Office Visits Preventing Emergency Room Visits: Evidence from the Flint Water Switch

Shooshan Danagoulian, Department of Economics, Wayne State University Daniel Grossman, Department of Economics, West Virginia University David Slusky, Department of Economics, University of Kansas

November 18, 2019

Abstract

Emergency department visits are costly to providers and to patients. We use the Flint water crisis to test if an exogenous increase in office visits reduced avoidable emergency room visits. In September 2015, citizens in Flint became aware of increased lead levels in their drinking water, resulting from the switch from Lake Huron to the Flint River. Using Medicaid claims for 2013-2016, we find that this information shock increased the share of enrollees with lead tests by 1.7 percentage points. Additionally, it increased office visits immediately following the information shock, then decreased them afterwards. This led to a reduction of 4.9 preventable, non-emergent, and primary care treatable emergency room visits per 1000 eligible children (8.2%). This decrease is present in shifts from emergency room visits to office visits across several common conditions. Our results suggest following lead tests, children were more likely to receive care from the same clinic and that establishing care reduces the likelihood a parent will take their child to receive care at the emergency room for conditions treatable in an office setting. Our results are potentially applicable to any situation in which individuals are induced to seek more care in an office visit setting.

JEL Codes: Medicaid; Lead; Environmental Regulation; Emergency Care

Keywords: H75, I12, I18, J13, Q53, Q58

Affiliations:

Danagoulian (corresponding author): Department of Economics, Wayne State University, 656 W. Kirby St. 2905 Faculty Administrative Building, Detroit, MI 48202, fr4523@wayne.edu

Grossman: Department of Economics, West Virginia University, 1601 University Ave, Box 6025, 411 College of Business and Economics, Morgantown, WV, 26506-6025, daniel.grossman@mail.wvu.edu

Slusky: Department of Economics, University of Kansas, 1460 Jayhawk Blvd., 415 Snow Hall, Lawrence, KS, 66045, david.slusky@ku.edu

Acknowledgements: We thank Glenn Copeland of the Vital Records and Health Statistics Division and Matthew Schneider of the Medicaid Division of the Michigan Department of Health and Human Services, for providing vital statistics data and facilitating receipt of Medicaid claims data linked to vital records. We gratefully acknowledge Phillip Levy and Office of Vice President of Research of Wayne State University for financial support. We also thank Antony Hsu, Michael Morrisey, Michael Kofoed, Valentina Duque, and Janet Currie, and conference participants at AHEC 2018, APPAM 2018, ASHEcon 2019, and the NBER Summer Institute Children's Meetings 2019, and seminar participants at Wayne State University, Kansas State University, and University of Michigan for their helpful comments and suggestions.

Introduction

Emergency departments (ED) are structured to diagnose and treat emergent conditions. As such, they may be an expensive alternative to primary care, to both the individual patient and to the health care system. For many individuals, they are the only option for healthcare, with those who lack access to primary care substituting to ED care (Grumbach, Keane, and Bindman 1993). Many of these individuals are of low socio-economic status, and may be eligible for Medicaid. Those who are of low socio-economic status and seeking care for their children are almost certainly eligible for Medicaid. While multiple studies have demonstrated that expanded access to Medicaid increases emergency room usage (Taubman et al. 2014; Nikpay et al. 2017), no study has been able to isolate the causal link between increased primary care and emergency room usage for those who are already eligible for Medicaid.

In this paper, we exploit an exogenous shock to primary care (measured by office visits) resulting from the Flint water contamination. On April 25, 2014, under state-appointed emergency management, the city of Flint switched its water source from Lake Huron to the Flint River. This new source needed to be treated with strong disinfectants, which made it substantially more corrosive than the old water, leaching lead out of the existing Flint water delivery system into residential water (Masten et al. 2016). However, during the period in which water was sourced from the Flint River, local officials stressed that the city water was safe for consumption. Despite warnings and boil advisories in August and September 2014, and an EPA violation for exceeding the organic chemical thresholds in the water in December 2014, the high level of lead content in the water was largely unconfirmed until September 2015. We use this last date as the start of the

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¹ Some argue that ED visits increase, while others argue the increase is simply a shift in payer case mix (see e.g. (Antwi et al., 2015; Finkelstein et al., 2016; Sommers et al., 2016; Sommers & Simon, 2017).

"treatment" period for our analysis, because it represents the point at which city officials first issued a lead advisory in the face of a preponderance of evidence that Flint's drinking water was hazardous to its residents' health.²

The goals of this paper are twofold. First, we establish the extent to which the knowledge of the water problems affected health care receipt. Then, we examine whether a change in primary care use causes a reduction in, or a change in the distribution of, ED visits towards fewer visits that are either treatable or preventable through primary care visits.

Whether, and to what extent, environmental disasters increase medical expenditures of the affected population remains an open empirical question. We determine the amount of medical services received by individuals in the affected areas before, during, and after a water change and a revelation of exposure to contaminated water. We find that Medicaid enrollees in Flint received lead tests at rates nearly 50 percent higher than enrollees from control cities following the information shock. The share of enrollees with any office visit increased by 4 percent and 11 percent, respectively, in the first two quarters immediately following the shock, before decreasing afterwards. Emergency department visits for preventable, non-emergent, and primary care treatable (which we aggregate as "avoidable") decreased by 4.9 visits per 1000 eligible children (8.2%). This decrease is present in shifts from ED visits to office visits across several common conditions. Flint residents who received lead tests were 15 percentage points (24%) more likely to visit the same clinic at which they received their lead test within 3 months. This suggests that establishing care at a specific clinic or with a given physician is associated with a decreased likelihood of receiving care in an ED for a condition that is treatable in an office setting.

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² We also estimate a flexible time form specification using two time periods -- January to August 2015, and September 2015 to December 2016 – with similar findings. These results are presented in Appendix H.

Total payments per Medicaid recipient-month increased by \$12 despite this shift away from ED care, constituting an additional \$550,000 per year (\$12 times 3800 enrollees residing in Flint) in costs to the Medicaid system for the Flint population aged 0 to 3. This constitutes a sizable portion of the total projected saving to the city (\$2.5 million) for medical care for a small portion of the population.

We contribute to several literatures in this study including those investigating lead exposure, the Flint water crisis, the unintended consequences of environmental or informational shocks on healthcare, and the substitutability of healthcare sources for emergency care. We discuss each in turn below.

Prior to the 1980's lead was used extensively in household paint and plumbing, particularly in the lining and soldering joints of copper pipes to help avoid leaks. As the toxicity of lead became better understood, however, such materials have been banned from new housing. Communities with older housing, such as those in Flint, are particularly vulnerable to lead contamination due to lack of investment in new plumbing.

Chronic exposure to lead has significant health consequences. High levels of lead in the bloodstream are associated with cardiovascular problems, high blood pressure, and developmental impairment affecting sexual maturity and the nervous system (ATSDR 2007; Zhu et al. 2010). Newer research, however, shows adverse outcomes at low levels of exposure, as well (Canfield et al. 2003; Jusko et al. 2008; Lanphear et al. 2005; Menke et al. 2006; Navas-Acien et al. 2007; Tellez-Rojo et al 2006). Reports from Flint suggest that children's blood lead levels increased within a few months following the water change (Hanna-Attisha et al. 2016; Zahran et al. 2017) while fertility rates dropped substantially (Grossman and Slusky 2019).

We also contribute to the literature investigating unintended consequences of environmental and informational shocks. While these unintended consequences are generally negative, this is not always the case. Deryugina and Molitor (2019) find that Medicare beneficiaries displaced by Hurricane Katrina who moved to lower mortality areas had lower mortality rates following the disaster.

In our study, we find children are induced by the information shock to go see a primary care physician. The likelihood of seeking preventive care and access to primary care physicians are positively correlated with household income (Sommers et al. 2017; Pitts et al. 2010). Others have attempted to study the causal effect of primary care on ED visits by incentivizing patients to visit their primary care physician (Bradley et al. 2012; 2018; Bradley and Neumark 2017) and by temporally increasing Medicaid reimbursements (Polsky et al. 2015; Candon et al. 2018; Decker 2018; Neprash et al. 2018; Alexander and Schnell 2019). The effects of these interventions depend on the insurance status of the participants. Using an RCT design, Bradley et al. (2018) find that those receiving the cash incentive are more likely to see a primary care physician and less likely to have a preventable ED visit. However, they find no change in overall costs due to an increase in outpatient visits. We build on this research by investigating a plausibly exogenous information shock to explore a similar research question in a quasi-experimental setting.

Lastly, this paper is essentially about creating a linkage to the healthcare system in the form of having a source of usual care. Children who have a usual source of care are more likely to receive preventive care, have higher quality of care, and are less likely to use the ED (Ettner 1996; Xu 2002; Starfield and Shi 2004; Paustian et al. 2014). This informational shock induces parents to take children for a lead test and provides them with a potential alternate place of service to the ED. Previous work has focused on the partial Medicaid expansion to study the effects of gaining

insurance coverage on emergency department and primary care usage by low income individuals (e.g., Sharma et al. 2016; Sommers et al. 2016; Gingold et al. 2017; Jacobs, Jenney, and Selden 2017; Klein et al. 2017; McConville et al. 2018; Ladhania et al. 2019; Pickens et al. 2019), or how the availability of retail clinics affects both primary care and emergency department utilization (e.g., Ashwood et al. 2016; Alexander, Currie, and Schnell 2019). Our paper differs from these in multiple ways. First, our population of interest is eligible for Medicaid throughout this time period, and so there were no formal coverage expansions. Second, we know of no major changes in clinics locations or availability in this time period. Third, our affected population experiences an information shock that results in additional primary care usage. It is this variation on which we focus our analyses.

The remainder of the paper proceeds as follows. First, we summarize the events surrounding the Flint water contamination. Next, we discuss the data and the methods used to identify changing utilization of medical services. The following section presents results. We then discuss our findings in the context of the Flint contamination, and conclude.

Background on the Flint Water Switch

In spring 2013, as part of an effort to reduce the budget of a city under emergency management, the state-appointed manager of Flint ordered the city water supply to be switched to the Flint River by April 25, 2014 (Kennedy 2016). It was previously sourced by the Detroit Water and Sewerage Department (DWSD) from Lake Huron. The switch was intended to be a temporary measure until a proposed pipeline could be completed to supply Flint with water from Lake Huron independently. The Flint Water Service Center (FWSC), however, was ill-equipped to supply adequate quality water to the city. It had not supplied the city since 1967, and was not given a

sufficient transition period to build up materials, facilities, and expertise to do so (Masten et al. 2016).

