Perspectives on the Future of Asset Pricing: Macroeconomics and Monetary Policy

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- Large announcement effects (macro and monetary policy news) Ai and Bansal (2018); Jarocinski and Karadi (2019)
- Structural change (slowing growth, rising profit shares, growing inequality, low real rates)
- ► Global financial crisis showed important feedback effects between financial markets and real economy.

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- Structural change (slowing growth, rising profit shares, growing inequality, low real rates)
- ► Global financial crisis showed important feedback effects between financial markets and real economy.
- Question not whether macroeconomy matters but how.

Does the Macroeconomy Matter?

- Many possible reasons evidence is mixed: models gross simplifications; data are mismeasured and limited in terms of what is covered, estimation tools restrictive, information sets unobserved.
- ▶ But, most above work based on **aggregate data**. Can also ask: are we looking at the wrong XS moments?
- ▶ Is the representative agent framework too much of an abstraction?

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- ▶ Is the representative agent framework too much of an abstraction?
 - XS tail risks: Schmidt (2016); Constantinides and Ghosh (2017)
 - ▶ Profit/labor shares: Danthine and Donaldson (2002); Favilukis and Lin (2016); Favilukis and Lin (2013); Favilukis and Lin (2015); Gomez (2016); Farhi and Gourio (2018); Lettau and Ludvigson (2013); Greenwald, Lettau, Ludvigson (2013, 2019)l Lettau, Ludvigson, and Ma (2019).

Topic 1: Distributional Shifts

- Profound distributional shifts in macroeconomy.
- Growing evidence heterogeneity matters.
- ► What are the consequences for stock prices and other asset values of redistributive shocks and redistributive trends?

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 - Empirical Evidence suggests monetary policy shocks have transitory effects (Christiano, Eichenbaum, and Evans (2005)).
- So why do monetary policy announcements seem to matter so much for stock prices if macro models and macro evidence imply transitory effects of policy?
 - Asset pricing has little to say...
- ► Are we looking at the wrong aspects of monetary policy? Are macro models wrong? Both?

Topic 2: Monetary Policy and Asset Values

How, why, and to what degree does monetary policy affect the stock market and other assets?

Topic 1: Distributional Shifts

- Survey of Consumer Finances (SCF)
- ▶ Definitions of income categories for the *i*th household:
 - ▶ $Wag_{i,t} \equiv$ annual wage and salary income.
 - ► $Cap_{i,t} \equiv$ sum of business, interest, dividend, realized capital gains, and pension fund income.
 - Let $Oth_{i,t} \equiv \text{transfers}$ and social security income.
 - ► Total log income $y_{i,t} \equiv \log(Wag_{i,t} + Cap_{i,t} + Oth_{i,t})$.
- ▶ Stock wealth *i*th household: any non-zero direct+indirect holdings.

Stock wealth is highly concentrated.

Panel A: Percent	of Stock We	alth, sorted	l by Stock	Wealth	
Percentile of Stock Wealth	1989	1998	2004	2013	2016
< 70%	0.01%	1.30%	1.35%	0.84%	0.98%
70 - 85%	3.12%	7.42%	7.41%	5.92%	5.81%
85 - 90%	4.19%	6.45%	6.70%	6.17%	5.46%
90 - 95%	11.16%	11.28%	13.26%	12.67%	11.89%
95 - 100%	81.54%	73.93%	71.21%	74.54%	75.86%
Panel B: Percent	of Total Inco	me, sortec	l by Stock	Wealth	
Percentile of Stock Wealth	1989	1998	2004	2013	2016
< 70%	43.64%	42.29%	40.76%	37.64%	35.61%
70 - 85%	17.58%	18.81%	17.43%	16.31%	15.72%
85 - 90%	7.36%	7.48%	7.74%	7.52%	6.79%
90 - 95%	8.13%	8.48%	9.83%	10.92%	10.75%
95 - 100%	23.28%	22.94%	28.28%	27.62%	31.12%
Panel C: S	Stock Marke	t Participa	tion Rates		
	1989	1998	2004	2013	2016
Raw (rpr)	31.7%	49.3%	49.7%	48.8%	52.1%
Wealth-weighted	12.3%	18.1%	20.1%	18.6%	18.5%

account, social security, and transfers/others. The wealth-weighted participation rate is calculated as value weighted ownership = $0.01 \left(w_{99-100\%}\right) + 0.04 \left(w_{95-99\%}\right) + (\textit{rpr} - 0.05) \left(1 - w_{5\%}\right) + (1 - \textit{rpr}) (0)$, where $w_{x\%}$ is the proportion of stock market wealth owned by top x% stockowners.

