# Learning about academic ability and the college drop-out decision

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Abstract: While assumptions about how agents update subjective beliefs in response to the arrival of new information play a central role in models of decision-making, little empirical evidence exists that can inform these modelling assumptions. In this paper we use unique data to examine how college students from low income families update their beliefs about academic ability and to examine the role that learning about ability and a variety of other factors play in the in the college drop-out decision.

### I. Introduction

While assumptions about how agents update subjective beliefs in response to the arrival of new information play a central role in both theoretical and empirical models of decision-making, very little empirical evidence exists that can inform these modelling assumptions; Manski (2004) writes that there exists a "critical need for basic research on expectations formation." In this paper, we use unique, new data to examine how students from low income families update their beliefs about academic ability after the time of college entrance and to provide direct evidence about the role that learning about ability and a variety of other factors plays in the college drop-out decision.

A natural first step towards providing evidence about expectations formation is to measure revisions to expectations about a factor of interest and to relate these revisions to observed events that provide an agent with new information about the factor. Unfortunately, even this measurement exercise, which requires knowledge of an agent's expectations about the factor of interest both before (prior beliefs) and after (posterior beliefs) the arrival of the new information, poses a challenge. The fundamental issue is that identifying expectations about the factor at any particular point in time is difficult using standard data on observed choices because a particular behavior is potentially consistent with multiple characterizations of preferences and expectations (Manski, 2002, 2004). In response to this fundamental identification issue, researchers in economics have recently paid much closer attention to the virtues of eliciting self-reports of subjective probabilities using carefully worded survey questions (Dominitz, 1998; Dominitz and Manski, 1996, 1997).\(^1\)
Nonetheless, the use of these data to measure revisions to expectations has remained quite limited. One reason for this is that most standard data sources have not traditionally provided expectations information.\(^2\) However, also problematic is that standard survey designs, which typically involve contacting respondents at most once a year, often do not have the flexibility to elicit beliefs about the factor of interest at the specific points in time

<sup>&</sup>lt;sup>1</sup>For early work in economics see Juster (1966).

<sup>&</sup>lt;sup>2</sup>Manski (2004) notes that this has changed rather dramatically very recently and describes surveys that currently elicit probabilistic information about expectations.

that might be ideal given the timing of information arrival in a particular context. As a result, the small amount of research that has measured revisions to expectations (see, for example, Dominitz, 1998; Dominitz and Manski, 2003; Dominitz and Hung 2003; Delevande, 2006; Lochner, 2007; Madeira, 2007) has often involved the collection of original data, sometimes employing an experimental design which allows the elicitation of both prior and posterior information in a single survey session.

Further, as described in Manski (2004), the difficulty of providing evidence about expectations formation becomes more acute if one wishes to move beyond the measurement of revisions towards an understanding of why particular revisions are observed:

Research that measures revisions to expectations and associates them with observed event realizations can be informative. However, I think that understanding expectations formations will also require intensive probing of persons to learn how they perceive their environments and how they process such new information as they may receive. Large-scale population surveys such as the HRS or SEE are not amenable to investigations of this type...

Our interest in understanding how students form expectations about academic ability is motivated, generally, by the well-recognized policy importance of understanding the substantial college drop-out rates of students from low income families and, more specifically, by the suggestion that the college drop-out decision is best viewed within an economic model in which students learn about their academic ability after matriculation (Manski, 1989, Altonji, 1993, Carneiro et al. 2005, Cunha et al. 2005). Our work is made possible by unique, new longitudinal data from the Berea Panel Study (BPS) which involves surveying low income students at Berea College approximately ten times each year. The BPS holds particular appeal for studying expectations formation because we initiated the survey with the specific goal of collecting information that a theoretical learning model indicates is useful for understanding the drop-out decision. The work here benefits greatly from the complete flexibility we had with respect to both the content of our survey instruments and the timing of the collection of these instruments.

Our approach to providing evidence about expectations formation is in the spirit suggested in the quote above from Manski (2004). We begin by measuring how students revise expectations and how these revisions relate to observed events. Specifically, we measure how the mean of a student's subjective grade distribution

changes between the time of college entrance and the beginning of the second semester, and we relate this change to new information that the student receives in the form of his first semester grade point average. We then pose our desire to understand why particular revisions are observed as a question of why revisions vary across students who have the same subjective mean at the time of college entrance and observe the same first semester grade information. To examine this question, we take advantage of survey questions that were designed in the spirit of the suggestion that information about how students "perceive their environments" is necessary in order to gain a better understanding of expectations formation. One survey question asks students to quantify their perceptions about the importance of various possible reasons for the gap between their expected first semester grade performance and their actual first semester grade performance. Another question asks student to quantify, in probabilistic form, their perceptions about the role that randomness plays in the determination of grades. To provide some structure to our analysis, we seek guidance from the prominent Bayesian model under which the survey questions above have a precise interpretation. However, as discussed in detail later, it is not our objective to provide a formal test of any particular model of learning.

Our ultimate goal is to provide findings of direct relevance to policymakers. As one example, considerable recent attention has been paid to understanding the role that the option value of schooling plays in the college entrance decision and college drop-out decision (Manski, 1989; Altonji, 1993; Cunha et al. 2005; Stange, 2007). This option value arises because, when uncertainty exists about ability at the time of entrance, students benefit from a system in which they decide sequentially (e.g., on a semester-by-semester basis) whether or not to stay in college as uncertainty is resolved. As such, the estimated importance of the option value will depend critically on how a researcher characterizes subjective beliefs about ability at the time of entrance. In the absence of direct evidence about subjective beliefs, researchers have frequently invoked a Rational Expectations (RE) assumption in which individuals' beliefs about a particular outcome coincide with the actual distribution from which that outcome is drawn (Das and van Soest, 2000). Consistent with the notion that this somewhat arbitrary assumption may can be problematic, we find that, at the time of college entrance, students

tend to substantially underestimate the possibility of poor grade performance.<sup>3</sup> This finding suggests that students are likely to perceive (at the time of entrance) that the option value of schooling is substantially smaller than would be predicted using a Rational Expectations assumption.

However, from a policy standpoint, the largest contribution of this work is made possible by the fact that, because we designed the BPS to provide direct evidence about all avenues through which theory suggests that attrition could arise, the data potentially allow the most comprehensive, direct look at drop-out to-date.<sup>4</sup> Given that respondents are interviewed approximately ten to twelve times each year, with the first survey taking place immediately after students arrive on campus in their freshmen year, the benefits of the data are especially prominent if one is interested in quantifying the role of learning per se.

Our findings about beliefs at the time of entrance highlight perhaps the most important potential use of expectations data - using these data to reduce the reliance of arbitrary assumptions (e.g., Rational Expectations) in the estimation of behavioral models.<sup>5</sup> However, because the estimation of a full learning model of the dropout decision is beyond the scope of this paper, we estimate simple models of drop-out that take advantage of the unique nature of our data. We estimate both models in which independent variables are derived from a student's stock of information at the time of the drop-out decision and models in which independent variables are derived from the amount that the student has learned by the time of the drop-out decision. We describe

<sup>&</sup>lt;sup>3</sup>This is consistent with previous work, including Madeira (2007) who examines the evolution of parental beliefs about the academic ability of their elementary, middle-school, and high school children using a survey question of the form "Please guess the marks that your child will get in Reading and Mathematics on the first report card this fall: Excellent, Good, Satisfactory, Unsatisfactory."

<sup>&</sup>lt;sup>4</sup>While the literature on college drop-out across disciplines is immense, many fundamental questions remain unanswered. In the most referenced work across disciplines, Tinto (1975) suggests that whether a students drops out is strongly related to the student's degree of academic and social integration. However, while this theory has received much attention, Draper (2005) writes "It is less clear whether there is much direct empirical support for it, and certainly it is hard to find direct empirical tests of and challenges of it." The problem is primarily related to data limitations; it is often hard to even tell exactly what information is being used to proxy for social and academic integration (Draper, 2005; Beekhoven, 2002).

<sup>&</sup>lt;sup>5</sup>The use of expectations data in the estimation of behavioral models is quite limited. Wolpin (1999), van der Klaauw (2000), and van der Klaauw and Wolpin (forthcoming) take advantage of expectations data in a different manner than what would be suggested by the data collected here - they show that self-reported expectations about future outcomes of interest can be used to increase the precision of estimators in models that make standard assumptions about beliefs.

the theoretical relationship between the two types of models, noting that the theory related to the latter (learning) models implies that they must contain information about how far from the margin of indifference (between being in and out of school) a student is at the time of college entrance. Our results indicate that learning about ability plays a particularly important role in the drop-out decision. Thus, although students may not likely anticipate a substantial option value at the time of entrance, the actual value of a higher education system in which students can make decisions sequentially is likely to be quite large. This dichotomy would, in essence, be assumed away by a standard Rational Expectations assumption. Among other possible factors of importance, our results show that, while students who find school to be unenjoyable are unconditionally much more likely to leave school, this effect arises to a large extent because these students also tend to receive poor grades.

While interactions with respondents make us confident that students were comfortable with the survey questions we used to elicit expectations, it is worth noting that it will never be possible to directly examine how accurately self-reported expectations data represent a person's true beliefs. Instead, confidence in the usefulness of this sort of data is best accumulated by examining, as in Manski (2004), its performance across a variety of substantive contexts. As such, by and large, we take as our starting point that useful information about subjective beliefs can be elicited from carefully worded survey questions. Nonetheless, our findings provide some of the strongest evidence to-date for this starting point. Simple theory related to the drop-out decision suggests that both a person's actual grade point average in the first year and the person's beliefs about future grade performance at the end of the first year should be important predictors of whether a person returns to college after the first year. We find that this theoretical implication is satisfied when we measure beliefs about future grade performance directly using self-reported expectations data. However, this theoretical implication is not satisfied when we construct beliefs about future grade performance using a standard Rational Expectations assumption.

In Section II we describe the Berea Panel Study. Section III examines beliefs at the time of entrance with a discussion of what a departure from Rational Expectations implies about the option value of schooling. Section IV examines expectations formation. In Section V we examine empirical models of the drop-out

decision. In Section VI we pay careful attention to students at the bottom of the grade distribution with a focus on examining whether these students understand the reasons for their poor grade performance. We conclude in Section VII.

### II. The Berea Panel Study and background work

Berea College is located in central Kentucky and operates under a mission of providing an education to students from low income families. The BPS baseline surveys were administered to students in the first BPS cohort (the 2000 cohort) immediately before they began their freshman year classes in the fall of 2000 and were administered to students in the second BPS cohort (the 2001 cohort) immediately before they began their freshman year classes in the fall of 2001. The baseline surveys took advantage of recent advances in survey methodology (see e.g., Barsky et al., 1997, Dominitz, 1998, and Dominitz and Manski, 1996 and 1997) to collect information about students' expectations towards uncertain future events and outcomes, including academic performance, that could influence the drop-out decision. Substantial follow-up surveys were administered at the beginning and end of each subsequent semester to provide information about how expectations change over time, and time-use information was collected four times each semester using the 24-hour time-diaries shown in the Appendix A (Stinebrickner and Stinebrickner, 2004; Stinebrickner and Stinebrickner, forthcoming A).

Berea offers a full tuition (and a room and board) subsidy to all entering students. Stinebrickner and Stinebrickner (2003) (hereafter referred to as S&S) found that, despite the fact that the direct costs of students at Berea are approximately zero (and perhaps negative), fifty percent of students do not graduate. S&S (forthcoming B) found that, in the aggregate, difficulties borrowing money to pay for consumption during school also do not play a particularly important role at Berea. Thus, the large majority of attrition at Berea should be attributed to factors unrelated to short-term credit constraints. As suggested in Manski (1989), Altonji (1993), Carneiro et al. (2005), and Cunha et al. (2005) learning about ability plays a central role in other prominent explanations for drop-out, and the BPS was designed specifically to allow an examination of this

<sup>&</sup>lt;sup>6</sup>S&S(2003) find that these drop-out rates are generally similar to those of low income students elsewhere.

possibility.

We focus on the 2001 BPS cohort because certain questions of interest for this paper were added after the 2000 cohort completed its first year at Berea. Because our interest in learning requires that we observe expectations about grade performance at two different points in time, we focus on students in the 2001 cohort who answered both our baseline survey and the survey that took place immediately before the beginning of the second semester. Three hundred seventy-five of 420 (89%) matriculating students in this cohort completed our baseline survey. Three hundred twenty-five of these students were still in school at the beginning of the second semester and answered the survey which took place at that time.

# III. Beliefs at the time of entrance: evidence about the validity of Rational Expectations and the importance of the option value of schooling

We observe each student's first semester grade point average measured on a four point scale,  $GPA_{Ii}$ , directly from administrative data. We elicit each student's subjective beliefs about the distribution of  $GPA_{Ii}$  using survey Question A.2 (Appendix A) which was administered immediately before the start of classes in the first semester. Paying close attention to methodological suggestions in Dominitz (1998) and Dominitz and Manski (1996, 1997), the question asks each student to report the "percent chance" that his  $GPA_{Ii}$  will fall in each of a set of mutually exclusive and collectively exhaustive categories, conditional on an expected level of study effort for the first semester,  $Expected\ STUDY_{Ii}$ , that is elicited in Question A.1 (Appendix A).

