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# Stock Market Listing, Investment, and Business Groups: How Firm Structure Impacts Investment?

Joseph J. French,<sup>a</sup> University of Northern Colorado Ryosuke Fujitani,<sup>b</sup> Hitotsubashi University Yukihiro Yasuda,<sup>c</sup> Hitotsubashi University

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## Abstract

We provide the first large sample comparison of investment by Japanese listed and unlisted public firms. Empirical results show listed firms invest more than comparable unlisted companies. Our findings suggest that the role of listing in alleviating financial constraints is more important than potential underinvestment due to myopic managerial behavior. However, we find that the positive relationship between listing and investment is primarily driven by standalone firms. Further analysis confirms that as the number of subsidiaries in a business group (i.e., firms with subsidiaries) increases the positive impact of listing on investment declines. In contrast, listed standalone firms invest more and are more sensitive to investment opportunities than unlisted companies. Additional results show that listing more positively impacts investment when a firm faces financial constraints. We also find a positive relationship between stock liquidity and investment for listed firms. Taken together, our results suggest that markets play an important role in easing financial constraints and preventing managerial shirking both of which increase investment. Finally, we show for all firms that higher levels of ownership by financial institutions, board members, and foreign investors facilitates higher levels investment.

#### JEL classification: G20, G31, G39, F3

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<sup>&</sup>lt;sup>a</sup> Monfort College of Business, University of Northern Colorado, Campus Box 128, Greeley Colorado 80639, USA. joseph.french@unco.edu

<sup>&</sup>lt;sup>b</sup> Graduate School of Commerce and Management, Hitotsubashi University, 2-1 Naka Kunitachi Tokyo 186-8601, Japan. cd171002@g.hit-u.ac.jp

<sup>&</sup>lt;sup>c</sup> Graduate School of Business Administration, Hitotsubashi University, 2-1 Naka Kunitachi Tokyo 186-8601, Japan y.yasuda@r.hit-u.ac.jp

## **1. Introduction**

Does being listed on a stock exchange impact long-term-investment? Existent literature suggests it does. On one hand, listing may reduce long-term investment. It is often argued that listing induces short-term pressures on management which leads to short-termism which may reduce long-term investment (Stein, 1989). On the other hand, listing reduces the cost of capital as firms have access to public equity markets, which should increase investment of listed firms relative to unlisted ones. Recent empirical literature is mixed. Asker et al. (2015) show that short-term market pressures cause lower investment levels for listed firms relative to unlisted firms. However, Bakke, Jens, and Whited (2012) find that stock market listing increases investment. The authors argue that the liquidity benefit for public companies induces them to invest more than private companies. Gilje and Taillard (2016) show that private firms are less responsive to investment opportunities compared to public firms. They suggest that access to capital is important in explaining differences between investment of public and private firms.

We build on this literature. In this study, we conduct a large sample comparison of investment behavior by large Japanese listed and unlisted companies from 2001-2017. Existing research primarily focuses on differences between public and private firms. However, these comparisons are susceptible to confounding effects due to different disclosure requirements. While listed firms are required to comply with the listing criteria of the stock exchange and regulatory agencies, generally private firms are required to disclose at a much lower level.

We overcome this problem by taking advantage of institutional characteristics of the Financial Instruments Exchange Act (J-FIEA, hereafter) of Japan. J-FIEA mandates all firms whose equity satisfies several criteria to file financial statements that are the Japanese counterpart of 10K filing (J 10-K, hereafter)<sup>1</sup>. Interestingly, the firms that are required to file financial statements include not only public listed firms but also many unlisted firms. This unique institutional feature allows us to understand the role of short-termism and capital constraints on the investment levels between listed and unlisted firms without confounding effects due to different disclosure requirements. Additionally, since our sample of unlisted firms are required to report consolidated financial statements and full format of J 10-K, we can collect the data on governance and ownership structure that are at the same level as listed firms.

We find that listed firms invest more than their unlisted counterparts. This result suggests that listing status alleviates financial constraints and increases investment in Japan. Our findings contrast with the hypothesis that market pressure exerted on listed firms induce short-termism and reduces investment relative to unlisted firms (Asker et al., 2015). Thus, our results indicate that a simple comparison between public and private firms might be misleading when examining the effects of listing status on firm investment behavior.

Next, we investigate how firm structure impacts investment behavior of listed and unlisted firms. We divide our sample into two sub-samples: 1) Firms that are members of business groups, and 2) standalone firms. Our results show that the positive relationship between listing and investment is primarily driven by standalone firms. We do not find any significant evidence on the difference in investment levels between listed and unlisted business groups. Further analysis confirms that as the number of subsidiaries in a business group increases the positive impact of listing on investment declines. In contrast, standalone firms invest more and are more sensitive to investment opportunities than unlisted firms. These findings suggest that listing relaxes the financing constraint and allows for greater investment and investment sensitivity for listed standalone firms.

<sup>&</sup>lt;sup>1</sup> We describe the institutional details of the requirements in A1.

In a series of additional tests, we show that the positive role of listing on investment is particularly important for financially constrained firms. Providing evidence that higher levels of investment by listed firms is not due to overinvestment created by potential agency problems.

In most cases, we show that higher ownership by financial institutions and foreign investors increases investment, potentially mitigating financial constraints in the case of financial institutions. Alternatively, these types shareholders may monitor management, which reduces managerial shirking and thereby increases investment. Indeed, we find that the positive impact of foreign investors on investment is greater for listed firms. Finally, we examine the role of stock liquidity on listed firm investment. We show that liquidity enhances the positive relationship between listing and investment, particularly in the business group subsample. This finding suggests that liquid stocks act as a monitor of management reducing shirking and producing higher levels of investment. Additionally, liquid stocks make raising additional capital easier, which also facilitates investment. Overall, our results have implication for understand corporate investment throughout the world, give the prevalence of business groups in many developed and developing economies.

Our research contributes to two strains of literature. This first relates to differences between financial decisions of public and private firms (i.e., the counterparts of listed and unlisted firms in our context). For example, Asker et al. (2015) find that compared with private firms, public firms invest less and are less sensitive to changes in investment opportunities. Orihara (2014) confirms the same qualitative results of Asker et al. (2015) in Japanese counterparts. Focusing in UK private companies, Brav (2009) finds that private firms rely on debt financing extensively, and thereby have higher leverage ratios and avoid external financing. Bigelli and Sancez-Vidal (2012) investigate cash holdings of Italian private firms. They find that higher cash holdings for smaller private firms that are characterized as being younger, riskier and financially constrained.

The current paper is also related to the literature on the impact of market pressures on investment. Public firms have better access to capital markets than private firms. However, short-term market pressures, such as quarterly earnings announcements exert pressure on management, which may distort investment. Many authors argue, that too much focus on short-term profits or stock price by public firms distort investment decisions and cause firms to forego positive value creating investments.<sup>2</sup> These pressures cause public firms to invest less than comparable private firms which are not subject to market pressure. However, using a sample of firms from the natural gas industry, Gilje and Taillard (2016) show that access to external capital is most relevant for explaining differences in investment between public and private firms. In a study focusing on Japan, Ikeda et al (2017) find evidence consistent with the 'quiet life hypothesis' or that managers of public firms avoid making difficult investment decisions when they are protected from the disciplinary effects of capital markets, which may also lead public firms to underinvest.

Empirical evidence in Orihara (2017) suggests that the liquidity market monitoring tradeoff of listing has heterogeneous effects on a firm's investment, depending on the nature of the firm. Furthermore, using private firms as a control group for the treatment group of listed firms, Ueda et al. (2019) show that listing on stock exchange mitigates financial constraints Whereas the current study examines the effects of listing status by using unlisted (public) firms, not private firms<sup>3</sup>.

 $<sup>^{2}</sup>$  Morck et al. (1990) is a good review on this line of research. See also Shleifer and Vishney (1990) and Stein (1989) for examples.

<sup>&</sup>lt;sup>3</sup> Ueda et al. (2019) argues that listing status of a firm barely changes over time in Japan. In this sense, using Japanese data has another advantage for this type of study because the selection bias between listed and private is always challenging topic in US studies.

The remainder of the paper proceeds as follows. Section 2 reviews relevant institutional background, in Section 3 our hypotheses are developed, Section 4 describes our sample and presents our empirical methods, Section 5 discusses the empirical results, and Section 6 provides a summary of observations and directions for further research.

# 2. Institutional Background

Identifying the effects of listing status on investment is a challenging task for the following reasons: 1) Unlisted firms are not contained in most databases,<sup>4</sup> and 2) Disclosure requirements for listed and unlisted firm are different. Listed firms are required to comply with the disclosure and legal criteria of the stock exchange, while unlisted firms generally have softer disclosure requirements. For instance, unlisted firms do not need to follow Regulation Fair Disclosure in the U.S. (Farre-Mensa, 2017)<sup>5</sup>. Thus, even if the data was available, the differences in the disclosure levels causes serious confounding effect problems.

To address these issues, we take the advantage of the institutional features of the Japanese Financial Instruments Exchange Act (J-FIEA, hereafter). Article 24 of the J-FIEA mandates firms to report audited financial statements, if a firm (a) issues securities listed in a financial instruments exchange, (b) issue securities publicly offered, or (c) issue unlisted securities held by more than one thousand investors.<sup>6</sup>

Figure 1 describes the definition of listed, quasi-private, and purely-private firms in the context of existing literature. The *X-axis* represents the strictness of mandatory disclosure

<sup>&</sup>lt;sup>4</sup> One important exception is Sageworks, which follows approximately 40,000 U.S. unlisted firms, and several studies use it to investigate unlisted firms (e.g., Asker et al., 2015; Farre-Mensa 2017).

<sup>&</sup>lt;sup>5</sup> Anecdotal discussion reports that required reporting for listed Japanese firms is more than 100 pages. Whereas unlisted Japanese firms (*Jigyou Houkokusho*) only has approximately 15 pages of required reporting, even if the firm is large (e.g., Hankyu Corporation whose total assets is more than approximately 1 trillion US dollars).

