Gendered Language

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Abstract

Languages use different systems for classifying nouns. Gender languages assign many — sometimes all — nouns to distinct sex-based categories, masculine and feminine. We construct a new data set, documenting this property for more than 4,000 languages which together account for more than 99 percent of the world’s population. We find a robust negative cross-country relationship between prevalence of gender languages and women’s labor force participation. In four countries in Sub-Saharan Africa and in India, we show that educational attainment and female labor force participation are lower among those whose native languages use grammatical gender.

JEL codes: J16, Z10, Z13

Keywords: grammatical gender, language, gender, linguistic determinism, labor force participation, educational attainment, gender gaps

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1 Introduction

Language structures thought. All human beings use language to articulate their ideas and communicate them to others. Yet, the world’s languages show tremendous diversity in terms of their structure and vocabulary. Different languages obviously use different words to describe the same concept, but they also organize the relationships between concepts in remarkably different ways. Because languages are so diverse and language is so fundamental to thought, some scholars have argued that the language we speak may limit the scope of our thinking. Benjamin Lee Whorf, one of the original proponents of this theory of *linguistic determinism*, famously argued that it was difficult for humans to think about ideas or concepts for which there was no word in their language (Whorf 2011[1956]a).

Though specious anecdotes about obscure languages abound, cognitive scientists have largely refuted the strongest forms of Whorf’s hypothesis (Boroditsky, Schmidt, and Phillips 2003). Nonetheless, there is mounting evidence for weaker forms of linguistic determinism: the languages we speak shape our thoughts in subtle, subconscious ways. For example, implicit association tests show that bilinguals display different subconscious attitudes when tested in their different languages (Ogunnaike, Dunham, and Banaji 2010, Danziger and Ward 2010). Russian speakers are better able to visually distinguish shades of blue than English speakers because Russian makes an obligatory distinction in shades that English does not (Winawer, Witthoft, Frank, Wu, Wade, and Boroditsky 2007). Differences in language structure also influence our behavior in the economic realm. Chen (2013), for instance, demonstrates that speakers of languages that demarcate the future as separate from the present (e.g. English) save less than those whose languages make no such distinction (e.g. German).

Several recent papers explore the link between language and gender roles. As Alesina, Giuliano, and Nunn (2013) note, views of the appropriate role for women in society differ markedly across cultures. Languages also vary in their treatment of gender. At one extreme, languages such as Finnish and Swahili do not mark gender distinctions in any systematic
nouns are not categorized as either masculine or feminine; and the same first, second, and third person pronouns are used for males and females. Many languages distinguish between human males and females by using different pronouns: for example, “he” and “she” in English. Some languages go even further, extending the gender distinction to inanimate nouns through a system of grammatical gender. For example, languages such as Spanish and Italian partition all nouns — even inanimate objects — into distinct gender categories. This feature of language forces gender into every aspect of life. For a speaker of a gender language, gender distinctions are salient in every thought and utterance: the space of words is divided into distinct masculine and feminine spheres, and one must constantly reference this mental partition to produce grammatically correct speech.

Does grammatical gender shape (non-grammatical) gender norms? Does it impact women’s participation in economic life? Writing nearly 100 years ago, Benjamin Lee Whorf argued that the existence of linguistic gender categories likely made other gender divisions appear more natural (Whorf 2011[1956]b), though he did not provide any empirical evidence that this was the case. However, recent work by social scientists supports his claim. For example, seemingly arbitrary grammatical gender distinctions do influence our subconscious thoughts, imbuing inanimate nouns with masculine or feminine attributes (e.g. strength or beauty) in line with their assigned grammatical gender category (Boroditsky, Schmidt, and Phillips 2003). Pérez and Tavits (forthcoming) show that Estonian/Russian bilinguals are more supportive of gender equality when interviewed in (non-gender) Estonian than in (gender) Russian.

Whether this pattern extends beyond specific cases has been difficult to assess empirically. In the economic realm, one recent study of immigrants to the United States shows that those who grew up speaking a gender language are more likely to divide household tasks along gender lines (Hicks, Santacreu-Vasut, and Shoham 2015), while another demonstrates that female labor supply is lower among immigrants who speak a gender language at home (Gay, Hicks, Santacreu-Vasut, and Shoham 2017). These analyses make use of the most comprehensive existing data source on languages, the World Atlas of Language Structures.
(WALS). The WALS documents whether a language employs grammatical gender, but only for a fraction of the world’s languages. Using it alone, analysis within Africa or Asia — where widely-spoken indigenous languages differ in their grammatical gender structure — is nearly impossible. Cross-country analysis using the WALS relies on the assumption that missing data on the native languages of half the world’s population is ignorable, yielding a set of bounding and clustering problems that severely hamper inference. Progress on this research topic demands a new source of data.

We provide new evidence to support the hypothesis that grammatical gender shapes views of women’s role in society and directly impacts women’s labor force participation. To do this, we construct a data set characterizing the grammatical gender structure of 4,346 living languages, expanding the number of languages for which systematic data on grammatical gender is available by almost a factor of ten. We draw on a range of data sources including language textbooks, historical records, academic work by linguists, and — in a small number of cases — firsthand accounts from native speakers and translators; using these data sources, we generate a measure of the grammatical gender structure of each of the languages in our data set. Taken together, these languages account for 6.44 billion people, or over 99 percent of the world population.

We use these data in two ways. First, we calculate — for every country in the world — an estimate of the proportion of the population whose native language is a gender language. We are able to account for more than 90 percent of the estimated population in all but three countries. In our first piece of analysis, we explore the cross-country relationship between grammatical gender and women’s labor force participation, women’s educational attainment, and gender attitudes among both men and women. We then complement our cross-country analysis by estimating the individual-level association between grammatical gender and women’s participation in economic life in countries where both gender and non-gender languages are indigenous and widely spoken. We do this within-country analysis separately in two contexts: using Afrobarometer data from four African countries (Kenya,
Niger, Nigeria, and Uganda) and, separately, using the India Human Development Survey, which covers 33 Indian states.

Our cross-country analysis suggests a robust negative relationship between grammatical gender and female labor force participation. Our preferred specification suggests that grammatical gender is associated with a 12 percentage point reduction in women’s labor force participation and an almost 15 percentage point increase in the gender gap in labor force participation. These associations are robust to the inclusion of a wide range of controls including suitability for the plough. Taken at face value, our coefficient estimates suggest that gender languages keep approximately 125 million women around the world out of the labor force. Following the approach suggested by Altonji, Elder, and Taber (2005) and Oster (2017), we estimate that unobservable country-level characteristics would need to be 1.44 times more correlated with treatment than observed covariates to fully explain the apparent impact of grammatical gender on the level of female labor force participation; unobserved factors would need to be 3.24 times more closely linked to treatment to explain the impact of grammatical gender on the gender gap in labor force participation.

We find a far more muted cross-country relationship between grammatical gender and women’s educational attainment. This may be due to the fact that the average within-country gender gap in educational attainment is much smaller than the gender gap in labor force participation — since many wealthy countries have no gender gap in educational attainment, particularly at the primary school level. The prevalence of gender languages is negatively associated with the gender gap in primary school completion after controlling for continent fixed effects, but the estimated relationship is only marginally statistically significant.

Using data from the World Values Survey (WVS), we show that grammatical gender predicts support for traditional gender roles. The coefficient estimate is large in magnitude, suggesting that differences in language could explain the entire gap in gender attitudes between Ukraine (at the 55th percentile of WVS countries in terms of support for gender equality) and Trinidad and Tobago (at the 80th percentile). As Whorf might have hypoth-
esized, gender languages are associated with greater support for traditional gender roles among both men and women.

Though our analysis uses much richer country-level data on grammatical gender than has previously been available, mis-measurement of our independent variable of interest is still a potential concern. We would typically expect measurement error in the independent variable to bias the estimated association toward zero, but the interval nature of our measure of the country-level prevalence of grammatical gender (when the gender structure of the native language is not known for the entire population) can also lead to invalid inference. Using a bounding technique proposed by Imbens and Manski (2004), we show that our results are robust to correcting for the censored nature of our independent variable of interest.

A more serious inference concern arises from the fact that languages are not independent. Within a language family, individual tongues have evolved in parallel over many centuries. While this slow process of language development may help to address potential concerns about reverse causality, it complicates statistical inference. Intuitively, language characteristics are assigned at the level of a “cluster,” but countries draw from many different “clusters.” In relation to grammatical gender, these “clusters” do not exactly coincide with any existing categorical tier of language families. We address this issue by implementing a permutation test that respects both the distribution of languages across countries and the observed pattern of variation in treatment (i.e. grammatical gender) across and within language families. We cluster languages at the highest level of the language tree where we do not observe variation in grammatical gender. Generating 10,000 hypothetical assignments of grammatical gender across the 203 clusters so generated allows us to calculate permutation-test p-values indicating the likelihood that the association between grammatical gender and our outcomes of interest would be as strong as the observed relationship under the null hypothesis — given the structure of the language tree, the observed variation in grammatical gender across languages, and the distribution of languages across countries. Results suggest that the strong association between grammatical gender and women’s labor
force participation is not spurious.

To further assess the likelihood of a causal link between gender languages and women’s involvement in economic life, we examine the individual-level association between grammatical gender and women’s labor force participation and educational attainment in two parts of the world where both gender and non-gender languages are indigenous and widely spoken: Sub-Saharan Africa and India. Combining our language data with (i) Afrobarometer surveys from Kenya, Nigeria, Niger, and Uganda and (ii) the India Human Development Survey, we show that — within countries — grammatical gender is associated with larger gender gaps in educational attainment and labor force participation in two distinct contexts. Women whose native language is a gender language obtain less education and are less likely to be in the labor force than women whose native language is not a gender language, even after controlling for interactions between gender (i.e. the indicator for being female) and religious affiliation. The approach suggested by Altonji, Elder, and Taber (2005) and Oster (2017) suggests that unobservable characteristics are unlikely to explain the relationship. Thus, gender languages appear to reduce women’s labor force participation and lower their educational attainment in both Sub-Saharan Africa and India.

The rest of this paper is organized as follows. Section 2 introduces the concept of grammatical gender and surveys recent research on its impacts. Section 3 presents a theoretical framework illustrating the channels through which grammatical gender might lead to larger gender gaps in educational attainment and labor force participation. Section 4 provides an overview of our data sources, including the data we have compiled on the grammatical structure of more than 4,000 languages. Section 5 presents our cross-country analysis, and Section 6 presents individual-level, within-country analysis. Section 7 discusses causality. Section 8 concludes.
2 Grammatical Gender

Many languages partition the set of all nouns into mutually exclusive categories. Membership in these categories, which are typically referred to as either genders or noun classes (Corbett 1991, Aikhenvald 2003), can be manifest in several ways. Members of a noun class may be semantically related, or they may be linked by morphology. For example, members of the KI-/VI- class in Swahili often begin with ki- in the singular and vi- in the plural — e.g. “chair” is *kiti* and “chairs” is *viti*. However, though semantic and morphological regularities are a common characteristic of noun classes, they are not required. Instead, membership in a specific noun class is defined based on agreement: class must be reflected in the conjugation of associated words within the noun phrase or predicate in grammatically correct speech (Aikhenvald 2003). In Swahili, for example, the noun class determines the prefixes used to modify adjectives, verbs, demonstratives, and other parts of speech. So, “these new chairs” is *viti vipya hivi*, while “these new teachers” is *walimu wapya hawa* because the word “teacher” is part of the M-/WA- class rather than the KI-/VI- noun class.

Nouns are said to belong to the same agreement class if, “given the same conditions, they will take the same agreement form” (Corbett 1991, p. 148), where the relevant “conditions”

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3 There is some debate among linguists as to whether agreement rules that do not involve elements of the noun phrase or the predicate can form the basis of a noun class system — specifically, linguists disagree as to whether requiring “anaphoric agreement” between nouns and associated pronouns constitutes a system of grammatical gender (Corbett 1991, Aikhenvald 2003). Corbett (1991) argues that there is no fundamental distinction between pronominal agreement and other forms of grammatical agreement; he consequently classifies languages that (only) require pronominal agreement (e.g. English) as gender languages in his work (Corbett 2013a, Corbett 2013b, Corbett 2013c). Aikhenvald (2003) agrees that there is no fundamental distinction between pronominal agreement and other forms of grammatical concordance, but advocates the use of the traditional definition of grammatical gender to avoid confusion. She also suggests restricting the use of the term “grammatical gender” to systems of noun classification involving a relatively small number of categories that include masculine and feminine. Since our focus is on the links between grammatical gender and non-grammatical gender norms, we adopt her terminology to avoid confusion. Employing the traditional definition of grammatical gender also facilitates the use of data from a wide range of linguistic and anthropological sources, since many historical sources distinguish between grammatical gender (which involves the assignment of nouns to gender categories) and systems that mark natural/human gender morphologically.