Column 1 of Table 1A shows the subjective probabilities of first semester grade point average,  $GPA_{Ii}$ , from Question A.2 averaged over the 325 students in our sample. Column 2 of Table 1A shows the proportion of students in the sample whose actual  $GPA_I$  falls in each category. Comparing Column 1 and Column 2 of Table 1A reveals that, on average, individuals are too optimistic about their grade performance. For example,

<sup>&</sup>lt;sup>7</sup>The response rate on the survey at the beginning of the second semester was approximately 94%. This implies that approximately half of the 50 students who answered the baseline survey but did not answer the survey at the beginning of the second semester had left school by the beginning of the second semester. While it would also be of interest to know how the beliefs about academic ability changed for these students (who are not in the sample used in this paper) between the time of entrance and the time of departure, the period between the beginning of the second semester and the beginning of the second year is the period with the highest amount of attrition.

on average, individuals believe that the probability of obtaining a  $GPA_{Ii}$  between 3.5 and 4.0 is .401 while, in reality, only .234 of students actually receive a  $GPA_{Ii}$  in this range. Similarly, on average, individuals believe that the probability of obtaining a  $GPA_{Ii}$  of less than 2.0 is .037 while, in reality, .141 of students actually receive a  $GPA_{Ii}$  in this range. A chi square goodness-of-fit test rejects the null hypothesis that the distribution in Column 2 is obtained by sampling from the distribution in Column 1at all traditional levels of significance (chi square statistic = 139.8 with 5 d.f.)<sup>8</sup>. Hence, the evidence does not support a Rational Expectations assumption.

The result that students assign too little probability to the poor academic outcomes arises as a combination of the fact that the mean of a student's subjective  $GPA_{Ii}$  distribution is, on average, biased upwards and the fact that, on average, students are more certain about their first semester grade point average than would be suggested by a Rational Expectations assumption. With respect to the former, we compute an approximate value of  $E(GPA_{Ii})$  for each person i from survey Question A.2 by assuming that the grade density is uniform within each of the mutually exclusive and collectively exhaustive grade categories. Referring to this approximate mean as  $prior\_mean_i$ , the second to last entry in Column 1 of Table 1A shows that the average  $prior\_mean_i$  for the 325 students in our sample is 3.22 while the last entry of Column 2 shows that the average actual  $GPA_{Ii}$  is 2.88. With respect to the latter, the last entry in Column 1 of Table 1A shows that, on average, the standard deviation of a student's subjective  $GPA_{Ii}$  distribution is .532 while the last entry of Column 2

<sup>&</sup>lt;sup>8</sup>It may be desirable to combine the intervals [1.0, 2.0) and [0.0,1.0) before performing the Chi square test but this has no effect on the conclusions from the test.

The need to approximate is a disadvantage of using our information about the entire grade distribution to compute the expected value. Nonetheless, because we are confident that students are comfortable answering our Question A.2 about the entire grade distribution and because we find it appealing to compute expected values ourselves (rather than relying on students to use an appropriate definition of "expected" grade performance), we believe that, on net, the use of this question is desirable for our purposes. Hereafter, we largely ignore the presence of approximation error in the construction of this variable.

 $<sup>^{10}</sup>$ A standard normal test rejects the null that the average  $prior\_mean_i$  is the same as the average actual  $GPA_{Ii}$  at all traditional significant levels (test statistic = 8.05).

shows that the standard deviation of actual  $GPA_{li}$  is .784. 1112

Tables 1B and 1C, which show the results in Table 1A separately by whether a student's high school grade point average,  $HSGPA_i$ , is above or below the median in our sample, show that the RE assumption may be particularly problematic for some subgroups. The last rows of Table 1B show that the difference between the average  $prior\_mean_i$  in the sample and the average actual  $GPA_{li}$  in the sample is .11 for students whose high school grade point average is above the median (3.275 vs. 3.164). However, the last rows of Table 1C show that this difference is .58 for students whose high school grade point average is below the median (3.161 vs. 2.579).<sup>13</sup>

One possible definition of the option value of college attendance is the difference between the discounted expected utility of entering college under the current system in which a student decides sequentially (e.g., on a semester-by-semester basis) whether or not to remain in college and the discounted expected utility of entering college under the counterfactual in which a student who enters college must precommit to remaining in school until graduation.<sup>14</sup> Thus, the option value will be lower if the student is more certain at the time of college entrance that he will graduate from college. The strong relationship consistently found in previous literature (and in Section V of this paper) between academic performance and college drop-out suggests that students who obtain high grades are typically well above the margin of indifference between being in and out

<sup>&</sup>lt;sup>11</sup>As above, we compute this latter number by assuming that the grade density is uniform within each of the mutually exclusive and collectively exhaustive grade categories in Question A.2.

<sup>&</sup>lt;sup>12</sup>While the evidence in the previous paragraph suggests that  $prior\_mean$  is far from a perfect predictor of first semester grade point average, we do find that  $prior\_mean$  contains information about actual grade performance. In particular, in results not shown, regressing  $GPA_{Ii}$  on  $prior\_mean_i$  yields an estimated coefficient (standard error) of .394 (.147) for  $prior\_mean_i$ . As can be seen in Table 2, the correlation between  $GPA_{Ii}$  and  $prior\_mean_i$  is .147.

 $<sup>^{13}</sup>$ For students above the median, a standard normal test rejects the null hypothesis that the average  $prior\_mean_i$  is the same as the average actual  $GPA_{Ii}$  at only significant levels greater than .134 (t-statistic = 1.50) while, for students below the median, this null hypothesis is rejected at all traditional significance levels (t-statistic = 8.00). For students above the median, a chi square test easily rejects the null hypothesis that the distribution in Column 2 is obtained by sampling from the distribution in Column 1 at all traditional significance levels (chi square statistic = 83.5 with 5 d.f.). A chi square test rejects the same null hypothesis for students whose high school grade point average is above the median at a significance level of .10 but not at .05. (Chi square statistic=10.0 with 5 d.f.).

<sup>&</sup>lt;sup>14</sup>This definition of the option value focuses on uncertainty about ability (Stange, 2007). One could also include in this definition the continuation value of starting school which is present even if no uncertainty exists about ability (Heckman, Lochner, and Todd, 2006; Heckman and Navarro, 2007).

of college. Then, our finding that even students who have relatively high actual probabilities of performing poorly believe that bad grade outcomes are very unlikely suggests that students may be substantially more certain than they should be that they will graduate from college. This suggests that students may perceive the option value to be substantially lower than what would be suggested by a Rational Expectations assumption.

From an additional survey question on the baseline BPS survey, we find additional, direct evidence that students are likely to perceive the option value to be substantially lower than what would be suggested by a Rational expectations assumption:

**Question B** What is the percent chance that you will eventually graduate from Berea College?

While approximately 60% of entering students in this cohort eventually graduate from Berea, on average, students believe that there is an 86% chance that they will graduate from Berea. <sup>15</sup> In Section V we discuss what simple theory implies about the usefulness of this question in empirical models that examine the effect of learning on college drop-out. We find empirical evidence that the answers to this question contain valuable information.

Recent evidence that families of potential college students tend to substantially overestimate the direct costs of college has led to a concern that inaccurate information may lead to a situation where too few students choose to enter college (NCES, 2003). The evidence here raises the very real possibility that more accurate information about academic performance in college could lead to a situation where some students who are currently entering might decide not to enter.<sup>16</sup>

# **IV.** Expectations formation

To study expectations formation, we pay close attention to the quote in the Section I from Manski (2004). We first measure revisions to expectations by comparing beliefs at the time of entrance to beliefs at a

<sup>&</sup>lt;sup>15</sup>The graduation rates at Berea increased slightly from the earlier periods studied in S&S (2003).

<sup>&</sup>lt;sup>16</sup>It is impossible to know from our data what effect accurate information would have on people who have chosen not to enter college. It is very possible that some of these people are incorrectly pessimistic about their college academic performance.

later point in time (Section IV.A). We then relate these revisions to new observed event information (Section IV.B). Finally, we take advantage of unique survey questions which elicit information about how students perceive their environments in an attempt to understand differences in revisions across students (Section IV.C).

IV.A. Measuring revisions to expectations: Comparing beliefs at the beginning of the 1<sup>st</sup> and 2<sup>nd</sup> semesters.

We examine revisions to expectations by comparing a student's subjective beliefs about  $GPA_{Ii}$  from Section III to the student's subjective beliefs about second semester grade point average,  $GPA_{2i}$ . The latter beliefs were elicited using Question A.4 (Appendix A) which was administered immediately before the start of classes in the second semester. The question asks each student to report the "percent chance" that his  $GPA_{2i}$  will fall in each of a set of mutually exclusive and collectively exhaustive categories, conditional on an expected level of study effort in the second semester,  $Expected\_STUDY_{2i}$ , that is elicited in Question A.3 (Appendix A). We discuss the issue of course difficulty in more detail later, but note here that, due in large part to a substantial number of required core courses, overall course difficulty is generally similar between the first and second semesters.

Table 3 shows the subjective probabilities of  $GPA_{2i}$  from Question A.4 averaged over the 325 students in our sample. In much of the remainder of this paper we focus our attention by examining revisions to mean expectations. We compute an approximate value of  $E(GPA_{2i})$  for person i from survey Question A.4 in the same way as described in Section III. Referring to this approximate mean as  $posterior\_mean_i$ , the second-to-last entry of Table 3 shows that  $posterior\_mean_i$  has an average value in our sample of 3.14. Comparing, in the first Column of Table 4, this value to the average  $prior\_mean_i$  of 3.22 from Section III reveals that, on average, individuals are significantly revising their beliefs about grade performance between the start of the first semester and the start of the second semester. Further, this comparison tends to mask the full degree of updating that is taking place at the individual level since some individuals revise their beliefs upwards while others revise their beliefs downwards; on average, the absolute value of the difference between  $prior\_mean_i$  and

<sup>&</sup>lt;sup>17</sup>A test of the null hypothesis that the average mean in the population is the same between the two semesters is rejected at traditional significance levels (p-value .0017).

posterior mean, is .289.

# IV.B. Relating revisions of expectations to new information: Is there evidence that both $GPA_{Ii}$ and $prior\_mean_i$ influence $posterior\_mean_i$ ?

The new event information that we consider is a person's first semester grade point average,  $GPA_{Ii}$ . Stratifying the sample by  $GPA_{Ii}$  we find sensible patterns of updating by  $GPA_{Ii}$  group. For example, in the second column of Table 4 we show the average  $prior\_mean_i$  and the average  $posterior\_mean_i$  for students in the top-third of the actual  $GPA_{Ii}$  distribution. In the third column of Table 4 we show the average  $prior\_mean_i$  and the average  $posterior\_mean_i$  for students in the bottom third of the actual  $GPA_{Ii}$  distribution. While the difference in the average  $prior\_mean_i$  between the groups is relatively small, .08, the difference in the average  $posterior\_mean_i$  between the groups becomes substantial, .50, after the groups observe large differences in first semester grade performance. <sup>18</sup>

More generally, we find strong evidence that  $prior\_mean_i$  and  $GPA_{li}$  are both important predictors of posterior  $mean_i$  when we use our entire sample to estimate a regression of the form

(1) posterior mean<sub>i</sub> = 
$$\beta_0 + \beta_1$$
 prior mean<sub>i</sub> +  $\beta_2$  GPA<sub>1i</sub> +  $\mu_i$ 

by Ordinary Least Squares (OLS). Specifically, Column 1 of Table 5 shows that the point estimates (standard errors) for  $\beta_1$  and  $\beta_2$  are .396 (.051) and .245 (.019), respectively. A test of the null hypothesis that  $\beta_1$ =0 and a test of the null hypothesis that  $\beta_2$ =0 are each overwhelmingly rejected with the tests having t-statistics of 7.76 and 12.89, respectively. In the remainder of this section we focus primarily on understanding why particular revisions are observed. In Section VII we return to the question of whether the quality of forecasts improve over time.

### IV.C. Understanding heterogeneity in revisions across students

In this section we pose our desire to understand why certain revisions are observed as a question of why substantial heterogeneity exists in  $posterior\_mean_i$  for students who have the same  $prior\_mean_i$  and  $GPA_{li}$ . We stress that it is not our objective to provide a formal test of any specific model of learning. Indeed, as will

<sup>&</sup>lt;sup>18</sup>A test of the null hypothesis that the difference in the average posterior means between the groups is the same as the difference in the average prior means between the groups is overwhelmingly rejected.

be discussed in detail later, doing so will always be extremely difficult using real-world data. <sup>19</sup> Moreover, doing so runs somewhat counter to perhaps the strongest rationale for collecting survey data on expectations - that it allows researchers estimating behavioral models to move away from somewhat arbitrary, untestable assumptions about the learning process in favor of relatively flexible specifications for the learning process. Nonetheless, given that little is understood about the process governing expectations formation in the real world, we turn to a particularly prominent model of learning - the Bayesian model - to obtain some suggestions about the type of environmental perceptions that might influence the manner in which a person updates his beliefs. We stress that there are other theories of learning, not studied here, that might also provide valuable insight. As such, our work is best viewed as one step towards an understanding of expectations formation.

In Section IV.C.1 and IV.C.2 we describe a textbook Bayesian model of learning and revisit the results of Section IV.B after applying some modifications suggested by the discussion. In Sections IV.C.3 and IV.C.4 we examine whether survey questions motivated by the theory in Section IV.C.1 can provide information about how individuals perceive their environment that is useful for understanding heterogeneity in updating.

## IV.C.1. A Simple Bayesian Learning Framework

## A textbook model

Suppose that student i's grade point average on a 4.0 scale in a particular semester t is determined by the sum of a constant  $\theta_i$  that represents his person-specific, permanent academic ability (or academic preparation that is essentially permanent in nature at time of college entrance) and a random variable  $\varepsilon_i^{20}$ ;

<sup>&</sup>lt;sup>19</sup>Providing evidence about whether agents update in, for example, a Bayesian manner has been the focus of experimental work.

<sup>&</sup>lt;sup>20</sup>Although we tend to refer to  $\theta_i$  as ability, we note that it is not necessarily literally true that changes in beliefs about  $\theta_i$  entirely represent changes in beliefs about ability. For example, a person who enters school with a belief that the mean of the school's grade distribution is higher than it actually is may reduce his beliefs about his own mean GPA after the time of entrance even if he has not learned anything about his ability relative to other students at the school. However, what seems important is that, if misperceptions about average grades at the school exist at the time of entrance, they are not likely to be overly large. Thus, students of perhaps the most policy interest - those who perform very poorly - are almost certainly learning primarily about something other than their relative position in the grade distribution when they see grades that are much lower than expected. Regardless, if one wanted to be technically correct,  $\theta_i$  would be referred to as mean GPA for person i, and for much of this paper the distinction is not of particular importance.