Source: Survey of Consumer Finance 1989-2016. Total income includes wage, business income, interest and dividend, capital gains, pension

Fraction of total income earned by top stockowners trending up

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Participation rates flat for 20 years

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 On a wealth-weighted basis participation rates far lower and falling since 2004

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Explaining the cross section of total income

Model:
$$y_i = \alpha + \beta^{\text{Age}} \cdot \text{Age}_i + \sum_j \beta_j^{\text{Edu}} \cdot \text{Edu}_{i,j} + \sum_j \beta_j^{\text{Wag}} \cdot \text{Wag}_{i,j} + \sum_j \beta_j^{\text{Sto}} \cdot \text{Sto}_{i,j} + \epsilon_i$$

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Full model R ²	0.65	0.61	0.63	0.67	0.70	0.71	0.71	0.70	0.67	0.72
Effect of: (partia	al R^2)									
Age	0.01	0.02	0.04	0.02	0.01	0.01	0.00	0.01	0.01	0.01
Education	0.08	0.05	0.04	0.02	0.03	0.02	0.02	0.02	0.01	0.02
Wages	0.29	0.26	0.29	0.29	0.23	0.25	0.18	0.27	0.23	0.23
Stock Wealth	0.19	0.23	0.25	0.29	0.37	0.39	0.44	0.38	0.37	0.43

Note: The first row reports the R^2 for the full model with explanatory variables $x = (\mathrm{Age}_i, \mathrm{Edu}_{i,j}, \mathrm{Wag}_{i,j}, \mathrm{Sto}_{i,j})$. The subsequent rows report partial R^2 statistics for the contribution of different $x_k \in x$. The partial $R^2_{x_1|x}$ gives the percent reduction in unexplained variation achieved from including x_1 when another set of variables x is already in the model. In rows 2 to 5, x_1 is one of the variables named in the first column and $x = \mathrm{all}\ x_k \neq x_1$. $\mathrm{Edu}_{i,j}$ is a set of dummy variables that equals 1 if the i-th household education is $\{\mathrm{No}\ \mathrm{education}, \mathrm{High}\ \mathrm{-school}, \mathrm{College}\ \mathrm{or}\ \mathrm{better}\}$. $\mathrm{Wag}_{i,j}(\mathrm{Sto}_{i,j})$ is a set of dummy variables that equals 1 if the i-th household falls in the j-th percentile range where $j = \{0.25, 25-50, 50.75, 75-90, 90.99, 99-100\}$ The sample size N = 47, 776 households. Source: Survey of Consumer Finances 1989-2016 and author's calculations.

Age unimportant; education of declining importance.

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▶ 1989: Wages far more important than stock ownership

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▶ 2016: Stock ownership far more important than wages

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Stock ownership increasingly important determinant of income

Explaining the cross section of total income

$$\textbf{Model:}\ y_i = \alpha + \beta^{\text{Age}} \cdot \text{Age}_i + \sum_j \beta_j^{\text{Edu}} \cdot \text{Edu}_{i,j} + \sum_j \beta_j^{\text{Wag}} \cdot \text{Wag}_{i,j} + \sum_j \beta_j^{\text{Sto}} \cdot \text{Sto}_{i,j} + \epsilon_i$$

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Stock Market v.s Broader Economy

► ME= Total value of market equity of the NFCS.



Notes: ME: Nonfinancial Corporate Sector Stock Value. E: Nonfinancial Corporate Business After-Tax Profits. GDP & C: Current Dollars GDP and personal consumption expenditures. NVA: Net Value Added of Nonfinancial Corporate Sector. The sample spans the period 1952:Q1-2017;Q4.

Stock Market v.s Broader Economy

▶ ME relative to 3 different measures of agg. economic activity is at or near post-war high, and has *trended up*.



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Stock Market v.s Broader Economy

▶ Notably, ME/E not near post-war high.