<sup>&</sup>lt;sup>6</sup> We describe the details of the requirements in A1.

requirements. In general, the disclosure requirements for listed firms are stricter than those for purely private firms. The reference point dividing purely-private, and listed and quasi-private firms represents the minimum requirements of disclosure for listed firms. Quasi-private firms do not go public, but the Japanese Financial Instrument Exchange Act requires that these firms disclose at the same level as listed firms.

We refer to these "quasi-private firms" as "unlisted firms" hereafter. Since these unlisted firms must report consolidated financial statements and full format of J 10-K (Japanese equivalent to U.S 10k), we are able to collect comparable financial, governance and ownership structure data to those of listed firms. In addition, we are able to avoid the confounding effects caused by the different disclosure requirement levels.

# [Figure 1]

This definition of a listed firm in this study does not necessarily coincide with a "public firm" in the previous literature<sup>7</sup>. Thus, quasi-private firms could be regarded as public firms. In our study, listed firms must be traded on public exchanges consistent with the definitions of Katz (2009) and Farre-Mensa (2010).

Several U.S. studies use quasi-private firms as the counterpart of unlisted firms, although they call them private firms. For instance, Gao et al. (2013) find that listed firms hold more cash reserves than quasi-private firms. They attribute the difference to the higher agency problems in listed firms. Acharya and Xu (2017) show that the private firms perform more intense innovation investment and see better innovation performance. Using quasi-private (unlisted) firms and listed

<sup>&</sup>lt;sup>7</sup> For example, Minnis and Shroff's (2017) define a private firm "as one with capital (e.g., debt or equity) that is not traded in a secondary market (p. 475)". Their definition of public firm includes not only listed firms, but also firms issuing equities and bonds that are traded in the over-the-counter market.

firms, other recent studies investigate the differences in CEO compensation (Gao and Li, 2015), CEO turnover (Gao et al., 2017), and innovation strategy (Gao et al., 2018).

## **3. Hypothesis Development**

Does listing status impact investment? Do standalone firms invest differently than business groups? There is a small but growing literature looking into these questions and we develop our hypotheses considering this literature.

Public ownership has benefits and costs. The dark-side of stock market listing indicates that listing induces short-term market pressures, which leads to myopic behavior by managers. Early theoretical work by Stein (1989) provides insights on why listed and unlisted firms invest differently. According to Stein's model, managers attempt to mislead markets about the value of their firm. In an effort to do so, they forsake some positive net present value (NPV) investments to increase current earnings. Stein shows that even when facing efficient markets managers continue to act myopically. Indeed, Graham, Harvey, and Rajgopal (2005) show that foregoing positive NPV projects can boost current earnings and potentially the stock price by reducing depreciation expenses and other project start-up costs. Since unlisted firms do not face market pressure to meet earnings targets, we expect that unlisted firms will invest more than comparable listed firms as there would not no incentive for them to forego value creating projects. An influential paper by Asker et al. (2015) shows that in the United States short-termism distorts investment behavior of public listed firms. They find that public firms invest less and are less responsive to investment opportunities when compared to private firms.

Agency theory provides insights on the expected relationships between investments by unlisted vs. listed firms. It is generally assumed that agency conflicts are lower for private firms comparted to their public counterparts (Jensen, 1989). Bhide (1993) argues that highly concentrated ownership and illiquidity incentivizes owners of private firms to monitor management. These arguments suggest that private firms are subject to less agency costs and therefore should invest more in most situations. Note that we isolate this effect in our experiment because we compare investment between listed and unlisted firms. They are both public firms and we can control the different ownership structure. Nonetheless, these arguments suggest that listed firms are subject to myopic pressures and therefore should invest less in most situations. Additionally, Boot et al. (2008) argue that listing creates uncertainty in ownership exposing management to uncertainty regarding shareholder intervention. This uncertainty may also impact managerial investment behavior in listed firms.

While one school of thought postulates that market pressures induces short-termism of publicly traded firms, another line of argumentation suggests that markets monitor management and ensures that pubic firms invest optimally. Evidence of the positive impact of market monitoring exists. Mangena and Tauringana (2007) find positive relationships between corporate governance, foreign ownership, and liquidity (a proxy for market monitoring) in Southern Africa. In an empirical country study of China, Tang and Wang (2011) examine the cross-sectional relation between corporate governance and firm liquidity. They find strong evidence of the positive governance-liquidity relationship. Their findings imply that increased market monitoring improves governance and valuation. Given this evidence, there are reasons to believe that the market monitoring of public firms may prod managers' act in the best interests of shareholder. If these arguments are correct we may uncover a positive relationship between listed firms and investment levels as listed firms are subject to market monitoring and therefore could underinvest. Furthermore, listing reduces the cost of obtaining funds by broadening the investor base.

Based on these conflicting arguments, we construct the following hypothesis.

*H1: Listed firms engage in more investment than unlisted firms.* 

Does firms structure impact the costs and benefits of listing? Agency theory again can provide some guidance. Like Ikeda et al (2017), we consider Hicks' (1935) quite life hypothesis as extended by Bertrand and Mullainathan (2003). According to the Bertrand and Mullainathan's model managers of firms who are protected from hostile takeovers or pressure from unfriendly shareholders are subject to more agency conflicts. Hence managers of these firms will prefer the 'quite life' and invest less than those firms which are subject to takeover threats and/or market monitoring. Firms that are members of a business group are more difficult to acquire than standalone firms. Therefore, we expect that publicly traded standalone firms are subject to a greater probability of takeover, which should give managers an incentive to invest more to avoid unfriendly takeovers. On the other hand, unlisted standalone firms do not face takeover threats and hence may not invest as much as corresponding listed standalone firms.

Another reason to expect listed standalone firms to invest more than unlisted standalone enterprises is due to financial constraints and lack of internal capital markets (Fazzari, Hubbard, and Petersen, 1988). We suspect that these reasons could explain some of the differences between investment among business groups and standalone firms.

Based on these conflicting arguments, we construct the following hypotheses.

*H2:* The impact of listing status on investment behavior is more important for standalone firms relative to business group firms.

# 4. Sample and Methodologies

#### 4.1 Data and Sample Selection

Our sample consists of listed and unlisted firms reporting their financial statements under Japanese accounting standard from March 2000 through April 2017. Since several variables are used in their lagged form, the observations from March 2000 through February 2001 are excluded from the main analyses. We exclude financial firms (Nikkei Medium Classification Industry Code 47-52), and winsorize each variable falling in the top or bottom 1%. All the data are obtained from Nikkei NEEDS Financial Quest 2.0 (FQ, hereafter).

To specify unlisted firms, we use the following procedure. First, we collect the entire financial statement data and the firm's security exchange ID from FQ. We identify firms without a security exchange ID as unlisted firms. Second, we exclude the firms without information on ownership and firms without cash flow statements. Several unlisted firms report financial statement based on JCA, but not J FIEA. These unlisted firms neither disclose J 10-K form nor cashflow statements. Another important difference is that firms are only required to report consolidated financial statements even if a firm has subsidiaries. To control for these differences in disclosure requirements, we specify firms reporting financial statement based on JCA and we exclude those firms. We describe the details and legal framework on financial statement disclosure in Appendix A.

#### 4.2 Empirical Methodologies

To test the effects of being listed on investment behavior, we follow Asker et al (2015):

$$investment_{it} = \alpha_l \ listed_{it} + \Gamma z + fe + \varepsilon_{it} \tag{1}$$

where the dependent variable is corporate investment (*investment*). We use the following four measures of investment: 1)  $\Delta ppe$  computed as the growth of property plant and equipment plus depreciation and impairment from the prior period, 2) *capex* is capital expenditure reported in

footnote of Form 10-K, 3) tang + int is the cash outflow from the purchase of tangible and intangible assets, and 4) capex + rd is the sum of capital expenditure and R&D expenditure. All investment variables are scaled by beginning-of year sum of tangible and intangible assets.

The key explanatory variable is an indicator taking a value of one for a listed firm (*listed*) and zero otherwise. We include several control variables (*z*), industry and year fixed effects (*fe*). The control variables include return on assets (*roa*) and firm age (*age*). We also consider and include several ownership structure variables: shareholding of financial institutions (*sh.financial*); of foreign investors (*sh.foreign*); of top 10 shareholders (*sh.top10*); and of board members (*sh.directors*). We expect positive coefficients on *roa* as better performing firms invest at greater levels (Fazzari et al. 1988)<sup>8</sup>. We predict the negative coefficients of *age* based on business life-cycle hypothesis. Details on the variable definitions are summarized in Table A1. The subscripts *i*, *t*, and *I* depicts firm *i*, year *t*, and industry *I*, respectively. Standard errors are clustered at the firm level.

To investigate the differences in the effects of listing status between business groups and standalone firms, we decompose the business group subsample into two components: 1) Firms that are members of business groups (Business Group), and 2) standalone firms (Standalone) and reestimate Equation (1) for these two sub-sample with all controls.

## 5. Empirical Results

### **5.1 Descriptive Statistics**

Table 1 presents the summary statistics. Descriptive statistics are reported separately for the entire sample, listed firms, and unlisted firms. These univariate results demonstrate significant differences among listed and unlisted firms. Univariate comparisons of investment proxies show

<sup>&</sup>lt;sup>8</sup> Some prior works cast doubt on this interpretation on investment sensitivity to performance (e.g., Kaplan and Zingales 1997). Also, Bushman, Smith and Zhang (2012) show that the positive relationship between investment and performance is mechanically observed.

that for all four measures of firm investment that firms with subsidiaries have higher levels of investment. Mean and median values for listed firms demonstrate significantly greater level of investment suggesting that unlisted firms are potentially capital constrained. Several firm characteristics are significantly different among listed firms and unlisted firms. Listed firms have more growth opportunities and are more profitable. Additionally, listed firms are younger and larger. While, unlisted firms have less cash and have more leverage.

Regarding ownership structure, listed firms have greater ownership percentages by financial institutions and foreign investors. Listed firms have higher ownership by top 10 shareholders which could include financial institutions and foreign investors. Interestingly, listed firms also have higher ownership levels by board members. In general, summary statistics demonstrate significant differences between levels of investment and firm characteristics between listed and unlisted firms.