4 Corbett (1991) states: “The existence of gender can be demonstrated only by agreement evidence. . . Evidence taken only from the nouns themselves, such as the presence of markers on the nouns, does not of itself indicate that a language has genders (or noun classes); if we accepted this type of evidence, then we could equally claim that English had a gender comprising all nouns ending in -ion.” Thus, though many nouns within a class may share particular prefixes or suffixes, it is the requirement that other parts of speech (particularly elements of the noun phrase or the predicate) conjugate or inflect appropriately that distinguishes noun classes from other phonological or orthographic partitions of the set of all nouns.
are linguistic and typically relate to number and case.

Systems of noun classification differ widely across languages, and not all languages have such a system. One of the most common bases for a system of noun classification is biological sex: (some) female humans and some other nouns are assigned to one category, while (some) male humans and some other nouns are assigned to a different category (Corbett 1991, Aikhenvald 2003, Hellinger 2003). Following Aikhenvald (2003) and Hellinger and Bußman (2003), we refer to systems which assign nouns, including some inanimate nouns, to agreement classes that are based on biological sex as grammatical gender; we refer to languages characterized by such systems of grammatical gender as gender languages. Spanish is a prominent example of a gender language: all Spanish nouns are either masculine or feminine, and both definite articles and adjectives must be consistent with a noun’s gender. So, for example, “the white house” is

\[
\text{la casa blanc-a} \\
\text{the.FEM house white-FEM,}
\]

because “house” is feminine, but “the white horse” is

\[
\text{el caballo blanc-o} \\
\text{the.MASC horse white-MASC}
\]

because “horse” is masculine. A Spanish speaker must therefore maintain a mental map that assigns each noun to one of these two distinct gender categories.

Systems of grammatical gender differ along several dimensions. Gender languages differ in the extent of agreement across parts of speech, and the extent to which the gender distinc-

\footnote{Almost all languages also distinguish between singular and plural, but this is not typically treated as a system of noun classification because the singular and plural forms are treated as two variants of the same noun.}

\footnote{Swahili, for example, has noun classes which determine agreement, but it is not a gender language because none of the Swahili noun classes relates to biological sex in any way.}

\footnote{Moreover, grammatical gender is only one of several ways that grammatical rules can make human gender distinctions salient. For instance, though typically not classified as a gender language, English employs a system of pronominal agreement — different third-person singular pronouns are used for male and female humans and, in some cases, male and female animals (Aikhenvald 2003, Boroditsky, Schmidt, and Phillips 2003, Hellinger and Bußman 2003, Kilarski 2013). Female pronouns have also traditionally been used to refer to ships and other large transportation vessels. Because pronouns agree with the natural gender of animate nouns, Corbett (1991) classifies English as a gender language with a strictly semantic system of noun classification (i.e. a system of grammatical gender based only on biological gender). Such systems of pronominal agreement based on the biological gender of animate referents (rather than the grammatical
tion represents a complete partition of the set of all nouns. Languages such as Spanish — with only two sex-based noun classes — are at one end of this spectrum. In such languages, every inanimate noun must be classified as either feminine or masculine. Languages such as German display a weaker form of grammatical gender because some objects are classified as neither feminine nor masculine. Intuitively, one might think that the partition of nouns into two dichotomous genders suggests that other aspects of the universe should also be so organized (for example, into male and female household tasks). In systems that assign objects (i.e. nouns) without natural gender to gender categories, there is also the question of what the observed grouping signals about the relative status of women and men. Though the rules used to assign nouns to different classes are often phonological (e.g. Spanish nouns that end in “o” are typically masculine), many languages assign some nouns to the feminine gender using semantic guidelines that have a certain cultural intelligibility. For example, dangerous objects are feminine in the Australian language Dyirbal (Lakoff 1987), while one linguist studying the Siberian language Ket suggested that certain small animals were feminine “because they are of no importance to the Kets” (Corbett 1991, p. 19).

8In many languages, the grammatical gender of inanimate objects reflects stereotypes about the physical distinctions between males and females. For example, in his discussion of the major Indo-Aryan languages (Bengali, Gujarati, Hindi, Marathi, Oriya, Panjabi, and Sindhi), John Beames (1875) notes: “In all the five languages which have gender expressed, the masculine is used to denote large, strong, heavy, and coarse objects; the feminine weak, small, and fine ones” (p. 148). In the Papuan language Manangu, inanimate objects that are long or thin are masculine, while those that are short or round are feminine (Aikhenvald 2003).

9No one knows exactly why grammatical gender systems arose in some language families and not in others. Janhunen (1999) hypothesizes that a single innovation in an ancient West Asian language brought grammatical gender into the Indo-European language family, but grammatical gender arose in indigenous language families on every continent. It is, of course, impossible to fully rule out the possibility that some aspect of culture contributed to the emergence of grammatical gender in certain ancestral languages. That said, since language structures evolve over centuries, even millennia, present-day gender attitudes cannot have had a causal impact on modern grammatical structures. Moreover, we have a relatively good understanding of the process through which grammatical gender was lost from certain widely spoken Indo-European languages; this evidence does not suggest a causal relationship between gender norms and the loss of grammatical gender. For example, McWhorter (2005) argues that the influx of Scandinavian adults into the community of English speakers contributed to the loss of grammatical gender, as an imperfect grasp of inflectional agreement paradigms is common among non-native speakers. This “contact hypothesis” may also explain why grammatical gender is typically absent from Creole languages (McWhorter 2005, Muhleisen
Whether grammatical gender distinctions influence (non-grammatical) gender attitudes is an empirical question, but the idea that they might is not new. Whorf, for example, argued that gender distinctions in language might make a gendered division of labor seem more natural, suggesting that viewing the world through the lens of a gender language would create “a sort of habitual consciousness of two sex classes as a standing classificatory fact in our thought-world” (Whorf 2011[1956], p. 69). This argument — which Whorf advanced without offering any empirical evidence to support it — has been controversial, to say the least. However, recent work in psychology and political science shows that grammatical gender shapes our subconscious attitudes in subtle and surprising ways. For example, Boroditsky, Schmidt, and Phillips (2003) conducted a study — in English — of native speakers of Spanish and German (all of whom were fluent in English); participants in the study were asked to provide (English) adjectives to describe pictures of objects that had been chosen because they had opposite grammatical genders in Spanish and German. Subjects tended to choose adjectives that aligned with the grammatical gender of the noun in their native language. For example, native German-speakers described a picture of a bridge (which is feminine in German) as “beautiful” and “elegant” while native Spanish-speakers described the same (masculine in Spanish) bridge as “big” and “dangerous” (Boroditsky, Schmidt, and Phillips 2003). Thus, the results suggest that grammatical gender shapes the way we think about inanimate objects without inherent biological gen-

and Walicek 2010). However, the reduction and simplification of languages resulting from an influx of non-native speakers is not restricted to the loss of grammatical gender (and has no inherent relationship to societal gender norms): McWhorter (2005) argues that the contact hypothesis also explains why Swahili is one of the few Bantu languages that is not tonal. Kastovsky (1999) proposes a complementary explanation, arguing that the English case-number-gender agreement system was, in essence, made precarious by its own complexity and the absence of reliable morphological rules that could be used to predict agreement classes; in this context, small changes in pronunciation could lead to the conflation of declensional paradigms and their subsequent loss. Aikhenvald (2003) points to a similar process of declensional conflation and subsequent gender loss in Bengali and Persian, and to a parallel loss of the neuter gender in French. Thus, the existing evidence tends to suggest that grammatical gender is most often lost through an interplay between linguistic factors (e.g. sound change, similarity between agreement paradigms) and the arrival of large numbers of non-native speakers within a linguistic community.

10His argument echoes earlier work by Durkheim and Mauss (1963), who highlighted the parallels between culture-specific systems for classifying humans and those used for classifying other aspects of reality. Describing the extension of the clan system of one group of native Australians to the universe of animals and inanimate objects, they wrote: “The reasons which led to the establishment of the categories have been forgotten, but the category persists and is applied, well or ill, to new ideas” (p. 21).
der. Grammatical gender also appears to shape gender attitudes — even within individuals. Pérez and Tavits (forthcoming) conduct a survey experiment with Estonian/Russian bilinguals, randomizing the language in which they are interviewed. They show that bilinguals who are interviewed in Russian (a gender language) are less supportive of gender equality than those who are interviewed in (non-gender) Estonian, even though interview languages were randomly assigned.¹¹

Recent work also suggests that the influence of grammatical gender extends into the economic realm. Using the World Atlas of Language Structures (WALS), a comprehensive data set on the grammatical structure of more than 500 languages, a number of authors have examined the links between grammatical gender and economic and political outcomes. For example, Mavisakalyan (2015) and Shoham and Lee (2017) use the WALS to examine the cross-country association between grammatical gender and gender inequality in the labor force. Santacreu-Vasut, Shoham, and Gay (2013) show that countries where the national language uses a sex-based system of grammatical gender are less likely to implement gender quotas for political office, while Santacreu-Vasut, Shenkar, and Shoham (2014) find that those countries also have relatively fewer women in corporate leadership positions. Hicks, Santacreu-Vasut, and Shoham (2015) show that immigrants to the United States assign tasks within the household along gendered lines if they grew up speaking a gender language; no such difference is found among immigrants who came to the U.S. before the age of language acquisition, or among the children of immigrants.¹² Importantly, these findings suggest that one’s native language plays a particularly crucial role in shaping one’s views on the appropriate role for women in society.

¹¹There is also evidence that pronominal gender impacts the salience of gender distinctions. Guiora (1983) finds that children who grow up speaking Hebrew, English, or Finnish come to understand their own biological genders at different ages; those who grow up using different pronouns for males and females become aware of their own natural gender earlier. As discussed above, English has a system of pronominal gender while Finnish does not. Hebrew also uses a dichotomous system of grammatical gender (all nouns are either masculine or feminine), and male and female Hebrew-speakers must use grammatically correct verb forms, for example, that reflect their natural gender. Hebrew also uses different second-person pronouns for males and females.

¹²In related work, Gay, Hicks, Santacreu-Vasut, and Shoham (2017) find that female immigrants to the United States exhibit lower labor market participation (working fewer hours, fewer weeks, etc.) if they speak a gender language at home.
These analyses suffer from the incompleteness of the WALS. Using it alone, within-country analysis of data from Africa or Asia is not feasible. Cross-country regressions require researchers to calculate country-level averages of a variable (the grammatical gender structure of one’s native language) that is missing for half the world’s population. One of the most cautious approaches to this missing data problem is to use Manski-style bounds, but doing so yields upper and lower bounds which contain almost the entire support of conceivable values. Moreover, the absence of data also limits the extent to which one can correct for the non-independence of languages within families while maintaining adequate statistical power. Robust inference requires an expanded data set on linguistic structures.

3 Conceptual Framework

Existing work examining the empirical relationship between grammatical gender and women’s involvement in economic life has not formally specified the potential causal pathway. In this section, we outline a stylized model that illustrates how grammatical gender — which may predispose us to think of things as either masculine or feminine — could induce gender disparities in education and labor force participation. The model is inspired by Whorf’s suggestion that a grammatical gender system makes the partition of the non-linguistic world into masculine and feminine domains appear more natural. We formalize this intuition by introducing a psychic cost $\phi > 0$ that a person who has grown up speaking a gender language experiences when she (resp. he) enters a domain dominated by the opposite sex.

We endogenize the definition of masculine and feminine domains by assuming that a domain (e.g. a school, the workforce, etc.) is masculine (resp. feminine) whenever the proportion of women (resp. men) in that domain falls below some threshold $\lambda \in [0,1]$. Thus, when the proportion of women in, say, the workforce is below $\lambda$, the work world is perceived as a masculine domain — so, women face a psychic cost when they choose to work outside the home. Symmetrically, if the proportion of women in the workforce exceeds $1-\lambda$, the workforce would be perceived as a feminine domain, and men would face

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13 We discuss these clusters further in Section 5.5.2.
a psychic cost when they chose to work. Equilibrium requires that each individual make a rational choice about whether or not to enter a domain conditional on the cost structure that results from the realized distribution of genders across each domain.

3.1 Education

We consider a simple model of educational attainment where students attend school whenever the expected benefits exceed the immediate costs. The net return to education depends on ability and may also differ across genders. We formalize the set-up as follows, first without grammatical gender and then introducing it. Girl i’s ability is given by $\gamma_i > 0$, where $\gamma \sim F_\gamma$ (for some smooth, etc. function $F_\gamma$). Let $R_g(\gamma_i)$ denote the net return to schooling for a girl with ability level $\gamma_i$. Without loss of generality, we assume that $R_g(\cdot)$ is net of any monetary costs of attending school. The return to education is increasing in ability: $R_g'(\gamma_i) > 0$. In the absence of grammatical gender, a girl will attend school whenever $R_g(\gamma_i) > 0$. As a result, there exists $\gamma^*$ such that $R_g(\gamma^*) = 0$, and a proportion $1 - F_\gamma(\gamma^*)$ of girls (all those with $\gamma_i \geq \gamma^*$) attend school.