Notes: ME: Nonfinancial Corporate Sector Stock Value. E: Nonfinancial Corporate Business After-Tax Profits. GDP & C: Current Dollars GDP and personal consumption expenditures. NVA: Net Value Added of Nonfinancial Corporate Sector. The sample spans the period 1952:Q1-2017;Q4.

Distributional Trends and Market Trends

Greenwald, Lettau, and Ludvigson (2019)

- ► Equity priced in the model by a representative *shareholder*, akin to wealthy household or large institutional investor.
- ▶ Remaining agents just **supply labor**, play no role in asset pricing.
- Shareholder preferences subject to a shock alters **appetite for risk**.
- Estimate full dynamic model that incorporates time variation in:
 - Expect. growth of rewards generated from productive activity
 - How rewards are apportioned between shareholders and labor
 - Equity risk premium and expected future path of short rates in nearand long-term
- ► Apply model to data on the corporate sector over period 1952:Q1-2017:Q4.

Market's rise: 43% since 1989 and 19% over full sample attributable to reallocation of rewards to shareholders without changing size of rewards.

	Pan	el: Market Ec	quity
Contribution	1952-2017	1952-1988	1989-2017
Total	1405.81%	151.23%	477.34%
Earnings Share, s_t	18.65%	-23.4%	42.54%
$s_{LF,t}$	17.04%	-21.65%	37.90%
$S_{HF,t}$	1.52%	-1.75%	4.64%
Risk Premium	25.74%	20.51%	24.41%
$x_{LF,t}$	25.69%	20.83%	24.31%
$x_{HF,t}$	0.05%	-0.32%	0.10%
Risk-Free Rate	2.16%	-8.52%	8.48%
$\delta_{LF,t}$	2.11%	-8.57%	8.35%
$\delta_{HF,t}$	0.05%	0.05%	0.13%
Real PC Output Growth	53.54%	111.41%	24.57%

Notes: The table presents the growth decompositions for the real value of market equity (top panel) or the market equity-output ratio (bottom panel). The persistence parameter of the risk price is set to its baseline value of 0.85. The sample spans the period 1952/Q1-2017/Q4. Source: Greenwald, Lettau, and Ludvigson (2019).

► Market's rise: 24% since 1989 and 26% over full sample attributable to declining risk-premium.

	Pan	el: Market Ec	uity
Contribution	1952-2017	1952-1988	1989-2017
Total	1405.81%	151.23%	477.34%
Earnings Share, s_t	18.65%	-23.4%	42.54%
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▶ Other components since 1989: much smaller role for $r_{f,t}$

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Economic Growth contributes **just 25**% since 1989; 54% over full sample.

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▶ 1952-1988: Δy_t explained 111% of market's rise. But...

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► That 37 year period created *less than half* wealth created in 29 years since 1989.

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Unanswered Questions

- ► Central observation: persistent rise in equity values over last 30-40 years *relative* to aggregate economy.
- ▶ Why have factors shares changed so persistently? Will these trends continue? Do the reasons for the changes matter? How is it related to broader trends in inequality and economic growth?
- Suggestion: trends, heterogeneity, and economic inequality could be playing a central role in asset pricing.

Topic 2: Monetary Policy and Asset Values

Motivation: Asset Valuations & Monetary Policy

Estimate $cay^{MS} = \alpha_{\tilde{\xi}_t} + \epsilon_t^{CAY}$, & $MPS_t^{MS} = \mu_{\tilde{\xi}_t} + \epsilon_t^{MPS}$, $\tilde{\xi}_t$ 2-state Markov switching system with synchronized regimes.

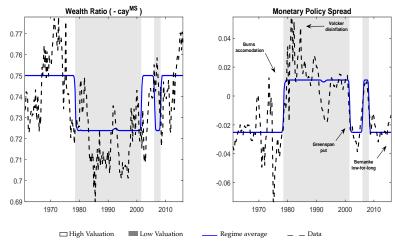


Figure plots the wealth ratio (-cay) and the monetary policy spread $MPS_t \equiv FFR_t$ - Expected Inflation $_t - r_t^*$. r^* is from Laubach and Williams (2003). The sample spans 1961:Q1-2017:Q3. Source: Bianchi, Lettau, and Ludvigson (2016).

