omega_{kib}} \right)^{\omega_{ib}}$$

$$V_b^M = \prod_{i \in \Omega_b} \left(\frac{\lambda_{ib}^M}{\lambda_{ib}^U} \right)^{\frac{\omega_{ib}}{\sigma_{ib}-1}} \times \prod_{i \in \Omega_b} \left[\left(\frac{E_{ib}^M / E_{ib}^U}{EPI_{ib}^M} \right)^{\frac{\omega_{ib} \theta_{ib}}{\theta_{ib}-1}} \right]$$

Parameter Estimation

- GMM estimation for σ_{ib} as Broda and Weinstein (2006,2010) [▶ Details](#)
 - mean 9.29, std.dev. 3.42
- Given σ_{ib} estimates, we use the Engel curve to estimate ε_{kib} as Comin et al. (2020). [▶ Details](#)
 - mean 0.82, std.dev. 1.79

Conclusion

- Price level ratio between Mexico and the US:

$$NH^M = \Theta^M \times ICP^M$$

$$0.72 = 0.90 \times 0.80$$

- $\Theta^M = S^M \times Q^M \times V^M$
 - Sampling bias: $S^M = 0.82$
 - Mexicans buy a larger share of items at stores where they are cheaper.
 - Quality bias: $Q^M = 1.45$
 - Low quality products in Mexico matched to high quality products in US.
 - Engel-curve Variety bias: $V^M = 0.75$
 - A significant presence of US brands in the Mexican market gives more variety to Mexican consumers.
- Real non-durable consumption in Mexico relative to US is 10 percent higher than previously estimated.

- Follow the procedures followed by ICP 2011
- Use the categories matched between Nielsen and ICP
 - 1 Select a single item i in country j
 - 2 Aggregate across stores using expenditure weights
 - 3 Estimate:

$$\log \bar{p}_{ib}^c = \eta_i^c + \eta_b^c + \varepsilon_{ib}^c$$

where \bar{p}_{ib}^c is the price of item i belonging to heading b in country c

- 4 The estimated PPP for basic heading b and country is: $\bar{p}_b^c = \exp(\eta_b^c)$.

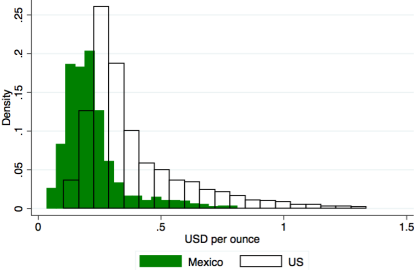
Fact 1: Sampling [◀ Back](#)

Average number of shopping trip per week is 5 in Mexico and 1 in the US.

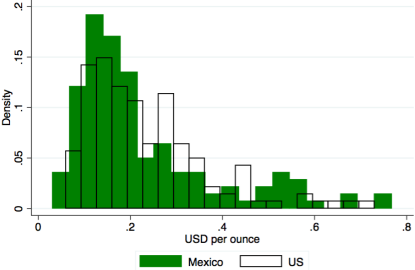
Fact 2: Quality, example of cheese

[◀ Back](#)

The distribution of prices in the US has a higher mean and a longer right tail, but these patterns are attenuated when we compare common goods.



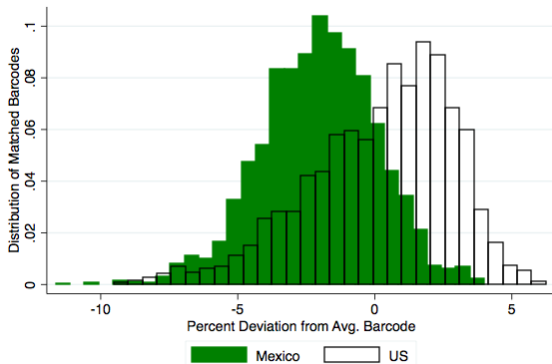
(a) All Products



(b) Overlapping Products

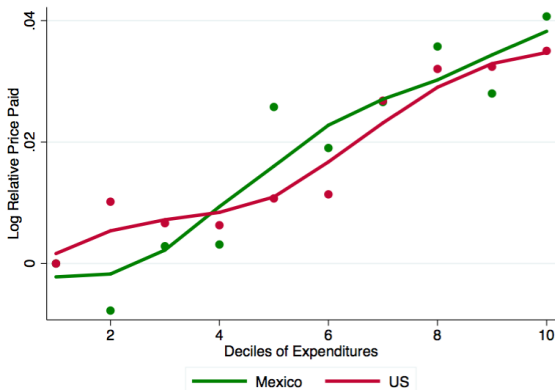
Fact 3: Variety ◀ Back

Mexican households spend less on overlapping products.