The setup is symmetric for boys. Boy i’s ability is given by $\beta_i > 0$, where $\beta \sim F_\beta$. In the absence of grammatical gender, a boy with ability level $\beta_i$ will attend school whenever $R_b(\beta_i) > 0$. There exists $\beta^*$ such that $R_b(\beta^*) = 0$, and all boys with $\beta_i \geq \beta^*$ attend school.

With equal numbers of girls and boys in the population, girls represent proportion

$$P_{girls} = \frac{1 - F_\gamma(\gamma^*)}{2 - F_\beta(\beta^*) - F_\gamma(\gamma^*)}$$

of students enrolled in school. The model is symmetric: if $F_\gamma = F_\beta$ and $R_g(\cdot) = R_b(\cdot)$, then $\lambda^* = \frac{1}{2}$ and $\gamma^* = \beta^*$.

When grammatical gender predisposes us to view domains as either masculine or fem-

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14 To focus on the key implications of the model, we assume that those who did not grow up speaking gender languages do not experience such psychic costs — though, of course, they may experience other social or other emotional costs when entering environments where they do not fit in. One could easily extend our model to consider the possibility that these costs exist for everyone but might be larger when partitioning the world in masculine, feminine, and potentially neutral spheres appears more natural.
inine, there are three possible equilibria: school can be either masculine, neutral (non-gendered), or feminine. In the masculine equilibrium (if it exists), boys attend school whenever \( R_b (\beta_i) \geq 0 \), but girls only attend if \( R_g (\gamma_i) \geq \phi \) — for girls, going to school entails a psychic cost because they perceive school as a masculine domain. An equilibrium exists if the set of children who would attend school conditional on the distribution of psychic costs associated with that equilibrium yields a gender composition (of students) consistent with that equilibrium. So, for example, it is possible for school to be a masculine domain in equilibrium if the set of students who would attend school when girls face a psychic cost but boys do not skews sufficiently male to keep the proportion of girls in the student body below \( \lambda \).

As we show in the Online Appendix, at least one of the three possible equilibria always exists. More interestingly, multiple equilibria are often possible, but both welfare and human capital attainment are highest in the gender-neutral equilibrium. In this context, policies such as single-sex schools can improve welfare and increase human capital by allowing girls (or boys) to attend school without the psychic costs associated with entering an environment that is perceived as the domain of the opposite sex.\(^{15}\) Other policies that increase the net return to education — for example, eliminating school fees or making education compulsory (which introduces costs for non-attendance) — can have indirect effects on female enrollment by changing the expected proportion of girls who attend school. If these policies bring the expected ratio of girls to boys closer to parity, the gendered equilibrium may cease to exist. Moreover, when multiple equilibria are possible, such policies have the potential to nudge a society from one feasible equilibrium to another.

### 3.2 Labor Force Participation and the Division of Household Tasks

Next, we consider the decision problem facing two parents who maximize their consumption while caring for their children. Again, we assume that the ability of girl/female/woman/mother \( i \) is characterized by \( \gamma_i \sim F_\gamma \) and the ability of boy/male/man/father is characterized by

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\(^{15}\) Fryer and Levitt (2010) discuss the prevalence of same-sex schools in middle Eastern countries where gender gaps in educational attainment are small but gender gaps in labor force participation persist.
\( \beta_i \sim F_{\beta, 16} \) \( \gamma \) and \( \beta \) both have continuous support between 0 and some finite maxima, \( \beta^{\max} \) and \( \gamma^{\max} \).

A household maximizes consumption:

\[
C = w_{\text{mom}} L_{\text{mom}} + w_{\text{dad}} L_{\text{dad}} - w_{\text{nanny}} H_{\text{nanny}}
\]

where \( w_{\text{mom}} = \gamma_i \) indicates the wage that a mom of ability \( \gamma_i \) earns if she works outside the home, \( w_{\text{dad}} = \beta_i \) the wage that a dad of ability \( \beta_i \) earns if he works outside the home, and \( w_{\text{nanny}} \) represents the market wage paid to nannies. We assume that nannies are female, and that they are young women who would not be included in the adult labor force if they were not employed as nannies (for example, au pairs, older sisters).\(^{16}\) Both mom and dad have one unit of time which they allocate to either work outside the home or childcare:

\( H_{\text{mom}} + L_{\text{mom}} = 1 \) and \( H_{\text{dad}} + L_{\text{dad}} = 1 \). One unit of adult time must be spent caring for the child: \( H_{\text{mom}} + H_{\text{dad}} + H_{\text{nanny}} = 1 \).

First, consider the case where there are no gendered domains. A household will hire a nanny to take care of the children whenever both the mother and the father are both able to earn more than the nanny’s wage — i.e. when \( \beta_i \geq w_n \) and \( \gamma_i \geq w_n \). When \( \gamma_i < w_n \) and \( \gamma_i \leq \beta_i \), the mother stays home while the father works. When \( \beta_i < w_n \) and \( \gamma_i > \beta_i \), the father stays home while the mother works. Panel A of Figure 1 illustrates this partition of the space of possible parental ability levels.

Let \( f_{\beta, \gamma}(\beta, \gamma) \) denote the joint distribution of \( \beta \) and \( \gamma \). In the absence of gendered domains, define \( P^*_{\text{mom}} \) as the proportion of households where the mother stays at home:

\[
P^*_{\text{mom}} = \int_{\beta=0}^{\beta=w_n} \int_{\gamma=0}^{\gamma=w_n} f_{\beta, \gamma}(\beta, \gamma) + \int_{\beta=w_n}^{\beta=\beta^{\max}} \int_{\gamma=0}^{\gamma=w_n} f_{\beta, \gamma}(\beta, \gamma).
\]

\(^{16}\)For obvious reasons, using the subscripts \( m \) and \( f \) to distinguish between male and female adults who are also mothers and fathers might be confusing.

\(^{17}\)While this assumption is realistic in a range of contemporary and historical settings, it also serves a purpose by increasing the likelihood that the home environment is a predominantly a feminine domain. Other ways of achieving the same goal (for example, endogenizing fertility and making it costly for women to enter the work force when children are very young) make the model too realistic to be useful.
In other words, $P_{mom}^*$ is the integral of $f_{\beta,\gamma} (\beta, \gamma)$ over the “mom at home” region in Figure 1. $P_{dad}^*$ and $P_{nanny}^*$ can be defined analogously:

$$P_{dad}^* = \int_{\beta=0}^{\beta=\gamma} \int_{\gamma=0}^{\gamma=w_n} f_{\beta,\gamma} (\beta, \gamma) + \int_{\beta=w_n}^{\beta=\gamma} \int_{\gamma=0}^{\gamma=w_n} f_{\beta,\gamma} (\beta, \gamma)$$

(4)

and

$$P_{nanny}^* = \int_{\beta=w_n}^{\beta=\gamma_{max}} \int_{\gamma=w_n}^{\gamma=\gamma_{max}} f_{\beta,\gamma} (\beta, \gamma).$$

(5)

For any $f_{\beta,\gamma} (\beta, \gamma)$, $P_{mom}^* + P_{dad}^* + P_{nanny}^* = 1$ since households must either have mom, dad, or a nanny at home with the children. Since all households have exactly one person at home, the proportion of homes where a woman takes care of the children is $P_{mom}^* + P_{nanny}^*$.

The proportion of women in the (out-of-the-home) workforce is:

$$\frac{P_{dad}^* + P_{nanny}^*}{1 + P_{nanny}^*}$$

(6)

since households with a nanny at home send both a man and a women into the workforce. $(P_{mom}^*, P_{dad}^*, P_{nanny}^*)$ is an equilibrium in a trivial sense, since every household optimizes and individual (household) optima are not strategically interdependent.
When individuals are predisposed to view domains as gendered (so \( \lambda \) and \( \phi \) play a role in decision-making), the equilibrium described above is one of nine that might exist. Home and work can both be either masculine, neutral (non-gendered), or feminine. Each of the nine candidate equilibria is a pair \( HW \) where \( H \in \{M, N, F\} \) characterizes the "home" domain and \( W \in \{M, N, F\} \) characterizes the "work" domain. So, the NN equilibrium would be one in which neither home nor work is a gendered domain, whereas the FM equilibrium would be one in which home is a feminine domain and work is a masculine domain.

The NN equilibrium, if it exists, is characterized by the same pattern of observed in the absence of grammatical gender (as shown in Panel A of Figure 1): both parents work whenever \( \gamma_i > w_n \) and \( \beta_i > w_n \), and the parent who would earn the lower wage stays home with the child otherwise. Hence, \( P_{NN}^{dad} = \hat{P}_{dad}^{*} \), \( P_{NN}^{mom} = \hat{P}_{mom}^{*} \), and \( P_{NN}^{nanny} = \hat{P}_{nanny}^{*} \). However, when domains can be gendered, this is only an equilibrium when

\[
\lambda < P_{NN}^{mom} + P_{NN}^{nanny} < 1 - \lambda
\]  

and

\[
\lambda < \frac{P_{NN}^{dad} + P_{NN}^{nanny}}{1 + P_{NN}^{nanny}} < 1 - \lambda.
\]  

(7)  

(8)

In other words, the equilibrium proportion of women taking care of children (i.e. households where a female takes care of the child) and the proportion of women in the (out-of-the-home) workforce must both fall between \( \lambda \) and \( 1 - \lambda \) for a neutral equilibrium — in which neither home nor work is a gendered domain — to exist. It is apparent that this becomes less likely when \( \lambda \) is close to one half and the scope for non-gendered domains is limited.

Next, consider the FN equilibrium, where home is a feminine domain but work is neither masculine nor feminine. If such an equilibrium exists, a man who stays home with his children will experience a psychic cost of \( \phi > 0 \). Total household utility if the father
provides childcare is therefore given by

\[ C = \gamma_i - \phi. \] (9)

In this equilibrium, a household will hire a nanny whenever \( \beta_i > w_n - \phi \); the father will stay home whenever \( \beta_i < \gamma_i - \phi \) and \( \beta < w_n - \phi \); and the mother will stay home whenever \( \beta_i > \gamma_i - \phi \) and \( \gamma_i < w_n \). As Panel B of Figure 1 illustrates, two types of men who do not work in the NN equilibrium will work in the FN equilibrium. Men whose wives work (because \( \gamma_i > w_n \)) will now enter the workforce whenever their ability (\( \beta_i \)) falls between \( w_n - \phi \) and \( w_n \). Men will also work whenever \( \gamma_i - \phi < \beta_i < \gamma_i < w_n \); their higher-ability wives will stay home because childcare (“women’s work”) entails a psychic cost for men when relatively few men stay home. Both of these changes lower the average ability level among those in the labor force.

In the Online Appendix, we characterize the feasible equilibria in greater detail and demonstrate that at least one equilibrium always exists. As in the case of educational attainment, multiple equilibria are possible, and the ability level of the labor force is always highest in the NN equilibrium, where neither home nor work is a gendered domain.

### 3.3 Implications of the Model

The model does not demonstrate that grammatical gender necessarily predicts lower educational attainment and labor force participation among women than men. Instead, we formalize Whorf’s intuition that grammatical gender predisposes us toward the view that men and women should exist in separate domains. We have kept the model as symmetric as possible while still recognizing the empirical fact that women and girls do all of the birthing and most of the childcare work in every human society ever studied (Lancy 2015). Because of its symmetry, our model allows for the possibility that gendered equilibria could exist in which boys attain less education than girls and in which men are less likely to work outside the home than women – though such equilibria are unlikely when the returns to education
are higher for males and childcare is costly.

The key prediction of the model is that the minimum equilibrium level of girls’ educational attainment and women’s labor force participation is lower with grammatical gender than without. This motivates the empirical tests presented in the rest of the paper. However, the model also demonstrates that grammatical gender is more a nudge than a constraint. When \( \lambda < 0.5 \), gender-neutral equilibria become possible.\(^{18}\) Hence, many societies where women’s labor force participation is still very low have the potential to move to a more equitable equilibrium very rapidly. The model also highlights the potential for policy responses that work around the subtle influence of the tendency to partition the world into sex-specific domains by creating female-centric spaces in the modern sector as many Middle Eastern countries have done to improve girls’ educational outcomes.

4 Data

We compile a new data set characterizing the gender structure of more than 4,000 living languages. Together, the languages that we classify account for over 99 percent of the world’s population. As discussed below, we collate data from a range of academic publications, pedagogical materials (e.g. language textbooks), and historical sources. The downside of this approach is that there may be measurement error at the language level: while many sources explicitly state that a language either does or does not use a system of grammatical gender, we cannot always be certain that the same precise definition of grammatical gender is being used across sources.\(^{19}\) The strength of our approach is that we are able to characterize the grammatical structure of thousands of languages accounting for almost all of the world’s population.