# Table 1

Table 2 presents summary statistics separately for business groups (i.e., firms with subsidiaries) and standalone firms. Univariate comparisons of investment proxies show that for most measures of firm investment (except  $\Delta ppe$ ) firms with subsidiaries have lower levels of investment than standalone firms. Standalone firms have more investment opportunities, are more profitable, younger, and smaller. Business group firms have less cash and higher leverage. Lower leverage for standalone firms might indicate face financial constraints due to limited borrowing capacity. Regarding ownership structure, financial institutions tend to hold higher percentages of ownership in business group firms, while director ownership appears to be more important in standalone firms.

# Table 2

#### **5.2 Regression Results**

Table 3 reports the results of equation (1) for various firm level investment proxies separately for business group firms and standalone firms to understand the impact of subsidiaries on investment by listed and unlisted firms. The coefficient of *listed* is positive and statistically significant across all investment measures, indicating that listed firms invest more than unlisted firms. Asker et al. (2015) show that short-term pressures cause public firms to invest less than comparable private firms. Our findings suggest that the role of listing in alleviating financial constraints is more important than potential underinvestment due to myopic managerial behavior. We note that this result is in contrast to Orihara (2017) who confirms the same qualitative results of Asker et al. (2015) in Japanese counterparts. Our results indicate that simple comparisons between public and private firms might be misleading when examining the effects of listing status on firm investment behavior.

The coefficient of  $pred_q$  and roa are positive and statistically significant, suggesting that firms with better investment opportunities and performance invest more. In contrast, the coefficient of age and size are negative, indicating that both firm age and firm size have negative impact on investment. These results are generally consistent with existing literature. Regarding the coefficients that are related to liquidity and financial constraints, our results show that firms with more cash or less leverage engage in more investment.

Taking advantage of our unique data on ownership structure, we report the impact of ownership structure on firm level investment. Higher levels of financial institution, foreign, and director ownership positively impacts investment. We suspect that that firms with stable ownership such as cross-shareholding in Japanese context tend to avoid making difficult investment decisions when they are protected from the disciplinary effects of capital markets. This is consistent with the 'quiet life hypothesis' by Ikeda et al (2017).

# Table 3

In the next phase of our analysis, we estimate equation 1 separately for firms that a member of a business group and standalone firms. Table 4 presents the results. Columns 1 to 4 show the results for business groups. The coefficients on *listed* are insignificant, indicating that there is no significant difference in investment levels between listed and unlisted firms which are members of business groups. We will explore more on this in later section. We note that coefficients of firm characteristics are consistent with our earlier analysis with the entire sample.

Columns 5 to 8 contain the results for the standalone subsample of firms. The coefficients on *listed* are significant across all investment proxies, suggesting that listed standalone firms invest at higher levels than their unlisted standalone counterparts. We suspect that this result stems from financial constraints faced by unlisted standalone firms. Unlisted standalone firms have limited access to capital, relative to listed firms. Listing relaxes this financial constraint and allows for greater investment by listed standalone firms. Columns 5 and 6 show significant positive coefficients on percentage owned by financial institutions. Financial constraints faced by standalone firms is mitigated when financial institutions own a higher percentage of the firm. Additionally, shareholding of foreign investors has positive impact on investment although it is marginally significant. The result may stem from external governance mechanisms encouraging greater investment.

# Table 4

To further investigate the structure of business groups on investment, we include *ln\_subs*, and an interaction term between *listed* and *ln\_subs*, where *ln\_subs* is the natural log of the number of subsidiaries of each business groups. Table 5 contains these estimations. The coefficients on *listed* are positive and statistically significant, indicating that listing status has positive impact on

investment once the number of subsidiaries is controlled. Interestingly, we find that the interaction terms of *listed* with  $ln\_subs$  are negative and statistically significant. This suggests that as business groups get larger, management becomes more sheltered from market discipline and investment declines consistent with the enjoying the quiet life hypothesis. In contrast, the coefficient on  $ln\_subs$  is positive and significant, suggesting that unlisted business groups invest more as the number of subsidiaries increases.

# Table 5

### 6. Extensions

#### 6.1. Selection bias

In this section, we conduct robustness checks and extent our analysis. Since listing is a choice by managers, our analyses might reflect selection bias. Pagano et al. (1998) argue that the multiple determinants of IPO might simultaneously affect the decision to go public and corporate investment decisions. To ensure that our results are robust to this type of selection bias, we use matching strategies and Heckman's Treatment Effect Model (TEM).

Following Asker et al. (2015) and Acharya and Xu (2017), we identify a matched unlisted firm for each listed firm. We estimate a propensity score matching by using firm size (*size*) for each industry and year. Caliper-based nearest neighborhood matching is used to identify an unlisted firm for each listed firm. We employ nearest-neighbor matching and drop observations with propensity scores outside the common support to ensure high match quality.

We also conduct several alternative matching procedures. First, to control for other corporate fundamentals, we use leverage (*lev*), cash holding (*cash*), and sales growth (*sg*). Second, we consider whether a firm belongs to a business group. This choice stem from the idea that listing status between business group firms and standalone firms may be driven by the choice to form a business group. By identifying a corresponding matched sample for each industry-year-business group, we control for the observable effects between business group firms and standalone firms. In addition, by using the TEM approach, we control for unobservable difference between listed and unlisted firms. Following Acharya and Xu (2017), we estimated inverse-mill's ratio which corrects for selection bias and is estimated by the following equation<sup>9</sup>:

$$Pr(listed_{it} = 1) = F(\phi_l \ln_{sales_{it-1}} + \phi_2 sg_{it-1} + \phi_3 roa_{it-1} + \phi_4 lev_{it-1} + \varepsilon_{it})$$
(3)

We add the inverse Mill's ratio (*mills*) in the right-hand side of equations (1) and (2). Table A3 represents the results of the first stage model (3). The regression model is estimated separately for all firms, business group firms, and standalone firms.

Table 6 present the results using the matched sample in columns 1 to 3, the matched sample with inverse mills ratio in columns 4 to 6, and the matched sub-sample for standalone firms and business group firms in columns 7 to 9, respectively. The coefficients on *listed* are positive and significant, confirming that listing status has a positive impact on investment. Additionally, note that the coefficient of *listed* in column 7 is now positive and statistically significant at the 10 percent level. Listing status has positive impact on investment even for business groups with matched samples are employed. However, the impact of listing in column 8 is small when compared with that of column 7. We conduct t-tests for differences in coefficients and reject the null hypothesis that the coefficients have the same impact. The results are consistent with those in Table 5, and column 9 confirms the results in Table 5. Overall, the results do not change when the Mill's

<sup>&</sup>lt;sup>9</sup> Industry q is another potentially important determinant (Pagano et al. 1998). When including industry q in the first stage model and note that the results do not change.

ratio is included in the estimations, suggesting that the main findings in Section 5 are robust to the selection bias on being listed<sup>10</sup>.

# Table 6

#### 6.2. Investment Sensitivity and Financial Constraint

Our results suggest that listing status enhances corporate investment, especially for standalone firms, through market monitoring and by reducing financial constraints. However, an alternative explanation suggests that agency problems within listed firms produces overinvestment in the standalone sub-sample. To test our main results are consistent with our expectation, we run two additional tests.

First, we add interaction terms of indicator variable of listed firm (*listed*) with two proxies of investment opportunities (*pred\_q*) and one proxy of financial constraint(*roa*)<sup>11</sup> in the right had side of equation (1). Equation (4) presents the empirical specification:

$$investment_{it} = \gamma_1 \ listed_{it} + \gamma_1 \ listed_{it} \times pred\_q_{it} + \gamma_1 \ listed_{it} \times roa_{it} + \Gamma z + fe + \varepsilon_{it}$$
(4)

Second, we examine the impact of listing status and financial constraints. We expect that listed firms invest more than unlisted counterpart when firms are faced with financial constraints.

<sup>&</sup>lt;sup>10</sup> We also examine cases where firms are consistently listed or unlisted across our sample years (i.e., balanced panel data) and control for IPO firms and delisted firms. We confirmed that qualitative results do not change and are available upon request.

<sup>&</sup>lt;sup>11</sup> We note that there are some debates on the appropriateness of the proxy of financial constraints. For example, Asker argues that "prior work shows that standard proxies for investment opportunities are not, as neoclassical theory predicts, a sufficient statistic for investment and that ROA correlates positively with investment. The latter is often interpreted as a sign of financing constraints (Fazzari, Hubbard, and Petersen 1988), though some disagree (Kaplan, and Zingales 1997)." Although we are uncertain of the debate on the interpretation, we follow the previous literature by using in the interaction term.

To test this prediction, we add financial constraint proxies (*constraint*) and its interaction with listing status indicator (*listed*). Equation (5) presents the empirical specification:

$$investment_{it} = \delta_1 \ listed_{it} + \delta_2 \ listed_{it} \times constraint_{it} + \delta_3 \ constraint_{it} + \Gamma z + fe + \varepsilon_{it}$$
(5)

where *constraint* is a proxy of financial constraint. Following long research stream on financial constraint, we use four proxies for financial constraints following existing literature: no payout indicator (*no\_payout*), no bond access indicator (*no\_bacc*), small firms (*small*), and Hadlock-Pierce index (*hp*) (Fazzari et al., 1988; Hadlock and Pierce, 2010; Farre-Mensa and Ljungqvist, 2016). For firm size (Hadlock-Pierce index), we consider the first (fifth) quartile of each measurement as the financial constraint firms.

Panel A of Table 7 presents the result of sensitivity to investment opportunities for standalone firms. The coefficients on *listed* suggest that unlisted standalone firms are less sensitive to changes in investment opportunities. In contrast, the coefficient of cross-term of listed with *pred\_q* is positive and statistically significant except in column 4, suggesting that listed standalone firms have greater access to lower cost capital and hence can be nimbler in their investment decisions. We also show that unlisted firms' investment is less sensitive to ROA than unlisted standalone firms (columns 3 and 4). These results suggest that unlisted standalone firms face financing constraints and cannot increase investment in response to opportunities as much as their listed counterparts.