Motivation: Asset Valuations & Monetary Policy

Wealth ratio, MPS in 5 subperiods fluctuate around distinct means

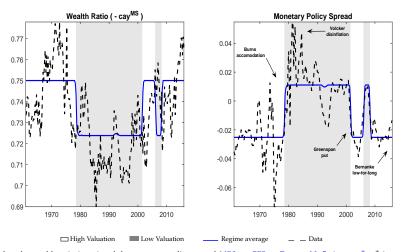


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Motivation: Asset Valuations & Monetary Policy

► High wealth ratio subperiods coincide with Low MPS subperiods.

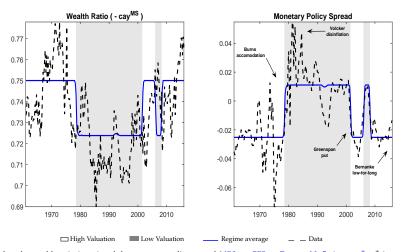


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Long-lasting Effects of Monetary Policy?

- Estimates suggest lower frequency co-movements of policy rates and asset valuations.
- Can this plausibly be due to monetary policy?
 - Macro models imply only short-run effects of monetary policy "shocks" on real variables.
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- Evidence: values of long-term financial assets respond to the actions and announcements of central banks.
 - ► Cochrane and Piazzesi (2002); Hanson and Stein (2015); Gertler and Karadi (2015); Gilchrist, López-Salido, and Zakrajšek (2015); Boyarchenko, Haddad, and Plosser (2016); Jarocinski and Karadi (2019).
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 - Hard to make sense of findings unless something associated with announcement is expected to have a persistent influence on real economy.
- ▶ Most above discuss **information** and **risk-premia** channels.
- I want to focus on different channel: real short-term interest rates.

Bianchi, Lettau, and Ludvigson (2016)

Prototypical New Keynesian Model (Galí (2015)) w/ 3 Gaussian shocks (in green) but w/ 3 modifications:

$$\begin{aligned} y_t &= \delta y_{t-1} - \sigma \left[i_t - \phi \pi_t - (1 - \phi) \, \overline{\pi}_t - r \right] + d_t \\ \pi_t &= \overline{\pi}_t + \frac{\kappa}{1 - \beta \phi} \left[y_{t-1} - y_{t-1}^* \right] \end{aligned} \qquad \qquad \begin{aligned} &\text{Phillips curve} \\ i_t - \left(r + \pi_{\xi_t}^T \right) &= \left(1 - \rho_{i,\xi_t} \right) \left[\psi_{\pi,\xi_t} \left(\pi_t - \pi_{\xi_t}^T \right) + \psi_{\Delta y,\xi_t} \left(y_t - y_{t-1} \right) \right] \\ &+ \rho_{i,\xi_t} \left[i_{t-1} - \left(r + \pi_{\xi_t}^T \right) \right] + \sigma_i \varepsilon_{i,t} \end{aligned} \qquad \qquad \end{aligned} \qquad \qquad \begin{aligned} &\text{Policy Rule} \\ \overline{\pi}_t &= \left[1 - \gamma^T \right] \underbrace{ \left[\overline{\pi}_{t-1} + \gamma \left(1 - \phi \right)^{-1} \left(\pi_t - \phi \pi_{t-1} - (1 - \phi) \, \overline{\pi}_{t-1} \right) \right]}_{\text{constant gain learning}} + \gamma^T \pi_{\xi_t}^T \end{aligned}$$

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 - 2. **Learning and adaptive expectations:** Agent learns about *trend* inflation $\overline{\pi_t}$ using constant-gain learning (Malmendier and Nagel (2015)) + backward-looking rule to form expectations of y_{t+i} .
 - 3. **Perceived** $\overline{\pi}_t \neq \pi_t^T$: Agents don't directly observe π_t^T and/or CB announcements not viewed as fully credible or informative.