The shortcomings of the new facility became quickly apparent following the switch. Initially, the water was underchlorinated resulting in water boil advisories issued in July and August of 2014 to counteract E. coli and coliform bacteria detected in the water supply. While chlorine levels were adjusted throughout the summer months to address the bacterial presence, levels of corrosion inhibitors were not. In October 2014, the General Motors engine plant in Flint switched to an alternate source of water, because of the effect the corrosiveness of the water was having on its engine parts.

During this time the water supply was highly corrosive, leading to red water and discoloration throughout the water system, and an unusually large number of water main breaks (Masten et al. 2016). The heavily chlorinated water corroded at the lining of city and residential pipes, leaching lead into the water supply.

The first high lead measurements in the city were detected in February 2015. City authorities assured residents that these measurements were outliers and that the water was safe to drink. By August 2015, Marc Edwards at Virginia Polytechnic Institute and State University had analyzed 120 samples from Flint homes, finding that 20% of samples exceeded the EPA threshold of 15 µg/L action level. In September 2015, city authorities acknowledged the widespread lead contamination of the water supply, issued a lead advisory, and switched back to Lake Huron water treated by DWSD on October 16, 2015.³

The timeline of the water contamination presents an interesting challenge to our analysis. While the water supply was switched in April 2014, and the first high lead measurements were

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³ A more detailed history of Flint and the timeline of the water contamination is presented in Appendix A.

disclosed in February 2015, residents did not have confirmation of the contamination until September 2015. While other studies measure the effect of exposure to lead contamination, we focus on the behavioral response to the knowledge of the contamination. Thus, our analysis focuses on medical utilization after September 2015 as the "treatment" period.

Data

Through an agreement with the Michigan Department of Health and Human Services (MDHHS), we link vital records for all children born in Michigan in 2013-2015 with their Medicaid claims files (both fee-for-service and managed care) for any enrollees in the sample. This unique dataset has several advantages. First, the dataset includes geocoded maternal residential address at the time of birth. Second, it contains birth certificate information about parental-demographic characteristics (e.g., race and educational attainment), and important measures of health at birth (e.g., birthweight, APGAR score, gestational age). Third, the Medicaid data is at the claim level, with detailed information regarding the setting of care, every procedure, test, and visit. Fourth, the Medicaid data also includes payment information for all fee-for-service visits, allowing us to extrapolate payments to the managed-care claims, for complete cost information. These data include Medicaid claims for the years 2013-2016.

We classify Medicaid claims data using the New York University Emergency Department (NYU ED) visit severity algorithm.⁶ For this, researchers reviewed ED records from the 1990s and

⁴ We apply cost information for managed care claims by matching procedure codes with payment made for fee-for-service procedures. Therefore, our cost estimates represent the upper range of costs to Medicaid, though are closer in line with costs for those who are privately insured.

⁵ We have received approval to supplement our analysis with Medicaid claims data covering 2017 and 2018 once the data become available.

⁶ https://wagner.nyu.edu/faculty/billings/nyued-background

categorized diagnosis codes (that did not include any alcohol, drug, injury, or mental health elements) into the following mutually exclusive, collectively exhaustive categories (Billings, Parikh, Mijanovich 2000; validated by Ballard et al. 2010):

- Emergent, ED care needed, and non-preventable (e.g., appendicitis)
- Emergent, ED care needed, but would have been preventable given adequate previous non-emergency care (e.g., diabetes, asthma)
- Emergent, care needed within 12 hours, but primary care would suffice (e.g., heartburn, eye pain)
- Non-emergent, care within 12 hours unnecessary (e.g., rubella, sunburn, jaw pain)

Many diagnoses do not always fall into the same category. For example, based on other details on the discharge record, out of 100 cases of:

- Croup: 57% are emergent and non-preventable, 19% are primary care treatable, and 24% are non-emergent.
- Cough: 12% emergent and non-preventable, 24% are primary care treatable, and 65% non-emergent
- Acute tonsillitis: 6% emergent but preventable, 28% primary care treatable, and 66% non-emergent.

Finally, some diagnoses could not be assigned to a category and so are listed as "unclassified".7

Methodology

This research allows us to track the use of medical services by children born in Flint between 2013 and 2015,⁸ from birth through age 3 and compare them to similarly aged children born elsewhere in the state of Michigan. A priori, we expect to identify higher incidence of adverse

⁷ Results which incorporate a "patch" that captures and classifies a share of uncategorized diagnosis codes (Johnston et al. 2017) are presented in Appendix G.

⁸ Given the result of Grossman and Slusky (2019) that the Flint water switch affected fertility rates, one might be concerned about compositional changes driving our results. In Appendix D, we limit our analysis to the sample of children born before April 2014 (and so unaffected by the fertility effects of the water switch) and find comparable results.

health outcomes, increased use of primary care, and increased costs for patients and insurers because of care received following the informational shock in the form of the city of Flint announcing a potential increase in lead in their water source.

Since the data are observational, we adjust for the differences between residents of Flint and those in the rest of the state. We follow the estimation method used by Grossman and Slusky (2019) which compares Flint to a subset of other large cities in Michigan. We focus exclusively on Michigan because we have complete Medicaid data for this state. Because we are interested in the behavioral response to information shocks as well as changes in water quality, we focus on September 2015, when Flint first released a public lead advisory. ^{9,10}

We employ a difference-in-differences empirical strategy presented below:

$$Outcome_{ict} = \alpha + \beta_1 Flint * After_{ct} + \beta_2 X_{ict} + \alpha_c + \delta_t + \varepsilon_{ict}$$
 (1)

in which Outcome is the medical service or procedure for individual i in city c at time t aggregated over the calendar month. Flint*After is a binary variable equal to 1 for Flint after the contamination or information shock and 0 otherwise. We include binary variables for the city in which an individual lived at the time of birth, α_c , which controls for time-invariant characteristics of a city, and year and month of service as well as year and month of birth fixed effects, δ_t , which control for general trends and seasonality in receipt of a given medical service. These fixed effects subsume the main effects for Flint and After. X_{ict} are individual-level characteristics or characteristics of a given city that vary over time. A potential confounder in our study is that the state of Michigan expanded Medicaid coverage through the ACA in 2014. To the extent that this

⁹ Mona Hanna-Attisha, a Flint pediatrician, held a press conference to announce her findings of a substantial increase in children with high blood-lead levels in September 2015, while Marc Edwards of Virginia Tech released his team's findings of high lead levels in Flint households in August 2015. Flint switched off Flint River water on October 16, 2015.

¹⁰ In Appendix C, we show that the results are robust to starting the treatment period in January 2016.

expansion affected all parts of Michigan equally, this will be captured by the time fixed effects. ¹¹ We investigate the percentage of the sample reporting: any lead test; any office visit; any ED visit; and any claims. We also investigate the total number of claims, and the total payments made. Standard errors are clustered at the city level to allow for serial correlation (Abadie, Athey, Imbens, and Wooldridge 2017). Additionally, we use wild bootstrap methods to adjust our inference because we only have one treated area (Cameron, Gelbach, Miller 2008).

We use a modified version of the above equation to investigate the impact of the water switch on different types of ED visits, defined by the NYU algorithm. For each category, at the individual-month level, we construct a per capita outcome variable by summing the fractional shares of each claim in that category. For example, if an individual had two discharges in a given month, one that was 20% preventable with primary care and the other 70% preventable with primary care, we assign a value of 0.9. Anyone without an ED claim in that category (or with no ED claims at all) receives a value of 0.

While coding those with no claims as having zero visits in a linear specification may bias the results (as some of the individuals would ideally have a negative number of emergency room visits), this bias would be toward zero, and so we consider our set up to be a lower bound on the true effect. We establish our intuition for this setup with three thought experiments. First, imagine that all ED visits are 100% preventable with primary care. Then, to estimate the reduction in per capita ED visits results from a shock to primary care, one would assign 0 to those without an ED visit, and the number of visits to anyone with an ED visit.

¹¹ This issue is further mitigated in that the ACA expansion largely affected adults and did not change federal poverty level coverage thresholds for those aged 0 to 3.

Second, now imagine that some ED visits are 100% non-preventable. These visits should not be affected by the primary care shock, and so the individuals with only these visits should still be assigned a value of 0 for the outcome variable. (They could still be used for a falsification test.)

Finally, consider our actual situation in which certain diagnoses are sometimes preventable and sometimes not. We only care about the preventable parts for our primary estimate, and so in aggregate we can add up the preventable shares of each one to get the outcome variable.

We estimate the elasticity of substitution by comparing relative magnitudes of the effect of the Flint water contamination shock on ED visits and primary care visits relative to their respective means.¹²

A final note is that the NYU ED algorithm is designed for the entire population, and not specifically for children. This is a known limitation of the algorithm, recognized by its developers (Billings, Parikh, Mijanovich 2000). However, lacking a child-specific algorithm, we consider this a valid starting point for our analysis.

A potential challenge to our identification is that the estimated differences could be attributed to the emergency management in Flint which began in December 2011, rather than the water contamination. To rule out the existence of trend in outcomes of interest prior to September 2015, as well as to explore its dynamics month-to-month, we estimate an extended form of specification (1) where the time period is disaggregated into monthly indicators:

$$Outcome_{ict} = a + \sum_{j} \beta_{1j} Flint * Month_{cj} + \beta_2 X_{ict} + \alpha_c + \delta_t + \varepsilon_{ict}$$
 (2)

where $Flint * Month_{cj}$ is a monthly indicator for an individual residing in Flint, and β_{1j} estimates the difference in month j between Flint and control cities with respect to September 2015.

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¹² Appendix J describes in detail the standard procedures for lead tests. While there is a plausible concern that parents are bringing children to the ED for lead tests, our review of provider surveys reported in the literature suggests that this is rare and most children in these situations are referred to primary care.

Results

Before proceeding with the analysis, we use an event study specification to justify selection of September 2015 as the beginning of the treatment period. Figure 1 shows results for event study, showing differences in monthly lead tests for children born in Flint compared to control cities. ^{13,14} Each point shows the difference in number of lead tests for children born in Flint compared to control cities with respect to September 2015. ¹⁵ The whiskers on each estimate provide the 95% confidence interval. The graph shows a clear rise in lead tests after September 2015, with a sharp peak between January and February 2016. The graph also shows no significant trend prior to September 2015, suggesting that despite ongoing speculation, the announcement of elevated residential tests by city authorities marked the beginning of the change in behavior with respect to health care receipt for their children among residents in Flint.

Table 1 shows summary statistics and unadjusted difference-in-differences estimates. In Panel A, we see minimal changes in the demographic characteristics in our sample population. Following the information shock, receipt of any lead test nearly doubles in Flint compared to a

¹³ Giving that one might be concerned by the increased volatility in the lead test results before May 2014, Appendix F shows consistent results starting the pre-period in May 2014.