Panel B presents the results of impact of financial constraints on investment. Except in column 2 (bond access), the coefficients on the cross-terms are positive and significant, confirming that more financially constrained listed firms tend to be more sensitive to investment opportunities. We also conduct Chow-test between financial constrained and unconstrained firms. The results are

contained in Appendix A2 and show that all the coefficients for financial constrained firms are positive and significant. In contrast, all the coefficients for unconstrained firms are insignificant. These results indicate that listing may alleviate financially constrained firms and facilitate investment.

# Table7

#### **6.3.** Ownership Structure

Taking advantage of our unique data on ownership structure of unlisted firms, we investigate the effects of ownership structure on the positive impacts of listing on investment in Table 8. We show that higher percentage ownership by financial institution or top 10 shareholders (i.e., stable ownership) reduces the positive impact of listing status on investment levels. In contrast, higher levels of stock ownership by foreign investors has a more positive impact on investment. These results suggest that foreign ownership intensify the market pressure of being listed while financial institutions or large stable ownership tends to protect management from the discipline of financial markets allowing them to enjoy a quieter life.

# Table 8

## 6.4. Listing Status and Liquidity

We re-estimate equation (2) replacing the listing dummy variable with proxies of stock liquidity. Maug (1998) derives and model that suggest that more liquid equity markets support better corporate governance in equilibrium. Support for Maug's argument is provided in several empirical studies. Using a sample of U.S. firms, Chung, Elder, and Kim (2010) construct a corporate governance index and examine the impact of corporate governance on share liquidity. Their results indicate that time-varying liquidity, measured by spreads and price impact is explained by their time-varying corporate governance index. They argue more market monitoring helps reduce information asymmetry between insiders and outside investors. Admati and Pfleiderer (2009) maintain that liquidity may help discipline management, mitigate agency problems, and thus improve firm performance. Khanna and Sonti (2004) assert that liquidity simulates the entry of informed traders who make prices more informative to other shareholders, thereby improving firms' operating performance and stock prices. Fang, Noe, and Tice (2009) find empirical support for this argument. Amihud and Levi (2019) show that stock liquidity enhances corporate investment through decreasing cost of equity.

We use *liquidity* as a proxy of each firm's stock liquidity that proportionally reflects the impacts of listing status. *Liquidity* is the negative value of Amihud's (2002) illiquidity measure and is defined in Appendix A1. The coefficients of these variables only exist for listed firms. Thus, the coefficients of liquidity variables are the same as those of cross-term of *listed* with these liquidity variables. Column (1) of Table 9 presents the positively significant relation between liquidity and corporate investment. In Columns (2) and (3), we decompose the sample into two subsamples, business group firms and standalone firms. The coefficient on *liquidity* is positively significant only for business group firms, but insignificant for standalone firms. The results imply that stock liquidity encourages more efficient investment by increasing the market monitoring of management, which helps to overcome managerial shirking. Furthermore, higher levels of liquidity make more it easier to raise additional capital which may also help to reduce financial constraints.

# Table 9

## 7. Conclusion

What is the impact of listing on corporate investment? Using an extensive database on Japanese listed and unlisted firms over various market cycles (2001-2017), we contribute to a small

but growing literature on public vs. private investment patterns. Our unique approach allows for a nuanced understanding of listed and unlisted firm investment without confounding effects due to different disclosure requirements.

We find that listed firms invest more than their unlisted counterparts. This result suggests that the role of listing in alleviating financial constraints is more important than potential underinvestment due to short-term market pressure in Japan. Our analysis reveals that the positive relationship between listing and investment is primarily driven by standalone firms. The positive relationship between listing and investment is weak for firms which are members of business groups. Furthermore, as the number of subsidiaries in a business group increases the positive impact of listing on investment declines. In contrast, standalone firms invest more and are more sensitive to investment opportunities than unlisted firms. We also find that the positive relationship between listing and investment is greater for financially constrained firms. We also examine the role of stock liquidity on listed firm investment behavior. We show that liquidity enhances the positive relationship between listing and investment, particularly in the business group subsample. This finding suggests that liquid stocks act as a monitor of management reducing shirking and producing higher levels of investment. Additionally, liquid stocks make raising additional capital easier, which also facilitates investment. In most cases, we show that higher ownership by financial institutions and foreign investors increases investment, potentially mitigating the financial constraints (in the case of financial institutions) or acting as monitors of management (in the case of foreign investors), which reduces shirking and thereby increases investment. Overall, our results have implications for understanding corporate investment throughout the world, give the prevalence of business groups in many developed and developing economies.

Our findings have several important implications. First, we demonstrate that the investment decision between listed and unlisted firms is a tradeoff between the benefits of lower cost capital from public financing and short-term market pressures that can induce myopic behavior. Second, we show that business structure matters for understanding corporate investment patterns. Future research may explore the role of internal capital markets on investment in business groups. For example, an alternative interpretation of our results for business groups is that unlisted firms that are parents of business groups are not capital constrained as they can access internal capital markets. Therefore, they tend to invest at the same level as similar listed firms who are exposed to short-term market pressures. This direction may be an important next step in understanding the investment behavior of listed and unlisted business groups firms.

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## **Figures and Tables**



## Figure 1 Japanese legal framework on financial reporting:

This diagram describes the classification across listed, unlisted (quasi-private), and purely private firms. *X* axis presents the strictness of mandatory disclosure required by laws. In general, the disclosure requirements for listed firms are stricter than those for purely private firms. The reference point dividing purely-private, and listed and quasi-private firms represents the minimum requirements of disclosure for listed firms. Quasi-private firms do not go public, but Japanese Financial Instrument Exchange Act requires these firms the same disclosure as listed firms, i.e.: they are required to disclose information at the same strict level as the listed firms.

## Table 1 Descriptive Statistics (Listed v.s. Unlisted):

This table presents the descriptive statistics on all variables used in main analyses. Columns (1) and (2) report the descriptive statistics of listed and unlisted firms, respectively. Columns (3) report the difference in each variable between listed and unlisted firms. \*\*\* indicates significance at the 1% level using a two-tailed test. All observations falling in the top or bottom 1 % with respect to each variable are winsorized. All variables are defined in Table A1.

	(1) Listed firms $(n=20.046)$			(2) Unlisted firms $(n-2, 182)$			(3) Listed -			
	(II-39,940) mean	median	sd	mean	p50	sd	mean		median	
∆ppe	0.1447	0.0922	0.2198	0.0851	0.0417	0.1800	0.0597	***	0.0505	***
capex	0.1508	0.0947	0.1968	0.0786	0.0345	0.1398	0.0722	***	0.0603	***
tan+int	0.1625	0.0994	0.2246	0.0975	0.0466	0.1853	0.0650	***	0.0528	***
capex+rd	0.2244	0.1365	0.3315	0.1151	0.0406	0.2729	0.1093	***	0.0960	***
pred_q	1.0867	0.9816	0.4670	1.0310	0.9409	0.3700	0.0556	***	0.0408	***
roa	0.4668	0.2363	1.1183	0.1947	0.1074	0.7109	0.2721	***	0.1289	***
age	3.8043	3.9890	0.5951	3.9570	4.0775	0.5441	-0.1527	***	-0.0886	***
size	10.3786	10.2507	1.5213	9.5776	9.7393	1.6732	0.8010	***	0.5114	***
cash	1.6230	0.4987	4.1899	0.9392	0.2472	3.2736	0.6838	***	0.2515	***
lev	0.2115	0.1773	0.1857	0.2893	0.2757	0.2409	-0.0778	***	-0.0983	***
sh_financial	0.1861	0.1627	0.1307	0.0740	0.0451	0.0873	0.1121	***	0.1176	***
sh_foreign	0.0766	0.0305	0.1028	0.0116	0.0000	0.0620	0.0650	***	0.0305	***
sh_top10	0.5081	0.5010	0.1606	0.4574	0.4869	0.2837	0.0507	***	0.0142	***
sh_directors	0.0975	0.0310	0.1363	0.0648	0.0134	0.1120	0.0327	***	0.0177	***

## Table 2 Descriptive Statistics (Business group firms vs. Standalone firms):

This table presents the descriptive statistics on all variables in main analysis, comparing the statistics between business group firms and standalone firms. Columns (1) and (2) report the descriptive statistics of business group firms and standalone firms, respectively. Columns (3) report the difference in each variable between group and standalone firms. \*, \*\*, \*\*\* indicate significance at the 10, 5, 1% levels using a two-tailed test, respectively. All observations falling in the top or bottom 1 % with respect to each variable are winsorized. All variables are defined in Table A1.

All										
	(1) Business						(3) Business			
	Group			(2) Standalone			Group -			
	(n=35,819)			(n=6,310)			Standalone			
	mean	median	sd	mean	median	sd	mean		median	
∆ppe	0.1425	0.0939	0.2135	0.1371	0.0607	0.2437	0.0054	*	0.0332	***
capex	0.1449	0.0951	0.1833	0.1596	0.0674	0.2506	-0.0148	***	0.0277	***
tan+int	0.1557	0.0994	0.2094	0.1786	0.0761	0.2886	-0.0230	***	0.0233	***
capex+rd	0.2128	0.1366	0.2978	0.2529	0.0974	0.4698	-0.0401	***	0.0392	***
pred_q	1.0784	0.9768	0.4517	1.1144	0.9951	0.5196	-0.0360	***	-0.0183	***
roa	0.4236	0.2323	0.9825	0.6175	0.2064	1.6141	-0.1939	***	0.0259	***
age	3.8440	4.0254	0.5903	3.6317	3.7612	0.5796	0.2123	***	0.2642	***
size	10.5677	10.4514	1.4906	9.0278	9.0491	1.0959	1.5399	***	1.4023	***
cash	1.3927	0.4687	3.6064	2.6935	0.5809	6.3035	-1.3008	***	-0.1122	***
lev	0.2212	0.1910	0.1869	0.1830	0.1091	0.2019	0.0382	***	0.0819	***
sh_financial	0.1955	0.1746	0.1323	0.0941	0.0780	0.0826	0.1013	***	0.0967	***
sh_foreign	0.0795	0.0334	0.1039	0.0376	0.0046	0.0825	0.0420	***	0.0288	***
sh_top10	0.4973	0.4899	0.1615	0.5521	0.5685	0.2035	-0.0548	***	-0.0786	***
sh_directors	0.0873	0.0245	0.1286	0.1436	0.0762	0.1605	-0.0563	***	-0.0517	***

#### Table 3 Investment level: Listed firms vs. Unlisted firms:

This table presents the results of regression model (1) by regressing investment on listed status dummy (*listed*) and other control variables. Control variables include predicted q (*pred\_q*), return on assets (*roa*), firm age (*age*), firm size (*size*), cash holding (*cash*), leverage (*lev*), and shareholding of financial institutions (*sh\_financial*), foreign investors (*sh\_foreign*), top 10 investors (*sh\_top10*), and board members (*sh\_directors*). We also control for year and industry fixed effects. Standard errors in parentheses are calculated clustered by the firm level. \*, \*\*, \*\*\* indicate significance at the 10, 5, 1% levels, respectively, using a two-tailed test. All variables are defined in Table A1.