Estimation: Model w/ Persistent Non-neutrality

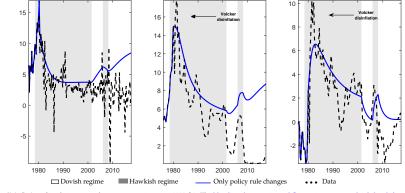
- Estimation: Implied expected inflation process must be consistent with survey data.
- ► All MS parameters $(\pi_{\xi_t}^T, \psi_{\pi,\xi_t}, \psi_{\Delta y,\xi_t})$ freely estimated, could in principle show no regime changes in policy rule.
- ▶ **Data**: GDP growth, inflation, federal funds rate, mean of Michigan survey of inflation expectations.

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- ▶ **Data**: GDP growth, inflation, federal funds rate, mean of Michigan survey of inflation expectations.
- ► Result: large regime shifts. High valuation regimes are Dovish m.p. regimes; low valuation regimes are Hawkish m.p. regimes.
- **Dovish regime** has a high π^T and low activism on inflation deviations from target relative to activism on growth.
- ► **Hawkish regime** has a low π^T and high activism on inflation deviations from target relative to activism on growth.

Inflation

A prominent episode. Start economy in 1980:Q1 (peak π) and shut down all est. Gaussian shocks.

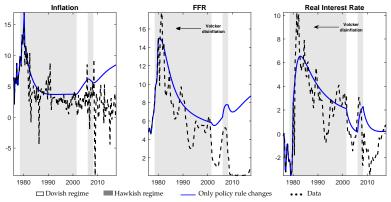


FFR

Real Interest Rate

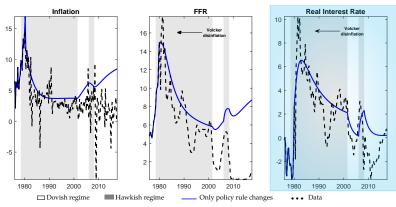
The Volcker disinflation. Simulation starts the economy as it was in 1980:Q1. The blue line shows estimated fluctuations generated *only* by shifts in the policy rule. Dovish regime has a high π^T and low activism on inflation deviations relative to activism on growth. Hawkish regime has a low π^T and high activism on inflation deviations relative to activism on growth. Real interest rate is the difference between the FFR and expected inflation based on the model solution. The sample spans 1961:Q1-2017-Q3. Source: Bianchi, Lettau, and Ludvigson (2016).

▶ **Blue line**: estimated movements *purely the result* of behavior of monetary authority (inflation target, activism).



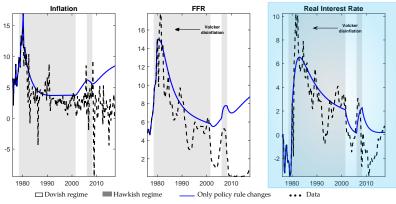
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Large, low frequency movements in real rate, and **most of its downward trend since 1980** attributable to changes in policy rule.



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▶ Mon. policy *shocks* have transitory effects.



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Dovish Monetary Policy and Reaching for Yield?

- ▶ Use estimation to identify component of real interest rates, RIR_t^{MP} , driven *only* by changes in policy rule.
- ▶ **Find**: When RIR_t^{MP} declines, equity risk-premia decline, consistent with a reach-for-yield (RFY).
- Suggestion: Persistently low or high interest rate environments, associated with shifts in *conduct* of monetary policy may have economic effects quite different from monetary policy "shocks".

Unanswered Questions

- ▶ We aren't first to suggest monetary policy *shocks* may not be most important for asset prices (e.g., Cochrane and Piazzesi (2002)).
- ➤ Why does the CB change the conduct of monetary policy? (Possibly in *reaction* to markets): (Cieslak and Vissing-Jorgensen (2017)).
- ▶ More evidence needed to understand how asset markets interact with *systematic conduct* of monetary policy.

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- ▶ We also aren't first to suggest low rates, possibly driven by unconventional monetary policy, => reach for yield (e.g., Rajan (2006); Farhi and Tirole (2012); Rajan (2013); Stein (2013); Hanson and Stein (2015); Di Maggio and Kacperczyk (2015); Choi and Kronlund (2015). But focus has been on heavily intermediated asset clases, rather than equity.
 - ► Model of preferences for retail investors and households?

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 - Model of preferences for retail investors and households?
- ▶ What is the role, if any, of banking sector in transmission of monetary policy to asset values? (Important starts: Drechsler, Savov, and Schnabl (2014); Piazzesi and Schneider (2015); Piazzesi, Rogers, and Schneider (2018)).