\(^{18}\) There are several ways of making this precise. For example, the closer \( \lambda \) is to zero, the wider the range of returns to education, wages, and joint distributions of ability that support a gender-neutral equilibrium; when ability levels and the returns to education do not differ by gender, any value of \( \lambda < 0.5 \) permits a gender-neutral equilibrium for education. The situation in the labor market is more complicated, both because of the presence of nannies and because of matching in the marriage market, upon which we have not imposed any structure.

\(^{19}\) Indeed, even recent work by linguists does not always agree on the definition of grammatical gender — see Corbett (1991) and Aikhenvald (2003) for discussion.
4.1 Building a Grammatical Gender Data Set

Data on the world’s native languages comes from the *Ethnologue*, a comprehensive database of over 7,000 languages (Lewis, Simons, and Fennig, eds., 2016). Combining the *Ethnologue* data with information on the grammatical gender structure of the world’s languages allows us to construct an estimate of the fraction of each country’s population that speaks a gender language as their native language. Of the 7,457 languages included in the *Ethnologue* database, we drop languages that are extinct or have no native speakers, sign languages, and dying languages that had fewer than 100 native speakers when last assessed by *Ethnologue* researchers. This leaves 6,190 languages. Together, these languages account for an estimated 6.50 billion native speakers. Of these, we successfully identify academic or historical sources characterizing the gender structure of native languages accounting for 6.44 billion native speakers (or more than 99 percent of the total).

Data on the gender structure of languages comes from a range of sources. Three of the best known are: the World Atlas of Language Structures (WALS), which characterizes the noun classification system of 525 languages; George L. Campbell’s *Compendium of the World’s Languages* (Campbell 1991); and George Abraham Grierson’s eleven-volume *Linguistic Survey of India* (Grierson 1903a, 1903b, 1904, 1905, 1907, 1908, 1909, 1916, 1919, 1921), which was compiled between 1891 and 1921 and covers more than 300 South Asian languages and dialects. Additional data on the grammatical gender structures of languages comes from academic articles and teaching materials focused on individual languages. We also collected first-person accounts from native speakers for a small number of relatively undocumented languages (e.g. Fiji Hindi and Rohingya). Detailed information on the range of sources (including quotes used to characterize each language’s grammatical gender) is provided in our (Online) Data Construction Appendix.

For each mother tongue in the *Ethnologue* database, we code two variables characterizing the language’s grammatical gender structure. First, we create an indicator for using any system of grammatical gender. We code a language as a gender language if it meets two criteria: first, the language must use a system of noun classes that includes masculine
and feminine as two of the possible categories; second, the masculine and feminine categories must include some inanimate objects — i.e. assignment to the gender noun classes should not be based exclusively on the biological sex (or human gender) of the referents.

Second, whenever possible, we also code an indicator for dichotomous gender languages (e.g. Spanish) that assign all nouns to either the masculine or the feminine noun class.

We successfully classify 4,346 languages which together account for more than 99 percent of the world’s population. We classify all but four of the 383 languages with more than one million native speakers, and we are able to confirm the gender structure using two independent data sources for 324 of these large languages. We are able to account for more than 99 percent of the population in 171 of 193 countries, and we account for less than 95 percent of the population in only eight countries: Eritrea (94.5 percent of native speakers coded), the Iran (93.7 percent), Ethiopia (92.6 percent), the Laos (90.2 percent), Timor-Leste (90.0 percent), Cameroon (89.1 percent), Chad (75.4 percent), and Papua New Guinea (32.0 percent).

Figure 2 characterizes the distribution of gender languages around the world. While many countries are dominated by either gender or non-gender languages, there is considerable within-country variation in Canada and the United States, Sub-Saharan Africa, South Asia, and the Andean region of South America. Across all countries, we estimate that approximately 38.6 percent of the world’s population speaks a gender native language.

4.2 Other Sources of Data

Additional data for our cross-country analysis comes from several sources. Data on labor force participation, income, and population come from the World Bank’s World Development Indicators database. We use data on labor force participation in 2011, which is available for 178 countries. We also use data on primary and secondary school comple-

\(^{20}\)As discussed above, linguistic sources do not always use the same implicit definition of grammatical gender. For example, the phrase “marks gender” can be used to indicate either grammatical gender or a more limited system of indicating the gender of a human referent. Since many linguistic sources explicitly distinguish between grammatical gender and lexical marking of human/animate gender, we only use sources that indicate whether inanimates are classed in terms of nominal gender.
tion from the Barro-Lee Educational Attainment Data Set (Barro and Lee 2013), which is available for 142 countries. Data on gender attitudes comes from the World Values Survey and is available for 56 countries (World Values Survey Association 2015). Finally, we take several country-level geographic controls (average precipitation and rainfall plus suitability for the plough) from Alesina, Giuliano, and Nunn (2013). These data are available for 173 countries.

Data for our individual-level analysis comes from two sources. For African countries, we use the nationally-representative Afrobarometer Surveys (Afrobarometer Data 2016). Afrobarometer surveys have been conducted in 36 African countries and are representative of the voting age population within each country. We use data from four countries where gender and non-gender languages are indigenous and widely spoken: Kenya, Niger, Nigeria, and Uganda. Data for Niger is only available in Round 5 of the Afrobarometer (2011–2013). For the other three countries, four rounds of data are available: 2002–2003, 2005–2006, 2008–2010, and 2011–2013.\footnote{Kenya, Nigeria, and Uganda were also included in the first round of the Afrobarometer. However, that data set does not contain detailed information on native languages.} We successfully classify the grammatical gender structure of the native languages of 99.1 percent of respondents, yielding a data set of 26,546 respondents who speak 175 different native languages.

We replicate our within-country analysis for India using the India Human Development Survey (Desai, Dubey, and Vanneman 2015). The IHDS includes data on 76,351 household heads and their spouses living in 33 Indian states. We are able to classify the grammatical gender structure of the native language of 99.5 percent of IHDS respondents, yielding a data set of 75,966 observations.

5 Cross-Country Analysis

5.1 Empirical Strategy

In our cross-country analysis, we examine the association between women’s labor force participation and the proportion of a country’s population whose native language is a
gender language, $Gender_c$. Our main empirical specification is an OLS regression of the form:

$$LFP_c = \alpha + \beta Gender_c + \delta c + \lambda X_c + \varepsilon_c$$

where $LFP_c$ is women’s labor force participation in country $c$ (in 2011), $Gender_c$ is the proportion of the population of country $c$ whose native language is a gender language, $\delta c$ is a vector of continent fixed effects, $X_c$ is a vector of of country-level geography controls, and $\varepsilon_c$ is a conditionally mean-zero error term.\footnote{Standard errors are clustered at the language level (by the most widely spoken language within each country).}

Our main outcome of interest is women’s labor force participation. However, we do not wish to conflate gender differences in labor market participation with structural factors that impact labor force participation among both men and women. To rule out this possibility, we include specifications where the outcome variable is the gender difference in labor supply, i.e. women’s labor force participation minus men’s labor force participation.\footnote{As a robustness check, we report specifications that use the ratio of women’s labor force participation to men’s labor force participation as the outcome variable (see Online Appendix Table A1).}

We also examine two other outcome variables related to gender norms: women’s educational attainment and gender attitudes. Our analysis of educational outcomes parallels our analysis of labor force participation. We examine rates of primary and secondary school completion among women and differences between women’s and men’s completion rates. In our analysis of WVS data on gender attitudes, we construct an index of gender attitudes by taking the first principal component of the eight WVS questions on gender roles. Since we are considering attitudes rather than behaviors, we do not report gender differences; instead we compare attitudes by gender to test whether grammatical gender shapes the views of traditional gender roles among both men and women.

\footnote{As discussed further below, our results are also robust to the inclusion of additional contemporaneous controls such as log GDP per capita and population. However, such controls might be directly impacted by gender norms and women’s involvement in the labor force, creating a “bad controls” problem and biasing the coefficient of interest (Angrist and Pischke 2008, Acharya, Blackwell, and Sen 2016). We therefore focus on geographic controls — proportion tropical, precipitation, temperature, suitability for the plough, and an indicator for being landlocked — which are plausibly exogenous.}
5.2 Labor Force Participation

Figure 3 summarizes female labor force participation in the 178 countries for which data is available. The figure highlights the fact that women’s participation in economic life varies tremendously across countries: the women’s labor force participation rate ranges from 9 percent in the Yemen to 87 percent in Madagascar. Figure 3 suggests a negative relationship between the prevalence of gender languages and women’s involvement in the labor force. In the figure, darker bars indicate a higher prevalence of grammatical gender. It is clear that many of the countries with the lowest levels of women’s labor force participation and the largest gender gaps in labor force participation are those where gender languages are dominant.

We confirm the statistical significance of this relationship in a regression framework in Table 1. In the first three columns, the outcome variable is the average level of female labor force participation in country \( c \). We report a parsimonious specification with no controls in Column 1. Gender languages are negatively and significantly associated with lower levels of female labor force participation. The coefficient estimate suggests that women’s labor force participation is 13.83 percentage points higher in the absence of gender languages (p-value \( 2.29 \times 10^{-6} \)). Column 2 of Table 1 reports a specification that includes continent fixed effects; Column 3 also includes geographic controls (percentage tropical, average temperature and precipitation, an indicator for being landlocked, and suitability for plough agriculture). The coefficient of interest is negative and statistically significant in both specifications. Moreover, it remains reasonably similar in magnitude: when all of our geographic controls are included, the coefficient suggests that grammatical gender is associated with an 11.92 percentage point decline in women’s labor force participation (p-value \( 5.04 \times 10^{-4} \)).

In Columns 4 through 6 of Table 1, we replicate our analysis using the gender difference in labor force participation as the dependent variable. Gender languages are also associated with robust differences in women’s labor force participation relative to men.\(^{24}\) As shown in Online Appendix Table A1, we obtain similar results when we use the ratio of female labor force participation to male labor force participation as the outcome variable.
parsimonious specification with no controls (Column 4), we find that grammatical gender is associated with an 11.61 percentage point increase in the gender gap in labor force participation (p-value $6.22 \times 10^{-6}$). When we include continent fixed effects and country-level geography controls, the coefficient rises to suggest that grammatical gender is associated with a 14.66 percentage point increase in the gender difference in labor force participation (p-value $1.37 \times 10^{-5}$). Thus, the proportion of a country’s population whose native language is a gender language is a robust predictor of gender differences in labor force participation.

Taken at face value, our coefficient estimates suggest that grammatical gender might keep as many as 125 million women around the world out of the labor force.

### 5.3 Educational Attainment

Next, we examine the association between grammatical gender and women’s educational attainment. Education is a key determinant of wages; in many countries, gender differences in educational attainment translate into gender gaps in wages and economic empowerment (Grant and Behrman 2010). Nonetheless, gender gaps in primary and secondary school completion are not nearly as large as gender gaps in labor force participation. Across the 142 countries in the Barro-Lee data set, the average gender gap in primary school completion is only six percentage points and the average gender gap in secondary school completion is only four percentage points. This reflects the very high rates of primary school completion in many parts of the world: more than two thirds of the countries in the Barro-Lee data set have rates of primary school completion above 90 percent for both men and women.  

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25In the Online Appendix, we report a range of robustness checks, all of which suggest that the relationship between grammatical gender and female labor force participation is not driven by outliers or specification choices. In Online Appendix Table A2, we show that our main result is robust to the inclusion of a range of “bad controls” — intermediate outcomes that could themselves have been impacted by grammatical gender. As is well known, including such controls could bias the coefficient of interest, making it impossible to interpret (Angrist and Pischke 2008, Acharya, Blackwell, and Sen 2016). Nevertheless, we note that our main result is robust to the inclusion of controls for log GDP per capita, population, major world religions, and an indicator for post-Communist regimes. In Online Appendix Table A3, we demonstrate that our results hold when we drop each of the major world languages — Arabic, English, and Spanish. Finally, in Online Appendix Table A4, we include an additional variable for the proportion of a country’s population whose native language is a dichotomous gender language with only two noun classes (masculine and feminine). Results suggest that even weak forms of grammatical gender predict women’s (lack of) involvement in the labor force.
Moreover, many wealthy countries have compulsory schooling laws which reduce gender gaps in educational attainment.

In Table 2, we examine the cross-country relationship between grammatical gender and primary school completion. As expected, the relationship is positive and significant when continent controls are not included — reflecting the fact that primary school completion rates are highest in Europe, where gender languages are dominant. Once continent fixed effects are included, the estimated association is negative but not statistically significant. In Columns 4 through 6 of Table 2, we examine the relationship between grammatical gender and the gender gap in primary school completion. After including continent fixed effects, we find a negative relationship that is marginally statistically significant. Coefficient estimates suggest that grammatical gender is associated with a 3.72 percentage point increase in the gender gap in primary school completion (Table 2, Column 6, p-value 0.088).