	∆ppe	capex	tan+int	capex+rd
	(1)	(2)	(3)	(4)
listed	0.0217***	0.0359***	0.0254***	0.0244*
	(3.33)	(5.54)	(3.56)	(1.94)
pred q	0.0774***	0.0794***	0.0801***	0.0845***
1 –1	(12.88)	(13.98)	(12.12)	(9.07)
Roa	0.0156***	0.0118***	0.0116***	-0.0011
	(4.17)	(3.23)	(2.68)	(-0.12)
Age	-0.0467***	-0.0519***	-0.0579***	-0.0903***
	(-11.47)	(-11.79)	(-12.64)	(-10.14)
Size	-0.0128***	-0.0127***	-0.0187***	-0.0138***
	(-7.51)	(-7.52)	(-9.75)	(-4.33)
Cash	0.0095***	0.0053***	0.0155***	0.0313***
	(9.16)	(4.63)	(12.95)	(10.12)
Lev	-0.0473***	-0.0545***	-0.0482***	-0.0851***
	(-5.57)	(-6.25)	(-5.33)	(-5.54)
sh_financial	0.1131***	0.1279***	0.1382***	0.1488***
	(7.41)	(8.21)	(8.94)	(5.61)
sh_foreign	0.1193***	0.0920***	0.1341***	0.1096***
	(6.03)	(4.83)	(6.35)	(2.85)
sh_top10	0.0168*	0.0047	-0.0010	-0.0295
	(1.69)	(0.48)	(-0.09)	(-1.64)
sh_directors	0.0341**	0.0555***	0.0419**	0.0602*
	(2.15)	(3.09)	(2.30)	(1.79)
Observations	42,129	42,129	42,129	42,129
Year fixed effects	yes	yes	yes	Yes
Industry fixed effects	yes	yes	yes	Yes
clustered by	firm	firm	firm	Firm
Adj. R <sup>2</sup>	0.190	0.193	0.294	0.341

### Table 4 Investment level: Business group firms vs. Standalone firms:

This table presents the sub-sample results for Business group firms and standalone firms by regressing investment on listing indicator (*listed*) and other control variables (model (1)). While columns (1) - (4) represent the results using subsample of business group firms, Columns (5) - (8) represent the results using subsample of standalone firms. Control variables include predicted q (*pred\_q*), return on assets (*roa*), firm age (*age*), firm size (*size*), cash holding (*cash*), leverage (*lev*), and shareholding of financial institutions (*sh\_financial*), foreign investors (*sh\_foreign*), top 10 investors (*sh\_top10*), and board members (*sh\_directors*). We also control for year and industry fixed effects. t-statistics in parentheses are calculated based on standard errors obtained by clustering at the firm level. \*, \*\*, \*\*\* indicate significance at the 10, 5, 1% levels, respectively, using a two-tailed test. All variables are defined in Table A1.

	Business Group				Standalone			
	∆ppe	capex	tan+int	capex+rd	∆рре	capex	tan+int	capex+rd
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
listed	-0.0037	0.0023	-0.0061	-0.0027	0.0361**	0.0980 * * *	0.0870***	$0.0789^{**}$
	(-0.55)	(0.38)	(-0.89)	(-0.23)	(2.32)	(5.85)	(4.47)	(2.50)
pred_q	0.0752***	0.0756***	0.0768***	0.0831***	0.0899***	0.1002***	$0.0969^{***}$	0.0908 * * *
	(11.46)	(12.36)	(11.33)	(8.45)	(5.79)	(6.67)	(5.13)	(3.43)
Roa	0.0216***	0.0113**	0.0132**	0.0022	0.0002	0.0094	0.0064	-0.0074
	(4.89)	(2.55)	(2.57)	(0.22)	(0.03)	(1.54)	(0.86)	(-0.46)
Age	-0.0405***	-0.0432***	-0.0490***	-0.0817***	-0.0639***	-0.0894***	-0.0897***	-0.1266***
	(-9.79)	(-9.79)	(-10.82)	(-8.98)	(-5.45)	(-6.77)	(-6.50)	(-5.05)
Size	-0.0166***	-0.0137***	-0.0190***	-0.0159***	-0.0064	-0.0156***	-0.0340***	-0.0169
	(-9.55)	(-7.57)	(-9.64)	(-5.28)	(-1.13)	(-2.88)	(-5.02)	(-1.21)
Cash	0.0108***	0.0061***	0.0172***	0.0271***	0.0076***	0.0029*	0.0117***	0.0371***
	(8.18)	(3.96)	(11.34)	(7.66)	(4.79)	(1.94)	(6.38)	(7.30)
Lev	-0.0523***	-0.0559***	-0.0502***	-0.0996***	-0.0363*	-0.0741***	-0.0727***	-0.0797**
	(-5.58)	(-6.00)	(-5.13)	(-6.38)	(-1.92)	(-3.44)	(-3.28)	(-2.03)
sh_financial	0.1162***	0.1249***	0.1286***	0.1447***	0.0867	0.1612**	0.1423**	0.0494
•	(7.65)	(8.14)	(8.42)	(5.73)	(1.37)	(2.54)	(2.35)	(0.47)
sh_foreign	0.1238***	0.0947***	0.1292***	0.1360***	0.1051	0.1221*	0.1770**	-0.0161
- 0	(6.17)	(4.82)	(5.97)	(3.39)	(1.60)	(1.90)	(2.53)	(-0.12)
sh_top10	0.0046	-0.0059	-0.0152	-0.0389**	0.0257	0.0113	0.0367	-0.0155
_ *	(0.42)	(-0.56)	(-1.28)	(-2.16)	(1.13)	(0.47)	(1.45)	(-0.33)
sh_directors	0.0375**	0.0741***	0.0549***	0.0586*	-0.0058	-0.0179	-0.0223	0.0522
	(2.11)	(3.69)	(2.72)	(1.69)	(-0.18)	(-0.47)	(-0.59)	(0.68)
Observations	35,819	35,819	35,819	35,819	6,310	6,310	6,310	6,310
Year fixed effects	yes	yes	yes	yes	yes	yes	yes	yes
Industry fixed effects	yes	yes	yes	yes	yes	yes	yes	yes
clustered by	firm	firm	firm	firm	firm	firm	firm	firm
Adj. R <sup>2</sup>	0.204	0.189	0.300	0.307	0.188	0.230	0.302	0.439

#### Table 5 Business group structure and the size effects of listing status

This table presents the results with an interaction term between listing indicator (*listed*) and the size of business group ( $ln\_subs$ ) by using business group subsample. We regress investment on listed firm indicator (*listed*), its interaction with the logarithm of the number of subsidiaries ( $listed \times ln\_subs$ ) and other control variables. Control variables include the number of subsidiaries ( $ln\_subs$ ), predicted q ( $pred\_q$ ), return on assets (roa), firm age (age), firm size (size), cash holding (cash), leverage (lev), and shareholding of financial institutions ( $sh\_financial$ ), foreign investors ( $sh\_foreign$ ), top 10 investors ( $sh\_top10$ ), and board members ( $sh\_directors$ ). We also control for year and industry fixed effects. t-statistics in parentheses are calculated based on standard errors obtained by clustering at the firm level. \*, \*\*, \*\*\*\* indicate significance at the 10, 5, 1% levels, respectively, using a two-tailed test. All variables are defined in Table A1.

_	Business Group			
_	∆ppe	capex	tan+int	capex+rd
	(1)	(2)	(3)	(4)
listed	0.0379**	0.0422***	0.0289*	0.0537*
	(2.23)	(2.84)	(1.70)	(1.92)
listed×ln_subs	-0.0194***	-0.0187***	-0.0161**	-0.0263**
	(-2.59)	(-2.75)	(-2.27)	(-2.09)
ln_subs	0.0269***	0.0246***	0.0298***	0.0366***
	(3.56)	(3.51)	(4.14)	(2.79)
pred_q	0.0744***	0.0749***	0.0752***	0.0819***
	(11.29)	(12.22)	(11.03)	(8.29)
roa	0.0220***	0.0116***	0.0138***	0.0026
	(5.00)	(2.61)	(2.72)	(0.26)
age	-0.0398***	-0.0426***	-0.0478***	-0.0807***
	(-9.65)	(-9.68)	(-10.58)	(-8.89)
size	-0.0207***	-0.0169***	-0.0260***	-0.0214***
	(-9.32)	(-7.29)	(-10.25)	(-5.41)
cash	0.0108***	0.0060***	0.0171***	0.0271***
	(8.22)	(3.95)	(11.44)	(7.65)
lev	-0.0564***	-0.0591***	-0.0582***	-0.1053***
	(-5.98)	(-6.34)	(-5.91)	(-6.69)
sh_financial	0.1143***	0.1237***	0.1229***	0.1420***
	(7.46)	(7.98)	(8.02)	(5.56)
sh_foreign	0.1206***	0.0926***	0.1206***	0.1315***
	(6.01)	(4.69)	(5.63)	(3.32)
sh_top10	0.0075	-0.0036	-0.0100	-0.0348*
	(0.69)	(-0.34)	(-0.85)	(-1.95)
sh_directors	0.0363**	0.0731***	0.0527***	0.0569*
	(2.04)	(3.64)	(2.63)	(1.65)
Observations	35,819	35,819	35,819	35,819
Year fixed effects	yes	yes	yes	Yes
Industry fixed effects	yes	yes	yes	Yes
clustered by	firm	firm	firm	Firm
Adj. R <sup>2</sup>	0.205	0.190	0.302	0.308