References I

- AI, H., AND R. BANSAL (2018): "Risk preferences and the macroeconomic announcement premium," *Econometrica*, 86(4), 1383–1430.
- BANSAL, R., D. KIKU, AND A. YARON (2016): "Risks for the long run: Estimation with time aggregation," *Journal of Monetary Economics*, 82, 52–69.
- BIANCHI, F., M. LETTAU, AND S. C. LUDVIGSON (2016): "Monetary Policy and Asset Valuaton," http://www.econ.nyu.edu/user/ludvigsons/reg.pdf.
- BOYARCHENKO, N., V. HADDAD, AND M. C. PLOSSER (2016): "The Federal Reserve and market confidence," .
- Breeden, D., M. Gibbons, AND R. Litzenberger (1989): "Empirical Tests of the Consumption-Oriented CAPM," *Journal of Finance*, 44, 231–262.
- CAMPBELL, J. Y., AND J. MEI (1993): "Where do betas come from? Asset price dynamics and the sources of systematic risk," *The Review of Financial Studies*, 6(3), 567–592.
- CHEN, N.-F., R. ROLL, AND S. A. ROSS (1986): "Economic Forces and the Stock Market," *Journal of Business*, 59(3), 383–403.
- CHOI, J., AND M. KRONLUND (2015): "Reaching for yield or playing it safe? Risk taking by bond mutual funds," Available at SSRN: https://ssrn.com/abstract=2669465.
- CHRISTIANO, L. J., M. EICHENBAUM, AND C. L. EVANS (2005): "Nominal Rigidities and the Dynamic Effects of a Shock to Monetary Policy," *Journal of Political Economy*, 113(1), 1–45.
- CIESLAK, A., AND A. VISSING-JORGENSEN (2017): "The economics of the Fed put," Unpublished Manuscript, University of California at Berkeley.

References II

- COCHRANE, J. H., AND M. PIAZZESI (2002): "The fed and interest rates-a high-frequency identification," AEA Papers and Proceedings, 92(2), 90–95.
- CONSTANTINIDES, G. M., AND A. GHOSH (2017): "Asset pricing with countercyclical household consumption risk," *The Journal of Finance*, 72(1), 415–460.
- DANTHINE, J.-P., AND J. B. DONALDSON (2002): "Labour Relations and Asset Returns," *Review of Economic Studies*, 69(1), 41–64.
- DI MAGGIO, M., AND M. T. KACPERCZYK (2015): "The unintended consequences of the zero lower bound policy," Columbia Business School Research Paper.
- Drechsler, I., A. Savov, And P. Schnabl (2014): "A model of monetary policy and risk premia," National Bureau of Economic Research No. w20141.
- FARHI, E., AND F. GOURIO (2018): "Accounting for Macro-Finance Trends: Market Power, Intangibles, and Risk Premia," Discussion paper, National Bureau of Economic Research.
- FARHI, E., AND J. TIROLE (2012): "Collective moral hazard, maturity mismatch, and systemic bailouts," *The American Economic Review*, 102(1), 60–93.
- FAVILUKIS, J., AND X. LIN (2013): "The Elephant in the Room: The Impact of Labor Obligations on Credit Risk,"
- https://sites.google.com/site/jackfavilukis/WageCreditRisk.pdf.
- ——— (2015): "Wage Rigidity: A Quantitative Solution to Several Asset Pricing Puzzles," *The Review of Financial Studies*, 29(1), 148–192.