We observe an even more muted cross-country relationship between gender languages and secondary school completion (Table 3). After including continent fixed effects, the association between grammatical gender and female secondary school completion is never statistically significant, nor do we observe a statistically significant association between grammatical gender and the gender gap in secondary school completion. Thus, grammatical gender explains cross-country variation in female labor force participation, but does not explain most of the observed cross-country variation in women’s educational attainment.

5.4 Gender Attitudes

Our main measure of gender attitudes is an index that we construct by taking the first principal component of the eight World Values Survey (WVS) questions related to gender. In Figure 4, we plot the cross-country relationship between each of these questions and the proportion of a country whose native language is a gender language. The prevalence of gender languages predicts responses to seven of the eight WVS questions.

In Table 4, we confirm the association between the prevalence of gender languages and our summary index of gender attitudes in a regression framework. After controlling for
continent fixed effects and country-level geography, the coefficient estimate suggests that grammatical gender is associated with greater support for traditional gender roles. To put the coefficient magnitudes in context, the estimates indicate that grammatical gender alone could explain the gap in gender attitudes between Ukraine (at the 55th percentile) and Trinidad and Tobago (at the 80th percentile). Thus, the estimated association between grammatical gender and non-grammatical gender attitudes is both statistically and culturally significant.

If grammatical gender shapes gender attitudes, we would expect it to impact the beliefs of both men and women. In Table 5, we show that — as expected — we observe a negative association between the country-level prevalence of grammatical gender and gender attitudes among both women (Columns 1 through 3) and men (Columns 4 through 6). The association is always statistically significant after including continent fixed effects. Moreover, though the coefficient is slightly larger for men, we can never reject equality across genders. Thus, the cross-country evidence suggests that grammatical gender predicts gender differences in behavior (specifically, involvement in the labor force), but also predicts traditional gender attitudes among both men and women.

5.5 Robust Inference

In this section, we discuss two potential concerns with our cross-country analysis. First, as discussed above, we were unable to classify the gender structure of some languages. Though these language tend to be small (in terms of numbers of native speakers), they account for more than one percent of the population in 22 countries. In Section 5.5.1, we present estimation that adjusts for the interval nature of our independent variable of interest, the proportion of each country’s population whose native language is a gender language. In Section 5.5.2, we consider the fact that language structures may be correlated within language families, since modern tongues evolved from common ancestors (Roberts, Winters, and Chen 2015). To address the potential correlation within families while maximizing statistical power (by exploiting variation in grammatical gender both across and between
families), we introduce a permutation test based on the structure of the language tree.

5.5.1 Measurement Error

In our cross-country analysis, our independent variable of interest is the proportion of the population whose native language is a gender language. However, as discussed above, we are unable to find information on the grammatical structure of many of the world’s smaller languages. Though these unclassified languages account for less than one percent of the world population, they make up a substantial fraction of the population in a small number of countries (e.g. Chad and Papua New Guinea). Even in countries where we successfully classify the gender structure of almost everyone, our independent variable of interest is an interval rather than a point in 85 of 193 countries — because the proportion of native speakers whose languages we classify is less than one.

This is a case described by Horowitz and Manski (1998) as “censoring of regressors,” discussed further by Aucejo, Bugni, and Hotz (2017). Our analysis so far assumes that this missingness is ignorable. Without this assumption, however, we can still estimate worst-case bounds for the maximum and minimum possible values of the parameter of interest; following Imbens and Manski (2004), we can construct a confidence interval around these bounds.

We use numerical optimization to search the space of possible independent variable values to establish worst-case upper and lower bounds, $\hat{\beta}^u$ and $\hat{\beta}^l$, that would result from estimation of Equation 10. We then use the associated standard errors on these extrema to compute a confidence interval, employing a formula analogous to that of Equations 6 and 7 in Imbens and Manski (2004). A confidence interval with coverage probability $\alpha$ is equal to:

$$CI_\alpha = [\hat{\beta}^l - \bar{C} \cdot SE(\hat{\beta}^l), \hat{\beta}^u + \bar{C} \cdot SE(\hat{\beta}^u)]$$

We use MATLAB’s `fmincon` interior point algorithm, and confirm results using a simple hill-climbing algorithm in Stata.

26We use MATLAB’s `fmincon` interior point algorithm, and confirm results using a simple hill-climbing algorithm in Stata.
where $\bar{C}$ satisfies

$$CDF \left( \bar{C} + \frac{\hat{\Delta}}{\max(SE(\hat{\beta}^l), SE(\hat{\beta}^u))} \right) - CDF(-\bar{C}) = \alpha$$

for the CDF of Student’s t-distribution with the appropriate number of degrees of freedom.\footnote{Imbens and Manski do this using the normal distribution, but using the Student t-distribution yields a wider, more conservative confidence interval.}

Intuitively, the Manski and Imbens approach formalizes a method for shortening each end of the confidence interval relative to the union of the OLS confidence intervals around the worst-case point estimates, since the union would include the true parameter value with probability above 0.95 in either worst-case scenario.

In Table 6, we compare naïve OLS confidence intervals with the more conservative Imbens-Manski confidence intervals which adjust for censoring of the regressor of interest. As expected, confidence intervals widen slightly, but patterns of significance are unchanged: those confidence intervals that did not include zero in the naïve specification do not include zero after adjusting for censoring. This result is largely as expected since missing data problems are relatively minor in most countries. However, if one attempted the same bounding exercise without our data set, using only the data available in the World Atlas of Language Structures, the Imbens-Manski confidence intervals would always include zero. Thus, our data set allows for more robust inference than had previously been possible.

### 5.5.2 Non-Independence within Language Families

A more serious inference concern arises from the fact that languages are not independent. Different tongues evolve over time from a common ancestor. Grammatical structures vary both across and within language families. Roberts, Winters, and Chen (2015) consider a range of approaches to correcting for the non-independence of modern languages. Many approaches have the drawback that they are statistically less powerful than they could otherwise be because they ignore variation in grammatical structure either within or between language families.
We propose a permutation test approach based on the observed structure of the language tree, as documented by the *Ethnologue*. Specifically, we cluster together languages up to the highest tree level at which we observe no variation in our treatment of interest, grammatical gender. That is, we form the largest possible clusters that are homogeneous in terms of grammatical gender. Thus, for entire top-level language families that show no variation in gender structure (e.g., the Austronesian language family), we cluster at the language family level. In intermediate cases, we designate clusters at the highest level of the tree where we do not observe variation in grammatical gender (e.g., all Western Nilotic languages cluster together; they are only a branch within the Eastern Sudanic part of the Nilo-Saharan family, which itself contains a number of other such clusters by our definition). In cases where two languages that differ in their gender structure otherwise share the same classification path through the entire language tree, we cluster at the language level.

Figure 5 illustrates this approach for a hypothetical language family. All of the languages in the Group A branch in the figure are gender languages, so they are assigned to a single cluster. Similarly, all of the languages on the Group C branch are non-gender, so they also represent a single cluster. Within Group B, the B1 languages show language-level variation: Languages B1.1 and B1.2 share the same path for the entire language tree, but they differ in gender structure. Thus, within the B1 branch of this hypothetical tree, individual languages are assigned to unique clusters. Finally, the B2 languages are all gender languages, so they are assigned to a single cluster that is distinct from the B1 clusters. Thus, the hypothetical language tree presented in the figure is partitioned into six clusters, each representing a sub-tree within the language tree that shows no gender variation.

This approach defines a set of 203 clusters, 69 of which have grammatical gender. Having assigned all the languages to clusters in this manner, we conduct a permutation test by randomly generating alternative (hypothetical) allocations of gender structure that would be possible while holding fixed the structure of the treatment variation across the language tree and the number of clusters “treated” with grammatical gender (69 of 203). We use each such hypothetical assignment of treatments to create an associated country-level measure of
grammatical gender (which would be observed if treatments were assigned according to our hypothetical allocation rule, given the structure of the language tree and the distribution of languages across countries). We repeat this process 10,000 times, allowing us to estimate the likelihood that the observed associations between grammatical gender and outcomes are spurious, given the structure of the language tree, the correlation in treatment within language families, and the distribution of languages across countries.

In Table 7, we compare naïve OLS p-values to those that result from our permutation test. It is clear that appropriate clustering matters: permutation test p-values are substantially higher than the naïve OLS p-values. Nevertheless, the negative association between grammatical gender and women’s labor force participation is still statistically significant after adjusting for the non-independence of languages. Figure A1 in the Online Appendix illustrates the full distribution of coefficient estimates under the null, highlighting the small fraction that exceed the magnitude of the true estimated coefficients. The relationships between grammatical gender and (i) the gender gap in primary school completion and (ii) gender attitudes also remain marginally significant. Thus, our results do not appear to be driven by the correlation in grammatical structure observed within language families.

6 Within-Country Analysis

6.1 Empirical Strategy

Next, we explore the relationship between gender languages and women’s labor force participation at the individual level in two contexts where both gender and non-gender languages are indigenous: sub-Saharan Africa and India. There are seven African countries where between 10 and 90 percent of the population speaks a gender native language: Chad, Kenya, Mauritania, Niger, Nigeria, South Sudan, and Uganda. In these countries, both gender and non-gender languages are indigenous — in contrast to, for example, several countries in South America where non-gender indigenous languages and a gender colonial language are

\[28\] To provide an example of the variation within the Nilo-Saharan language family: Maasai is a gender language; Luo, however, is not.
both widely spoken. The same is true in India, where 62 percent of the population speaks a gender language as their mother tongue (Lewis, Simons, and Fennig, eds., 2016). Both the Dravidian language family and the Indo-Aryan branch of the Indo-European family include both gender and non-gender languages (Masica 1991, Krishnamurti 2001). Hence, both India and sub-Saharan Africa allow us to examine the relationship between grammatical gender and women’s outcomes while holding much of the cultural and institutional context constant.

We use two data sources in our within-country analysis: the Afrobarometer surveys (Afrobarometer Data 2016) and the India Human Development Survey (Desai, Dubey, and Vanneman 2015). Of the seven African countries listed above, we focus on the four that have been included in at least one round of the Afrobarometer survey: Kenya, Niger, Nigeria, and Uganda. Four rounds of data are available for Kenya, Nigeria, and Uganda, while only one round of data is available for Niger. Our sample includes 26,546 Afrobarometer respondents who speak 175 different languages. Our IHDS sample includes 75,966 household heads and their spouses living in 33 Indian states. IHDS respondents in our sample speak 61 distinct Indian languages.

Our individual-level analysis parallels our cross-country analysis. We consider two main outcomes: labor force participation (an indicator equal to one if a respondent either does some type of income-generating activity or is actively looking for a job) and education (indicators for having completed primary and secondary school). We report two regression specifications. First, we estimate the association between grammatical gender and each outcome of interest in a sample of (only) women, estimating the OLS regression equation:

$$Y_i = \alpha + \beta Gender_i + \gamma Z_i + \varepsilon_i$$ (13)

where $Y_i$ is the outcome of interest for woman $i$, $Gender_i$ is an indicator for having a gender

---

29 The first round of the Afrobarometer surveys did not include sufficiently detailed data on native languages for inclusion in our analysis. Our analysis includes data from Afrobarometer Rounds 2 through 5 for Kenya, Nigeria, and Uganda. Niger was only added to the Afrobarometer in Round 5; that round is included in our analysis.
language as one’s mother tongue, $Z_i$ is a vector of controls (age, age$^2$, and a set of religion dummies), and $\varepsilon_i$ is a mean-zero error term. In our analysis of the Afrobarometer data, we also include a country-by-survey-round fixed effects. As in our cross-country analysis, we wish to avoid confounding the impact of grammatical gender on women’s education and labor force participation with other cultural factors that might impact both outcomes for both men and women. To do this, we also report pooled OLS regressions that include data on both men and women. These take the form:

$$Y_i = \alpha + \beta Gender_i + \zeta Female_i + \mu Gender \times Female_i + \gamma Z_i + \varepsilon_i$$

(14)

where $Gender \times Female_i$ is an interaction between a female dummy and the indicator for being a native speaker of a gender language. In these specifications, we also include interactions between the $Female_i$ dummy and our age and religion controls. Throughout our analysis, we cluster standard errors by language.