## **Table 6 Robustness tests**

This table reports the matching results in treatment effect model. In all the regressions, control variables include return on assets (*roa*), firm age (*age*), firm size (*ln.tast*), cash holding (*cash*), and shareholding of financial institutions (*sh.financial*), foreign investors (*sh.foreign*), top 10 investors (*sh.top10*), and board members (*sh.directors*). We also control for industry and year fixed effects in all the models. t-statistics in parentheses are calculated based on standard errors obtained by clustering at the firm level. \*, \*\*, \*\*\* indicate significance at the 10, 5, 1% levels, respectively, using a two-tailed test. All variables are defined in Table A1.

				dep	endent variable: c	apex				
		Matching			Matching+TEM		Ν	$\begin{tabular}{ c c c c c c c c c c c c c c c c c c c$		
		year			year					
		+industry	year		+industry	year		year		
		+size	+industry		+size	+industry		+industry		
Matching criteria?	vear	+lev	+Business	vear	+lev	+Business		+Business		
	+industrv	+cash	Group	+industrv	+cash	Group		Group		
	+size	+sg	+size	+size	+sg	+size		+size		
		0			0		Standalana	Business	Business	
							Standalone	Group	Group	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	
listed	0.0657***	0.0209***	0.0449***	0.0572***	0.0247***	0.0250***	0 1 496***	0.0159*	0.0700***	
lisied	$(0.063)^{+++}$	(2.40)	(5.00)	(5.20)	(2.97)	(4.02)	(4.82)	(1.97)	(2.80)	
	(6.01)	(3.40)	(5.06)	(5.20)	(2.87)	(4.05)	(4.82)	(1.87)	(3.80)	
listea×in_subs									-0.0284	
he auto									(-3.34)	
in_subs									(1.96)	
mand a	0.0007***	0.0272	0.0720***	0.0000***	0.0262	0.0729***	0.1001***	0.0677***	(1.00)	
prea_q	(5, 62)	(1.25)	(2.07)	(5.55)	(1.20)	(2.04)	(2.82)	(2,20)	(2, 42)	
<b>D</b> og	(5.03)	(1.55)	(3.97)	(5.55)	(1.30)	(3.94)	(2.82)	(3.39)	(3.42)	
коа	0.0015	$0.0470^{****}$	0.0005	-0.0071	0.0396***	-0.0120	-0.0046	0.0007	-0.0006	
	(0.21)	(3.80)	(0.06)	(-0.85)	(2.99)	(-1.13)	(-0.30)	(0.05)	(-0.04)	
Age	-0.0590***	-0.0513***	-0.0498***	-0.0588***	-0.0528***	-0.049/***	-0.0610***	-0.026/***	-0.026/***	
	(-4.95)	(-4.34)	(-5.31)	(-4.94)	(-4.53)	(-5.37)	(-2.59)	(-3.06)	(-3.02)	
Size	-0.0016	-0.0025	-0.0054	-0.0145**	-0.0121**	-0.0232***	-0.0159	-0.0051	-0.0059	
	(-0.44)	(-0.73)	(-1.46)	(-2.39)	(-2.05)	(-3.63)	(-1.36)	(-1.28)	(-1.15)	
Cash	0.0091***	0.0026	0.0071***	0.0096***	0.0029	0.0081***	0.0057*	0.0058	0.0057	
	(3.85)	(0.98)	(2.59)	(3.99)	(1.09)	(2.82)	(1.65)	(1.37)	(1.34)	
Lev	-0.0493**	-0.0447*	-0.0666***	0.0137	-0.0002	0.0170	-0.0677	-0.0752***	-0.0721***	
	(-2.25)	(-1.92)	(-3.40)	(0.42)	(-0.01)	(0.56)	(-1.50)	(-3.62)	(-3.54)	
sh_financial	0.0352	0.0815*	0.0874*	0.0357	0.0809*	0.0855*	0.1390	0.0778*	0.1005**	
	(0.70)	(1.92)	(1.93)	(0.72)	(1.90)	(1.86)	(0.61)	(1.80)	(2.25)	
sh_foreign	-0.0708	0.0243	0.0095	-0.0408	0.0461	0.0585	0.0144	0.0910*	0.1280**	
	(-1.51)	(0.57)	(0.21)	(-0.82)	(1.00)	(1.21)	(0.09)	(1.85)	(2.45)	
sh_top10	-0.0016	-0.0045	-0.0056	-0.0075	-0.0086	-0.0123	-0.0039	-0.0150	-0.0176	
	(-0.07)	(-0.21)	(-0.28)	(-0.34)	(-0.39)	(-0.61)	(-0.08)	(-0.77)	(-0.90)	
sh_directors	0.0451	0.0403	0.0363	0.0404	0.0375	0.0336	-0.0316	0.0962**	0.0996**	
	(1.00)	(0.90)	(0.85)	(0.89)	(0.84)	(0.79)	(-0.34)	(2.25)	(2.34)	
mills				0.2340***	0.1863**	0.3302***				
				(2.79)	(2.13)	(3.69)				

Observations	4,252	3,435	3,570	4,252	3,435	3,570	755	2,815	2,815
Year fixed effects	yes								
Industry fixed effects	yes								
clustered by	firm								
Adj. R <sup>2</sup>	0.218	0.179	0.178	0.221	0.181	0.184	0.247	0.141	0.146

#### Table 7 Investment opportunities and financial constraints: Standalone firms

This table presents the results of additional analyses.

Panel A shows the results of regression model (4). The dependent variables are the proxies of corporate investment. The variable of interest is the interaction term between listing indicator (*listed*) and predicted q (*pred\_q*). Panel B presents the results of model (4) with the interaction term between listing indicator (*listed*) and several financial constraint proxies. Following prior studies, we use four financial constraint proxies. The first proxy is no payout indicator (*no\_payout*). The second is no bond access indicator (*no\_bacc*). The third is the first quintile of firm size (*small*). The final proxy is the 5<sup>th</sup> quintile of Hadlock-Pierce index (*hp*). In both panel, the control variables include return on assets (*roa*), firm age (*age*), firm size (*ln.tast*), cash holding (*cash*), and shareholding of financial institutions (*sh.financial*), foreign investors (*sh.foreign*), top 10 investors (*sh.top10*), and board members (*sh.directors*). We also control for year and industry fixed effects in all the models. t-statistics in parentheses are calculated based on standard errors obtained by clustering at the firm level. \*, \*\*, \*\*\* indicate significance at the 10, 5, 1% levels, respectively, using a two-tailed test. All variables are defined in Table A1.

	Standalone			
	∆ppe	capex	tan+int	capex+rd
	(1)	(2)	(3)	(4)
listad	-0.0525*	0.0126	-0.0509	0.0805
usieu	(-1.88)	(0.46)	(-1, 24)	(1.35)
listed x pred a	0.0827***	0.788***	0 1355***	0.0183
usieu~preu_q	(4.02)	(3.30)	(4.02)	(0.34)
listedyrad	0.0172	0.0125	0.0603*	0.0974*
usieu×10u	(0.97)	(0.67)	(1.05)	(1.02)
nred a	(-0.97)	(-0.07)	(-1.95)	(-1.92)
preu_q	(0.80)	(1.26)	-0.0212	(1.44)
	(0.89)	(1.30)	(-0.04)	(1.44)
rou	0.0135	(1, 14)	(2.05)	(1.60)
	(0.89)	(1.14)	(2.03)	(1.09)
age	-0.0634***	-0.0880****	-0.0908***	-0.1323****
-1	(-5.38)	(-0./1)	(-0.08)	(-5.28)
size	-0.0066	-0.0159***	-0.0334***	-0.0145
1	(-1.16)	(-2.94)	(-5.06)	(-1.05)
casn	0.0076***	0.0029*	0.0118***	0.03/3***
,	(4.85)	(1.94)	(6.43)	(7.47)
lev	-0.0361*	-0.0/32***	-0.0//4***	-0.0948**
	(-1.90)	(-3.42)	(-3.58)	(-2.49)
sh_financial	0.0798	0.1548**	0.1291**	0.0424
	(1.27)	(2.45)	(2.14)	(0.40)
sh_foreign	0.1036	0.1211*	0.1706**	-0.0277
	(1.56)	(1.88)	(2.46)	(-0.21)
sh_top10	0.0203	0.0067	0.0235	-0.0295
	(0.90)	(0.28)	(0.93)	(-0.63)
sh_directors	-0.0073	-0.0193	-0.0249	0.0514
	(-0.23)	(-0.50)	(-0.66)	(0.67)
Observations	6,310	6,310	6,310	6,310
Year fixed effects	yes	yes	yes	yes
Industry fixed effects	yes	yes	yes	yes
clustered by	firm	firm	firm	firm
Adj. R <sup>2</sup>	0.190	0.232	0.307	0.443

#### Panel A: Sensitivity to investment opportunities

#### **Panel B: Financial constraints**

	Standalone			
	capex			
	(1)	(2)	(3)	(4)
listed	0.0666***	0.0576*	0.0605***	0.0750***
listed×no_payout	0.0568**	(100)	(2172)	(0.00)
no_payout	-0.0370**			
listed×no_bond	(2)	0.0402		
no_bond		-0.0618*		
listed×small		(101)	0.0490**	
small			-0.0586**	
listed×hp				0.1142*** (2.61)
hp				-0.0463 (-1.07)
pred_q	0.1020*** (6.72)	0.1002*** (6.70)	0.1009*** (6.68)	0.0993*** (6.63)
roa	0.0101* (1.68)	0.0098 (1.62)	0.0094 (1.55)	0.0095
age	-0.0889*** (-6.82)	-0.0891***	-0.0898***	-0.0609***
size	-0.0140*** (-2.58)	-0.0169*** (-3.12)	-0.0212*** (-2.79)	-0.0162*** (-2.99)
cash	0.0027*	0.0029*	0.0029*	0.0027*
lev	-0.0812***	-0.0878*** (-4.04)	-0.0740***	-0.0748***
sh_financial	0.1638** (2.58)	0.1686*** (2.68)	0.1692*** (2.67)	0.1863*** (2.91)
sh_foreign	0.1123*	0.1263*	0.1273**	0.0973
sh_top10	0.0079	0.0122	0.0065	0.0228
sh_directors	-0.0160 (-0.41)	-0.0197 (-0.52)	-0.0146 (-0.38)	-0.0330 (-0.87)
Observations	6,310	6,310	6,310	6,310
Year fixed effects	yes	yes	yes	yes
Industry fixed effects	yes	yes	yes	yes
clustered by	firm	firm	firm	firm
Adj. R <sup>2</sup>	0.231	0.232	0.231	0.237

#### Table 8 Effects of ownership structure

This table presents the results of model (1) by regressing capital expenditure (*capex*) on the interaction terms between listing status and several ownership variables. Control variables include return on assets (*roa*), firm age (*age*), firm size (*ln.tast*), cash holding (*cash*), and shareholding of financial institutions (*sh.financial*), foreign investors (*sh.foreign*), top 10 investors (*sh.top10*), and board members (*sh.directors*). We also control for year and industry fixed effects in all the models. t-statistics in parentheses are calculated based on standard errors obtained by clustering at the firm level. \*, \*\*, \*\*\* indicate significance at the 10, 5, 1% levels, respectively, using a two-tailed test. All variables are defined in Table A1.