¹⁴ The primary control cities are the other most populous cities in Michigan (Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, and Wyoming). Appendix K shows comparable results using alternative smaller control cities that have a history of high levels of lead in their drinking water (Detroit, Grand Rapids, Kalamazoo, Lansing, Wyoming, Battle Creek, Port Huron, Hamtramck, and Saginaw; see Urban 2018).

¹⁵ The American Academy of Pediatrics (AAP) recommends that children be tested for lead levels at ages 1 and 2, and suggests lead screening for older children who have not been tested. During the period covered by this study, the AAP changed its recommendation to venous blood draws for testing, noting that finger-prick sample testing yielded a high rate of false positives. The AAP lists Lead Screening in Children using CPT Code 83655 which we use in our analysis. The code does not allow for differentiation between finger-prick or venous blood tests. Because venous tests are more difficult to administer, this may introduce more heterogeneity among children who receive the test in Flint compared to other cities after the contamination became known. Though we would like to account for different methods of testing, we are not able to do so.

small increase in comparison areas. The unadjusted difference-in-differences results show a 1.6 percentage point increase in lead tests among Flint children compared to others, but much more modest changes in other types of health care, except ED visits which we discuss in greater detail below. We also see a small increase in payments in Flint compared to other cities. It is important to note that the after period (9/2015-2016) is much shorter than the lookback period (2013-9/2015). In Panel B, we find that ED visits that are non-preventable do not change, but we find decreases in all three of the preventable or non-emergent categories.

Main Results

Table 2 shows our primary differences-in-differences results. Using September 2015 as the treatment date (when the independent evidence of increases in lead exposure became public), the likelihood of receiving any lead test increased by 1.7 percentage points (pp), a 49 percent increase. We estimate a small, statistically insignificant decrease in the share of individuals having any office visits. Interestingly, given our results below, we see a slight, marginally statistically significant decrease in the share of children with an ED visit. This is possibly because ED visit is a heterogeneous measure including unclassified visits, which could be dampening the power of our analysis. Any claims and any payments increase by 1.8 and 1.7 pp, respectively. These are essentially a 4 percent increase in both categories. In panel B we examine the number per capita rather than an indicator for any receipt. The results are unsurprisingly quite similar for lead claims as individuals likely only receive one lead test. However, vaccinations demonstrate a potential positive spillover effect of receiving primary care for other services, with vaccinations increasing 1.3 pp (16%) in Flint compared to other areas following the information shock (Carpenter and Lawler 2019). Finally, claims per capita increase by 7.3, while overall payments increase \$12, which is only marginally statistically significant. This suggests that even if lead tests and their

related office visits are substituting for ED visits, they may be doing so at such a low rate that overall healthcare spending increases.

Table 3 contains results using the per capita measures of ED visits calculated using the method described above. We find no change in the number of non-preventable ED visits. For each of the other three types, our estimates indicate a decrease of between 1 and 2 visits per thousand enrollees per month, though the estimate for visits that were treatable in primary care is only marginally statistically significant on its own. We create two composite metrics: (1) PC sensitive, a combination of primary care treatable and non-emergent; and (2) Avoidable, a combination of primary care preventable, primary care treatable, and non-emergent. Each show strongly statistically significant decreases in ED visits per capita. The information shock in Flint is associated with nearly 5 fewer avoidable ED visits per 1000 enrollee-months, a decrease of 8.2 percent.

We perform event study analyses on office visits, avoidable ED visits, and payments in Table 4 and Figure 2. The motivation for this analysis is that we find a very large increase in lead tests only at specific times, most notably September 2015 and January 2016, in Figure 1. To test our hypothesis that this increase in lead tests should also increase office visits, we separate our results by post-information shock quarter. Our results suggest that immediately following the information shock, office visits increase by 0.9 pp (4%), while they increase by 2.7 pp (11%) in the first quarter of 2016. Office visits decrease in the last two quarters of treatment in Flint compared to control areas. Avoidable ED visits initially remain constant, but then decrease substantially and statistically significantly for the rest of the treatment period. This suggests an initial increase in office visits having a prolonged effect on ED visits. One way to explain these results is that this initial increase in office visits created a link between the patient (and his or her

parents) and the healthcare system. We explore this idea in more detail in the mechanism section.

To explore whether the substitution between ED and office visits was driven by specific lead exposure concerns, we selected the most common Clinical Classification Software (CCS)¹⁶ categories in the ED prior to September 2015. CCS categories were developed by the Healthcare Cost and Utilization Project of the Agency for Healthcare Research and Quality to classify ICD-9 diagnoses and procedures into clinically meaningful categories.

For this analysis, we identified the 10 most commonly occurring CCS categories in the ED corresponding to claims prior to September 2015 with diagnoses which the NYU algorithm classifies as avoidable.¹⁷ These CCS categories encompass over 86% of all avoidable claims in the ED and are listed in Table 5. We then aggregated claims to the person-month-CCS category, so that for each individual in our data, we have monthly use indicators, now split by CCS category. We excluded all individuals with no claims in the CCS category in the month. As with the personmonth analysis, we sum the NYU Algorithm indicators for preventable and non-preventable care in the ED. We re-estimate our specification for two venues of care: office visits (all diagnoses in each CCS), and ED (only avoidable shares as defined above).

We present results from this analysis in two formats. Figure 3 shows coefficient estimates by CCS category for any office visits (Panel A), avoidable ED visits (Panel B), and a scatterplot by category (Panel C). Table 6 then tests the hypothesis that in each CCS category the increase in office visits is mirrored in a decrease in avoidable ED visits. First looking at Figure 3, we see that in 6 of 9 CCS categories, office visits (Panel A) increase, with 5 of those 6 increases being

¹⁶ https://www.hcup-us.ahrq.gov/toolssoftware/ccs/ccs.jsp

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¹⁷ For this classification, we limited claims to those with diagnoses any avoidable component, then identified the 10 most common CCS within that subsample of claims. We chose to focus on CCS because classifying diagnoses is too specific and not sufficiently informative. This also allows us to impute avoidability of CCS based on these most common diagnoses, but by including all diagnoses in a given CCS we avoid defining this category too narrowly.

statistically significant. Preventable ED visits (Panel B), on the other hand, decline in 6 of 9 categories. Comparing specific CCS categories, we particularly notice a sharp increase in office visits and a decrease in preventable ED visits for skin and subcutaneous tissue infections. Abdominal pain is another category with a sharp increase in office visits and a decrease in preventable ED visits, as is gastritis and duodenitis. Comparing the any office visit and avoidable ED visits by condition (Panel C), we see a clear negative, linear relationship between the two results, with a greater increase in the share of children having any office visits causing a greater reduction in avoidable ED visits per capita.¹⁸

In Table 6 we present the results of a chi square test that compares the estimated change in office visits to that of preventable ED visits, (H₀: $\beta_{OfficeVisits} = -\beta_{AvoidableEDVisits}$), by CCS category. The chi square test fails to reject the null in any category, suggesting that, indeed, the increase in office visits is statistically indistinguishable from the decrease in preventable ED visits. *Mechanisms*

To test the potential mechanisms for changing medical utilization, we use individuals' episodes of care to explore choices in primary and ED care following the administration of lead test. Our main results suggest that the contamination has increased awareness of primary care through increased interaction with a physician or clinic. To examine this further, our analysis focuses on treatments received in the three months following a lead test to identify changing trends in utilization in Flint after September 2015.

The results of medical utilization in the post-lead test period are reported in Table 7. Here, the sample is limited to visits in the 3 month period following a lead test (columns (1)-(4)), and visits for patients with any lead tests (column (5)). We find statistically significant increases in

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¹⁸ Appendix E shows a similar relationship between total office visits and avoidable ED visits.

likelihood of indicators of established care: 3.7pp increase in immunization, 2.9pp increase in well-visit, 8pp increase in seeing the same provider, 14.8pp increase in using the same clinic. These results strongly suggest that following the information shock, those who received a lead test living in Flint were more likely to continue to receive regular care from the same clinic. It follows from this result that should a child become ill after having received a lead test, a parent would likely take the child back to the clinic at which he or she received this test to see the same physician. Parents who had not taken their child to receive a lead test would be more likely to take their sick child to the ED.

Pre-Trend Analysis

To test the validity of our specification, as well as to discern monthly trends of our analysis, we estimate the event study proposed in equation (2). The estimates of β_{1j} are presented graphically in Figure 4; each point represents the difference in outcome between Flint and control cities relative to September 2015. Panel (A) shows results for number of claims, Panel (B) is any office visits, and Panel (C) reflect preventable ED visits. All three panels show that, despite seasonal variation, there is no discernable trend in these outcomes prior to September 2015, validating our use of difference-in-differences estimation method. Furthermore, we note a sustained increase in claims and office visits in the treatment period.

Robustness Checks

We include several robustness checks in the appendices. This includes stratifying our sample to children born to black mothers (Table B1), children in fee-for-service Medicaid (Table B2), and children in managed care plans (Table B3). Our estimates follow a similar pattern and we continue to see a sizable decrease in avoidable ED visits in each of these samples. We also start

treatment in January 2016, instead of September 2015 (Table C1) and limit sample to the cohort of children born before April 2014 to avoid potential bias from endogenous fertility (Table D1). We also start the pre-period in May 2014 (Table F1), and use the "patched" NYU Algorithm (Table G1), per Johnson et al. (2017). Additionally, to test the sensitivity of our findings to the treatment period we estimated a flexible form specification, with two treatment periods – January to August 2015 and September 2015 to December 2016 - with qualitatively and quantitatively similar results, reported in Appendix H. This test confirms the validity of informational shock as we find all effects starting in September 2015, and not before.

One also might be concerned that the 49 percent increase in lead tests shown in our main results was coming from individuals who were not representative of the broader population. Table I1 addresses this concern by looking at unadjusted changes in demographic variables of those receiving a lead test in our difference-in-differences framework and finding no effect.

Finally, one might also be concerned that choosing comparison cities by population results in non-comparable group with regards to lead levels. Appendix K repeats our main analysis using an alternate control group of cities (including smaller ones) that have a history of high levels of lead (Urban 2018), and finds comparable results.

Discussion

The results in Figure 2 show that definitive public information about the Flint water led to a 4 percent increase in office visits in quarter 4 of 2015, the first treatment quarter. The associated change in office visits increased to 11 percent in the first quarter of 2016. From Table 3, column (6), we find a decrease of 4.9 visits per thousand, which on a mean of 59 per thousand represents an 8% decrease. Dividing the percent change in avoidable ED visits by the percent change in office

visits provides us with an estimate of elasticity of substitution between primary and ED care varying between -0.5 and -2.2.