(2) 
$$GPA_{ti} = \theta_i + \varepsilon_{ti}$$
.

For ease of exposition, we begin by describing the model under assumptions about  $\epsilon_{ti}$  that would be standard in a textbook version of the learning model - that  $\epsilon_{ti}$  represents semester randomness in grades that is purely idiosyncratic in nature and is not observed directly by students. However, we stress that it is the reality that this assumption should not be taken literally which makes it difficult to provide formal tests of any particular learning model using real-world data. We discuss this issue in detail in the next subsection. For now, we begin by thinking of  $\epsilon_{ti}$  as capturing semester-specific random "luck" and note that, as will be discussed later, we use the term "luck" in some of our survey questions.

In order to decrease the dimension of the learning problem, we follow previous work in assuming that students beliefs about  $\varepsilon_{ii}$  are fixed for all t.<sup>21</sup> Thus, all learning takes place because a student is uncertain about his academic ability,  $\theta_i$ . The person's beliefs about  $\theta_i$  at the beginning of the first semester (t=1) are characterized by a distribution that we refer to simply as the "prior" distribution. We refer to the mean and variance of this prior distribution of ability for person i simply as the prior mean and prior variance. After the first semester, the student observes his first semester grade point average  $GPA_{Ii} = \theta_i + \varepsilon_{1i}$  which represents a noisy signal of academic ability. This noisy signal (hereafter often referred to simply as the "signal") is used in conjunction with the prior distribution of  $\theta_i$  and knowledge of the distribution of  $\varepsilon_{ti}$  to generate the distribution that characterizes the person's beliefs about  $\theta_i$  at the beginning of the second semester (t=2). Hereafter, we refer to this t=2 distribution simply as the "posterior" distribution. We refer to the mean and variance of this distribution of ability for person i simply as the posterior mean and posterior variance respectively.

From our survey questions it is not possible to identify the entire prior distribution or the entire posterior distribution of  $\theta_i$ . For example, because survey question A.2 (Appendix A) elicits information about  $GPA_{Ii} = \theta_i + \epsilon_{1i}$ , an attempt to identify the prior distribution of  $\theta_i$  from this question is confounded by the presence of  $\epsilon_{1i}$ .

<sup>&</sup>lt;sup>21</sup>These beliefs may or may not be correct. What is important is that students are not learning about the distribution of  $\epsilon_{ti}$ .

with the problem being that a student's beliefs about  $\varepsilon_{1i}$  are not observed.<sup>22</sup> Similarly, because survey question A.4 (Appendix A) elicits information about  $GPA_{2i} = \theta_i + \varepsilon_{2i}$ , an attempt to identify the posterior distribution of  $\theta_i$  from this question is confounded by the presence of  $\varepsilon_{2i}$ . However, under the assumption that students believe that  $E(\varepsilon_{ti})=0$ , we can identify the prior mean and posterior mean from the survey questions A.2 and A.4 since this assumption implies that  $E_{t=1}(GPA_{1i})=E_{t=1}(\theta_i)$  and  $E_{t=2}(GPA_{2i})=E_{t=2}(\theta_i)$ .<sup>23</sup> The construction of  $E(GPA_{1i})$  and  $E(GPA_{2i})$  from the survey questions was described in Sections III and IV, respectively, with these variables being referred to in these sections as *posterior mean<sub>i</sub>* and *posterior mean<sub>i</sub>*, respectively.

Thus, while one would be interested, most generally, in how the entire posterior distribution of ability evolves, we focus our attention (in part out of necessity) by attempting to understand how the posterior mean evolves. Under Bayesian updating, a convenient linear form arises for the posterior mean under the assumption that a student's prior distribution of  $\theta_i$  is normally distributed and the assumption that the student believes that  $\epsilon_{ti}$  is normally distributed. Specifically, under these assumptions and the assumption that the student believes that  $E(\epsilon_{ti})=0$ :

(3a) 
$$posterior\_mean_i = W_{1i} \cdot prior\_mean_i + W_{2i} \cdot GPA_{1i}$$

(3b) 
$$= prior_mean_i + W_{2i} \cdot (GPA_{li} - prior_mean_i)$$

$$= GPA_{li} - W_{li} \cdot (GPA_{li} - prior\_mean_i),$$

where, denoting  $\sigma^2_{\ \epsilon i}$  to be the variance of  $\epsilon_{ti}$ ,

(3d) 
$$W_{1i} = \frac{\sigma_{\epsilon i}^{2}}{\sigma_{\epsilon i}^{2} + prior\_variance_{i}}, \quad W_{2i} = \frac{prior\_variance_{i}}{\sigma_{\epsilon i}^{2} + prior\_variance_{i}}.$$

<sup>&</sup>lt;sup>22</sup>One might think about estimating the distribution of  $\varepsilon_{li}$  to use in place of beliefs about  $\varepsilon_{li}$ . However, assuming that beliefs correspond to the true distribution is not particularly appealing given that the objective of this paper is to examine what can be learned directly from the data without making such assumptions. Further, beliefs constructed in this way would be contradicted directly by the data; for some individuals, the variance of  $\theta_i + \varepsilon_{li}$  elicited in survey question A.2 would be smaller than the estimated variance of  $\varepsilon_{li}$ .

<sup>&</sup>lt;sup>23</sup>Roughly speaking, the assumption is that, on average, luck is neither good nor bad. Grades are bounded on the [0.0, 4.0] grade interval. While this reality does suggest that certain distributions for ε cannot be literally correct, it does not by itself directly lead to a contradiction of the assumption that a person believes that  $E(ε_{ti})=0$ . Nonetheless, we stress that this is an assumption made largely for convenience with the objective of keeping the subsequent analysis and interpretation as simple/transparent as possible.

Thus, in the Bayesian model, heterogeneity exists in  $posterior\_mean_i$  for students with the same  $GPA_{Ii}$  and  $prior\_mean_i$  because different individuals assign different weights to  $GPA_{Ii}$  and  $prior\_mean_i$ . Specifically, equations (3a) and (3d) indicate that a person puts more (less) weight on his prior mean (first semester grades) if he thinks that the signal is noisy and puts more (less) weight on his first semester grades (prior mean) if he is uncertain about his ability at the time of entrance.

### Difficulties in this context with the textbook model

Unfortunately, in any real-world application, the possibility of model misspecification will make it very difficult to construct credible, formal tests of whether individuals update in a strictly Bayesian manner. Perhaps most obviously, it is natural to question functional form assumptions, for example whether normality is a good approximation or whether linearity is appropriate for the grade equation (2).

However, as alluded to after equation (2), there will also typically be plenty of scope for questioning other assumptions about  $\varepsilon_{ii}$ . For example, implicit in the textbook model associated with equation (2) is an assumption that a student does not observe elements of  $\varepsilon_{Ii}$  directly when reporting his prior mean or forming his update and does not observe elements of  $\varepsilon_{2i}$  directly when reporting his posterior mean.<sup>24</sup> In practice this assumption will almost certainly be violated, in part because students are likely to sometimes observe random elements of  $\varepsilon$  and in part because some elements of  $\varepsilon$  are likely to be determined directly by the choices of individuals.

If a researcher knew which elements of  $\varepsilon$  were known to the individual and could measure these elements, then the researcher could adjust equation (2) to reflect this information. While it is important to be cognizant of the reality that it will never be possible to observe everything contained in a person's information set at a particular time, our data are helpful in this respect. Most notably, the data are unique in providing information about both expected study effort and actual study effort that S&S(forthcoming A) shows is perhaps the most important input in the grade production function. As a result, we can modify equation (2) to explicitly model the role that  $STUDY_{ii}$ , the average number of hours a person studies per day in semester t, plays in the

<sup>&</sup>lt;sup>24</sup>Also implicit is an assumption that  $\varepsilon_{ti}$  is uncorrelated across time.

determination of grades, thereby removing the influence of effort from the unobservable  $\varepsilon$ :

(4) 
$$GPA_{ti} = \alpha_i STUDY_{ti} + \theta_i + \varepsilon_{ti}$$
,

where  $\alpha_i$  is the causal effect of studying one additional hour per day on academic performance for person i.

Subsequently, we focus most of our attention on results based on this specification which includes study effort. However, as described in the next section, including study effort in our specifications requires that we move away from the simplest OLS estimator. As a result, we also show results using the most parsimonious possible model without study effort (equation 2).<sup>25</sup>

### IV.C.2. Revisiting the results from Section IV.B in a model with study effort

The goal of Sections IV.C.3 and IV.C.4 is to provide some understanding of why heterogeneity exists in the posterior mean for people with the same prior mean and the same first semester grade point average. For the most parsimonious model without study effort (equation 2), the starting point is equation (1) of Section IV.B. In order to form our starting point for the model with study effort (equation 4), we examine a modification of equation (1) that takes into account study effort.

The motivation for taking into account study effort can be seen directly in two observations. First, a simple comparison of  $prior\_mean_i$  to  $posterior\_mean_i$  would produce a misleading view of what a person has learned about his academic ability if the number of hours that the person expects to study per day in semester  $t(Expected\_STUDY_{ii})$  changes between the beginning of the first semester (t=1) and the beginning of the second semester (t=2). Second, a simple comparison of  $GPA_{Ii}$  to  $prior\_mean_i$  would provide misleading information about the accuracy of a person's prior beliefs about his ability if a person's study effort in the first semester was different than what he expected at the beginning of the first semester.

What is needed are modified measures of the prior mean, posterior mean, and first semester grade performance in which study effort is held constant. Here we hold effort constant at  $Expected\_STUDY_{li}$ , the study effort on which a person conditions when answering the question from which we construct  $prior\_mean_i$ 

<sup>&</sup>lt;sup>25</sup>Our data could also be used to deal with other possible concerns about  $\varepsilon_{ii}$  that are similar in spirit. For example, because we have transcript data for all students, we could, in theory, adjust equation (4) further to take into account information about course difficulty.

(Question A.2). Our modified measure  $posterior\_mean_i^*$  represents the posterior mean that the student would have reported if he had planned to study  $Expected\_STUDY_{li}$  rather than  $Expected\_STUDY_{2i}$  in the second semester, and our modified measure  $GPA_{li}^*$  represents the first semester grade point average the person believes he would have received if he had studied  $Expected\_STUDY_{li}$  rather than  $STUDY_{li}$  in the first semester. Assuming that students believe that study effort affects grade performance in the manner described in equation (4):

(5) 
$$posterior\_mean_i^* = posterior\_mean_i - \alpha_i (Expected\_STUDY_{2i} - Expected\_STUDY_{1i})$$

(6) 
$$GPA_{li}^* = GPA_{li} - \alpha_i (STUDY_{li} - Expected\_STUDY_{li})$$
.

For simplicity we assume that  $\alpha_i$  is homogeneous and use an estimate of .36 that is obtained in S&S (forthcoming A) from an identification strategy in which we exploit exogenous variation in a student's study effort created by whether his randomly assigned roommate brought a video game to school. Expected\_STUDY<sub>1i</sub> and Expected\_STUDY<sub>2i</sub> are obtained from survey Question A.1 and Question A.3 (Appendix A), respectively. They have means (standard deviations) of 3.47 (1.32) and 3.07 (1.27), respectively, for the sample of 291 students who have valid observations for both of the expected-study variables. STUDY<sub>1i</sub> represents the average number of hours that a person actually studies per day over *all* days in the first semester and is not fully observed. What is observed is the number of hours that students studied on (up to) four particular days during the first semester on which we administered the 24-hour time diaries (Appendix A) mentioned in Section II. Defining N(i) to be the number of time diaries (out of four) that person i completed in the first semester and letting  $S_{ji}$ , j=1,..., N(i) represent the N(i) observed daily study amounts for person i, a noisy proxy  $ST\hat{UDY}_{1i}$  of  $STUDY_{1i}$  can be constructed as:

(7) 
$$ST\hat{U}DY_{1i} = \frac{1}{N(i)} \sum_{i=1}^{N(i)} S_{ji}$$
.

In this subsection, the objective will be to reestimate equation (1) after replacing  $posterior\_mean_i$  and  $GPA_{li}$  with the modifications  $posterior\_mean_i^*$  and  $GPA_{li}^*$  that account for study effort. In the next two

<sup>&</sup>lt;sup>26</sup>In S&S (forthcoming A) we find that a person's study effort is strongly influenced by the instrument and find strong support in favor of the validity of the instrument by using a wide variety of survey questions that were designed specifically to document possible violations of the exogeneity condition.

subsections, the goal will be to examine heterogeneity in weighting by expanding this specification to include additional interaction terms involving  $posterior\_mean_i^*$  and  $GPA_{Ii}^*$ . The difficulty that arises in these exercises is that the use of  $ST\hat{UD}Y_{1i}$  directly in the construction of  $GPA_{Ii}^*$  (equation 6) leads to an errors-in-variables problem in the independent variable  $GPA_{Ii}^*$ . This problem is not easily corrected using textbook methods both because the number of completed time diaries varies across individuals and because the specifications in the next two sections include multiple regressors that are functions of  $GPA_{Ii}^*$ . Further, the problem is potentially important; in the estimation of grade equations, S&S (2004) found that using  $ST\hat{UD}Y_{Ii}$  directly in place of  $STUDY_{Ii}$  can lead to a non-trivial attenuation bias.

To address this issue, we take a Maximum Likelihood (MLE) approach suggested by S&S(2004). This approach deals with the measurement error issue under the assumption that  $S_{ji}$  is given by the permanent/transitory process

$$(8) S_{ii} = \mu_i + \nu_{ii}.$$

The permanent component  $\mu_i$  represents the average amount that person i studies per day and the transitory component  $\nu_{ji}$  represents a daily deviation from this average amount. We assume that in the population  $\mu_i$   $\sim N(C,\sigma^2_{\mu})$ . We assume that  $\nu_{ji}$  is independent across both j and i and that  $\nu_{ji} \sim N(0,\sigma^2_{\nu})$ . Then, as described in detail in Appendix B, analogous to the MLE's derived in the missing data literature, the likelihood contribution for person i is the joint probability of *posterior\_mean\_i\** and the observed daily study amounts  $S_{1i}$ ,..., $S_{N(i)i}$ .