6.2 Results

We summarize our regression results in Figure 6 (regression results are presented in Online Appendix Tables A5 through A12). Panel A presents results on women’s labor force participation. In the Afrobarometer data, we see a negative and statistically significant relationship between grammatical gender and both levels of and gender differences in labor force participation. Coefficient estimates are broadly similar in the Indian data, particularly the estimates of gender differences in labor force participation. However, the relationship is not statistically significant after clustering at the language level. Turning to primary school completion (Panel B of Figure 6), we see that grammatical gender is negatively and significantly related to both rates of primary school completion and the gender difference in primary school completion in both Sub-Saharan Africa and India. Coefficient estimates suggest that having a gender mother tongue is associated with more than a 10 percentage point decline in the likelihood that a woman completed primary school. We see a more
muted association between grammatical gender and secondary school completion (Panel C of Figure 6), though results still suggest a negative and statistically significant relationship in both the African and the Indian data — particularly after controlling for completion rates among men within the same ethnomelinguistic group. Thus, in both Africa and India, we see that the cross-country pattern is largely replicated within country, even when restricting attention to indigenous languages that differ in terms of their grammatical gender structure.

7 Causality

The analysis presented thus far documents a strongly negative cross-country relationship between grammatical gender and women’s labor force participation, and shows that it is robust to a permutation test that addresses the potential non-independence of languages. We also document a positive cross-country relationship between grammatical gender and traditional gender attitudes, and a marginally statistically significant cross-country relationship between grammatical gender and the gender gap in primary school completion. We then find that the negative associations between grammatical gender and women’s educational attainment and labor force participation are replicated within four African countries and within India. The caveat, of course, is that all of these are correlations, and not necessarily causal relationships.

In most cases, whether a language has retained grammatical gender is driven by idiosyncracies of history far-removed from outcomes of interest in this paper. For example, scholars believe that English lost grammatical gender because its complex declensional agreement system eroded over time, in part because of the influx of Scandinavian immigrants (who learned English as a second language in adulthood) into the linguistic community (McWhorter 2005, Kastovsky 1999). So, English did not lose grammatical gender because of changes in gender norms in pre-Norman England. Nevertheless, gender languages are not randomly assigned. The observed correlations may be driven by some unobserved causal
factor that is correlated with both language and gender norms.

To assess whether the observed correlation is likely to represent a causal link between language and our outcomes of interest, we follow the approach suggested by Altonji, Elder, and Taber (2005) and further refined by Oster (2017)\(^{30}\). Under the assumption that the relationship between the outcome variables, treatment, and the observed controls is similar to the relationship between the outcomes, treatment, and unobserved controls, this approach relates changes in coefficient magnitudes as controls are added to changes in the observed \(R^2\). Intuitively, omitted variable bias is assumed to be proportional to changes in regression coefficients as controls are added; however, these changes must be scaled by changes in the \(R^2\) — adding controls that do not explain the outcome variable does little to address concerns about omitted variable bias.

Following the procedures outlined by Oster (2017), we estimate two measures of coefficient stability. These additional statistics are calculated using the results from two OLS regressions: (i) a bivariate regression of an outcome of interest on grammatical gender, which generates a coefficient of interest, \(\hat{\beta}\), and an associated \(\hat{R}^2\); and (ii) a multivariate regression of the same outcome on grammatical gender plus a set of controls, which generates a second OLS coefficient, \(\tilde{\beta}\), and an associated \(\tilde{R}^2\).

In this framework, \(\delta^*\) is the proportional selection coefficient. Given the empirical relationship between the outcome, the treatment, and the observed controls, \(\delta^*\) indicates how much more correlated with treatment the unobservables would need to be in order to explain the entire association between treatment and the outcome of interest. If \(\delta^* > 1\), then an observed empirical relationship is relatively robust in that unobservables would need to be more correlated with treatment than observables to explain the association. A second parameter of interest is \(\beta^*\). It indicates the likely causal impact of grammatical gender on an outcome of interest under the assumption that \(\delta^* = 1\) (i.e. assuming that the covariance structure is the same for observables and unobservables).

\(^{30}\)An alternative approach would be to try and identify a suitable instrument for grammatical gender. However, recent work suggests that conventional approaches may overstate the precision of 2SLS estimates, leading to invalid inference (Young 2018). Thus, OLS with caution may be an equally reasonable approach.
Coefficient stability results are presented in Table 8. Cross-country results are presented in Panel A. Results indicate that our estimates of the impact of grammatical gender on women’s labor force participation are unlikely to be driven by selection alone. Unobservables would need to be 1.44 times more correlated with treatment (than observables) to explain the observed link between grammatical gender and the level of women’s labor force participation; unobservables would need to be 3.24 times more correlated with treatment to explain the gender gap in labor force participation. Thus, the analysis suggests that gender languages reduce women’s labor force participation in both absolute and relative terms.

Our individual-level analysis is presented in Panel B (Afrobaromater data) and Panel C (IHDS data for India) of Table 8. In all cases, the Oster (2017) approach suggests that the empirical relationship between grammatical gender and outcomes of interest is unlikely to be driven by selection on unobservables.

Thus, the coefficient stability approach supports the hypothesis that grammatical gender has a causal impact on women’s labor force participation and, in India and parts of Sub-Saharan Africa, women’s educational attainment. Nevertheless, this approach — like instrumental variables — relies on fundamentally untestable assumptions. Though modern gender attitudes could not plausibly have impacted the grammatical structure of language, we cannot fully rule out the possibility that cultural factors shaped both grammatical structure and gender norms. As in all studies of history and culture, it is not possible to run experiments and relevant sample sizes are fairly small; some measure of caution about causal claims is therefore warranted.

8 Conclusion

Using a new data set on the grammatical gender structure of more than 4,000 languages, we document a robust negative association between gender languages and women’s labor force participation. At the country level, an increase in the proportion of the population whose native language is a gender language is associated with lower female labor force
participation and — perhaps more importantly — larger gender differences in labor force participation. Using data from the World Values Survey, we show that grammatical gender also predicts support for traditional gender roles. The prevalence of gender languages is also related to gender gaps in primary school completion, though the association is only marginally statistically significant.

Focusing on five countries where both gender and non-gender languages are indigenous and widely spoken (India, Kenya, Niger, Nigeria, and Uganda), we show that a similar pattern holds within countries. Speaking a gender native language is associated with lower labor force participation and educational attainment among women, both in absolute terms and relative to men from the same ethnolinguistic group. Both our cross-country and our individual-level regressions are robust to the inclusion of controls that could not plausibly have been impacted by treatment; if one is willing to assume that the relationship between unobserved omitted factors, treatment, and the outcomes of interest is similar to the observed relationship between controls, treatment, and the outcomes of interest, our estimates suggest that grammatical gender has a large negative impact on women’s labor force participation.

Our results are consistent with research in psychology, linguistics, and anthropology suggesting that languages shape patterns of thought in subtle and subconscious ways. Languages are a critical part of our cultural heritage, and it would be inappropriate to suggest that some languages are detrimental to development or women’s rights. However, languages do evolve over time; and the direction of their evolution is shaped by both individual choices (for example, whether to use gendered pronouns like “he” or “she” or gender-neutral alternatives such as “they”) and conscious decisions by government agencies (e.g. the Académie Française) and other thought leaders (e.g. major newspapers and magazines). Our results suggest that individuals should reflect upon the social consequences of their linguistic choices, as the nature of the language we speak shapes the ways we think, and the ways our children will think in the future.
References


Figure 2: The Distribution of Gender Languages

The figure shows the percentage of the native speakers in each country whose native language is a gender language (i.e. the fraction of Ethnologue native speakers whose native language uses a system of grammatical gender). The figure assumes that missing data (on 0.8 percent of all native speakers worldwide) is ignorable.

Figure 3: Cross-Country Variation in Female Labor Force Participation

The figure plots the level of female labor force participation (top panel) and the gender difference in labor force participation (bottom panel) by country. Darker bars indicate countries with a higher proportion of native speakers of gender languages.
The figure summarizes the results from a series of regressions of (country-level averages of) responses to World Values Survey (WVS) questions on the proportion of a country’s population whose native language is a gender language. We present the results for all eight WVS questions related to gender attitudes. Responses to all eight questions are coded so that the answer most consistent with traditional gender norms (involving separate roles for men and women) is equal to 1 and the response most consistent with gender equality is equal to 0. Each regression is estimated via OLS and includes continent fixed effects. The outcome in the first row is the average response to the question “When a mother works for pay, the children suffer” (agreement is coded as a 1, disagreement as a 0). The outcome variable in the second row is the average response to the statement “When jobs are scarce, men should have more right to a job than women.” In the third row, the outcome variable is based on the statement “On the whole, men make better political leaders than women.” In the fourth row, the outcome variable is based on the statement “On the whole, men make better business executives than women do.” In the fifth row, the outcome variable is based on the statement “Being a housewife is just as fulfilling as working for pay;” agreement was coded as 0 and disagreement was coded as 1. In the sixth row, the outcome variable is based on the statement “If a woman earns more money than her husband, it’s almost certain to cause problems.” In the seventh row, the outcome variable is based on the statement “A university education is more important for a boy than for a girl.” In the last row, the outcome variable is based on the statement “Having a job is the best way for a woman to be an independent person;” in this case, disagreement was coded as 1 and agreement was coded as 0.
Figure 5: Assignment to Clusters for the Permutation Test

Figure illustrates a hypothetical language family. Gender languages and branches of the tree that include only gender languages are boxed and printed in red. Languages are assigned to clusters at the highest level of the language tree that shows no variation in grammatical gender.
Figure 6: Within-Country Variation in Grammatical Gender

Panel A: Labor Force Participation

Panel B: Primary School Completion

Panel C: Secondary School Completion
Table 1: Cross-Country OLS Regressions of Labor Force Participation

<table>
<thead>
<tr>
<th></th>
<th>LFP_f</th>
<th>LFP_f - LFP_m</th>
</tr>
</thead>
<tbody>
<tr>
<td>Dependent variable:</td>
<td>OLS (1)</td>
<td>OLS (2)</td>
</tr>
<tr>
<td>Specification:</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Proportion speaking gender language</td>
<td>-13.83 (2.80)</td>
<td>-17.67 (3.52)</td>
</tr>
<tr>
<td>Continent Fixed Effects</td>
<td>No [p &lt; 0.001]</td>
<td>Yes [p &lt; 0.001]</td>
</tr>
<tr>
<td>Country-Level Geography Controls</td>
<td>No</td>
<td>No [p &lt; 0.001]</td>
</tr>
<tr>
<td>Observations</td>
<td>178</td>
<td>178</td>
</tr>
<tr>
<td>( R^2 )</td>
<td>0.15</td>
<td>0.25</td>
</tr>
</tbody>
</table>

Robust standard errors are clustered by the most widely spoken language in all specifications; they are reported in parentheses. P-values are reported in square brackets. LFP\_f is the percentage of women in the labor force, measured in 2011. LFP\_f - LFP\_m is the gender difference in labor force participation — i.e. the difference between female and male labor force participation, again measured in 2011. Geography controls are the percentage of land area in the tropics or subtropics, average yearly precipitation, average temperature, an indicator for being landlocked, and the Alesina et al. (2013) measure of suitability for the plough.
### Table 2: Cross-Country OLS Regressions of Primary School Completion

<table>
<thead>
<tr>
<th>Specification</th>
<th>PRI&lt;sub&gt;f&lt;/sub&gt;</th>
<th>OLS</th>
<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
<th>PRI&lt;sub&gt;f&lt;/sub&gt; - PRI&lt;sub&gt;m&lt;/sub&gt;</th>
<th>OLS</th>
<th>(4)</th>
<th>(5)</th>
<th>(6)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Proportion speaking gender language</td>
<td>14.79</td>
<td>-4.72</td>
<td>-6.71</td>
<td>1.21</td>
<td>-3.87</td>
<td>-3.72</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(5.83)</td>
<td>(4.44)</td>
<td>(4.40)</td>
<td>(2.14)</td>
<td>(2.04)</td>
<td>(2.16)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Continent Fixed Effects</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Country-Level Geography Controls</td>
<td>No</td>
<td>No</td>
<td>Yes</td>
<td>No</td>
<td>No</td>
<td>Yes</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Observations</td>
<td>142</td>
<td>142</td>
<td>142</td>
<td>142</td>
<td>142</td>
<td>142</td>
<td></td>
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</tr>
<tr>
<td>R&lt;sup&gt;2&lt;/sup&gt;</td>
<td>0.06</td>
<td>0.53</td>
<td>0.61</td>
<td>0.00</td>
<td>0.18</td>
<td>0.20</td>
<td></td>
<td></td>
<td></td>
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</tr>
</tbody>
</table>

Robust standard errors are clustered by the most widely spoken language in all specifications; they are reported in parentheses. P-values are reported in square brackets. PRI<sub>f</sub> is the rate of primary school completion among adult women. PRI<sub>f</sub> - PRI<sub>m</sub> is the gender difference in primary school completion. Geography controls are the percentage of land area in the tropics or subtropics, average yearly precipitation, average temperature, an indicator for being landlocked, and the Alesina et al. (2013) measure of suitability for the plough.