Figure 3 then breaks this result down by common diagnosis classifications that are often avoidable. In addition to our results not being driven by one or two conditions, we generally see a negative relationship between the magnitude of the effect on office visits for a particular condition and the magnitude of the effect on avoidable ED visits for that same condition. For upper respiratory infections, skin and subcutaneous tissue infections, abdominal pain and gastritis and duodenitis, we find precisely estimated and opposite effects. A chi-squared test of parity between the magnitudes of the estimated coefficients (H₀: $\beta_{OfficeVisits} = -\beta_{AvoidableEDVisits}$) yields statistically insignificant results suggesting we cannot reject the null that these estimates are of equal magnitude just oppositely signed. We are hesitant in interpreting this test as it may lack specificity to reject our null hypothesis. However, this lends credence to our postulation that increased office visits are preventing avoidable ED visits.

Despite this substitution from potentially avoidable ED visits to office visits, we also find a statistically significant increase in total Medicaid spending. We attribute this to relative frequency of each type of visits; given the vast difference in the share of enrollees with any ED visit (0.091) in a given month vs. any office visit (0.265), the absolute increase in office visits and associated testing costs more than the savings from prevented ED visits. However, the total Medicaid spending amounts do not take into account that ED visits have other costs (e.g., stress, lost time, lost sleep, increased risk of complication, medical error, and infection) that may still make this substitution welfare improving. For this thought experiment, we assume that these other costs are larger for ED visits than for more common office visits.

Nevertheless, it is worth estimating the aggregate impact of healthcare costs following the water switch and comparing that to the proposed savings from the water switch. The 1.31 million enrollee months for the entire 2013-2016 data corresponds to 62,258 enrollees. Of those, approximately 3800 enrollees reside in Flint in the treatment period. Taking the \$12/month coefficient from Table 2, and multiplying it by 3800 and by 12 months provides an estimate of approximately \$550,000 additional Medicaid spending per year.

Flint city officials estimated that the water switch could save the city \$2.5 million a year.¹⁹ This means that Michigan Medicaid alone spent more than 20% of the projected savings on Flint enrollees between the ages of 0 and 3. This does not include the future costs of any resulting health conditions, the current health costs of individuals above the age of 3, or productivity losses of the Flint labor force.²⁰

Conclusion

As the intensity of exposure to environmental pollutants decreases with improved regulation and control, health outcomes and subsequent treatments associated with them will decrease. This, however, does not negate the burden imposed by such contaminations on communities, as the anxiety and uncertainty associated with such exposure increase, among other things, utilization of all medical services. This paper contributes by identifying opportunities inherent when such environmental disasters increase awareness of health and health care.

¹⁹ https://www.mlive.com/news/flint/index.ssf/2015/01/flints_dilemma_how_much_to_spe.html

²⁰ Future analysis will include testing the impact of the information shock on the flow of eligible Medicaid patients enrolling in Medicaid. We will also incorporate CMS provider data to enable study of heterogeneity on the physician side.

Our findings show that residents of affected communities often turn to health care providers for guidance on appropriate response. Because the population studied here is low-income Medicaid-covered young children, our findings directly benefit communities and policymakers attempting to determine what to emphasize (e.g., education, screening, remediation) to counteract potential negative health effects of lead exposure in early childhood. In addition, to the extent that we find medically inefficient care, these results can provide additional information for physicians to address parental anxiety about the possible long-term effects of exposure to lead.

Furthermore, the Flint water switch led to increases in lead tests and associated office visits and gives us a unique opportunity to study the substitution between office visits and potentially avoidable ED visits. While we find suggestive evidence of substitution, we do not find overall healthcare cost savings. Rather, health care costs in Flint increase. These results are specific to a cohort aged 0 to 3 years old. Thus, these results may not be generalizable to the general public.

This work documents the important role information can play in improving health care receipt for a disadvantaged population. Increased lead tests prompted parents to seek care for their children at the same clinic from which they received a lead test and reduced the likelihood of going to the ED for avoidable conditions. This may have important ramifications for any situation in which individuals are exogenously induced to seek care more often in primary care settings.

References

- Abadie, A., Athey, S., Imbens, G., and J. Wooldridge. 2017. When Should You Adjust Standard Errors for Clustering? NBER Working Paper No. 24003.
- Abadie, A, A. Diamond, and J. Hainmueller. 2010. Synthetic Control Methods for Comparative Case Studies: Estimating the Effect of California's Tobacco Control Program. *Journal of the American Statistical Association* 105(490): 493–505.
- Alexander, D., J. Currie, and M. Schnell. 2019. Check up before you check out: Retail clinics and emergency room use. Forthcoming, *Journal of Public Economics*.
- Alexander, Diane, Molly Schnell. 2019. The Impacts of Physician Payments on Patient Access, Use, and Health. NBER Working Paper No. 26095.
- Alsan, M. and M. Wanamaker. 2018. Tuskegee and the Health of Black Men. *Quarterly Journal of Economics* 133(1): 407–455.
- Ashwood, J. S; M. Gaynor; C.M. Setodji; R.O. Reid; E. Weber, and A Mehrotra. 2016. Retail Clinic Visits for Low-Acuity Conditions Increase Utilization and Spending. *Health Affairs* 35(3): 449–55.
- ATSDR (Agency for Toxic Substances and Disease Registry). 2007. Toxicological Profile for Lead. Case No. 7439–92–1. Atlanta, Georgia: ATSDR. Available at: http://www.atsdr.cdc.gov/toxprofiles/tp13.pdf.
- Ballard, D.W. et al. 2010. Validation of an Algorithm for Categorizing the Severity of Hospital Emergency Department Visits. *Med Care* 48(1): 58–63.
- Billings, J, Parikh, N, and Mijanovish, T. 2000. Emergency Room Use: The New York Story. *The Commonwealth Fund Issue Brief* 1–11.
- Bradley, C. J., S. O. Gandhi, D. Neumark, S. Garland, and S. M. Retchin. 2012. Lessons for coverage expansion: a Virginia primary care program for the uninsured reduced utilization and cut costs. *Health Affairs* 31(2):350–9.
- Bradley, C. J., and D. Neumark. 2017. Small Cash Incentives Can Encourage Primary Care Visits By Low-Income People With New Health Care Coverage. *Health Affairs* 36(8):1376–1384.
- Cameron, A. C., Gelbach, J. B., and D. L. Miller. 2008. Bootstrap-Based Improvements for Inference with Clustered Errors. *Review of Economics and Statistics* 90:3, 414–427
- Candon, M., S. Zuckerman, D. Wissoker, B. Saloner, G.M. Kenney, K. Rhodes, and D. Polsky. 2018. Declining Medicaid Fees and Primary Care Appointment Availability for New Medicaid Patients. *JAMA Intern Medicine* 178(1):145–146.

- Carpenter, C.S., and E. Lawler. 2019. Direct and Spillover Effects of Middle School Vaccination Requirements. *American Economic Journal: Economic Policy* 11(1): 95–125.
- Decker, S. 2018. No Association Found Between the Medicaid Primary Care Fee Bump and Physician Reported Participation in Medicaid. *Health Affairs* 37 (7): 1092–1098
- Ettner, S.L. 1996. The timing of preventive services for women and children: the effect of having a usual source of care. *American Journal of Public Health* 86(12): 1748–54.
- Gingold, DB, R. Pierre-Mathieu, B. Cole, A. Miller, and J. Khaldun, 2017. Impact of the Affordable Care Act Medicaid Expansion on Emergency Department High Utilizers with Ambulatory Care Sensitive Conditions: A Cross-Sectional Study. *American Journal of Emergency Medicine* 35(5): 737–742.
- Grossman, D. and D. Slusky. 2019. Lead in the Water and Birth Outcomes: The Case of Flint. Forthcoming, *Demography*.
- Grumbach, K., D. Keane, and A. Bindman. 1993. Primary Care and Public Emergency Department Overcrowding. *American Journal of Public Health* 83(3): 372–378.
- Hanna-Attisha M., J. LaChance, R.C. Sadler, and A. Champney Schnepp. 2016. Elevated Blood Lead Levels in Children Associated with the Flint Drinking Water Crisis: A Spatial Analysis of Risk and Public Health Response. *American Journal of Public Health* 106: 283–290.
- Jacobs, P.D., G. Kenney, and T. Selden. 2017. Newly Eligible Enrollees In Medicaid Spend Less And Use Less Care Than Those Previously Eligible. *Health Affairs* 36(9): 1637–1642.
- Johnston, K.J., L. Allen, T.A. Melanson, and S.R. Pitts. 2017. A "Patch" to the NYU Emergency Department Visit Algorithm. *Health Services Research* 52(4): 1264–1276.
- Klein, E., et al. 2017 The Effect of Medicaid Expansion on Utilization in Maryland Emergency Departments. *Annals of Emergency Medicine* 70(5): 607–614.
- Ladhania, R., A. Haviland, A. Venkat, R.l Telang, and J. Pines. 2019. The Effect of Medicaid Expansion on the Nature of New Enrollees' Emergency Department Use. Forthcoming, Medical Care Research and Review.
- Masten, S.J., S.H. Davies, and S.P. Mcelmurry. 2016. Flint Water Crisis: What Happened and Why? *Journal of the American Water Works Association* 108(12): 22–34.
- McConville, S., M. Raven, S. Sabbagh, and R. Hsia. 2018. Frequent Emergency Department Users: A Statewide Comparison Before And After Affordable Care Act Implementation. *Health Affairs* 37(6): 881–889.
- Neprash, H. T., A. Zink, J.Gray, and K.Hempstead. 2018. Chronic Care, Prescription Drugs & More Physicians' Participation In Medicaid Increased Only Slightly Following Expansion. *Health Affairs* 37(7): 1087–1091