Non-homotheticity [◀ Back](#)

Within category of products, richer households buy more expensive products.



Sampling Bias: $\text{cov}(\omega_b, \ln(\bar{p}_b^c))$ [← Back](#)

$$\omega_{ib} = \alpha + \beta \ln(\bar{p}_{ib}^c) \times 1\{c = \text{Mexico}\} + \lambda^c + \theta_b + \varepsilon_{ib}^c$$

	(1)	(2)	(3)	(4)
$\ln(\bar{p})$	-0.010 (0.071)	-0.044 (0.051)	-0.010 (0.071)	-0.039 (0.059)
$\ln(\bar{p}) \times \text{Mexico}$	-0.002 (0.095)	0.004 (0.023)	-0.002 (0.095)	-0.006 (0.056)
Observations	58	58	58	58
R-squared	0.001	0.775	0.001	0.775
Basic Heading	N	Y	N	Y
Country	N	N	Y	Y

Sampling Bias: $\text{cov}(\phi_{ib}^c, \ln(\bar{p}_{ib}^c))$ ◀ Back

$$\phi_{sib}^c = \alpha + \beta \ln(\bar{p}_{sib}^c) \times 1\{c = \text{Mexico}\} + \theta_s^c + \varepsilon_{sib}^c$$

	(1)	(2)	(3)	(4)
$\ln(\bar{p})$	-0.000 (0.000)	-0.000*** (0.000)	-0.000*** (0.000)	-0.000*** (0.000)
$\ln(\bar{p}) \times \text{Mexico}$	-0.002*** (0.000)	-0.001*** (0.000)	-0.001*** (0.000)	-0.001*** (0.000)
Observations	764,419	764,419	761,751	761,750
R-squared	0.028	0.030	0.212	0.212
Store	N	N	Y	Y
Country	N	Y	N	Y

- Feenstra (1994), Broda and Weinstein (2006, 2010)
- Double-difference log UPC expenditure shares and UPC pricing rule over time and relative to the largest UPC within each firm.

$$\Delta^{u,t} \ln S_{kibt} = (1 - \sigma_{ib}) \Delta^{u,t} \ln P_{kibt} + \omega_{kibt}$$

$$\Delta^{u,t} \ln P_{kibt} = \frac{\delta_{ib}}{1 + \delta_{ib}} \Delta^{u,t} \ln S_{kibt} + \kappa_{kibt}$$

$$\omega_{kibt} = [\Delta^t \ln \phi_{kibt} - \Delta^t \ln \phi_{\underline{kibt}}] \text{ and } \kappa_{kibt} = [\Delta^t \ln a_{kibt} - \Delta^t \ln a_{\underline{kibt}}]$$

- Orthogonality of the double-differenced demand and supply shocks defines a set of moment conditions:

$$G(\beta_g) = E_T[\mathbf{v}_{kibt}(\beta_g)] = 0$$

where $\beta_g = [\sigma_{ib}, \delta_{ib}]'$ and $\mathbf{v}_{kibt} = \kappa_{kibt} \omega_{kibt}$.

- We proceed with GMM.

- Given σ_{ib} estimates, we use the Engel curve to estimate ε_{kib} as Comin et al. (2020):

$$\begin{aligned} \ln \frac{s_{kibt}^h}{s_{\mathbf{K}ibt}^h} - (1 - \sigma_{ib}) \ln \frac{p_{kibt}^h}{p_{\mathbf{K}ibt}^h} \\ = (\varepsilon_{kib} - 1) \left(\ln \frac{E_{ibt}^h}{p_{\mathbf{K}ibt}^h} + \frac{1}{(1 - \sigma_{ib})} \ln s_{\mathbf{K}ibt}^h \right) + \psi_t^h + \varepsilon_{kibt}^h \end{aligned}$$

where \mathbf{K} is the benchmark barcode, which corresponds to the largest selling barcode in each item, and ψ_t^h is the set of fixed effects and controls.

Parameter Estimation [◀ Back](#)

	mean	std. dev.	10th-percentile	median	90th-percentile
σ_i	9.29	3.42	5.61	8.73	12.49
ε_{kib}	0.82	1.79	-1.58	0.89	2.99
θ_{ib}	0.02	0.13	-0.13	0.03	0.16
E_{ib}^M / E_{ib}^U	0.72	0.60	0.14	0.53	1.47