### Table 3: Cross-Country OLS Regressions of Secondary School Completion

<table>
<thead>
<tr>
<th>Specification</th>
<th>SEC&lt;sub&gt;f&lt;/sub&gt;</th>
<th>OLS</th>
<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
<th>SEC&lt;sub&gt;f&lt;/sub&gt; - SEC&lt;sub&gt;m&lt;/sub&gt;</th>
<th>OLS</th>
<th>(4)</th>
<th>(5)</th>
<th>(6)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Proportion speaking gender language</td>
<td>14.52</td>
<td>-1.63</td>
<td>0.43</td>
<td>0.48</td>
<td>0.72</td>
<td>-0.86</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(5.77)</td>
<td>(4.22)</td>
<td>(3.70)</td>
<td>(1.93)</td>
<td>(2.31)</td>
<td>(2.35)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Continent Fixed Effects</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Country-Level Geography Controls</td>
<td>No</td>
<td>No</td>
<td>Yes</td>
<td>No</td>
<td>No</td>
<td>Yes</td>
<td></td>
<td></td>
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<tr>
<td>Observations</td>
<td>142</td>
<td>142</td>
<td>142</td>
<td>142</td>
<td>142</td>
<td>142</td>
<td></td>
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</tr>
<tr>
<td>R&lt;sup&gt;2&lt;/sup&gt;</td>
<td>0.06</td>
<td>0.47</td>
<td>0.67</td>
<td>0.00</td>
<td>0.07</td>
<td>0.10</td>
<td></td>
<td></td>
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<td></td>
</tr>
</tbody>
</table>

Robust standard errors are clustered by the most widely spoken language in all specifications; they are reported in parentheses. P-values are reported in square brackets. SEC<sub>f</sub> is the rate of secondary school completion among adult women. SEC<sub>f</sub> - SEC<sub>m</sub> is the gender difference in secondary school completion. Geography controls are the percentage of land area in the tropics or subtropics, average yearly precipitation, average temperature, an indicator for being landlocked, and the Alesina et al. (2013) measure of suitability for the plough.
### Table 4: Cross-Country OLS Regressions of Gender Attitudes

<table>
<thead>
<tr>
<th>Specification:</th>
<th>Gender Attitudes Index</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>OLS (1)</td>
</tr>
<tr>
<td></td>
<td>OLS (2)</td>
</tr>
<tr>
<td></td>
<td>OLS (3)</td>
</tr>
<tr>
<td>Proportion speaking gender language</td>
<td>-0.03 (0.05) [0.576]</td>
</tr>
<tr>
<td></td>
<td>-0.11 (0.03) [p &lt; 0.001]</td>
</tr>
<tr>
<td></td>
<td>-0.12 (0.04) [0.002]</td>
</tr>
<tr>
<td>Continent Fixed Effects</td>
<td>No</td>
</tr>
<tr>
<td>Country-Level Geography Controls</td>
<td>No</td>
</tr>
<tr>
<td>Observations</td>
<td>56</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.01</td>
</tr>
</tbody>
</table>

Robust standard errors clustered by most widely spoken language in all specifications. The Gender Attitudes Index is constructed by taking the first principal component of the 8 World Values Survey questions relating to gender norms (described in Figure 4) at the individual level, and then calculating the average of this index within a country. Numbers closer to 1 indicate more support for gender equality. Geography controls are the percentage of land area in the tropics or subtropics, average yearly precipitation, average temperature, an indicator for being landlocked, and the Alesina et al. (2013) measure of suitability for the plough.

### Table 5: OLS Regressions of Gender Attitudes Index — Women vs. Men

<table>
<thead>
<tr>
<th>Specification:</th>
<th>Attitudes among Women</th>
<th></th>
<th>Attitudes among Men</th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>OLS (1)</td>
<td>OLS (2)</td>
<td>OLS (3)</td>
<td>OLS (4)</td>
</tr>
<tr>
<td>Proportion speaking gender language</td>
<td>-0.02 (0.05) [0.714]</td>
<td>-0.09 (0.03) [0.002]</td>
<td>-0.10 (0.04) [0.012]</td>
<td>-0.04 (0.06) [0.508]</td>
</tr>
<tr>
<td>Continent Fixed Effects</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td>No</td>
</tr>
<tr>
<td>Country-Level Geography Controls</td>
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<td>Yes</td>
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<tr>
<td>Observations</td>
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<tr>
<td>$R^2$</td>
<td>0.00</td>
<td>0.69</td>
<td>0.73</td>
<td>0.02</td>
</tr>
</tbody>
</table>

Robust standard errors are clustered by the most widely spoken language in all specifications; they are reported in parentheses. P-values are reported in square brackets. The dependent variable is constructed by taking the first principal component of the 8 World Values Survey questions relating to gender norms (described in Figure 4) at the individual level, and then calculating the average of this index by gender (i.e. separately among men and women) within a country. Geography controls are the percentage of land area in the tropics or subtropics, average yearly precipitation, average temperature, an indicator for being landlocked, and the Alesina et al. (2013) measure of suitability for the plough.
Table 6: Robust Inference: Manski-Imbens Worst-Case 95-Percent Confidence Intervals

<table>
<thead>
<tr>
<th>Outcome</th>
<th>Naïve OLS CI</th>
<th>Imbens-Manski CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>Female labor force participation</td>
<td>$[-18.533, -5.305]$</td>
<td>$[-18.467, -5.013]$</td>
</tr>
<tr>
<td>Female primary school completion</td>
<td>$[-15.431, 2.010]$</td>
<td>$[-16.221, 1.673]$</td>
</tr>
<tr>
<td>Gender difference in primary school completion</td>
<td>$[-8.003, 0.559]$</td>
<td>$[-8.446, 0.432]$</td>
</tr>
<tr>
<td>Female secondary school completion</td>
<td>$[-6.901, 7.769]$</td>
<td>$[-8.261, 7.327]$</td>
</tr>
<tr>
<td>Gender difference in secondary school completion</td>
<td>$[-5.510, 3.799]$</td>
<td>$[-5.401, 3.746]$</td>
</tr>
<tr>
<td>Gender attitudes index</td>
<td>$[-0.193, -0.045]$</td>
<td>$[-0.194, -0.047]$</td>
</tr>
<tr>
<td>Gender attitudes index among women</td>
<td>$[-0.173, -0.022]$</td>
<td>$[-0.173, -0.023]$</td>
</tr>
<tr>
<td>Gender attitudes index among men</td>
<td>$[-0.214, -0.063]$</td>
<td>$[-0.215, -0.064]$</td>
</tr>
</tbody>
</table>

Confidence intervals estimated following procedures outlined in Section 5.5.1. For each outcome, the naïve confidence interval comes from the associated regression in a previous table. The Imbens-Manski worst-case confidence interval is calculated by finding the minimum and maximum possible point estimates of the relevant coefficient based on the interval nature of the dataset (without complete data on the grammatical structure of all languages, the right-hand-side variable—the fraction of a country’s population speaking a gender language—is only observed up to an interval in some cases), then by tightening the confidence interval for correct coverage following Imbens and Manski (2004).

Table 7: Robust inference: Language structure

<table>
<thead>
<tr>
<th>Outcome</th>
<th>Naïve OLS P-values</th>
<th>Permutation-based P-values</th>
</tr>
</thead>
<tbody>
<tr>
<td>Female labor force participation</td>
<td>0.00050</td>
<td>0.01520</td>
</tr>
<tr>
<td>Gender difference in labor force participation</td>
<td>0.00001</td>
<td>0.00810</td>
</tr>
<tr>
<td>Female primary school completion</td>
<td>0.13012</td>
<td>0.16920</td>
</tr>
<tr>
<td>Gender difference in primary school completion</td>
<td>0.08773</td>
<td>0.08820</td>
</tr>
<tr>
<td>Female secondary school completion</td>
<td>0.90692</td>
<td>0.92410</td>
</tr>
<tr>
<td>Gender difference in secondary school completion</td>
<td>0.71638</td>
<td>0.73140</td>
</tr>
<tr>
<td>Gender attitudes index</td>
<td>0.00225</td>
<td>0.05030</td>
</tr>
<tr>
<td>Gender attitudes index among women</td>
<td>0.01223</td>
<td>0.09620</td>
</tr>
<tr>
<td>Gender attitudes index among men</td>
<td>0.00063</td>
<td>0.03040</td>
</tr>
</tbody>
</table>

P-values estimated using 10,000 permutations, following procedures outlined in Section 5.5.2. For each outcome, the naïve p-value comes from the associated regression in a previous table. The permutation-based p-value is the fraction of permutations in which the magnitude of the estimated coefficient (from a hypothetical permutation of the gender indicator that respects the cluster structure of the language tree) exceeds the magnitude of the estimated coefficient in the true (non-permuted) data set. Distributions underlying first two rows are shown in Figure ??.
Table 8: Coefficient Stability

<table>
<thead>
<tr>
<th>OLS Coefficients</th>
<th>( \hat{\beta} )</th>
<th>( \tilde{\beta} )</th>
<th>( \beta^*(R_{max}, 1) )</th>
<th>( \delta^* )</th>
</tr>
</thead>
</table>

**Panel A. Cross-Country Regressions**

<table>
<thead>
<tr>
<th></th>
<th>( \hat{\beta} )</th>
<th>( \tilde{\beta} )</th>
<th>( \beta^*(R_{max}, 1) )</th>
<th>( \delta^* )</th>
</tr>
</thead>
<tbody>
<tr>
<td>Female labor force participation</td>
<td>-13.83</td>
<td>-11.92</td>
<td>-8.35</td>
<td>1.44</td>
</tr>
<tr>
<td>Gender difference in labor force participation</td>
<td>-11.61</td>
<td>-14.66</td>
<td>-17.87</td>
<td>3.24</td>
</tr>
<tr>
<td>Female primary school completion</td>
<td>14.79</td>
<td>-6.71</td>
<td>-19.40</td>
<td>( \delta &lt; 0 )</td>
</tr>
<tr>
<td>Gender difference in primary school</td>
<td>1.21</td>
<td>-3.72</td>
<td>-6.27</td>
<td>( \delta &lt; 0 )</td>
</tr>
<tr>
<td>Female secondary school completion</td>
<td>14.52</td>
<td>0.43</td>
<td>-9.69</td>
<td>0.05</td>
</tr>
<tr>
<td>Gender difference in secondary school</td>
<td>0.48</td>
<td>-0.86</td>
<td>-1.77</td>
<td>( \delta &lt; 0 )</td>
</tr>
<tr>
<td>Gender attitude index</td>
<td>-0.03</td>
<td>-0.12</td>
<td>-0.20</td>
<td>( \delta &lt; 0 )</td>
</tr>
<tr>
<td>Gender attitudes among women</td>
<td>-0.02</td>
<td>-0.10</td>
<td>-0.18</td>
<td>( \delta &lt; 0 )</td>
</tr>
<tr>
<td>Gender attitudes among men</td>
<td>-0.04</td>
<td>-0.14</td>
<td>-0.23</td>
<td>( \delta &lt; 0 )</td>
</tr>
</tbody>
</table>

**Panel B. Individual-Level Regressions — Afrobarometer Data**

<table>
<thead>
<tr>
<th></th>
<th>( \hat{\beta} )</th>
<th>( \tilde{\beta} )</th>
<th>( \beta^*(R_{max}, 1) )</th>
<th>( \delta^* )</th>
</tr>
</thead>
<tbody>
<tr>
<td>In labor force (Table A5, women only)</td>
<td>-0.24</td>
<td>-0.18</td>
<td>-0.13</td>
<td>2.11</td>
</tr>
<tr>
<td>Female \times in labor force (Table A6)</td>
<td>-0.17</td>
<td>-0.11</td>
<td>-0.06</td>
<td>1.86</td>
</tr>
<tr>
<td>Completed primary school (Table A7, women only)</td>
<td>-0.31</td>
<td>-0.22</td>
<td>-0.15</td>
<td>2.18</td>
</tr>
<tr>
<td>Female \times completed primary school (Table A8)</td>
<td>-0.12</td>
<td>-0.11</td>
<td>-0.10</td>
<td>4.64</td>
</tr>
<tr>
<td>Completed secondary school (Table A7, women only)</td>
<td>-0.19</td>
<td>-0.16</td>
<td>-0.14</td>
<td>3.47</td>
</tr>
<tr>
<td>Female \times completed secondary school (Table A8)</td>
<td>-0.06</td>
<td>-0.06</td>
<td>-0.06</td>
<td>6.01</td>
</tr>
</tbody>
</table>