- Nikpay, S., S Freedman, H Levy, and T Buchmueller. 2017. Effect of the Affordable Care Act Medicaid Expansion on Emergency Department Visits: Evidence From State-Level Emergency Department Databases. *Annals of Emergency Medicine* 70(2): 215–225.
- Paustian, M. L., Alexander, J. A., El Reda, D. K., Wise, C. G., Green, L. A. and Fetters, M. D. (2014), Partial and Incremental PCMH Practice Transformation: Implications for Quality and Costs. *Health Service Research* 49: 52–74.
- Pickens, G. et al. 2019. Changes In Hospital Service Demand, Cost, And Patient Illness Severity Following Health Reform. *Health Services Research* 54(4):739–751.
- Pitts, S. R., E. R. Carrier, E. C. Rich, and A. L. Kellermann. 2010. Where Americans get acute care: increasingly, it's not at their doctor's office. *Health Affairs* 29(9):1620–9.
- Polsky, D., M. Richards, S. Basseyn, D. Wissoker, G.M. Kenney, S. Zuckerman, K.V. Rhodes. 2015. Appointment Availability after Increases in Medicaid Payments for Primary Care. *New England Journal of Medicine* 372:537–545
- Powell, D. 2018. Imperfect Synthetic Controls: Did the Massachusetts Health Care Reform Save Lives? Working Paper. RAND Corporation Working Paper WR-1246.
- Sharma, A., S. Dresden, E. Powell, R. Kang, J. Feinglass. 2017. Emergency Department Visits and Hospitalizations for the Uninsured in Illinois Before and After Affordable Care Act Insurance Expansion, *Journal of Community Health* 42(3): 591–59
- Sommers, Benjamin, Robert Blendon, E. John Orav, Arnold Epstein, 2016. Changes in Utilization and Health Among Low-Income Adults After Medicaid Expansion or Expanded Private Insurance. *Journal of the American Medical Association* 176(10): 1501–1509.
- Sommers, B. D., B. Maylone, R. J. Blendon, E. J. Orav, and A. M. Epstein. 2017. Three-Year Impacts Of The Affordable Care Act: Improved Medical Care And Health Among Low Income Adults. *Health Affairs* 36(6):1119–1128.
- Starfield B., Shi L. 2004. The medical home, access to care, and insurance: a review of evidence. *Pediatrics* 113: 1493–1498.
- Taubman, S.L., H.L. Allen, B.J. Wright, K. Baicker, and A.N. Finkelstein. 2014. Medicaid Increases Emergency-Department Use: Evidence from Oregon's Health Insurance Experiment. *Science* 343(6186): 263–268.
- Urban, N. 2018. Beyond the Water: Lead Exposure in Michigan's Children. Data Driven Detroit. Accessed: 11/11/2019. https://datadrivendetroit.org/blog/2018/12/19/lead-exposure-in-michigans-children/
- Xu, K.T. 2002. Usual Source of Care in Preventive Service Use: A Regular Doctor versus a Regular Site. *Health Service Research*. 37(6): 1509–1529.

- Zahran, S., S.P. McElmurry, and R.C. Sadler. 2017. Four phases of the Flint Water Crisis: Evidence from blood lead levels in children. *Environmental Research* 157: 160–172.
- Zhu M., E.F. Fitzgerald, K.H. Gelberg, S. Lin, and C.M. Druschel. 2010. Maternal low–level lead exposure and fetal growth. *Environmental Health Perspectives* 118: 1471–1475.

Table 1: Summary Statistics

Panel A: Demographics and Primary Outcomes

	Before		Af	ter	Difference-in-	
	Flint	Other	Flint	Other	Differences	
Female	0.483	0.493	0.489	0.494	0.005*	
Black	0.610	0.533	0.617	0.536	0.0036	
Maternal Age	24.61	26.08	24.64	26.24	-0.1255***	
	(5.29)	(5.81)	(5.20)	(5.81)		
Any Lead Test	0.030	0.029	0.055	0.038	0.016***	
Any Office Visit	0.388	0.281	0.296	0.198	-0.008***	
Any ED Visit	0.104	0.091	0.091	0.082	-0.0037**	
# of Claims	3.814	3.766	2.335	2.305	-0.0177	
	(9.53)	(8.41)	(6.40)	(5.81)		
Payment	825.8	803.6	316.8	248.1	10.43	
	(3463.8)	(3447.4)	(1777.3)	(1985.1)		
Person Months	58927	762858	56549	751713		
Persons	3699	51091	3913	53914		

Panel B: Per Capita Emergency Department Visits by Type

	Before		Aft	Difference in Differences	
	Flint	Other	Flint	Other	
Non-Preventable	0.0093	0.0081	0.007	0.006	-0.0001
	(0.0674)	(0.0622)	(0.0561)	(0.0510)	
Preventable	0.0088	0.0061	0.0079	0.0071	-0.0019***
	(0.0630)	(0.0546)	(0.0603)	(0.0635)	
Primary Care Treatable	0.0378	0.0294	0.0324	0.026	-0.0020**
	(0.1604)	(0.1384)	(0.1491)	(0.1297)	
Non-Emergent	0.0275	0.0248	0.0239	0.0224	-0.0012*
	(0.1311)	(0.1249)	(0.1229)	(0.1199)	
PC Sensitive	0.0653	0.0541	0.0563	0.0484	-0.0032**
	(0.2489)	(0.2274)	(0.2306)	(0.2146)	
Avoidable	0.0742	0.0602	0.0643	0.0555	-0.0052**
	(0.2789)	(0.2498)	(0.2588)	(0.2405)	
Person Months	58927	762858	56549	751713	
Persons	3699	51091	3913	53914	

Table 2: Individual Level Difference-in-Differences Results for all Enrolled Children

	(1)	(2)	(3)	(4)	(5)	(6)
	Lead Claims	Office Visits	Vaccines	ED Visits	Claims	Payment
Panel A: Any						
Flint*After	0.017***	-0.003	0.005	-0.003*	0.018***	0.017***
	(0.001)	(0.006)	(0.003)	(0.002)	(0.004)	(0.004)
	[0.00]	[0.49]	[0.04]	[0.15]	[0.00]	[0.00]
R-squared	0.004	0.074	0.028	0.012	0.091	0.091
Dependent Variable Mean	0.035	0.249	0.064	0.088	0.461	0.46
Panel B: Number per	Capita					
Flint*After	0.017***	-0.027	0.013***	0.0010	0.073**	12.130*
	(0.001)	(0.036)	(0.004)	(0.002)	(0.032)	(6.646)
	[0.00]	[0.40]	[0.00]	[0.81]	[0.02]	[0.18]
R-squared	0.004	0.046	0.019	0.007	0.042	0.059
Dependent Variable Mean	0.039	1.058	0.083	0.138	3.067	349.042
Obs.	1,618,450	1,618,450	1,618,450	1,618,450	1,618,450	1,618,450
Number of enrollees	61,511	61,511	61,511	61,511	61,511	61,511
Number of Cities	16	16	16	16	16	16

Notes: Regressions at the at the enrollee-month level, for all eligible, enrolled children. Treated city is Flint. Control cities are Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, and Wyoming. Each coefficient is from a separate regression. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Total payment is trimmed to exclude top 1%. Robust standard errors are clustered at city level. Brackets contain wild bootstrapped p values.

*** p<0.01, ** p<0.05, * p<0.1

Table 3: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non- Preventable	Preventable	Primary Care Treatable	Non- Emergent	PC Sensitive	Avoidable
Flint*After	-0.00001	-0.0020***	-0.0019*	-0.0011**	-0.0030**	-0.0049***
rillit"Alter	(0.0003)	(0.0002)	(0.0009)	(0.0005)	(0.0010)	(0.0009)
	[1.00]	[0.01]	[0.28]	[0.01]	[0.06]	[0.01]
R-squared	0.004	0.003	0.008	0.006	0.01	0.01
Obs.	1,618,450	1,618,450	1,618,450	1,618,450	1,618,450	1,618,450
Number of enrollees	61,511	61,511	61,511	61,511	61,511	61,511
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.0072	0.0068	0.0284	0.024	0.0524	0.0592

Notes: Primary Care (PC) Sensitive visits include PC Treatable and Non-Emergent; Avoidable visits include Preventable, PC Treatable, and Non-Emergent. Regressions at the at the enrollee-month level, for all eligible, enrolled children. Treated city is Flint. Control cities are Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, and Wyoming. Each coefficient is from a separate regression. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Robust standard errors are clustered at city level. Brackets contain wild bootstrapped p values. *** p<0.01, ** p<0.05, * p<0.1

Table 4: Separate Coefficients Estimated for Each Post-Period Quarter

Panel A: Individual Level Difference-in-Differences Results for all Enrolled Children

	(1)	(2)	(3)	(4)	(5)	(6)
	Lead Claims	Office Visits	Vaccines	ED Visits	Claims	Payment
Panel A: Any						
Flint*After Qtr 1	0.013***	0.009**	0.003**	0.005**	0.025***	0.024***
	(0.001)	(0.004)	(0.001)	(0.002)	(0.004)	(0.005)
Flint*After Qtr 2	0.084***	0.027***	0.012***	-0.009***	0.033***	0.032***
	(0.001)	(0.004)	(0.002)	(0.001)	(0.006)	(0.006)
Flint*After Qtr 3	0.006***	0.0003	0.010**	-0.005**	0.016***	0.014***
	(0.001)	(0.006)	(0.003)	(0.002)	(0.004)	(0.004)
Flint*After Qtr 4	-0.012***	-0.024**	-0.001	-0.007**	0.007**	0.006*
	(0.003)	(0.010)	(0.004)	(0.003)	(0.003)	(0.003)
Flint*After Qtr 5	-0.007***	-0.033***	0.0003	-0.003	0.009	0.008
	(0.002)	(0.007)	(0.004)	(0.002)	(0.008)	(0.008)
R-squared	0.005	0.074	0.028	0.012	0.091	0.091
Dependent Variable Mean	0.035	0.249	0.064	0.088	0.461	0.46
Panel B: Number per 0	Capita					
Flint*After Qtr 1	0.013***	0.049*	0.013***	0.004	0.132***	23.742***
	(0.002)	(0.025)	(0.002)	(0.004)	(0.037)	(5.031)
Flint*After Qtr 2	0.086***	0.062**	0.015***	0.001	0.207***	26.440***
	(0.001)	(0.022)	(0.004)	(0.003)	(0.043)	(8.183)
Flint*After Qtr 3	0.004***	0.015	0.019***	-0.002	0.056	8.391
	(0.001)	(0.041)	(0.005)	(0.003)	(0.040)	(8.305)
Flint*After Qtr 4	-0.013***	-0.114*	0.011*	-0.004	0.024	-2.286
	(0.003)	(0.061)	(0.006)	(0.004)	(0.030)	(7.075)
Flint*After Qtr 5	-0.006**	-0.167***	0.013*	0.002	-0.013	-5.328
	(0.003)	(0.046)	(0.007)	(0.003)	(0.045)	(9.223)
R-squared	0.005	0.046	0.019	0.008	0.043	0.062
Dependent Variable Mean	0.039	1.058	0.083	0.138	3.067	349.042
Obs.	1,618,450	1,618,450	1,618,450	1,618,450	1,618,450	1,618,450
Number of enrollees	61,511	61,511	61,511	61,511	61,511	61,511
Number of Cities	16	16	16	16	16	16