The MLE results of estimating equation (1) by maximum likelihood after replacing *posterior\_mean<sub>i</sub>* with *posterior\_mean<sub>i</sub>*\* and  $GPA_{Ii}$  with  $GPA_{Ii}$ \* are shown in Column 2 of Table 5. Both the test of the null hypothesis that  $\beta_1$ =0 and the test of the null hypothesis that  $\beta_2$ =0 continue to be overwhelmingly rejected with t-statistics of 11.098 and 8.091, respectively.

# IV.C.3 Evidence of predictable heterogeneity in weighting: the role of perceptions about the reasons for observed grade performance

In this section we examine whether a survey question motivated by the theory in Section IV.C.1 can predict differences in the posterior mean for students with the same prior mean and same observed first semester grade performance. As discussed earlier, we focus on the Bayesian explanation that heterogeneity may exist

in the weights assigned to the prior mean and to the noisy signal.

When  $posterior\_mean_i^*$  is used in place of  $posterior\_mean_i$  and  $GPA_{Ii}^*$  is used in place of  $GPA_{Ii}$ , equation (3b) becomes

(9) posterior mean,\*

= 
$$prior_mean_i + W_{2i} \cdot (GPA_{1i} *- prior_mean_i)$$

= prior 
$$mean_i + W_{2i} \cdot [GPA_{1i} - \alpha(STUDY_{1i} - Expected STUDY_{1i}) - prior mean_i]$$
,

where the second line follows from substituting equation (6). Thus, in the model with study effort,  $W_{2i}$  can be viewed as the proportion of the  $GPA_{Ii}$  -  $prior\_mean_i$  gap that the student believes will persist into the future after the portion of the gap that arises because a person studies a different amount than expected in the first semester is removed.

This formulation motivated the wording of Question C (Appendix A) which, between the first and second semesters, elicited individual perceptions about the percentage of the  $GPA_{Ii}$  -  $prior\_mean_i$  gap that should be attributed to each of the following: better or worse than expected ability (Line A), better or worse than expected preparation (Line B), higher or lower than expected study effort (Line C), and better or worse than expected luck (Line D).<sup>27</sup> An estimator of  $W_{2i}$  (and hence  $W_{1i}$ ) can then be constructed if one is willing to make assumptions about which of the reasons on Line A, Line B, and Line D tend to be viewed by the student as persisting beyond the first semester and which of these reasons tend to be viewed as being transitory in nature. Line C plays no role because the last line of equation (9) explicitly takes into account the role of different than expected study effort.<sup>28</sup>

Assuming that better or worse than expected ability and better or worse than expected preparation tend

<sup>&</sup>lt;sup>27</sup>In order to keep this experimental question manageable, if a student performed worse (better) than expected, the survey question abstracts from the possibility that the student might have found expected ability to be better (worse) than expected or expected preparation to be better (worse) than expected or study effort to be higher (lower) than expected or luck to be better (worse) than expected.

<sup>&</sup>lt;sup>28</sup>This is beneficial because it is hard to know whether students would tend to think of higher or lower than expected study effort as being permanent or transitory (although in theory our data would be helpful for this issue because it contains  $Expected\_STUDY_{li}$ ,  $Expected\_STUDY_{2i}$ , and information related to  $STUDY_{li}$ . Of course, equation (9) does make a specific assumption about students' beliefs about the role that higher or lower than expected study effort plays in their better or worse than expected grade performance.

to be viewed as persistent and that better or worse than expected luck tends to be viewed as transitory yields the following estimators for the weights:

(10) 
$$\hat{W}_{2i} = \frac{\% lineA + \% lineB}{\% lineA + \% lineB + \% lineD}, \qquad \hat{W}_{1i} = 1 - \hat{W}_{2i}.$$

Because equation (10) does not use information on Line C, estimates of the weights cannot be constructed if a student assigns a percentage of 100 to line C. Restricting our attention to the 211 students who did not have a percentage of 100 on line C and who correctly recognized in Question C.1 whether they had performed better or worse than expected in the first semester, we find substantial heterogeneity in students' interpretations of the reasons for the  $GPA_{Ii}$ -prior\_mean<sub>i</sub> gap; the mean and standard deviation of  $\hat{W}_{2i}$  are .688 and .408, respectively.<sup>29</sup>

To see whether the estimates from equation (10) can succeed in predicting differences in the posterior mean for students with the same prior mean and same first semester grade performance we use the MLE approach described in Section IV.C.2 to estimate

(11) posterior\_mean<sub>i</sub>\*=
$$\beta_0 + \beta_1$$
 prior\_mean<sub>i</sub>+  $\beta_2$  GPA<sub>1i</sub>\*+  $\beta_3$  prior\_mean<sub>i</sub> x  $\hat{W}_{1i} + \beta_4$  GPA<sub>1i</sub>\* x  $\hat{W}_{2i} + u_i$ .

This specification is estimated for the 191 students of our 211 student subsample who also provided valid information about study effort that is needed to construct  $posterior\_mean_i^*$  and  $GPA_{Ii}^*$ . In the second column of Table 6 we find that the point estimates of  $\beta_3$  and  $\beta_4$  are of quantitatively important size. For example, a student who believes that his better (or worse) than expected performance is caused entirely by persistent factors ( $\hat{W}_{2i}=1$ ) would put 2.1 times as much weight on  $GPA_{Ii}^*$  as someone who believes that his better (or worse) than expected performance is caused entirely by transitory factors ( $\hat{W}_{2i}=0$ ). Further, although

<sup>&</sup>lt;sup>29</sup>Two hundred forty-six of 325 correctly identified whether they did better or worse than expected with the majority of the incorrect responses coming for individuals whose performance was very close to what they expected. Thirty-five individuals attributed 100% of the gap to study effort.

<sup>&</sup>lt;sup>30</sup>The first column of Table 6 recomputes the results in the second column of Table 5 for the 191 person subsample.

as expected  $\beta_3$  and  $\beta_4$  are not estimated particularly precisely, a test of the null hypothesis that  $\beta_3$ =0 is rejected at significance levels greater than .038 (t-statistic=2.077) and a test of the null hypothesis that  $\beta_4$ =0 is rejected at significance levels greater than .031 (t-statistic=2.160). The second column of Table 7 shows generally similar results when the model without study effort is estimated by OLS.<sup>31</sup>

#### IV.C.4 Understanding heterogeneity in weighting at a more primitive level

In the previous section, we found that heterogeneity in updating conditional on a student's prior mean and first semester grade point average can be predicted by students' perceptions about the reasons that grade performance was better or worse than expected (Question C). However, obtaining a more satisfactory understanding of updating requires that we predict the heterogeneity in updating on the basis of more primitive environmental parameters/beliefs since the ultimate objective of understanding updating might be to allow the creation of a flexible specification for the learning process that could be embedded in a behavioral model. Put another way, it would be valuable to understand why differences exist in perceptions about the reasons that grade performance was better or worse than expected (Question C).

Our interest in the Bayesian model is that it provides guidance about primitive environmental parameters/beliefs that might lead to heterogeneity in weighting. For example, equation (3d) shows that  $W_{1i}$  has a positive partial derivative with respect to  $\sigma_{\epsilon i}^2$  and  $W_{2i}$  has a negative partial derivative with respect to  $\sigma_{\epsilon i}^2$ . This implication – that students who believe that the signal is particularly noisy put more weight on their prior mean and less weight on the signal – motivated question D (Appendix A) which at the beginning of the first semester asked individuals about the importance of "luck" in the determinations of grades, where we have attempted to define luck to contain a wide range of transitory factors that would be contained in  $\epsilon$ . If "luck" is roughly synonymous with  $\epsilon$  (and under the previous assumption that students believe that  $E(\epsilon_i)=0$ ), the sum

<sup>&</sup>lt;sup>31</sup>This involves estimating equation (11) with *posterior\_mean<sub>i</sub>* and  $GPA_{Ii}$  in place of *posterior\_mean<sub>i</sub>*\* and  $GPA_{Ii}$ \*. A student who believes that his better (or worse) than expected performance is caused entirely by persistent factors ( $\hat{\mathbf{W}}_{2i}$ =1) would put 1.6 times as much weight on  $GPA_{Ii}$  as someone who believes that his better (or worse) than expected performance is caused entirely by transitory factors ( $\hat{\mathbf{W}}_{2i}$ =0). A test of the null hypothesis that  $\beta_3$ =0 is rejected at significance levels greater than .087 (t-statistic=1.722) and a test of the null hypothesis that  $\beta_4$ =0 is rejected at significance levels greater than .069 (t-statistic=1.830).

of the responses on lines D.2 and D.3 of question D would represent a person's beliefs about  $Pr(\varepsilon_{ii} > .25|\varepsilon_{ii} > 0)$ . Although this interpretation of the sum of the responses on lines D.2 and D.3 is certainly not literally correct, the sum should be strongly (positively) correlated with  $\sigma_{\varepsilon i}^2$  so, acknowledging the obvious abuse of notation, we refer to the sum as  $\hat{\sigma}_{\varepsilon i}^2$ . In Table 8 we regress  $\hat{W}_{ii}$  on  $\hat{\sigma}_{\varepsilon i}^2$  and find strong evidence of the predicted positive relationship, with a test of the null hypothesis of no relationship being rejected at significance levels greater than .007. Thus, as would be suggested by the Bayesian model, we find evidence that beliefs about the reasons for better or worse than expected performance from Question C (and hence the estimates of the weights that are created directly from these reasons) are related to views about how much noise exists in the grade signal.<sup>33</sup>

It is useful to examine the influence of our proxy for  $\hat{\sigma}^2_{\epsilon i}$  directly. The last column of Table 6 shows results from the regression

(12) 
$$posterior\_mean_i^* = \beta_0 + \beta_1 prior\_mean_i + \beta_2 GPA_{Ii}^* + \beta_3 prior\_mean_i \times \hat{\sigma}_{\epsilon i}^2 + \beta_4 GPA_{Ii}^* \times \hat{\sigma}_{\epsilon i}^2 + u_i.$$

Consistent with the finding that the weights  $\hat{W}_1$  and  $\hat{W}_2$  are related to  $\hat{\sigma}^2_{\epsilon i}$ , a test of the null hypothesis that  $\beta_3$ =0 is rejected at significance levels greater than .042 (t-statistic=2.029) and a test of the null hypothesis that  $\beta_4$ =0 is rejected at significance levels greater than .018 (t-statistic=-2.376). The last column of Table 7 shows even stronger results when we estimate, by OLS, the analog to equation (12) that does not take into account study

<sup>&</sup>lt;sup>32</sup>This measure has a mean (standard deviation) of .511 (.250) for the sample of 211 students in Table 6.

<sup>&</sup>lt;sup>33</sup>Some additional care is perhaps warranted. In this regression we have not held  $prior\_variance_i$  constant. If  $\sigma_{\epsilon i}^2$  and  $prior\_variance_i$  were positively correlated, then Equation (3d) shows that the unconditional correlation between  $W_{1i}$  and  $\sigma_{\epsilon i}^2$  might not necessarily be positive even though there would exist a well-defined, positive relationship between  $W_{1i}$  and  $\sigma_{\epsilon i}^2$  when  $prior\_variance_i$  is held constant. However, we find no evidence that  $\sigma_{\epsilon i}^2$  and  $prior\_variance_i$  are positively correlated. Specifically, constructing a measure of  $prior\_variance_i + \sigma_{\epsilon i}^2$  from the uncertainty that a person exhibits about  $GPA_{li} = \theta_i + \epsilon_{li}$  in survey Question A.2, we find no evidence of a relationship between this measure and our proxy for  $\sigma_{\epsilon i}^2$ . If there was no relationship between  $\sigma_{\epsilon i}^2$  and  $\sigma_{\epsilon i}^2$  would be negatively correlated.

effort.34

Thus, by appealing to simple theory to help construct survey questions, we are able to identify an underlying, structural, environmental belief (directly related to the amount of noise in the signal) that predicts heterogeneity in updating conditional on  $prior\_mean_i$  and  $GPA_{Ii}^*$ . Not only is this understanding of interest in and of itself, but it would also allow a more useful specification for the learning process to be embedded in a behavioral model of the drop-out decision.

# V. Direct evidence about the reasons for drop-out.

The outcome variable we use, *dropout*<sub>i</sub>, is an indicator of whether students in our sample (who were all enrolled at the beginning of the second semester and typically finished the second semester) leave school before the beginning of the second year. We find that 17% (56 out of 325) of the students in our sample do so.

The drop-out decision at the end of the second semester can be written

(13)  $dropout_i = 1 \text{ iff } dropout_i = E_2(V_N) - E_2(V_S) > 0$ ,

where  $E_2(V_S)$  is the expected present value of lifetime utility at the end of the second semester (t=2) of returning to college and  $E_2(V_N)$  is the expected present value of lifetime utility at the end of the second semester of entering the workforce.