	Capex				
	(1)	(2)	(3)	(4)	(5)
listed	0.0609***	0.0338***	0.0582***	0.0335***	0.0780***
	(6.64)	(5.10)	(4.75)	(4.50)	(6.16)
listed×sh financial	-0.3123***				-0.2802***
_	(-5.03)				(-4.21)
listed×sh_foreign		0.1577***			0.1605**
~ 0		(2.81)			(2.54)
listed×sh_top10			-0.0462**		-0.0530**
-			(-2.03)		(-1.99)
listed×sh_directors				0.0346	0.0536
				(0.68)	(0.91)
sh_financial	0.4317***	0.1272***	0.1253***	0.1283***	0.3975***
	(6.95)	(8.16)	(8.06)	(8.22)	(5.98)
sh_foreign	0.0983***	-0.0617	0.0927***	0.0922***	-0.0577
	(5.13)	(-1.13)	(4.86)	(4.84)	(-0.93)
sh_top10	0.0008	0.0058	0.0426**	0.0051	0.0463*
	(0.08)	(0.59)	(2.12)	(0.51)	(1.92)
sh_directors	0.0551***	0.0547***	0.0570***	0.0225	0.0048
	(3.07)	(3.05)	(3.18)	(0.46)	(0.09)
pred_q	0.0796***	0.0792***	0.0795***	0.0794***	0.0794 ***
	(14.01)	(13.92)	(14.00)	(13.97)	(13.96)
roa	0.0117***	$0.0118^{***}$	0.0118***	$0.0118^{***}$	0.0117***
	(3.21)	(3.23)	(3.24)	(3.22)	(3.22)
age	-0.0514***	-0.0521***	-0.0518***	-0.0518***	-0.0513***
	(-11.70)	(-11.81)	(-11.78)	(-11.74)	(-11.67)
size	-0.0131***	-0.0127***	-0.0129***	-0.0127***	-0.0133***
	(-7.74)	(-7.55)	(-7.56)	(-7.52)	(-7.82)
cash	0.0052***	0.0053***	0.0052***	0.0053***	0.0052***
	(4.61)	(4.65)	(4.58)	(4.63)	(4.57)
lev	-0.0562***	-0.0541***	-0.0554***	-0.0544***	-0.0565***
	(-6.47)	(-6.21)	(-6.33)	(-6.24)	(-6.48)
Observations	42,129	42,129	42,129	42,129	42,129
Year fixed effects	Yes	yes	yes	yes	yes
Industry fixed effects	Yes	yes	yes	yes	yes
clustered by	Firm	firm	firm	firm	firm
Adj. R <sup>2</sup>	0.194	0.193	0.193	0.193	0.194

#### **Table 9 Liquidity**

This table presents the results of extended model (1) by regressing capital expenditure (*capex*) on the interaction between listing indicator and the stock liquidity variables. The variables of stock liquidity are Amihud illiquidity measure taking negative (*liquidity*) and stock turnover (*turn over*).

Control variables include return on assets (*roa*), firm age (*age*), firm size (*ln.tast*), cash holding (*cash*), and shareholding of financial institutions (*sh.financial*), foreign investors (*sh.foreign*), top 10 investors (*sh.top10*), and board members (*sh.directors*). We also control for year and firm fixed effects in all the models. t-statistics in parentheses are calculated based on standard errors obtained by clustering at the firm level. \*, \*\*, \*\*\* indicate significance at the 10, 5, 1% levels, respectively, using a two-tailed test. All variables are defined in Table A1.

	capex						
	All Firms	Business Group	Standalone				
	(1)	(2)	(3)				
liquidity	2.0812***	2.9338***	0.0715				
	(2.87)	(3.40)	(0.04)				
pred_q	0.0530***	0.0468***	0.0747***				
· ·	(9.78)	(7.93)	(4.99)				
roa	0.0165***	0.0158***	0.0233***				
	(3.75)	(3.27)	(3.27)				
age	-0.0612**	-0.0543**	0.0126				
	(-2.38)	(-2.08)	(0.16)				
size	-0.0466***	-0.0415***	-0.0519*				
	(-6.53)	(-5.80)	(-1.82)				
cash	0.0143***	0.0165***	0.0086***				
	(9.30)	(9.53)	(3.17)				
lev	-0.1760***	-0.1838***	-0.1714***				
	(-10.23)	(-10.11)	(-3.09)				
sh_financial	0.1814***	0.1742***	0.1550				
	(6.84)	(7.24)	(1.15)				
sh_foreign	0.0687**	0.0734**	-0.1710				
	(2.37)	(2.50)	(-1.34)				
sh_top10	-0.0089	-0.0095	-0.0186				
	(-0.61)	(-0.62)	(-0.49)				
sh_directors	0.0526*	0.0648**	0.0087				
	(1.92)	(2.23)	(0.12)				
Observations	37,128	31,679	5,449				
Year fixed effects	yes	yes	yes				
Industry fixed effects	yes	yes	yes				
clustered by	firm	firm	firm				
Adj. R <sup>2</sup>	0.447	0.451	0.499				

# Appendix on

# "Stock Market Listing, Investment, and Business Groups: How Firm Structure Impacts Investment?"

Joseph J. French University of Northern Colorado

> Ryosuke Fujitani Hitotsubashi University

> and Yukihiro Yasuda Hitotsubashi University

#### A1. Disclosure requirement in Japan

The requirements on financial reporting in Japan are quite unique: institutional requirement is enforced for not only listed firm but also for unlisted ones that satisfy the specified conditions. This enables us to observe the effects of being listed by excluding the effects of requirement on the different level of firm disclosures. Figure A1 is the matrix showing the relationship across the requirements of the two acts on the disclosure, firm size, and firm condition (public or private). The X axis distinguishes listed firms and unlisted firms, and Y axis depicts firm classification of Corporate Act (J-CA, hereafter).

# [Figure A1]

J-CA and J-FIEA are the legal background on financial reporting in Japan. Required disclosure is different between J-CA and J-FIEA. J-CA requires Large Company to report annual audited financial statements, which do not include cash flow statement (Article 435 and 444). In addition, Large Company which is not mandated to report consolidated financial statements do not need to report consolidated financial statements. Upper area of the matrix depicts Large Company, and firms belonging to the area need to report audited financial statements but unnecessarily consolidated one. Small and Medium Companies are required to comply SME accounting standard, which is less strict and complex.

J-FIEA mandates firms satisfying conditions prescribed in Item 1 of Article 24: i.e. if the firms issue:

(i) Securities listed in a Financial Instruments Exchange (excluding Specified Listed Securities);(ii) Securities specified by a Cabinet Order as those of which the state of distribution can be regarded as being equivalent to Securities referred to in the preceding item (excluding Securities)

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specified by Cabinet Order as having equivalent distribution statuses to Specified Listed Securities);

(iii) Securities to whose Public Offering or Secondary Distribution the main clause of Article 4, paragraph (1), the main clause of Article 4, paragraph (2), the main clause of Article 4, paragraph (3), or the main clause of Article 23-8, paragraph (1) or (2) ap-plies (excluding those specified in the preceding two items); or

(iv) Securities (limited to share certificates, Rights in a Securities Investment Business, etc. that are deemed to be Securities pursuant to Article 2, paragraph (2), and other Securities specified by Cabinet Order) that are issued by the company, for which the number of holders on the last day of the relevant business year or on the last day of any of the business years that began within four years before the day on which the relevant business year began is at least the number specified by Cabinet Order (or, for Rights in a Securities Investment Business, etc. that are deemed to be Securities pursuant to Article 2, paragraph (2), if the number of holders on the last day of the relevant business year is at least the number specified by Cabinet Order) (excluding Securities specified in the preceding three items).

The "number" mentioned in Item (iv) is specifies in article 3-6 of Order for Enforcement of the Financial Instruments and Exchange Act as:

(4) The number specified by a Cabinet Order, referred to in Article 24 (1)(iv) of the Act, is 1000 (in cases where the Securities are Securities for Professional Investors, the number obtained by adding the number of Professional Investors calculated pursuant to the provisions of a Cabinet Office Ordinance to 1000).

# A2. Variable definitions

All variables are defined in Table A1.

### **Appendix Figures and Tables**



### c. No requirement on cash flow statement

## Figure A1 Japanese legal framework on financial reporting:

This diagram describes the classification across firms required to report financial statements by Japanese legal provisions on financial reporting. For simplicity, this diagram shows disclosure requirements of J-FIEA stricter and more specific than J-CA, and Large and Public firms need to report financial statements required by J-FIEA. *X axis* separates firms into listed and unlisted firms, and *Y axis* separates them into Large and Small and Medium Enterprises. J-CA requires Large Companies to report a) audited but b) unconsolidated financial statements. J-CA does not prescribe any requirement on cash flow statement. J-FIEA requires firms satisfying conditions prescribed in Article 24 to report a) audited b) consolidated financial statements, including c) cash flow statement. UDFs are the firms included in the blue shaded area: i.e., unlisted firms required to report financial statements by J-FIEA.