Notes: Regressions at the at the enrollee-month level, for all eligible, enrolled children. Treated city is Flint. Each coefficient is from a separate regression. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Robust standard errors are clustered at city level. *** p<0.01, ** p<0.05, * p<0.1

Panel C: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non- Preventable	Preventable	Primary Care Treatable	Non- Emergent	PC Sensitive	Avoidable
Flint*After Qtr 1	0.0011*	-0.0008***	0.0004	0.0022**	0.0025	0.0017
	(0.0006)	(0.0003)	(0.0010)	(0.0009)	(0.0017)	(0.0016)
Flint*After Qtr 2	0.0004***	-0.0004	-0.0014*	-0.0027***	-0.0040***	-0.0045***
	(0.0001)	(0.0003)	(0.0007)	(0.0008)	(0.0010)	(0.0011)
Flint*After Qtr 3	-0.0002	-0.0023***	-0.0006	-0.0028***	-0.0034***	-0.0057***
	(0.0003)	(0.0001)	(0.0007)	(0.0007)	(0.0009)	(0.0009)
Flint*After Qtr 4	-0.0008	-0.0030***	-0.0042***	-0.0027***	-0.0069***	-0.0099***
	(0.0005)	(0.0004)	(0.0014)	(0.0005)	(0.0017)	(0.0020)
Flint*After Qtr 5	0.0001	-0.0029***	-0.0016*	0.0004	-0.0012	-0.0041***
	(0.0004)	(0.0007)	(0.0008)	(0.0009)	(0.0013)	(0.0013)
R-squared	0.004	0.003	0.008	0.006	0.01	0.01
Obs.	1,618,450	1,618,450	1,618,450	1,618,450	1,618,450	1,618,450
Number of enrollees	61,511	61,511	61,511	61,511	61,511	61,511
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.007	0.007	0.028	0.024	0.052	0.059

Notes: Primary Care (PC) Sensitive visits include PC Treatable and Non-Emergent; Avoidable visits include Preventable, PC Treatable, and Non-Emergent. Regressions at the at the enrollee-month level, for all eligible, enrolled children. Treated city is Flint. Control cities are Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, and Wyoming. Each coefficient is from a separate regression. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Robust standard errors are clustered at city level. *** p<0.01, ** p<0.05, * p<0.1

Table 5: Top CCS Categories for Avoidable Claims in the ED

CCS	Description	% of Claims
126	Upper Respiratory Infection (URI)	48.51
133	Lower Respiratory Infection (LRI)	10.83
197	Skin and Subcutaneous Tissue Infection	7.16
128	Asthma	6.81
251	Abdominal Pain	3.65
83	Epilepsy; convulsions	3.82
222	Hemolytic Jaundice and Perinatal Jaundice	1.74
140	Gastritis and Duodenitis	1.39
107	Cardiac Arrest and Ventricular Fibrillation	1.28
125	Acute Bronchitis	1.25

Notes: Top 10 most frequently occurring CCS categories in claims for care identified as avoidable by the NYU Algorithm taking place in the ED prior to September 2015.

Table 6: Effect Comparison of Substitution Between Office Visits and Avoidable ED Visits By Category of Care.

Description	Any Offic	e Visits	Avoidab	le ED Visits		\mathbf{H}_0
	Coeff	Std. Err	Coeff.	Std. Err	Chi2	p>Chi2
All	0.027	0.004	-0.003	0.001	0.020	0.886
Upper Respiratory Infection (URI)	0.017	0.008	-0.044	0.014	0.030	0.852
Lower Respiratory Infection (LRI) Skin and Subcutaneous Tissue	-0.027	0.012	0.007	0.016	0.050	0.831
Infection	0.086	0.015	-0.036	0.012	0.220	0.636
Asthma	0.002	0.016	-0.031	0.014	0.020	0.877
Abdominal Pain	0.168	0.027	-0.086	0.024	0.110	0.743
Epilepsy; convulsions	-0.078	0.018	-0.001	0.020	0.060	0.811
Jaundice	0.058	0.024	-0.007	0.003	0.280	0.596
Gastritis and Duodenitis	0.100	0.024	-0.095	0.008	0.000	0.968
Acute Bronchitis	-0.020	0.010	-0.003	0.005	0.020	0.902

Note: H_0 : $\beta_{OfficeVisits} = -\beta_{AvoidableEDVisits}$. Each estimate comes from a separate regression at the at the enrollee-month level, for all children with claims in the specified CCS category. Treated city is Flint. Control cities are Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, and Wyoming. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Robust standard errors are clustered at city level.

Table 7: Use of Primary Care Following Lead Testing

	(1)	(2)	(3)	(4)
			Same	Same
	Immunization	Well	Provider	Clinic
Flint*After	0.0367*	0.0294***	0.0802**	0.1480***
	(0.0148)	(0.0050)	(0.0231)	(0.0295)
R-squared	0.0252	0.0242	0.24	0.3463
Obs.	313-5-		*	
Obs.	21,413	21,413	16,820	16,820
Number of Cities	16	16	16	16
Dependent Variable				
Mean	0.1668	0.2004	0.5383	0.6272

Note: Each column shows estimates for specification for care received within 91 days of a lead test. The dependent variables are: Immunization – immunization as primary reason for visit (CCS code 10); Well – well child visit (CCS code 255 and 256); Same provider – provider seen was the same (National Provider Identifier) as the one administering the lead test; Same clinic – clinic was the same (National Biller Identifier) as in the one billing for lead test. Specifications (1)-(4) limit observations to visits within 91 days of administration of lead test. Treated city is Flint. Control cities are Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, and Wyoming. All regressions control for female, maternal race and education, and include fixed effects for city, month, year, birth year, and birth month. Robust standard errors are clustered at city level.

Figure 1 Number of Lead Tests in Flint Compared to Control Cities

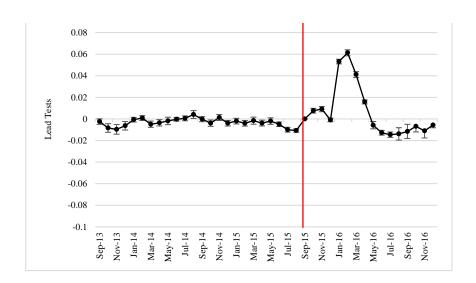


Figure 2: Quarterly Difference-in-Differences Effects in Office Visits, Avoidable ED Visits, and Payments

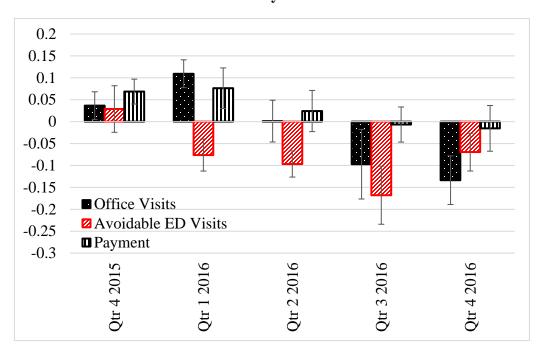
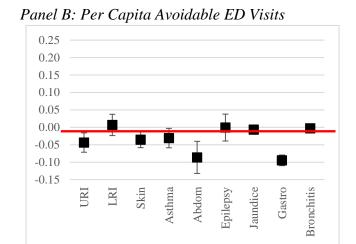
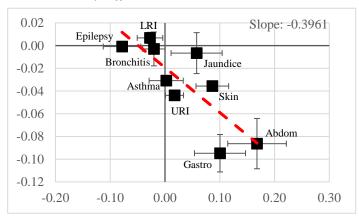


Figure 3: Changes in Outcome by Diagnosis Classification

Panel A: Any Office Visits 0.25 0.20 0.15 0.10 0.05 0.00 -0.05 -0.10 -0.15 URI LRI Skin Epilepsy Jaundice Asthma Abdom Gastro **Bronchitis**



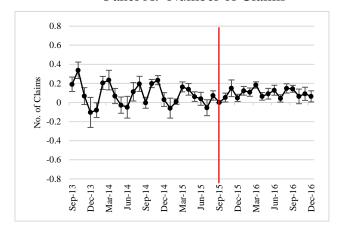
Panel C: Any Office Visits vs. Avoidable ED Visits



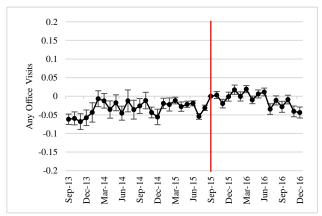
Notes: Each point is the estimate of a separate specification at the at the enrollee-month level, for all children with claims in the specified CCS category. Treated city is Flint. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Robust standard errors are clustered at city level. Whiskers show a 95% confidence interval. Panel C plots any office visit on the y-axis and per capita avoidable ED visits on the x-axis from panels A and B.

Figure 4: Adjusted Monthly Differences by Outcome

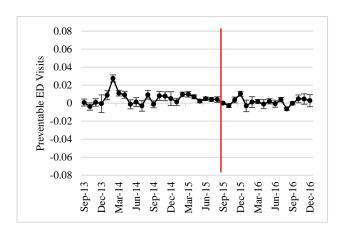
Panel A: Number of Claims



Panel B: Any Office Visits



Panel C: Preventable ED visit



Notes: Each graph represents estimation results from a separate specification. Each point represents the monthly difference between treated and control, adjusted for gender, maternal race, and maternal education. Treated city is Flint. Control cities are Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, and Wyoming. All regressions include fixed effects for city, birth year, and birth month. Robust standard errors are clustered at city level. Whiskers show a 95% confidence interval.

Appendix A. Background on Flint (adapted from Grossman and Slusky 2019)

Until 1967, Flint used the Flint River as a water source. The city had shortage concerns given its expanding population (Carmody 2016), and so began drawing water from Lake Huron through the Detroit Water and Sewerage Department (DWSD). In 2011, the Governor of Michigan appointed an Emergency Manager for the city to make fiscal decisions, given the city's precarious economic health (Longley 2011). At this time, DWSD water rates were rising (Zahran, McElmurry, and Sadler 2017). To avoid these higher rates, the Emergency Manager explored building a pipeline directly to Lake Huron (City of Flint 2015; Walsh 2014). However, the project would take more than two years to complete. In the interim, Flint would use water from the Flint River (beginning in April 2014), while Genesee County continued to work with DWSD (Carmody 2016).