While, in reality,  $dropout_i$  represents the solution to a non-trivial dynamic programming problem, the estimation of a structural model of the drop-out decision is beyond the scope of this paper. Instead we employ straightforward linear probability models, so that the contribution of the analysis here comes entirely from the unique nature of the data. Although we have a specific interest in understanding the importance of learning, note that Equation (13) motivates specifications in which variables related to a person's state at the end of the second semester, rather than what the person has learned during the year, enter as independent variables. In terms of state variables related to academic ability, a student's cumulative grade point average at the end of the second semester (GPA Cumulative<sub>i</sub>) influences  $E_2(V_s)$  because it represents a student's current stock of grades

<sup>&</sup>lt;sup>34</sup>In this case, a test of the null hypothesis that  $\beta_3$ =0 is rejected at significance levels greater than .018 (t-statistic=2.384) and a test of the null hypothesis that  $\beta_4$ =0 is rejected at significance levels greater than .004 (t-statistic=-2.907).

and a student's beliefs about the mean of his ability/grade distribution at the end of the second semester influences  $E_2(V_s)$  because it represents beliefs about the future flow of grades.<sup>35</sup> The former is essentially just an average of  $GPA_{li}$  and  $GPA_{2i}$  and is obtained directly from administrative data. For the latter, we construct a self-reported version  $EOY_{mean_i}$  from an end-of-the-second-semester survey question that is identical to Survey Questions A.2 and A.4 except that it elicited beliefs about grades in the fall term of the second year by asking students to "assume that you return to Berea and that the classes you take in the fall term are of equal difficulty as those you took this year."<sup>36</sup>

Two hundred sixty-eight students have legitimate values of both  $EOY\_mean_i$  and  $GPA\_Cumulative_i$ . We remove an additional six people who had  $GPA\_Cumulative_i < 1.5$ , and, therefore, were forced to leave school. Regressing  $dropout_i$  on  $GPA\_Cumulative_i$  in the first column of Table 9 indicates that  $GPA\_Cumulative_i$  is a statistically significant and quantitatively important predictor of  $dropout_i$  (t-statistic = -5.772). Regressing  $dropout_i$  on  $EOY\_mean_i$  in the second column of Table 9 indicates that  $EOY\_mean_i$  is a statistically significant and quantitatively important predictor of  $dropout_i$  (t-statistic = -5.615). More importantly, Column 3 of Table 9 shows that  $EOY\_mean_i$  continues to be significant predictor of  $dropout_i$  (t-statistic -2.349) when it is included in a specification that also includes  $GPA\_Cumulative_i$  (t-statistic -2.673). Thus, the results when we use beliefs about academic ability/grade performance from self-reported expectations data are consistent with the theoretical implication that actual grade performance and beliefs about future grade performance should both influence the drop-out decision.  $^{37}$ 

The potential value of the self-reported expectations data becomes particularly evident after viewing the results in Column 4 of Table 9 in which the self-reported *EOY\_mean<sub>i</sub>* is replaced by a Rational Expectations estimate *RE\_EOY\_mean<sub>i</sub>* which is constructed as described in Appendix C. Inconsistent with the simple theory,

<sup>&</sup>lt;sup>35</sup>While the entire belief distribution could influence decisions, we focus on the mean of this distribution for simplicity.

<sup>&</sup>lt;sup>36</sup>This variable is constructed in the same manner as *prior mean*, and *posterior mean*, in Sections III and IV.

<sup>&</sup>lt;sup>37</sup>One could perhaps construct a scenario in which only current grade performance matters. However, given that future grade performance will account for roughly 75% of a person's final grade point average, it seems only natural to believe that both current grades and beliefs about future grades should be important determinants of drop-out.

the RE results indicate that beliefs about future grade performance do not affect the drop-out decision after conditioning on  $GPA\_Cumulative_i$ . The estimated effect of  $RE\_EOY\_mean_i$  is quantitatively very small (.021) and of an unexpected sign. In addition, both the effect of  $RE\_EOY\_mean_i$  and the effect of  $GPA\_Cumulative_i$  are estimated imprecisely. Intuitively, this occurs because  $GPA\_Cumulative_i$  both enters the drop-out specification itself and is forced to play a central role in the construction of  $RE\_EOY\_mean_i$ .

Of relevance for thinking about various explanations for drop-out that might serve as alternatives to explanations involving academic ability/grade performance, in S&S (forthcoming B) we find that, although credit constraints likely influence the drop-out decision of some students at Berea, the large majority of attrition is unrelated to credit constraints. We focus most of our attention here on the most prominent remaining explanation - that students often leave school because they find it to be stressful or unenjoyable - although we do return later to a discussion of other possible explanations.<sup>38</sup> At the end of the second semester, the BPS elicited information about how enjoyable person i found being in school relative to being out of school using Survey Question E (Appendix A). To simplify the discussion we ignore the qualitative nature of the variable and treat this information as a quantitative, continuous variable EOY enjoyability, which, in practice, has five possible values (1-5) with higher values representing a situation in which school is less enjoyable. Column 5 of Table 9 shows that students who find school to be unenjoyable are much more likely to drop out, with the estimated effect of EOY enjoyability, on dropout, being significant at .001. However, Table 10 shows that EOY enjoyability, is very strongly correlated with both GPA Cumulative, (p-value .0002) and EOY mean, (pvalue < .0001). In Column 6 of Table 9 we see that adding EOY enjoyability, to the specification that includes EOY mean, and GPA Cumulative, (Column 3 of Table 9) leaves the estimated effects of EOY mean, and GPA Cumulative, essentially unchanged. However, the estimated effect of EOY enjoyability, decreases by fifty percent relative to the specification in which EOY enjoyability, is included by itself (Column 5), with the pvalue associated with a test of the hypothesis of no effect for EOY enjoyability, increasing from .001 (Column

<sup>&</sup>lt;sup>38</sup>Our data collection is related to the notion that a first pass at collecting measures of social integration in Tinto (1975) might utilize a question along the lines of "Are you enjoying being at university?" Draper (2005).

5) to .079 (Column 6). Thus, while the enjoyability of school does itself appear to have some effect on dropout, much of the drop-out of unhappy students appears to arise because these students also tend to have had poor academic performance and tend to expect to have poor academic performance in the future.

From an interpretation standpoint, an important, open question is whether, to some extent, bad grade performance causes school to be unenjoyable or vice-versa. If unhappiness at school causes bad grade performance, one might expect this to be revealed primarily in substantial decreases in effort (relative to what was expected at the time of entrance) for those who find school to be unenjoyable. A test of the null hypothesis that students with higher values of  $EOY\_enjoyability_i$  study the same amount as other students during the first year can be rejected at significance levels .085.<sup>39</sup> However, the relatively small size of the estimated effect implies that the effect of unhappiness on grade performance through this route would be quite small.<sup>40</sup> Thus, it seems likely that bad grade performance itself is an important underlying cause of dropout.

The results from Column 6 of Table 9 can be used to quantify the role that learning plays in the drop-out decision. The average predicted value of  $dropout_i$  is .126.<sup>41</sup> To examine the amount of drop-out that would have been present if students had learned nothing about their ability after the time of entrance, we compute the average predicted value of  $dropout_i$  under the counterfactual assumptions that  $EOY\_mean_i$  is equal to  $prior\_mean_i$  and that  $GPA\_Cumulative_i$  is equal to  $prior\_mean_i$ . Under these assumptions we find that the average predicted drop-out rate decreases by 34% to .083. To examine the amount of drop-out that would be present if students had learned nothing about the enjoyability of school, we compute the average predicted value of  $dropout_i$  under the counterfactual assumption that  $EOY\_enjoyability_i$  is equal to what was expected at the beginning of the year (as elicited by Question F in Appendix A and referred to as  $prior\_enjoyability_i$ ). Under

 $^{39}$ We find no differences in the amount that students expected to study at the beginning of the year by EOY enjoyability, group

 $<sup>^{40}</sup>$ A one unit increase in  $EOY\_enjoyability_i$  is associated with a .139 decrease in a student's average study hours per day in the first year.

<sup>&</sup>lt;sup>41</sup>Note that 33 of 262 (12.6%) of the 262 person subsample have *dropout*<sub>i</sub> equal to one. This percentage becomes 14.6% if we add back in the students who satisfied the condition of having legitimate values of both *EOY\_mean*<sub>i</sub> and *GPA\_Cumulative*<sub>i</sub> but had *GPA\_Cumulative*<sub>i</sub><1.5. This is still slightly lower than the drop-out rate for the full 325 person sample (17%), in part because a few students drop-out during the second semester.

this counterfactual, the average predicted drop-out rate decreases by 10% to .113.<sup>42</sup>

Thus, the results suggest that what a students learns after arriving at school plays an important role in the drop-out decision, with learning about ability being particularly prominent. Given our specific interest in the role of learning, it is worth exploring whether it is possible to estimate specifications that incorporate learning directly. To motivate the form of these specifications, rewrite equation (13):

(14) 
$$dropout_i = 1 \text{ iff } dropout_i *= E_0(V_N) - E_0(V_S) + [E_2(V_N) - E_2(V_S) - \{E_0(V_N) - E_0(V_S)\}] > 0.$$

The second term,  $[E_2(V_N)-E_2(V_S)-\{E_0(V_N)-E_0(V_S)\}]$ , captures what the student has learned, between the time of college entrance (t=0) and the end of the second semester (t=2), about the expected benefits of being in school. To allow comparability with our earlier results, we begin by focusing on learning about ability as measured by the difference in mean expectations  $EOY\_mean_i$ -prior\\_mean\_i and learning about the enjoyability of school as measured by  $EOY\_enjoyability_i$ - prior\\_enjoyability\_i. The presence of the first term,  $E_0(V_N)-E_0(V_S)$ , indicates that, in order for a specification in which learning enters directly to be sensible, it must also include information about how far each student was from the margin of indifference at the time of college entrance.<sup>43</sup> Since, all else being equal, students who are closer to the margin of indifference at the time of entrance will be less likely to graduate, we are able to take into account possible differences in this "initial condition" by including each student's perception about the probability that he will graduate at the time of college entrance (prob\_grad\_i) which we obtain from Question B in Section III.

The results in the first column of Table 11 indicate that there is information in the self-reported initial conditions variable  $prob\_grad_i$ ; students with higher values of  $prob\_grad_i$  are significantly less likely to drop out conditional on what they learn about their ability and the enjoyability of school with a test of the hypothesis

<sup>&</sup>lt;sup>42</sup>The analysis of this section has included only individuals who have *GPA\_Cumulative*<sub>i</sub> greater than 1.5. Including the remainder of the individuals and reestimating Column 6 of Table 9, we find that the drop-out rate would have decreased by 39% due to learning about ability. The decrease in the drop-out rate due to learning about the enjoyability of school is predicted to remain at 10%.

<sup>&</sup>lt;sup>43</sup>That is, whether a student ultimately ends up above the margin of indifference and out of school ( $dropout_i^* > 0$ ) depends on how far below the margin of indifference he was at the time of entrance (i.e. how negative  $E_0(V_N)$ - $E_0(V_S)$  is) and, given this, whether the amount he learns after college entrance pushes him over the margin of indifference.

that  $prob\_grad_i$  has no effect having a p-value of .039. The effect of learning about ability is highly significant with a test of the null hypothesis that  $EOY\_mean_i$ -prior\\_mean\_i has no effect on drop-out yielding a t-statistic of -4.225. The effect of learning about the enjoyability of school is just significant at 5%.

We can use this model to re-examine our earlier results which quantify the role that learning plays in the drop-out decision. As before, the average predicted value of *dropout*<sub>i</sub> is .127. Assuming no learning about ability (setting *EOY\_mean<sub>i</sub>-prior\_mean<sub>i</sub>* equal to zero for all students) leads to an average predicted drop-out value of .103. Assuming no learning about the enjoyability of school (setting *EOY\_enjoyability<sub>i</sub>* - *prior\_enjoyability<sub>i</sub>* equal to zero for all students) leads to an average predicted drop-out value of .117. In the second column of Table 11 we add the additional learning term *GPA\_Cumulative<sub>i</sub>-prior\_mean<sub>i</sub>* to account for learning about the stock of grades at the end of the first year. Using this model leads to an average predicted drop-out value of .089 when we assume that no learning about ability takes place (setting both *EOY\_mean<sub>i</sub>-prior\_mean<sub>i</sub>* and *GPA\_Cumulative<sub>i</sub>-prior\_mean<sub>i</sub>* equal to zero) and leads to an average predicted drop-out value of .119 when we assume that no learning about the enjoyability of school takes place. Thus, the results are very consistent with what was found using equation (13).

The BPS also collects information about other factors that could influence a student's state at the end of the first year. In the last column of Table 9 we add, to the specification in Column 6, variables representing a person's beliefs about the financial returns to schooling at the end of the year (EOY\_returns<sub>i</sub>), whether the student has a parent that lost a job during the year (parental\_job\_loss<sub>i</sub>), and a student's health on a four point scale where higher values represent better health (EOY\_health<sub>i</sub>).<sup>44</sup> A student's health and whether the student has a parent that lost a job may influence how enjoyable it is to be in school. Indeed, Table 10 shows a statistically significant correlation between EOY\_enjoyability<sub>i</sub> and EOY\_health<sub>i</sub> and between EOY\_enjoyability<sub>i</sub> and Parental\_job\_loss<sub>i</sub>. As a result, adding the additional variables makes it difficult to interpret the estimated

<sup>&</sup>lt;sup>44</sup>EOY\_returns<sub>i</sub> is the difference, at the end of the first year, between a student's beliefs about the median earnings he would receive if he graduated from college with a 3.0 grade point average and the student's beliefs about the median earnings he would receive if he left school immediately. EOY\_health<sub>i</sub> comes from Question G (Appendix A).

effect of *EOY\_enjoyability<sub>i</sub>*. However, the primary message from Column 6 of Table 9 - that previous academic performance and beliefs about ability/future grade performance play a central role in the drop-out decision - remains strong with the estimated effect of *EOY\_mean<sub>i</sub>* and *GPA\_Cumulative<sub>i</sub>* staying roughly the same when the new variables are added. In the last column of Table 11, we add variables representing what a person has learned about the financial returns to schooling during the academic year (*EOY\_returns<sub>i</sub>* - *prior\_returns<sub>i</sub>*), the change in the student's health during the year (*EOY\_health<sub>i</sub>* - *prior\_health<sub>i</sub>*), and the parental job loss variable (*parental\_job\_loss<sub>i</sub>*) to the specification in Column 1 of Table 11. Again we find that adding these variables has little effect on the estimated importance of learning about ability (*EOY mean<sub>i</sub>* - *prior\_mean<sub>i</sub>*).