References III

- ——— (2016): "Does wage rigidity make firms riskier? Evidence from long-horizon return predictability," *Journal of Monetary Economics*, 78, 80–95.
- GALÍ, J. (2015): Monetary policy, inflation, and the business cycle: an introduction to the new Keynesian framework and its applications. Princeton University Press.
- GERTLER, M., AND P. KARADI (2015): "Monetary policy surprises, credit costs, and economic activity," *American Economic Journal: Macroeconomics*, 7(1), 44–76.
- GHOSH, A., C. JULLIARD, AND A. P. TAYLOR (2016): "What is the Consumption-CAPM missing? An information-theoretic framework for the analysis of asset pricing models," *The Review of Financial Studies*, 30(2), 442–504.
- GILCHRIST, S., D. LÓPEZ-SALIDO, AND E. ZAKRAJŠEK (2015): "Monetary policy and real borrowing costs at the zero lower bound," *American Economic Journal: Macroeconomics*, 7(1), 77–109.
- GOMEZ, M. (2016): "Asset prices and wealth inequality," Unpublished paper: Princeton. http://www.princeton.edu/~mattg/files/jmp.pdf.
- GREENWALD, D., M. LETTAU, AND S. C. LUDVIGSON (2014): "Origins of Stock Market Fluctuations," National Bureau of Economic Research Working Paper No. 19818.
- ——— (2019): "How the Wealth Was Won: Factors Shares as Market Fundamentals," National Bureau of Economic Research Working Paper No. 25769.
- HANSON, S. G., AND J. C. STEIN (2015): "Monetary policy and long-term real rates," *Journal of Financial Economics*, 115(3), 429–448.

References IV

- HERSKOVIC, B., A. MOREIRA, AND T. MUIR (2019): "Hedging risk factors," *Available at SSRN 3148693*.
- JAROCINSKI, M., AND P. KARADI (2019): "Deconstructing monetary policy surprises: the role of information shocks," *American Economic Journal: Macroeconomics* forthcoming.
- KOIJEN, R. S., H. LUSTIG, AND S. VAN NIEUWERBURGH (2017): "The cross-section and time series of stock and bond returns," *Journal of Monetary Economics*, 88, 50–69.
- LAUBACH, T., AND J. C. WILLIAMS (2003): "Measuring the natural rate of interest," *Review of Economics and Statistics*, 85(4), 1063–1070.
- LETTAU, M., AND S. C. LUDVIGSON (2001): "Resurrecting the (C)CAPM: A Cross-Sectional Test When Risk Premia are Time-Varying," *Journal of Political Economy*, 109(6), 1238–1287.
- LETTAU, M., AND S. C. LUDVIGSON (2013): "Shocks and Crashes," in *National Bureau of Economics Research Macroeconomics Annual*: 2013, ed. by J. Parker, and M. Woodford, vol. 28, pp. 293–354. MIT Press, Cambridge and London.
- LETTAU, M., S. C. LUDVIGSON, AND S. MA (2019): "Capital Share Risk in U.S. Asset Pricing," *The Journal of Finance*, 74(4), 1753–1792.
- LEWELLEN, J. W., AND S. NAGEL (2006): "The Conditional CAPM Does Not Explain Asset Pricing Anomalies," *Journal of Financial Economics*, 82(2), 289–314.
- MALMENDIER, U., AND S. NAGEL (2015): "Learning from inflation experiences," *The Quarterly Journal of Economics*, 131(1), 53–87.

References V

- Parker, J., and C. Julliard (2004): "Ultimate Consumption Risk and the Cross-Section of Expected Returns," *Journal of Political Economy*, 113(1), 185–222.
- PIAZZESI, M., C. ROGERS, AND M. SCHNEIDER (2018): "Money and banking in a New Keynesian model," Discussion paper, Working Paper, Stanford.
- PIAZZESI, M., AND M. SCHNEIDER (2015): "Payments, credit and asset prices," Unpublished manuscript, Stanford University.
- RAJAN, R. G. (2006): "Has finance made the world riskier?," European Financial Management, 12(4), 499–533.
- ——— (2013): "A step in the dark: unconventional monetary policy after the crisis," *Andrew Crockett Memorial Lecture, BIS, Basel,* 23.
- ROUSSANOV, N. (2014): "Composition of wealth, conditioning information, and the cross-section of stock returns," *Journal of Financial Economics*, 111(2), 352–380.
- SAVOV, A. (2011): "Asset pricing with garbage," The Journal of Finance, 66(1), 177–201.
- SCHMIDT, L. (2016): "Climbing and falling off the ladder: Asset pricing implications of labor market event risk," *Available at SSRN* 2471342.
- STEIN, J. C. (2013): "Overheating in Credit Markets: Origins, Measurement, and Policy Responses," Lecture at "Restoring Household Financial Stability after the Great Recession: Why Household Balance Sheets Matter", Federal Reserve Bank of St. Louis, St. Louis, Missouri.