Flint had to treat the new water source, but did not use anti-corrosive inhibitors (Pieper et al. 2017; Olson et al. 2017). Flint citizens were concerned about the appearance and odor of the water but were repeatedly assured that it was safe to drink (City of Flint 2015a,b). While the city issued multiple boil advisories due to a positive fecal coliform tests and an EPA violation for excess trihalomethanes (TTHM) in the water (Fonger 2014a, 2014b; Adams 2014), Flint consistently reassured citizens the water was safe and that any issues would be fixed soon (City of Flint 2015a,b).

In the summer of 2015, a team led by Mark Edwards began independently testing Flint's water and in August reported much higher levels of lead than previously reported, due to extremely corrosive water (http://flintwaterstudy.org/wp-content/uploads/2015/10/Flint-Corrosion-Presentation-final.pdf). Mona Hanna-Attish, a Flint pediatrician, in September, 2015 reported a substantial increase in blood lead levels in children (Fonger 2015c; Hanna-Attish et al. 2016). This finally led the city to switch back to Lake Huron water on October 16, 2015 (Emery 2015).

Appendix Figure A1: Timeline of Important Events in Flint

1897: Flint passes ordinance that all connections with any water main be made with lead pipes (Masten et al. 2016)	1967- 2014: Flint receives water from Detroit Water and Sewerage Department (DWSD)	2011: Governor appoints Emergency Manager	2009-2013: Water rates (prices) consistently increase	March 2014: Flint and Genesee County plan own pipeline to Lake Huron	-	Aug – Sept 2014: Positive test for fecal coliform, first boil advisory
Oct 2014: Flint GM plant switches off Flint water supply because of engine corrosion.	Jan – Mar 2015: Emergency manager stresses wate is safe, refuse to return to DWSD	Dr. Ed indeper tests Fli r lead lev	int waterpressivels, 19 confinere annotation increase.	Hanna- sha holds serence ouncing eased s of child d lead	Oct 2015: Flint stops receiving water from Flint River.	Jan 2016: Michigan Governor apologizes on national television

Source: Adapted from Grossman and Slusky (2019)

Appendix B: Stratified Analysis

Table B1: Sample limited to children born to black mothers

Panel A: Individual Level Difference-in-Differences Results for all Enrolled Children

	(1) Any lead	(2) Any office	(3) Any	(4) Any ED	(5) # of	(6) Total
	claims	visit	vaccines	visit	claims	payment (\$)
Flint*After	0.017*** (0.001)	-0.007** (0.003)	0.013*** (0.002)	-0.009*** (0.001)	0.019 (0.026)	-14.655 (8.571)
R-squared	0.005	0.049	0.023	0.007	0.041	0.059
Obs.	876,835	876,835	876,835	876,835	876,835	866,961
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.035	0.197	0.066	0.106	3.04	350.767

Panel B: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non- Preventable	Preventable	Primary Care Treatable	Non- Emergent	PC Sensitive	Avoidable
Flint*After	-0.0005*** (0.0001)	-0.0031*** (0.0001)	-0.0021*** (0.0003)	-0.0039*** (0.0006)	-0.0060*** (0.0008)	-0.0091*** (0.0009)
R-squared	0.004	0.002	0.005	0.004	0.006	0.005
Obs.	876,835	876,835	876,835	876,835	876,835	876,835
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.0082	0.0089	0.0353	0.0292	0.0645	0.0734

Table B2: Sample limited to children in fee-for-service plans

Panel A: Individual Level Difference-in-Differences Results for all Enrolled Children

	(1)	(2)	(3)	(4)	(5)	(6)
	Any lead	Any office	Any	Any ED	# of	Total
	claims	visit	vaccines	visit	claims	payment (\$)
Flint*After	0.024** (0.009)	0.095*** (0.010)	0.020** * (0.005)	-0.019 (0.014)	0.305** * (0.104)	214.064*** (18.505)
R-squared Obs. Number of Cities	0.039	0.099	0.029	0.037	0.008	0.057
	170,150	170,150	170,150	170,150	170,150	166,327
	16	16	16	16	16	16
Dependent Variable Mean	0.041	0.382	0.102	0.138	6.318	911.996

Panel B: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non-	Preventable	Primary Care	Non-	PC	Avoidable
	Preventable	Tieventable	Treatable	Emergent	Sensitive	Avoidable
Flint*After	-0.0007	-0.0033	-0.0035	-0.0051	-0.0086	-0.0118
	(0.0010)	(0.0022)	(0.0041)	(0.0049)	(0.0090)	(0.0111)
R-squared	0.005	0.014	0.021	0.018	0.026	0.03
Obs.	170,150	170,150	170,150	170,150	170,150	170,150
Number of	16	16	16	16	16	16
Cities	10	10	10	10	10	10
Dependent	0.0123	0.0098	0.0427	0.0365	0.0793	0.0891
Variable Mean	0.0123	0.0090	0.0427	0.0303	0.0793	0.0071

Table B3: Sample limited to children in managed care plans

Panel A: Individual Level Difference-in-Differences Results for all Enrolled Children

	(1)	(2)	(3)	(4)	(5)	(6)
	Any lead	Any office	Any	Any ED	# of	Total
	claims	visit	vaccines	visit	claims	payment (\$)
Flint*After	0.035***	0.024***	0.015**	-0.021**	-0.05	13.503
	(0.009)	(0.008)	(0.006)	(0.010)	(0.094)	(18.065)
R-squared	0.027	0.103	0.044	0.042	0.029	0.074
•	576.010	576 010	576 010	576 010	576,21	564.055
Obs.	576,213	576,213	576,213	576,213	3	564,255
Number of Cities	16	16	16	16	16	16
Dependent	0.085	0.587	0.149	0.208	6.748	722.562
Variable Mean	0.083	0.387	0.149	0.208	0.748	122.302

Panel B: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non- Preventable	Preventable	Primary Care Treatable	Non- Emergent	PC Sensitive	Avoidable
Flint*After	-0.0005 (0.0003)	-0.0053** (0.0019)	-0.0036 (0.0024)	-0.0078* (0.0042)	-0.0114* (0.0065)	-0.0167* (0.0084)
R-squared Obs.	0.007 576,213	0.013 576,213	0.023 576,213	0.019 576,213	0.028 576,213	0.031 576,213
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.0166	0.0161	0.0672	0.0565	0.1238	0.1399

Appendix C: Alternative Treatment Starting Date

Table C1: Treatment starting in January 2016

Panel A: Individual Level Difference-in-Differences Results for all Enrolled Children

	(1)	(2)	(3)	(4)	(5)	(6)
	Any lead	Any office	Any	Any ED	# of	Total
	claims	visit	vaccines	visit	claims	payment (\$)
Flint*After Jan'16	0.016*** (0.002)	-0.008 (0.006)	0.005 (0.003)	-0.007*** (0.002)	0.051 (0.035)	3.638 (7.649)
R-squared Obs.	0.004	0.074	0.028	0.012	0.042	0.059
	1618450	1618450	1618450	1618450	1618450	1602669
Number of Cities Dependent Variable Mean	16	16	16	16	16	16
	0.035	0.249	0.064	0.088	3.067	349.042

Panel B: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non- Preventable	Preventable	Primary Care Treatable	Non- Emergent	PC Sensitive	Avoidable
Flint*After Jan '16	-0.0003 (0.0002)	-0.0020*** (0.0002)	-0.0020** (0.0007)	-0.0023*** (0.0006)	-0.0043*** (0.0009)	-0.0063*** (0.0009)
R-squared Obs.	0.004 1618450	0.003 1618450	0.008 1618450	0.006 1618450	0.01 1618450	0.01 1618450
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.0072	0.0068	0.0284	0.024	0.0524	0.0592

Appendix D: Fixed Birth Cohort

Table D1: Following the sample of children born before April 2014

Panel A: Individual Level Difference-in-Differences Results for all Enrolled Children

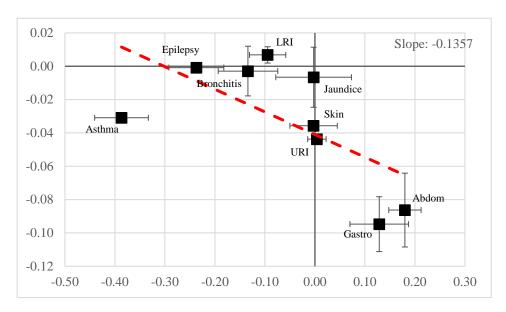
	(1)	(2)	(3)	(4)	(5)	(6)
	Any lead	Any office	Any	Any ED	# of	Total
	claims	visit	vaccines	visit	claims	payment (\$)
Flint*After	0.030***	-0.012	-0.003	-0.006**	0.111***	6.486*
	(0.004)	(0.007)	(0.004)	(0.003)	(0.035)	(3.481)
R-squared Obs.	0.007	0.075	0.028	0.013	0.05	0.08
	856217	856217	856217	856217	856217	849352
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.035	0.241	0.053	0.086	2.765	301.847

Panel B: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non- Preventable	Preventable	Primary Care Treatable	Non- Emergent	PC Sensitive	Avoidable
Flint*After	-0.0004 (0.0005)	-0.0026*** (0.0001)	-0.0046** (0.0016)	-0.0012 (0.0008)	-0.0058** (0.0022)	-0.0084*** (0.0021)
R-squared	0.005	0.003	0.009	0.007	0.01	0.011
Obs.	856217	856217	856217	856217	856217	856217
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.0072	0.007	0.0273	0.0235	0.0509	0.0579

Appendix E: Changes in Total Office Visits by Diagnosis Classification

Figure E1: Total Office Visits vs. Avoidable ED Visits



Notes: Each point is the estimate of a separate specification at the at the enrollee-month level, for all children with claims in the specified CCS category. Treated city is Flint. Control cities are Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, and Wyoming. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Robust standard errors are clustered at city level. Whiskers show a 95% confidence interval.

Appendix F: Pre-period Starting in May 2014 for ED Visits Type

Table F1: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non- Preventable	Preventable	Primary Care Treatable	Non- Emergent	PC Sensitive	Avoidable
Flint*After	-0.0001 (0.0003)	-0.0016*** (0.0001)	-0.0006 (0.0008)	-0.0016*** (0.0005)	-0.0022** (0.0010)	-0.0038*** (0.0009)
R-squared	0.004	0.003	0.009	0.006	0.01	0.01
Obs.	1398971	1398971	1398971	1398971	1398971	1398971
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.0067	0.007	0.0282	0.0237	0.0519	0.059

Appendix G: Results with "Patched" NYU Algorithm

Using Johnson et al. (2017) classification of uncategorized visits, we re-estimated specification (1) for ED visits. Results are presented in Table G1; though the significance of most estimates is lost, and the magnitudes are attenuated, the sign is consistent with our main results. We choose not to use this "patch" because the new classifications are not validated.