## VI. A study of poorly performing students

Our results indicate that what a student learns about his academic ability plays an important role in whether he drops out of school. Thus, from a policy standpoint it is important to understand to what extent students update their beliefs in a reasonable fashion. Here we study this issue by focusing on the group of students who are particular concern to policymakers - students with poor grade performance in the first semester. Specifically, in order to continue to take advantage of Question C of Appendix A (which examines how well a student performs relative to expected), we examine students who have values of  $GPA_{II}$ - $Prior\_Mean_{I}$  in the bottom third of our sample.<sup>45</sup> One hundred of the 109 students in the bottom third correctly recognized in Question C.1 that their first semester grades were worse than expected. On average, these 100 students have a  $GPA_{II}$ - $Prior\_Mean_{I}$  gap of -1.25 points. Table 12A uses Question C.3 to show how much of this gap these students believe, on average, should be attributed to worse than expected ability/preparation, worse than expected luck, and lower than expected study effort. Comparing the numbers in Table 12A to the numbers in Table 12B for students in the top third of the  $GPA_{II}$ - $Prior\_Mean_{I}$  distribution raises the possible concern that students in the bottom third may be incorrectly dismissing new information

<sup>45</sup>These students have a  $GPA_{Ii}$ - $Prior\_Mean_i$ <-.56. We note that, because most students tend to be very optimistic initially about their grade performance, being in the bottom of the  $GPA_{Ii}$ - $Prior\_Mean_i$  is essentially synonymous with being in the bottom of the  $GPA_{Ii}$  distribution; the average value of  $GPA_{Ii}$  is 2.012 for the former group and is 1.924 for the latter group.

about their ability; students in the bottom third attribute, on average, only 27% of the  $GPA_{Ii}$ -Prior\_Mean<sub>i</sub> to worse than expected ability/preparation whereas students in the top third attribute, on average, 67% of the  $GPA_{Ii}$ -Prior Mean<sub>i</sub> to better than expected ability/preparation.<sup>46</sup>

## VI.A. Are beliefs about the reasons for poor academic performance accurate?

In order to examine whether the beliefs in Table 12A are accurate we must determine what proportion of the average  $GPA_{Ii}$ - $Prior\_Mean_i$  gap should actually be attributed to each of the three possibilities. Because the decomposition exercise requires assumptions, the results should be viewed as an attempt to provide some rough evidence about whether, on average, students are generally able to comprehend the underlying reasons for their worse than expected performance.

#### The portion of the gap that should be attributed to lower than expected study effort

As seen in Table 12a, on average, students in the bottom third of the  $GPA_{Ii}$ - $Prior\_Mean_{Ii}$  distribution believe that the majority of their grade gap (55%) should be attributed to the fact that they did not study as hard as expected. Our results provide striking evidence that they are indeed correct in believing that their effort was substantially lower than expected; while Question A.1 indicates that, on average, students in the bottom third expected to study 3.98 hours per day, our time-diaries show that they actually studied only 3.07 hours a day. Further, as a bit of an aside, individuals in the top third of the  $GPA_{Ii}$ - $Prior\_Mean_{Ii}$ , who, on average, were much less likely to report that studying more than expected was an important reason for their better than expected grade performance (21%), did indeed have initial beliefs about studying that were much more accurate; while, on average, students in the top third expected to study 3.44 hours per day, in reality they studied 3.57 hours per day. As

<sup>&</sup>lt;sup>46</sup>In results not shown, we find that differences in updating between these groups is broadly consistent with this finding; students in the bottom third put significantly less weight on  $GPA_{Ii}$  than students in the top third in the construction of the posterior mean.

<sup>&</sup>lt;sup>47</sup>A test of the null hypothesis that the average actual amount studied is the same as the average expected amount studied leads to rejection at all traditional significance levels with a standard normal test statistic of 3.65.

<sup>&</sup>lt;sup>48</sup>A test of the null hypothesis that the average actual amount studied is the same as the average expected amount studied cannot be rejected at any traditional significance levels with a standard normal test statistic of .57.

As discussed earlier, in S&S (forthcoming A) we find that an extra hour of studying per day increases first semester grade point average by .36. Using this estimate, the fact that, on average, students in the bottom third studied .91 hours a day less than expected implies that (.91\*.36/1.25)%=26% of the average  $GPA_{Ii}$  -  $Prior\_Mean_{Ii}$  gap should be attributed to the fact that these students studied less than expected. This number is shown in Table 13.<sup>49</sup>

# The portion of the gap that should be attributed to worse than expected luck

The intuition underlying our approach for determining the portion of the gap that should be attributed to worse than expected luck is that a group of students who, on average, have bad luck in the first semester should see their average grades rebound in the second semester (after adjusting for study effort in the two semesters). Using our estimate of .36 for  $\alpha$  and referring to equation (4):

(15) 
$$GPA_{li} = .36Study_{li} + \theta_i + \epsilon_{li}$$

(16) 
$$GPA_{2i} = .36Study_{2i} + \theta_i + \varepsilon_{2i}$$
.

By assumption,  $E(\varepsilon_{1i})=0$  and  $E(\varepsilon_{2i})=0$ . Further, for this section we think of the transitory components  $\varepsilon_{1i}$  and  $\varepsilon_{2i}$  as representing short-term luck, in which case it is reasonable to assume that  $\varepsilon_{1i}$  is independent of  $\varepsilon_{2i}$ . The question of the degree to which students in the bottom third have bad luck in the first semester is then a question of how much the average value of  $\varepsilon_{1i}$  is less than zero for these students.

<sup>&</sup>lt;sup>49</sup>One possible explanation for the finding that individuals in the bottom third significantly overstate the role of effort is that students have perceptions about the causal effect of studying on academic performance that are not be the same as the estimated causal effect in S&S (forthcoming A). Survey question H (Appendix A) not only allows heterogeneity across students but also allows a particular student's belief about the effect of an additional hour of studying to vary with the number of hours that he is currently studying. We find that, on average, students believe that, if they had studied as much as they expected, their grade performance would be higher by .51. Using the information from Question H as if it is the truth (in place of our estimate of .36) and recomputing the numbers in Table 13, we find that the proportion of the gap that should actually be attributed to less effort than expected increases from 26% to 41% (with the perceived amount of 55% from Table 12A).

 $<sup>^{50}</sup>$ This is a non-trivial assumption. There would seem to be little doubt that things that could be characterized as short-term luck are likely to be an important component of the  $\epsilon$ 's. For example, fitting this description would be bad matches with teachers, sampling variation in test taking, inopportune short-term sicknesses etc. Examples of things that would not fit this description are, for example, the ability to choose courses and long-term sicknesses. The effect of course choice is mitigated to a large extent by the reality that many first-year courses are mandatory under a general studies curriculum and our data reveal very little evidence of long-term health problems that develop after the beginning of the first year. However, it is not possible to rule out other possible examples as well (e.g., problems with living arrangements that last a full year). As a result we do not believe that this assumption is literally true.

Differencing equation (16) from equation (15) and rearranging yields

(17) 
$$GPA_{1i} - GPA_{2i} - .36(Study_{1i} - Study_{2i}) = \varepsilon_{1i} - \varepsilon_{2i}$$
.

Thus, the left hand side of equation (17) represents the difference in a person's "luck" between the two semesters. Conditional on being in the bottom third of the  $GPA_{Ii}$ -Prior\_Mean<sub>Ii</sub> distribution, the average value of  $\varepsilon_{1i}$  may indeed be negative. However, under the assumption that  $\varepsilon_{1i}$  is independent of  $\varepsilon_{2i}$ , individuals in the bottom third have  $E(\varepsilon_{2i})=0$ . Thus, the intuition underlying our approach is that, if individuals in the bottom third have bad luck, on average, in the first semester, then we should see the average grades of this group rebound in the second semester (after taking into account any change study effort). Specifically, taking expectations in equation (17) and rearranging, we get an equation for the average first semester luck

(18) 
$$E(\varepsilon_{1i}) = E[GPA_{1i} - GPA_{2i} - .36(Study_{1i} - Study_{2i})].$$

Thus, an estimate of  $E(\varepsilon_{1i})$  for the bottom third can be computed as the sample average value of  $GPA_{1i} - GPA_{2i}$  -  $.36(STUDY_{1i} - STUDY_{2i})$  for individuals in the bottom third.

For students in the bottom third, there is evidence of the rebound that would indicate bad luck in the first semester. For this group, the sample averages of  $GPA_{Ii}$  and  $GPA_{2i}$  are 1.98 and 2.44, respectively, and the sample averages of  $STUDY_{Ii}$  and  $STUDY_{2i}$  are 3.07 and 2.99, respectively. These numbers imply that students in the bottom third have a sample average of -.49 for  $GPA_{Ii}$  - $GPA_{2i}$  -.36( $STUDY_{Ii}$  -  $STUDY_{2i}$ ). Thus, as shown in Table 13, bad luck accounts for (.49/1.25)%=39% of the average gap.

# The portion of the gap that should be attributed to worse than expected ability/preparation

The portion of the average gap that should be attributed, on average, to lower than expected ability/preparation is the residual between the average total gap and the portions that should be attributed to lower than expected effort and bad luck. Thus, as shown in Table 13, 100%-39%-26%=35% is the proportion of the average  $GPA_{li}$ -*Prior Mean*<sub>li</sub> gap that should be attributed to worse ability/preparation than expected.

# **Discussion**

A comparison of Table 12A to Table 13 indicates that, on average, this group of poorly performing students have a reasonably accurate perception of the extent to which the  $GPA_{Ii}$ - $Prior\_Mean$  gap should be

attributed to worse than expected ability/academic preparation (27% perceived vs. 35% actual). The fact that, if anything, individuals tend to understate the role of these presumably persistent factors should perhaps be reassuring to policymakers worried that students might leave school prematurely when things go badly. With respect to the remaining portion of the  $GPA_{Ii}$ - $Prior\_Mean$  gap, perhaps somewhat surprisingly, students tend to take much personal responsibility for their poor performance in the sense that they attribute a larger percentage of the gap than they should to lower than expected effort (55% perceived versus 26% actual) and a lower percentage of the gap than they should to bad luck (18% perceived versus 39% actual).

### V. B. Do forecasts improve for this group?

For students in the bottom third, Table 4 shows that the average value of  $Posterior\_Mean_i$  (2.923) is much closer to the average value of  $GPA_{2i}$  (2.457) than the average value of  $Prior\_Mean_i$  (3.267) is to the average value of  $GPA_{1i}$  (2.012).<sup>51</sup> Similarly, for this group, the average value of  $|Posterior\_Mean_i - GPA_{2i}|$  is .720 while the average value of  $|Prior\_Mean_i - GPA_{1i}|$  is .1.277.<sup>52</sup> The improved prediction occurs, in part, because the grades of this group improve in the second semester, but also because students in this group have significantly revised their beliefs about grade performance - a null hypothesis that the average value of  $Posterior\_Mean_i$  is the same as the average value of  $Prior\_Mean_i$  is rejected at all traditional significance levels. Approximately 1/3 of the gap that remains between the average value of  $Posterior\_Mean_i$  and the average value of  $GPA_{2i}$  can be attributed to the fact that, although students in this group do revise beliefs about how much they will study, they remain too optimistic; although, on average, students in this group have a value  $Expected\_STUDY_{2i}$  of 3.391 (down from an average  $Expected\_STUDY_{1i}$  of 3.978), the average value of  $STUDY_{2i}$  is 2.992.

#### VII. Conclusion

The paper shows that learning about ability plays a major role in the drop-out decision process. We

<sup>&</sup>lt;sup>51</sup>A test rejects, at all traditional significance levels, the null hypothesis that there is no difference between  $E(Posterior\ Mean_i)$ - $E(GPA_{2i})$  and  $E(Prior\ Mean_i)$ - $E(GPA_{2i})$ .

<sup>&</sup>lt;sup>52</sup>A test rejects, at all traditional significance levels, the null hypothesis that there is no difference between the average value of  $|Posterior\_Mean_i - GPA_{2i}|$  and the average value of  $|Prior\_Mean_i - GPA_{1i}|$ .

leave the estimation of a dynamic, learning model of the drop-out decision to future work, although the results here suggest the value of taking advantage of self-reported expectations data to reduce the reliance of assumptions in such a behavioral model.

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Table 1A

FIRST SEMESTER GPA - BELIEFS AND ACTUAL (n=325)

1st semester GPA interval	Subjective probability	Sample probability
[3.5,4.0]	.401 (.256)	.234
[3.0,3.5)	.329 (.175)	.302
[2.5,3.0)	.160 (.132)	.200
[2.0,2.5)	.072 (.073)	.123
[1.0,2.0)	.025 (.035)	.108
[0.0,1.0)	.012 (.022)	.033
<pre>prior_mean; Approximate subjective mean of GPA<sub>1i</sub> sample average (sample standard deviation)</pre>	3.220 (.292)	
Approximate subjective standard deviation of GPA <sub>1i</sub> sample average (sample standard deviation)	.532 (.195)	
Actual GPA <sub>li</sub>		2.879 (.784)
sample average (sample standard deviation)		

Note: First column shows average subjective probability (standard deviation) of having  $GPA_1$  in each category. Second column shows proportion of sample with actual  $GPA_{1i}$  in each category.

Table 1B

FIRST SEMESTER GPA - BELIEFS AND ACTUAL
High School Grade Point Average Greater Than Median (n=157)

1 <sup>st</sup> semester GPA interval	Subjective probability	Sample probability
[3.5,4.0]	.454 (.245)	.338
[3.0,3.5]	.310 (.158)	.369
[2.5,3.0)	.139 (.102)	.153
[2.0,2.5)	.064 (.060)	.089
[1.0,2.0)	.022 (.031)	.031
[0.0,1.0)	.010 (.020)	.019
prior_mean; Approximate subjective mean of GPA <sub>1i</sub>	3.275 (.266)	
sample average (sample standard deviation) Approximate subjective standard deviation of GPA <sub>Ii</sub>	.526 (.192)	
sample average (sample standard deviation)		
Actual GPA <sub>1i</sub>		3.164 (.661)
sample average (sample standard deviation)		

Note: See Note Table 1A

High School Grade Point Average not observed for 11 of 325 students in sample.