 Table A1 Variable definition:

This table describes the definition of variables used in analysis. All the data are collected from Nikkei NEEDS Financial Quest 2.0.

Variables	Definition
D	
Dependent variable	es The changes in plant, property, and equipment from the providus period plus depression
⊿ppe	and impairments scaled by the sum of tangible and intangible assets
caner	Capital expenditure reported in footnote of Form I 10-K scaled by the sum of tangible and
сарел	intangible assets.
<i>capex+rd</i>	Capital expenditure reported in footnote of Form J 10-K ( <i>capex</i> ) plus R&D expenditure
1	scaled by the sum of tangible and intangible assets.
tan+int	Cash outflow to purchase both tangible and intangible assets scaled by the sum of tangible
	and intangible assets.
Independent varial	bles and other variables
listed	Indicator which equals one if firm <i>i</i> is listed firm, zero otherwise.
roa	Operating income sum of tangible and intangible assets.
pred q	Following Campello and Graham (2013), predicted q is computed by the following regres-
	sion:
	$q = \eta_0 + \eta_1 sg + \eta_2 roa + \eta_3 net\_income + \eta_4 lev + fe + \varepsilon,$
	net_income is ordinary income, fe includes industry and year fixed effects, and the other
	variables are defined in this table. After estimating the model, we then use the regression
	coefficients to generate <i>predicted q</i> for each firm, both listed and unlisted firms.
age	Natural logarithm of Firm's age plus one.
size	Natural logarithm of total assets.
cash	Sum of cash, cash equivalent and short-term investment securities divided by sum of tan-
lay	gible and initial gible assets.
lev In salas	Natural logarithm of cales
sh financial	Common stock ownership percentage of financial intermediaries
sh foreign	Common stock ownership of foreign investors
sh.joreign	Common stock ownership of top 10 shareholders.
sh_director	Common stock ownership percentage of board members.
no navout	An indicator taking one if the firm did not new dividend or execute stock repurchase for
no_puyoui	three years before
no bacc	An indicator taking one if the firm did not issue bond for three years before
small	An indicator taking one if the firm belongs to the first quartile of firm size, zero otherwise.
hp	An indicator taking one if the firm belongs to the $5^{th}$ quartile of Hadlock-Pierce index, zero
1	otherwise. Hadlock-Pierce index is defined as:
	Hadlock-Pierce index: = $(-0.737 \times total\_asset) + (0.043 \times total\_asset^2) - (0.040 \times Age)$ ,
	where <i>total_asset</i> is total assets in the previous period, and <i>Age</i> is firm age.
liquid	Amihud illiquidity ( <i>illiq</i> ) taking negative. Amihud illiquidity is computed as:
1	$illiq = (1/d) \Sigma [/ret   / (vol \times price)]$
	ret represents daily stock returns, vol represents daily trading volume, price represents the
	stock price, and d represents the number of the dates of fiscal year. Thus, <i>liquidity</i> is:
	$liquid = (-1) (1/d) \Sigma[ ret /(vol \times price)]$
turn_over	The fiscal year average of trading volume scaled by total share outstanding.

## Table A2 Descriptive Statistics (Business group firms vs. Standalone firms):

This table presents the descriptive statistics on all variables in main analysis. Panel A and B compares the statistics of all firms (listed firms and unlisted firms) between bisiness group firms and standalone firms. Columns (1) and (2) report the descriptive statistics of listed and unlisted firms, respectively. Columns (3) report the difference in each variable between listed and unlisted firms. \*, \*\*, \*\*\* indicate significance at the 10, 5, 1% levels using a two-tailed test. Left hand side in each panel reports the statistics of firms with subsidiaries, and right-hand side reports those of standalone firms. All observations falling in the top or bottom 1 % with respect to each variable are winsorized. All variables are defined in Table A1.

Panel A: Busine	ess Group sub-sample									
	(1) Listed firms			(2) Unlisted			(3) Listed - Un-			
	(n=34,433)			Firms (n=1,386)			listed			
	mean	median	sd	mean	median	sd	mean		median	
∆ppe	0.1445	0.0956	0.2153	0.0917	0.0604	0.1538	0.0528	***	0.0352	***
capex	0.1471	0.0967	0.1851	0.0901	0.0617	0.1157	0.0569	***	0.0350	***
tan+int	0.1578	0.1007	0.2113	0.1016	0.0660	0.1441	0.0562	***	0.0346	***
capex+rd	0.2166	0.1396	0.3007	0.1184	0.0708	0.1910	0.0982	***	0.0688	***
pred_q	1.0828	0.9812	0.4557	0.9675	0.9013	0.3173	0.1153	***	0.0799	***
roa	0.4327	0.2368	0.9966	0.1990	0.1353	0.4711	0.2337	***	0.1015	***
age	3.8346	4.0073	0.5926	4.0781	4.1744	0.4732	-0.2435	***	-0.1671	***
size	10.5723	10.4599	1.4983	10.4520	10.2651	1.2803	0.1203	***	0.1948	**
cash	1.4188	0.4794	3.6461	0.7446	0.2776	2.3266	0.6742	***	0.2018	***
lev	0.2163	0.1854	0.1837	0.3435	0.3333	0.2221	-0.1272	***	-0.1479	***
sh_financial	0.1992	0.1785	0.1323	0.1030	0.0785	0.0923	0.0961	***	0.1000	***
sh_foreign	0.0823	0.0370	0.1044	0.0109	0	0.0579	0.0714	***	0.0370	***
sh_top10	0.4981	0.4894	0.1564	0.4767	0.4972	0.2561	0.0214	***	-0.0078	
sh_directors	0.0884	0.0245	0.1297	0.0612	0.0250	0.0925	0.0272	***	-0.0005	***

Panel B: Standalone sub-sample										
	(1) Listed firms			(2) Unlisted			(3) Listed - Un-			
	(n=5,513)			Firms (n=797)			listed			
	mean	median	sd	mean	median	sd	mean		median	
∆ppe	0.1463	0.0695	0.2459	0.0735	0.0150	0.2177	0.0728	***	0.0545	***
capex	0.1743	0.0803	0.2567	0.0585	0.0071	0.1723	0.1158	***	0.0731	***
tan+int	0.1914	0.0879	0.2927	0.0903	0.0157	0.2407	0.1011	***	0.0721	***
capex+rd	0.2736	0.1126	0.4784	0.1094	0.0087	0.3750	0.1642	***	0.1039	***
pred_q	1.1105	0.9837	0.5318	1.1415	1.0893	0.4251	-0.0310		-0.1056	***
roa	0.6797	0.2314	1.6755	0.1871	0.0394	0.9995	0.4926	***	0.1919	***
age	3.6151	3.7377	0.5758	3.7465	3.9318	0.5933	-0.1314	***	-0.1942	***
size	9.1682	9.1517	1.0242	8.0569	7.8610	1.0843	1.1113	***	1.2907	***
cash	2.8982	0.6690	6.5033	1.2777	0.2067	4.4471	1.6205	***	0.4623	***
lev	0.1813	0.1131	0.1951	0.1951	0.0512	0.2435	-0.0137	*	0.0619	*
sh_financial	0.1043	0.0904	0.0818	0.0235	0.0060	0.0453	0.0809	***	0.0844	***
sh_foreign	0.0411	0.0073	0.0837	0.0129	0	0.0687	0.0282	***	0.0073	***
sh_top10	0.5707	0.5765	0.1719	0.4238	0.3913	0.3236	0.1468	***	0.1853	***
sh directors	0.1541	0.0949	0.1607	0.0710	0.0060	0.1395	0.0831	***	0.0890	***

## Table A3 Treatment effect model first step:

This table reports the first-step regression results of treatment effect model for results of regression model (1): \*, \*\*, \*\*\* indicate significance at the 10, 5, 1% levels, respectively, using a two-tailed test. All variables are defined in Table A1.

The first stage model of treatment effect model					
	Probit Model Dependent variable = Listed				
<u>-</u>					
la salar	0.2190***				
in_sales	(0.0221)				
	(0.0221)				
gr.sales	0.2400***				
	(0.0573)				
roa	0.1173***				
	(0.0265)				
lev	-1.0100***				
	(0.1393)				
Constant	-0.3867*				
	(0.2044)				
Observations	52.038				
Pseudo R <sup>2</sup>	0.101				
ln_sales gr.sales roa lev Constant Observations Pseudo R <sup>2</sup>	$\begin{array}{c} 0.2189^{***} \\ (0.0221) \\ 0.2400^{***} \\ (0.0573) \\ 0.1173^{***} \\ (0.0265) \\ -1.0100^{***} \\ (0.1393) \\ -0.3867^{*} \\ (0.2044) \\ 52,038 \\ 0.101 \end{array}$				

#### Table A4 Listing status and financial constraints

This table presents the results of financial constraint analyses. We estimate Model (1) using subsamples. The subsamples are divided according to the variables of financial constraints. We report the difference in the coefficients between financial constrained and unconstrained subsample, and their F-value using Chow test.

All the models include the control variables: assets (*roa*), firm age (*age*), firm size (*ln.tast*), cash holding (*cash*), and shareholding of financial institutions (*sh.financial*), foreign investors (*sh.foreign*), top 10 investors (*sh.top10*), and board members (*sh.directors*). \*, \*\*, \*\*\* indicate significance at the 10, 5, 1% levels, respectively, using a two-tailed test. All variables are defined in Table A1.

	Payout	Bond Access	Size	HP
Constrained	0.1065***	0.0509***	0.1025***	0.1344***
	(6.82)	(6.62)	(6.05)	(4.59)
Unconstrained	0.0064	-0.0108	-0.0072	0.0081
	(1.02)	(-1.07)	(-0.63)	(1.11)
Difference	0.1001***	0.0617***	0.1097***	0.1263***
Chow test	1114.55	397.57	5008.37	56.89