Table G1: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non- Preventable	Preventable	Primary Care Treatable	Non- Emergent	PC Sensitive	Avoidable
Flint*After	0.0000 (0.0005)	-0.001*** (0.0001)	-0.0006 (0.0009)	-0.0006 (0.0004)	-0.0012 (0.0013)	-0.0022* (0.0012)
R-squared	0.0051	0.0027	0.0079	0.0056	0.0089	0.0088
Obs.	1,326,764	1,326,764	1,326,764	1,326,764	1,326,764	1,326,764
Number of enrollees	67,167	67,167	67,167	67,167	67,167	67,167
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.0121	0.0072	0.037	0.029	0.066	0.0738

Appendix H: Flexible form time indicator

Table H1: Flint dummy interacted with multiple dummies for multiple post-periods

Panel A: Individual Level Difference-in-Differences Results for all Enrolled Children

	(1)	(2)	(3)	(4)	(5)	(6)
	Any lead claims	Any office visit	Any vaccines	Any ED visit	# of claims	Total payment (\$)
Flint*Jan '15	-0.0031***	-0.0022	-0.0016	0.0015	-0.1813***	-22.66
	(0.0006)	(0.0061)	(0.0029)	(0.0024)	(0.0337)	(13.29)
Flint*Sept '15	0.0304***	0.0104	0.0070**	-0.0016	0.0513	42.86*
	(0.0008)	(0.0063)	(0.0030)	(0.0022)	(0.0429)	(21.02)
R-squared	0.0072	0.0683	0.027	0.0098	0.041	0.0374
F-Test	1703.3	7.1	15.8	8.7	62.6	29.7
Obs.	1,330,177	1,330,177	1,330,177	1,330,177	1,330,177	1,330,177
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.0328	0.0266	0.0685	0.091	3.336	630.191

Notes: *Flint*Jan '15* indicates enrollee-month observations in Flint between January and August 2015. *Flint*Sept '15* indicates enrollee-month observations in Flint between September 2015 and December 2016. F-statistic of joint significance of DiD coefficients reported. Regressions at the at the enrollee-month level, for all eligible, enrolled children. Treated city is Flint. Control cities are Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, and Wyoming. Each coefficient is from a separate regression. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Robust standard errors are clustered at city level. *** p<0.01, *** p<0.05, * p<0.1

Panel B: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non- Preventable	Preventable	Primary Care Treatable	Non- Emergent	PC Sensitive	Avoidable
Flint*Jan '15	0.0005**	-0.0005	-0.0029	0.0023***	-0.0006	-0.0012
	(0.0001)	(0.0003)	(0.0018)	(0.0004)	(0.0023)	(0.0026)
Flint*Sept '15	0.0004	-0.0015***	-0.0022	-0.0001	-0.0023	-0.0039**
	(0.0003)	(0.0001)	(0.0016)	(0.0004)	(0.0017)	(0.0017)
R-squared	0.004	0.003	0.0068	0.0052	0.008	0.0082
F-Test	9.04	48.81	1.58	13.79	6.84	19.59
Obs.	1,330,177	1,330,177	1,330,177	1,330,177	1,330,177	1,330,177
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.0077	0.0067	0.0295	0.0248	0.0544	0.0611

Notes: *Flint*Jan '15* indicates enrollee-month observations in Flint between January and August 2015. *Flint*Sept '15* indicates enrollee-month observations in Flint between September 2015 and December 2016. F-statistic of joint significance of DiD coefficients reported. Dependent variables: Primary Care (PC) Sensitive visits include PC Treatable and Non-Emergent; Avoidable visits include Preventable, PC Treatable, and Non-Emergent. Regressions at the at the enrollee-month level, for all eligible, enrolled children. Treated city is Flint. Control cities are Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, and Wyoming. Each coefficient is from a separate regression. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Robust standard errors are clustered at city level. *** p<0.01, *** p<0.05, * p<0.1

Appendix I: Patient Characteristics for Lead Tests

Table I1: Characteristics of Patients Receiving a Lead Test

	Before	e	After	After		
_	Flint	Other	Flint	Other	in- Differences	
Female	0.493	0.476	0.502	0.466	0.018	
Black	0.559	0.339	0.545	0.352	-0.027	
Maternal Age	24.766	26.842	25.067	27.32	-0.177	
	(5.456)	(5.797)	(5.186)	(5.715)		
Claims	1300	9511	56549	9063		

Note: *** p<0.01, ** p<0.05, * p<0.1 Standard deviation in parenthesis for non-dummy variables..Other cities are Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, and Wyoming. All regressions control for female, maternal race and education, and include fixed effects for city, month, year, birth year, and birth month. Robust standard errors are clustered at city level.

Appendix J: Lead Testing in the ED

While testing for blood lead level is possible in the ED, it is done so on suspicion of lead poisoning in anticipation of inpatient admission. Treatment for exposure to high levels of lead, warranting hospital admission, is chelation therapy. Our data does not include any claims for chelation therapy. Thus, we feel confident that admissions on suspected lead exposure did not occur in Flint during our period covered by our data.

Subacute lead exposure among children presents with symptoms which are nonspecific and may only involve irritability, difficulty concentrating, and fatigue. Most commonly, it is associated with constipation. Beyond admission, the recommended best practice for suspected exposure to lead is to remove the source of contamination, test for lead in outpatient setting, and follow up with a primary care provider. For a child brought by their parent to the ED on suspicion of lead poisoning, the providers may ascertain that the child is in no immediate danger and take a blood sample to send to an off-campus testing facility with results sent to a primary care provider for follow up. Alternatively, the provider may ascertain that the child is in no immediate danger and refer the parent to primary care for testing.

In an informal survey, 13 emergency physicians were asked: "A parent brings their child to the ED requesting a lead test. The child has no specific symptoms, maybe a mild rash or mild abdominal pain. No apparent urgency. Which would you do?" The responses were distributed as follows:

Order lead test in the ED: 2

Refer to primary care: 6

Test and follow up with primary care: 5

We, therefore, conclude that parents requesting a blood lead test for their child in the ED setting would be referred to primary care.

Reference:

Williams Saralyn. Heavy Metals and Iron Overdose. In: Mattu A and Swadron S, ed. CorePendium. Burbank, CA: CorePendium, LLC.

https://www.emrap.org/corependium/chapter/recGL1d1CsAmcMhdL/Heavy-Metals-and-Iron-Overdose. Updated November 7, 2019. Accessed November 7, 2019.

Appendix K: Alternative Control Groups

Table K1: Cities with Most Lead in Water in Michigan

Panel A: Individual Level Difference-in-Differences Results for all Enrolled Children

	(1)	(2)	(3)	(4)	(5)	(6)
	Any lead	Any office	Any	Any ED	# of	Total
	claims	visit	vaccines	visit	claims	payment (\$)
Flint*After	0.015***	-0.007	0.0001	-0.001	0.126**	22.066
	(0.001)	(0.005)	(0.003)	(0.001)	(0.039)	(12.054)
R-squared	0.005	0.062	0.023	0.009	0.043	0.059
Obs.	1,379,219	1,379,219	1,379,219	1,379,219	1,379,219	1,364,489
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.036	0.227	0.054	0.096	3.084	356.483

Notes: Regressions at the at the enrollee-month level, for all eligible, enrolled children. Treated city is Flint. Control cities are Detroit, Grand Rapids, Kalamazoo, Lansing, Wyoming, Battle Creek, Port Huron, Hamtramck, and Saginaw (Urban 2018). Each coefficient is from a separate regression. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Total payment is trimmed to exclude top 1%. Robust standard errors are clustered at city level.

Panel B: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non- Preventable	Preventable	Primary Care Treatable	Non- Emergent	PC Sensitive	Avoidable
Flint*After	0.0003	-0.002***	-0.001	-0.0004	-0.002***	-0.004***
	(0.0002)	(0.0001)	(0.001)	(0.001)	(0.0004)	(0.0005)
R-squared	0.005	0.002	0.007	0.005	0.008	0.008
Obs.	1,379,219	1,379,219	1,379,219	1,379,219	1,379,219	1,379,219
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.008	0.007	0.031	0.026	0.058	0.065

Notes: Regressions at the at the enrollee-month level, for all eligible, enrolled children. Treated city is Flint. Control cities are Detroit, Grand Rapids, Kalamazoo, Lansing, Wyoming, Battle Creek, Port Huron, Hamtramck, and Saginaw (Urban 2018). Each coefficient is from a separate regression. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Total payment is trimmed to exclude top 1%. Robust standard errors are clustered at city level. *** p < 0.01, ** p < 0.05, * p < 0.1

Table K2: Main sample including Pontiac and Muskegon

Panel A: Individual Level Difference-in-Differences Results for all Enrolled Children

	(1) Any lead claims	(2) Any office visit	(3) Any vaccines	(4) Any ED visit	(5) # of claims	(6) Total payment (\$)
Flint*After	0.017*** (0.001)	-0.003 (0.006)	0.005* (0.003)	-0.004** (0.002)	0.087*** (0.029)	11.221* (6.120)
R-squared	0.004	0.073	0.027	0.012	0.042	0.059
Obs.	1,738,325	1,738,325	1,738,325	1,738,325	1,738,325	1,721,319
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.035	0.251	0.063	0.089	3.091	349.352

^{***} p<0.01, ** p<0.05, * p<0.1

Panel B: Changes in Per Capita ED Visits by Type

	(1)	(2)	(3)	(4)	(5)	(6)
	Non- Preventable	Preventable	Primary Care Treatable	Non- Emergent	PC Sensitive	Avoidable
Flint*After	-0.00002	-0.002***	-0.002*	-0.001**	-0.003**	-0.005***
	(0.0003)	(0.0002)	(0.001)	(0.0004)	(0.001)	(0.001)
R-squared	0.004	0.003	0.008	0.006	0.009	0.009
Obs.	1,738,325	1,738,325	1,738,325	1,738,325	1,738,325	1,738,325
Number of Cities	16	16	16	16	16	16
Dependent Variable Mean	0.007	0.007	0.028	0.024	0.053	0.059

Notes: Regressions at the at the enrollee-month level, for all eligible, enrolled children. Treated city is Flint. Control cities are Control cities are Ann Arbor, Dearborn, Detroit, Farmington Hills, Grand Rapids, Kalamazoo, Lansing, Livonia, Rochester Hills, Southfield, Sterling Heights, Troy, Warren, Westland, Wyoming, Pontiac, and Muskegon. Each coefficient is from a separate regression. All regressions include fixed effects for city, claim year, claim month, birth year, and birth month. Total payment is trimmed to exclude top 1%. Robust standard errors are clustered at city level.