Table 1C

FIRST SEMESTER GPA - BELIEFS AND ACTUAL

High School Grade Point Average Less Than Median (n=157)

1st semester GPA interval	Subjective probability	Sample probability
[3.5,4.0]	.339 (.250)	.121
[3.0,3.5]	.355 (.188)	.236
[2.5,3.0)	.186 (.155)	.248
[2.0,2.5)	.080 (.083)	.153
[1.0,2.0)	.072 (.038)	.191
[0.0,1.0)	.017 (.024)	.051
<i>prior_mean<sub>i</sub></i> : Approximate subjective mean of GPA <sub>1i</sub> sample average (sample standard deviation)	3.161 (.301)	
Approximate subjective standard deviation of GPA <sub>1i</sub>	.538 (.198)	
sample average (sample standard deviation)		
Actual GPA <sub>1i</sub>		2.579 (.802)
sample average (sample standard deviation)		

Note: See Note Table 1A

High School Grade Point Average not observed for 11 of 325 students in sample.

Table 2

	Prior_Mean	$GPA_{I}$	Posterior_Mean	$GPA_2$
Prior_Mean	1.0	.147	.403	.177
$GPA_1$		1.0	.585	.605
Posterior_Mean			1.0	.394
$GPA_2$				1.0

Note: High School Grade Point Average not observed for 11 of 325 students in sample.

Table 3

SECOND SEMESTER GPA - BELIEFS (n=325)

1st semester GPA interval	Subjective probability	
[3.5,4.0]	.311 (.278)	
[3.0,3.5)	.365 (.205)	
[2.5,3.0)	.200 (.174)	
[2.0,2.5)	.084 (.101)	
[1.0,2.0)	.031 (.058)	
[0.0,1.0)	.009 (.025)	
posterior_mean <sub>i</sub> : Approximate subjective mean of GPA <sub>2i</sub>	3.140 (.357)	
sample average (sample standard deviation)		
Approximate subjective standard deviation of GPA <sub>2i</sub>	.477 (.194)	
sample average (sample standard deviation)		

**Table 4** Descriptive statistics for entire sample (n=325) and subgroups

Variable	Full Sample	Top third	Bottom third	Top third	Bottom third
variable		$GPA_I$	$GPA_I$	$GPA_{I}$ -prior_mean	GPA <sub>1</sub> -prior_mean
	mean (s.d.)	mean (s.d.)	mean (s.d.)	mean (s.d.)	mean (s.d.)
prior_mean	3.220(.292)	3.278 (.291)	3.199 (.295)	3.123 (.311)	3.267 (.267)
posterior_mean	3.140 (.357)	3.386 (.287)	2.886 (.303)	3.293 (.307)	2.923 (.306)
$GPA_{I}$	2.879 (.784)	3.650 (.225)	1.924 (.553)	3.541 (.354)	2.012 (.605)
$GPA_2$	2.929 (.772)	3.398 (.560)	2.380 (.783)	3.320 (.668)	2.457 (.786)

Each entry shows sample average (sample standard deviation)

Students in bottom third of GPA1 distribution have GPA1<2.64

Students in top third of GPA1 distribution have GPA1>3.30

Students in top third of GPA<sub>1</sub>-prior\_mean distribution have GPA<sub>1</sub>-prior\_mean>.09.

Students in bottom third of GPA<sub>1</sub>-prior\_mean distribution have GPA<sub>1</sub>-prior\_mean<-.56.

# DETERMINANTS OF POSTERIOR MEAN(column 1)

AND POSTERIOR MEAN\* (column 2)

Independent Variable	dependent variable posterior_mean estimate (std. error) n=325	dependent variable posterior_mean* estimate (std. error) n=291
Constant Prior_Mean GPA <sub>1</sub> GPA <sub>1</sub> *	1.158 (.166)** .396 (.051)** .245 (.019)**	.323 (.040)** .597 (.053)** .344 (.042)**
Study equation (8) C $\sigma^2_{\mu}$ $\sigma^2_{\nu}$	$R^2 = .445$	3.249 (.110)** 1.20 (.103)** 1.56 (.071)** Log Like =-1338.49

The first column is estimated by OLS and uses all sample observations. The second column is a measurement error model estimated by Maximum Likelihood (using approach of Section IV.C.2) and using all observations for which a person reports legitimate values of Expected\_STUDY<sub>1i</sub> and Expected\_STUDY<sub>2i</sub> \*significant at .10

<sup>\*\*</sup>significant at .05

DETERMINANTS OF Posterior Mean\* (Maximum Likelihood) estimate estimate estimate Independent (std. error) (std. error) (std. error) Variable n = 191n = 191n=191.661 (.388)\* .550 (.399) .814(.370)\*\* Constant .517 (.112)\*\* .494 (.115)\*\* .357 (.129)\*\* Prior\_Mean, .446(.064)\*\*  $GPA_{Ii}*$ .316 (.039)\*\* .176 (.068)\*\* .177 (.085)\*\*  $Prior\_Mean_i \times \hat{W}_{1i}$ .198 (.091)\*\*  $GPA_{li}*x \hat{W}_{2i}$  $\begin{array}{ccc} Prior\_Mean_i x \hat{\sigma}^2_{\epsilon i} \\ GPA_{Ii}^* x \hat{\sigma}^2_{\epsilon i} \end{array}$ .248 (.122)\*\* -.303 (.127)\*\* Study equation (8) C 3.524 (.124)\*\* 3.522 (.125)\*\* 3.527 (.125)\*\*  $\sigma^2_{\ \mu}$   $\sigma^2_{\ \nu}$ 1.272 (.121)\*\* 1.270 (.123)\*\* 1.260 (.122)\*\* 1.608 (.072)\*\* 1.609 (.072)\*\* 1.611 (.072)\*\* Log Like Log Like Log Like -1328.637-1326.037

The sample contains the 191 students who a) correctly recognized in Question C.1 whether they had performed better or worse than expected in the first semester; b) did not have a percentage of 100 on line c of Question C.2 (or C.3); c) provided legitimate information about Expected\_STUDY<sub>1</sub>, and Expected\_STUDY<sub>2i</sub>. The first column repeats the results in Column 2 of Table 5 for the smaller sample here. Estimation is by maximum likelihood (using approach of Section IV.C.2).

-1314.874

<sup>\*</sup>significant at .10

<sup>\*\*</sup>significant at .05

Table 7

DETERMINANTS OF Posterior Mean (OLS)

Independent Variable	estimate	estimate	estimate
independent variable	(std. error)		(std. error)
	n=211	n=211	n=211
Constant	1.203 (.223)**	1.181 (.223)**	1.352 (.223)**
Prior_Mean <sub>i</sub>	.380 (.068)**	.362 (.069)**	.221 (.089)**
$GPA_{Ii}$ (the noisy signal)	.242 (.022)**	.169 (.045)**	.387 (.057)**
$Prior\_Mean_i \times \hat{W}_{1i}$		.083 (.048)*	
$GPA_{li} \times \hat{W}_{2i}$		.101 (.055)*	
$Prior\_Mean_i \times \hat{\sigma}^2_{\epsilon i}$			.225 (.094)**
$GPA_{li} \times \hat{\sigma}^2_{\epsilon i}$			298 (.102)**
	$R^2 = .436$	$R^2 = .445$	$R^2 = .466$

The sample contains the 211 students who correctly recognized in Question C.1 whether they had performed better or worse than expected in the first semester and did not have a percentage of 100 on line c of Question C.2 (or C.3). Estimation is by Ordinary Least Squares.

<sup>\*</sup>significant at .10

<sup>\*\*</sup>significant at .05

Independent Variable

We are in the set of the set of

For a description of the sample see the note in Table 6.

**Table 9** DETERMINANTS OF DROPOUT:

The effect of a student's state at end of first year

		DETERM	IINANTS O	F dropout <sub>i</sub> (	OLS)		
Independent	estimate	estimate	estimate	estimate	estimate	estimate	estimate
Variable	(std.	(std.	(std.	(std.	(std.	(std. error)	(std. error)
	error)	error)	error)	error)	error)	n=261	n=260
	n=262	n=262	n=262	n=245	n=261	(6)	(7)
	(1)	(2)	(3)	(4)	(5)	, ,	
Constant	.671**	1.113**	1.024**	.647**	.015	.920**	.970**
Constanti	(.096)	(.176)	(.178)	(.173)	(.039)	(.195)	(.206)
$GPA\_Cumulative_i$	182**		114**	194**		105**	110**
	(.031)		(.042)	(.112)		(.042)	(.043)
$EOY\_Mean_i$		315**	177**			171**	156**
		(.056)	(.075)			(.076)	(.078)
$RE\_EOY\_Mean_i$				.021			
				(.154)			
school_enjoyability <sub>i</sub>					.062**	.033* (.019)	.026
					(.019)		(.019)
$EOY\_health_i$							019
•							(.031)
$EOY\_returns_i$							029
•							(.037)
parental_job_loss <sub>i</sub>							.077
·							(.061)
	$R^2 = .113$	$R^2 = .108$	$R^2 = .132$	$R^2 = .104$	$R^2 = .038$	$R^2 = .146$	$R^2 = .155$

Table shows results of linear probability models with dependent variable *dropout*<sub>i</sub>. Explanatory variables measure student's state at end of first year.

Sample is as described in Section V.

Column 4 has less observations since, as discussed in Appendix C, constructing *RE\_EOY\_Mean<sub>i</sub>* involves use of HSGPA and ACT which are missing for some observations.

<sup>\*</sup>significant at .10

<sup>\*\*</sup>significant at .05

**Table 10** Correlation Matrix (n=260)

Variable	$dropout_i$	GPA_ Cumulative;	EOY_ Mean;	EOY_ enjoyability;	parental_job _loss;	EOY_health <sub>i</sub>	EOY_returns <sub>i</sub>
dropout <sub>i</sub>	1.00	Cumulative;	meun;	enjoyaomiy;	·		
$GPA\_Cumulative_i$	339**	1.00					
EOY_Mean <sub>i</sub>	.0000 335**	.676**	1.00				
$EOY_{enjoyability_{i}}$	.0000 .195**	.0000 232**	251**	1.00			
parental_job_loss <sub>i</sub>	.001 .115* .063	.0002 048 .437	.0000 056 .364	.172**	1.00		
$EOY\_health_i$	107*	.089	.227**	.005 109*	103*	1.00	
EOY_returns <sub>i</sub>	.084 978	.148 .020	.0002 .070	.078 156**	.097 079	044	1.00
·	.204	.739	.260	.011	.119	.472	

The first number in each box is the sample correlation between two variables. The second number is a p-value from a test that the population correlation between two variables is zero.

<sup>\*</sup>significant at .10

<sup>\*\*</sup>significant at .05

**Table 11** DETERMINANTS OF DROPOUT: The effect of learning during the first year

DETERMINANTS OF dropout: (OLS)

Independent Variable	estimate	estimate	estimate
independent variable	(std. error)	(std. error)	(std. error)
	n=259	n=259	n=248
Constant	.302**	.280**	.272**
Constant	(.102)	(.101)	(.103)
$prob\_grad_i$ ,	242**	230**	228**
	(.116)	(.115)	(.117)
$EOY\_mean_i$ - $prior\_mean_i$ .	251**	107*	238**
	(.059)	(.060)	(.061)
$EOY\_enjoyability_i$ - $prior\_enjoyability_i$	.031	.023	.029
	(.19)*	(.018)	(.018)
$GPA\_Cumulative_i$ - $prior\_mean_i$		117**	
		(.044)	
EOY_health; - prior_health;			042
			(.033)
$EOY\_returns_i - prior\_returns_i$			004
			(.035)
$parental\_job\_loss_i$			.094
	<b>7</b> 2 006	T 2	(.063)
	$R^2 = .086$	$R^2 = .113$	$R^2 = .098$

Table shows results of linear probability models with dependent variable *dropout*<sub>i</sub>. Explanatory variables measure how much a person has learned during first year. sample as described in Section V.

<sup>\*</sup>significant at .10
\*\*significant at .05

## Percentages from Question C.3

100 individuals who have GPA1-PRIOR MEAN in bottom third and correctly indicated on Question C.1 that grades were lower than expected

Independent Variable	mean		
independent variable	(standard deviation)		
a+b) Worse than expected ability/preparation	27.1% (30.5)		
c) Lower than expected study effort	55.1% (35.1)		
d) Worse than expected luck	17.7% (27.2)		

## Table 12B

## Percentages from Question C.2

79 individuals who have GPA1-PRIOR MEAN in top third and correctly indicated on Question C.1 that grades were higher than expected

Independent Variable	mean (standard deviation)	
mucpenucht variable		
a+b) Better than expected ability/preparation	67.1% (34.3)	
c) Higher than expected study effort	21.4% (27.6)	
d) Better than expected luck	11.5% (25.0)	

#### Table 13

Estimates of actual importance of categories in Question O
100 individuals who have GPA1-PRIOR MEAN in bottom third and correctly indicated on
Question O that grades were lower than expected

me1Gan	
(standard deviation)	
35%	
26%	
39%	

## **Appendix A: Survey Questions**

**Question A.1**. During your first year of college, how many hours do you expect to spend in the following activities on an average **weekday** (Monday-Friday).

<u>Activity</u>	Weekday hours	<u>Avg</u>
1. Studying and Home	ework	
2. Sleeping		
3. School Athletics, C	lubs, other school act	tivities

**Question A.2.** We realize that you do not know exactly how well you will do in classes. However, we would like to have you describe your beliefs about the grade point average that you expect to receive in the first semester.

Given the amount of study-time you indicated in question A.1, please tell us the percent chance that your grade point average will be in each of the following intervals. That is, for each interval, write the number of chances out of 100 that your final grade point average will be in that interval.

Note: The numbers on the six lines must add up to 100.

<u>Interval</u>	Percent Chance (number of chances out of 100).	
[3.5, 4.00]		
[3.0, 3.49]	<del></del>	
[2.5, 2.99]		
[2.0, 2.49]	<del></del>	
[1.0, 1.99]		
[0.0, .99]		

Note: A=4.0, B=3.0, C=2.0, D=1.0, F

**Question A.3**. During the second semester, how many hours do you expect to spend in the following activities on an average **weekday** (Monday-Friday).

Activity _ Avg W	eekday hours _
<ol> <li>Studying and Homework</li> <li>Sleeping</li> <li>School Athletics, Clubs, other school activities</li> </ol>	

**Question A.4.** We realize that you do not know exactly how well you will do in classes. However, we would like to have you describe your beliefs about the grade point average that you expect to receive in the second semester.

Given the amount of study-time you indicated in question A.3, please tell us the percent chance that your grade point average will be in each of the following intervals. That is, for each interval, write the number of chances out of 100 that your final grade point average will be in that interval.

Note: The numbers on the six lines must add up to 100.

Note: A=4.0, B=3.0, C=2.0, D=1.0, F

<u>Interval</u>	Percent Chance (number of chances out of 100).		
[3.5, 4.00]			
[3.0, 3.49]	<del></del>		
[2.5, 2.99]			
[2.0, 2.49]			
[1.0, 1.99]			
[0.0, .99]			

#### **Question C.1.** Circle the one that is true

- a). I received grades in the Fall term that were higher than I had expected to get when I came to Berea.
- b). I received grades in the **Fall** term that were **lower** than I had expected to get when I came to Berea.

If you circled a), GO TO Question C.2 below.

If you circled b), GO TO Question C.3 below.

#### Question C.2. (Answer this question if you circled that grades better than expected in Question C.1.)

Please circle those reasons why you think you received grades in Fall term that were higher that you had expected.

A)	My ability is better than I thought it was when I came to Berea.	
B)	I am better prepared for Berea College than I thought I was when I came to Berea.	
C)	I studied harder than I had expected I would when I came to Berea.	
D)	I had better luck than I expected when I came to Berea in that those things that	
	influence grades but were out of my control turned out to be very much in my favor.	

Now consider the difference between the grades you received in Fall term and the grades you had expected. On the **lines to the right of the reasons**, write the percentage of this difference that you would attribute to each of the reasons you circled. (The items you did not circle should have zero percentage or be left blank.) **Note: The numbers on the lines should add to 100.** 

#### Question C.3. (Answer this question if you circled that grades are worse than expected in Question C.1)

Circle those reasons why you think you received grades in Fall term that were lower than what that you had expected.

A)	My ability is not as good as I thought it was when I came to Berea.	
B)	I am not as well prepared for Berea College as I thought I was when I came to Berea.	
C)	I did not study as hard as I thought I would when I came to Berea.	
D)	I had worse luck than I expected when I came to Berea in that those things that	
	influence grades but were out of my control turned out to be hurting my grades.	

Now consider the difference between the grades you received in Fall term and the grades you had expected. On the **lines to the right of the reasons**, write the percentage of this difference that you would attribute to each of the reasons you circled. (The items you did not circle should have zero percentage or be left blank.) **Note: The numbers on the lines should add to 100.** 

**Question D.** Your grades are influenced by your academic ability/preparation and how much you decide to study. However, your grades may also be influenced to some extent by good or bad luck which may vary from term to term and may be out of your control. Examples of "luck" may include 1) The quality of the teachers you happen to get and how hard or easy they grade; 2) Whether you happened to get sick (or didn't get sick) before important exams; 3) Whether a noisy dorm kept you from sleeping before an important exam; 4) Whether you happened to study the wrong material for exams; 5) Whether unexpected personal problems or problems with your friends and family made it hard to concentrate on classes.

We would like to know how important you think "luck" is in determining your grades in a particular semester. We'll have you make comparisons relative to a semester in which you have "average" luck. Average luck means that a usual number of things go right and wrong during the semester. Assume you took classes at Berea for many semesters.

#### GOOD LUCK IN A TERM MEANS THAT YOU HAVE BETTER THAN AVERAGE LUCK IN THAT TERM

#### Assume for this section that you are in a semester in which you have good luck

**D.1** In what percentage of semesters that you have good luck would good luck raise your grade point average (GPA) by between 0.00 points and 0.25 points compared to a semester in which you received "average" luck.

Note. (If you are taking four courses, good luck would raise your GPA by 0.25 points if good luck led to a full letter grade increase in one of your courses).

**D.2** In what percentage of semesters that you have good luck would good luck raise your grade point average (GPA) by between 0.26 points and 0.50 points compared to a semester in which you received "average" luck.

Note: (If you are taking four courses, good luck would raise your GPA by .50 points if good luck led to a full letter grade increase in two of your courses or a two letter grade increase in one of your courses).

**D.3** In what percentage of semesters that you have good luck would good luck raise your grade point average (GPA) by 0.51 or more points compared to a semester in which you received "average" luck. \_\_\_\_\_\_\_

Note: (For a student taking four courses, this would mean that good luck would lead to a full letter grade increase in three or more courses)

The numbers in the three spaces above in the good luck section should add up to 100 (because if you are in a semester where you have good luck, good luck must increase your grades by between 0 and .25 points, or by between .25 and .5 points, or by more than .5 points).

**Question E.** Circle the one answer that describes your beliefs at this time.

- 1. I believe that being in college at Berea will be much more enjoyable than not being in college.
- 2. I believe that being in college at Berea will be somewhat more enjoyable than not being in college.
- 3. I believe that I will enjoy being in college at Berea about the same amount as I would enjoy not being in college.
- 4. I believe that being in college at Berea will be somewhat less enjoyable than not being in college.
- 5. I believe that being in college at Berea will be much less enjoyable than not being in college.

**Question F.** Which of the following best describes your beliefs now? Circle the one best answer.

- 1. I believe that being in college at Berea is much more enjoyable than not being in college.
- 2. I believe that being in college at Berea is somewhat more enjoyable than not being in college.
- 3. I have enjoyed being in college at Berea about the same amount as I would have enjoyed not being in college.
- 4. I believe that being in college at Berea is somewhat less enjoyable than not being P in college.
- 5. I believe that being in college at Berea is much less enjoyable than not being in college.

Question G. How would you rate your current health? Poor Fair Good Excellent

**Question H.** For each of the following possible amounts that you might study, write down the percent chance that you will study that amount and the grade point average you expect to receive if you study that amount.

Number of Study Hours a Day	Expected Grade Point Average
0 hours a day	
1 hour a day	
2 hours a day	
3 hours a day	
4 hours a day	
5 hours a day	
6 or more hours a day	

# Survey #5 (Please complete both sides of this sheet)

Question A.

Reminders: Be sure to put an arrow (->) next to the time that it is right now. And label this arrow with the words YESTERDAY and START.

Beginning with the What were you doing box next to the arrow, fill in your activities starting 24 hours ago (yesterday) and ending right before you began completing this survey.

Please use the 13 words listed in BOLD on the right of this page to describe your activities.

Time Period	What were you doing?	Time Period	What were you doing?	LIST OF WORDS in bold
6:00 AM		6:00 PM		In Class
6:20 AM		6:20 PM		Attending class, attending labs,
6:40 AM		6:40 PM		attending required class
7:00 AM		7:00 PM	·	sessions
		7:20 PM		<b>-</b>
7:20 AM 7:40 AM		7:40 PM	4	Studying (Outside of
8:00 AM		8:00 PM		class time)
8:20 AM		8:20 PM		(refer to pg 2 for more details)
	·	-11		(Total to pg 2 for more details)
8:40 AM		8:40 PM		Athletics
9:00 AM		9:00 PM		(Intercollegiate or Intramural -
9:20 AM		9:20 PM		
9:40 AM		9:40 PM		games or practice)
10:00 AM		10:00 PM		CPI
10:20 AM	**************************************	10:20 PM		Clubs
10:40 AM		10:40 PM		
11:00 AM		11:00 PM		Exercising
11:20 AM		11:20 PM		_
11:40 AM		11:40 PM		Recreation
AFTERNOON		NIGHT		(reading which is unrelated to
12:00 вооп		12:00 midnight		courses, listening to music, watching movie, spending time
12:20 PM		12:20 AM		watching movie, spending time with friends, etc.)
12:40 PM		12:40 AM		widi filends, etc.)
1:00 PM		1:00 AM		
1:20 PM		1:20 AM		Shopping
1:40 PM		1:40 AM		
2:00 PM		2:00 AM		Eating
2:20 PM		2:20 AM		
2:40 PM		2:40 AM		Sleeping
3:00 PM		3:00 AM		*
3:20 PM		3:20 AM		Partying
3:40 PM		3:40 AM		
4:00 PM		4:00 AM		Personal
4:20 PM		4:20 AM		
4:40 PM		4:40 AM		Working (in Labor position)
5:00 PM		5:00 AM		
5:20 PM		5:20 AM		Other
5:40 PM		5:40 AM		(Please describe on your sheet)

## Appendix B

Intuitively speaking, if we knew the value of  $\mu_i$  for each person and the distribution of  $v_{ji}$ , we could integrate out the effect of the missing information in any outcome equation of interest. Our MLE takes into account that, while we do not know the value of  $\mu_i$  for each person i, the observed values of  $S_{ji}$ , when viewed through equation (7), provide evidence about the likelihood of different values of  $\mu_i$ . More specifically, analogous to the MLE's derived in the missing data literature, the likelihood contribution for person i,  $L_i$ , is the joint probability of *posterior\_mean\_i\**, and the daily study amounts  $S_{1i}$ ,..., $S_{N(i)i}$ . Noting that, under our permanent/transitory assumption in equation (8), each of the daily study amounts and *posterior\_mean\_i\** are independent conditional on  $\mu_i$ ,

 $L_i = \int g_1(S_{1i}|\mu_i) \cdots g_1(S_{N(i)i}|\mu_i) \ g_2(\textit{posterior\_mean}_i * |\mu_i) \ h(\mu_i) \ d\mu_i$ 

where the g's and h are density functions.

In practice, we assume normality for all densities of relevance, in which case the assumptions above about  $\mu_i$  imply that  $h \sim N(C, \sigma^2_{\mu})$  and equation (8) implies that  $g_1 \sim N(\mu_i, \sigma^2_{\nu})$  and. Defining N\* to be the total number of days in the first semester,  $\rightarrow \mu_i$  by the Law of Large Numbers as N\* becomes large. Then,  $g_2$  is normally distributed with a variance of  $\sigma^2_{\mu}$  (the variance of u in equation 4) and a mean of

 $\beta_0 + \beta_1 prior\_mean_i + \beta_2 [GPA_{Ii} - \alpha (\mu_i - Expected\_STUDY_{Ii})]$ .53

<sup>&</sup>lt;sup>53</sup>From equation (4), the mean is  $\beta_0 + \beta_1 prior\_mean_i + \beta_2 GPA_{Ii}^*$ . Substituting equation (6) yields  $\beta_0 + \beta_1 prior\_mean_i + \beta_2 [GPA_{Ii} - \alpha (STUDY_{Ii} - Expected\_STUDY_{Ii})]$ . Equation (10) is then obtained because  $STUDY_{Ii}^{-}$   $\mu_i$ 

## Appendix C. Computing RE\_EOY\_Mean

We construct  $RE\_EOY\_mean_i$  for each person under the assumption that individuals update in a Bayesian manner. Rewriting equation (3) to take into account that the observed noisy signal in this exercise comes from the grade point average for the first full year (instead of just the first semester),

(3a\*) 
$$RE\ EOY\ mean_i = W_{ii} * prior\ mean_i + W_{ii} * GPA\ Cumulative_i$$

What is necessary is to provide values of  $prior\_mean_i$ ,  $W_{1i}$ , and  $W_{2i}$ . The Rational Expectations assumption would imply that  $prior\_mean_i$  is the mean grade point average of students who are deemed to be similar person i. Here we assume that people are similar if they have the same sex, high school grade point average, and score on the American Achievement Test (ACT). Thus,  $prior\_mean_i$  can be constructed as a predicted value from a regression of grades on SEX, HSGPA, and ACT.

Modifying equation (3d) to take into account that  $GPA\_Cumulative_i = (GPA_{1i} + GPA_{2i})/2$ , the weights are given by

(3d\*) 
$$W_{1i}* = \frac{\frac{\sigma_{\epsilon i}^{2}}{2}}{\frac{\sigma_{\epsilon i}^{2}}{2} + prior\_variance_{i}}, \quad W_{2i}* = \frac{prior\_variance_{i}}{\frac{\sigma_{\epsilon i}^{2}}{2} + prior\_variance_{i}}.$$

Under the RE assumption,  $prior\_variance_i$  can be thought of as the amount of permanent heterogeneity that exists in grades across students that are deemed to be like person i.  $\sigma_{ei}^2$  can be thought of as the amount of transitory variation in grades across students that are deemed to be like person i. Then, it is natural to estimate  $prior\_variance_i$  and  $prior\_variance_i$  using a Random Effects estimator which takes advantage of grade performance in both the first semester and the second semester. Doing so as in the Table Appendix C below leads to an estimate for  $prior\_variance_i$  of .131 and an estimate for  $prior\_variance_i$  of .226. Then, the weights are .463 and .537, respectively.

## Grade Point Average

## estimate std. error

n=245	
.217 (.289) **	
099 (.067)	
.438 (.075)**	
.056 (.009)**	
.131	
.226	
	.217 (.289) **099 (.067) .438 (.075)** .056 (.009)** .131

Table contains estimates from a Random Effects Estimation of grade performance in the first and second semesters.

<sup>\*</sup>significant at .10

<sup>\*\*</sup>significant at .05