**Panel C. Individual-Level Regressions — India IHDS Data**

<table>
<thead>
<tr>
<th></th>
<th>( \hat{\beta} )</th>
<th>( \tilde{\beta} )</th>
<th>( \beta^*(R_{max}, 1) )</th>
<th>( \delta^* )</th>
</tr>
</thead>
<tbody>
<tr>
<td>In labor force (Table A9, women only)</td>
<td>-0.08</td>
<td>-0.07</td>
<td>-0.07</td>
<td>11.70</td>
</tr>
<tr>
<td>Female \times in labor force (Table A10)</td>
<td>-0.10</td>
<td>-0.08</td>
<td>-0.04</td>
<td>1.90</td>
</tr>
<tr>
<td>Completed primary school (Table A11, women only)</td>
<td>-0.14</td>
<td>-0.13</td>
<td>-0.12</td>
<td>12.14</td>
</tr>
<tr>
<td>Female \times completed primary school (Table A12)</td>
<td>-0.13</td>
<td>-0.12</td>
<td>-0.11</td>
<td>13.19</td>
</tr>
<tr>
<td>Completed secondary school (Table A11, women only)</td>
<td>-0.03</td>
<td>-0.02</td>
<td>-0.02</td>
<td>7.20</td>
</tr>
<tr>
<td>Female \times completed secondary school (Table A12)</td>
<td>-0.03</td>
<td>-0.03</td>
<td>-0.03</td>
<td>25.89</td>
</tr>
</tbody>
</table>

Parameters estimated following procedures outlined in Altonji, Elder, and Taber (2005) and Oster (2017). \( \hat{\beta} \) is the coefficient of interest from a bivariate regression. \( \tilde{\beta} \) is the coefficient from a regression that includes the full set of observable controls. \( \beta^*(R_{max}, 1) \) is the implied causal impact of grammatical gender on each outcome assuming a proportional selection coefficient (\( \delta \)) equal to 1 and a maximum \( R^2 \) equal to 1.3 times the \( R^2 \) from the regression with controls (Oster 2017). \( \delta^* \) is the proportional selection coefficient required to explain the observed relationship under the null hypothesis of no causal effect of grammatical gender on outcomes of interest.
A Additional Tables and Figures

Figure A1: Permutation Tests

Panel A: Female Labor Force Participation

Panel B: Gender Difference in Labor Force Participation
Table A1: Cross-Country Regressions of LFP Ratio

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>OLS (1)</th>
<th>OLS (2)</th>
<th>OLS (3)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Specification:</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Proportion speaking gender language</td>
<td>-0.16 (0.03)</td>
<td>-0.25 (0.04)</td>
<td>-0.18 (0.04)</td>
</tr>
<tr>
<td></td>
<td>[p &lt; 0.001]</td>
<td>[p &lt; 0.001]</td>
<td>[p &lt; 0.001]</td>
</tr>
<tr>
<td>Continent Fixed Effects</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Country-Level Geography Controls</td>
<td>No</td>
<td>No</td>
<td>Yes</td>
</tr>
<tr>
<td>Observations</td>
<td>178</td>
<td>178</td>
<td>178</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.13</td>
<td>0.37</td>
<td>0.44</td>
</tr>
</tbody>
</table>

Robust standard errors are clustered by the most widely spoken language in all specifications; they are reported in parentheses. P-values are reported in square brackets. LFP$_{ratio}$ is the ratio of the percentage of women in the labor force, measured in 2011, to the percentage of men in the labor force. Geography controls are the percentage of land area in the tropics or subtropics, average yearly precipitation, average temperature, an indicator for being landlocked, and the Alesina et al. (2013) measure of suitability for the plough.
Table A2: Cross-Country Regressions of LFP — Including “Bad” Controls

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>LFP&lt;sub&gt;f&lt;/sub&gt;</th>
<th>LFP&lt;sub&gt;f&lt;/sub&gt; - LFP&lt;sub&gt;m&lt;/sub&gt;</th>
</tr>
</thead>
<tbody>
<tr>
<td>Specification:</td>
<td>OLS (1)</td>
<td>OLS (2)</td>
</tr>
<tr>
<td>Proportion speaking gender language</td>
<td>-6.66 (2.80) [p &lt; 0.001]</td>
<td>-10.42 (2.84) [p &lt; 0.001]</td>
</tr>
<tr>
<td>Continent Fixed Effects</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Country-Level Geography Controls</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Observations</td>
<td>176</td>
<td>176</td>
</tr>
<tr>
<td>R&lt;sup&gt;2&lt;/sup&gt;</td>
<td>0.57</td>
<td>0.68</td>
</tr>
</tbody>
</table>

Robust standard errors are clustered by the most widely spoken language in all specifications; they are reported in parentheses. P-values are reported in square brackets. LFP<sub>f</sub> is the percentage of women in the labor force, measured in 2011. LFP<sub>f</sub> - LFP<sub>m</sub> is the gender difference in labor force participation — i.e. the difference between female and male labor force participation, again measured in 2011. Geography controls are the percentage of land area in the tropics or subtropics, average yearly precipitation, average temperature, an indicator for being landlocked, and the Alesina et al. (2013) measure of suitability for the plough. Bad controls are log GDP per capita (in 2011), log population (in 2011), and the percentage Catholic, Protestant, other Christian, Muslim, and Hindu (taken from Alesina et al. 2013), and an indicator for former communist countries.
<table>
<thead>
<tr>
<th>Specification:</th>
<th>Proportion speaking gender language</th>
<th>Continent Fixed Effects</th>
<th>Country-Level Geography Controls</th>
<th>Observations</th>
<th>$R^2$</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Dependent variable:</strong></td>
<td><strong>LFP$_f$</strong></td>
<td><strong>LFP$_f$ – LFP$_m$</strong></td>
<td><strong>LFP$_f$</strong></td>
<td><strong>LFP$_f$ – LFP$_m$</strong></td>
<td><strong>LFP$_f$</strong></td>
</tr>
<tr>
<td><strong>Omitted Language:</strong></td>
<td>ARABIC</td>
<td>ENGLISH</td>
<td>SPANISH</td>
<td>ARABIC</td>
<td>ENGLISH</td>
</tr>
<tr>
<td><strong>OLS</strong></td>
<td>(1)</td>
<td>(2)</td>
<td>(3)</td>
<td>(4)</td>
<td>(5)</td>
</tr>
<tr>
<td>Proportion speaking gender language</td>
<td>-6.18</td>
<td>-12.33</td>
<td>-10.10</td>
<td>-9.09</td>
<td>-15.31</td>
</tr>
<tr>
<td>(3.56)</td>
<td>(3.84)</td>
<td>(3.87)</td>
<td>(3.52)</td>
<td>(3.59)</td>
<td>(3.39)</td>
</tr>
<tr>
<td>[0.085]</td>
<td>[0.002]</td>
<td>[0.010]</td>
<td>[0.011]</td>
<td>[p &lt; 0.001]</td>
<td>[0.001]</td>
</tr>
<tr>
<td>Continent Fixed Effects</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Country-Level Geography Controls</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Observations</td>
<td>159</td>
<td>167</td>
<td>160</td>
<td>159</td>
<td>167</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.21</td>
<td>0.34</td>
<td>0.37</td>
<td>0.31</td>
<td>0.49</td>
</tr>
</tbody>
</table>

Robust standard errors are clustered by the most widely spoken language in all specifications; they are reported in parentheses. P-values are reported in square brackets. LFP$_f$ is the percentage of women in the labor force, measured in 2011. LFP$_f$ – LFP$_m$ is the difference between male and female labor force participation in 2011. Geography controls are the percentage of land area in the tropics or subtropics, average yearly precipitation, average temperature, an indicator for being landlocked, and the Alesina et al. (2013) measure of suitability for the plough.
Table A4: Cross-Country Regressions of LFP — Weak vs. Strong Gender Categories

<table>
<thead>
<tr>
<th>Specification:</th>
<th></th>
<th></th>
<th></th>
<th></th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td>Dependent variable:</td>
<td>LFP&lt;sub&gt;f&lt;/sub&gt;</td>
<td>LFP&lt;sub&gt;f&lt;/sub&gt; - LFP&lt;sub&gt;m&lt;/sub&gt;</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>OLS</td>
<td>OLS</td>
<td>OLS</td>
<td>OLS</td>
<td>OLS</td>
<td>OLS</td>
</tr>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
<td>(3)</td>
<td>(4)</td>
<td>(5)</td>
<td>(6)</td>
</tr>
<tr>
<td>Proportion speaking (any) gender language</td>
<td>-6.66</td>
<td>-8.10</td>
<td>-7.19</td>
<td>4.29</td>
<td>-5.59</td>
<td>-5.77</td>
</tr>
<tr>
<td></td>
<td>(2.54)</td>
<td>(3.63)</td>
<td>(3.91)</td>
<td>(1.65)</td>
<td>(4.23)</td>
<td>(4.34)</td>
</tr>
<tr>
<td></td>
<td>[0.010]</td>
<td>[0.027]</td>
<td>[0.068]</td>
<td>[0.010]</td>
<td>[0.189]</td>
<td>[0.185]</td>
</tr>
<tr>
<td></td>
<td>(4.78)</td>
<td>(3.86)</td>
<td>(4.16)</td>
<td>(3.54)</td>
<td>(4.19)</td>
<td>(4.53)</td>
</tr>
<tr>
<td></td>
<td>[0.029]</td>
<td>[0.003]</td>
<td>[0.116]</td>
<td>[p &lt; 0.001]</td>
<td>[p &lt; 0.001]</td>
<td>[0.007]</td>
</tr>
<tr>
<td>Continent Fixed Effects</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Country-Level Geography Controls</td>
<td>No</td>
<td>No</td>
<td>Yes</td>
<td>No</td>
<td>No</td>
<td>Yes</td>
</tr>
<tr>
<td>Observations</td>
<td>178</td>
<td>178</td>
<td>178</td>
<td>178</td>
<td>178</td>
<td>178</td>
</tr>
<tr>
<td>R&lt;sup&gt;2&lt;/sup&gt;</td>
<td>0.19</td>
<td>0.27</td>
<td>0.33</td>
<td>0.30</td>
<td>0.46</td>
<td>0.50</td>
</tr>
</tbody>
</table>

Robust standard errors are clustered by the most widely spoken language in all specifications; they are reported in parentheses. P-values are reported in square brackets. LFP<sub>f</sub> is the percentage of women in the labor force, measured in 2011. LFP<sub>f</sub> - LFP<sub>m</sub> is the gender difference in labor force participation — i.e. the difference between female and male labor force participation, again measured in 2011. Geography controls are the percentage of land area in the tropics or subtropics, average yearly precipitation, average temperature, an indicator for being landlocked, and the Alesina et al. (2013) measure of suitability for the plough.
Table A5: OLS Regressions of African Women’s Labor Force Participation

<table>
<thead>
<tr>
<th>Specification:</th>
<th>OLS (1)</th>
<th>OLS (2)</th>
<th>OLS (3)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Native language is a gender language</td>
<td>-0.24 (0.05)</td>
<td>-0.20 (0.04)</td>
<td>-0.18 (0.04)</td>
</tr>
<tr>
<td>Country-Wave Fixed Effects</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Individual Controls</td>
<td>No</td>
<td>No</td>
<td>Yes</td>
</tr>
<tr>
<td>Observations</td>
<td>13154</td>
<td>13154</td>
<td>13154</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.04</td>
<td>0.07</td>
<td>0.10</td>
</tr>
</tbody>
</table>

Robust standard errors clustered at the language level. The dependent variable is an indicator for being in the labor force (either working for a wage, self-employed, or actively seeking employment). Data is from Afrobarometer Rounds 2 through 5. The analysis includes data from Kenya, Niger, Nigeria, and Uganda; Niger was only added to the Afrobarometer in Round 5, while the other countries appear in all four rounds. Individual controls are age and age-squared and indicators for being identifying as Muslim, Catholic, Protestant, or another religion.
Table A6: OLS Regressions of Gender Differences in Labor Force Participation in Africa

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>In Labor Force</th>
</tr>
</thead>
<tbody>
<tr>
<td>Specification:</td>
<td>OLS</td>
</tr>
<tr>
<td></td>
<td>(1)</td>
</tr>
<tr>
<td></td>
<td>OLS</td>
</tr>
<tr>
<td></td>
<td>(2)</td>
</tr>
<tr>
<td></td>
<td>OLS</td>
</tr>
<tr>
<td></td>
<td>(3)</td>
</tr>
<tr>
<td>Female × gender language</td>
<td>-0.17</td>
</tr>
<tr>
<td></td>
<td>(0.05)</td>
</tr>
<tr>
<td></td>
<td>[0.001]</td>
</tr>
<tr>
<td>Native language is a gender language</td>
<td>-0.08</td>
</tr>
<tr>
<td></td>
<td>(0.02)</td>
</tr>
<tr>
<td></td>
<td>[p &lt; 0.001]</td>
</tr>
<tr>
<td>Female</td>
<td>-0.08</td>
</tr>
<tr>
<td></td>
<td>(0.02)</td>
</tr>
<tr>
<td></td>
<td>[p &lt; 0.001]</td>
</tr>
<tr>
<td>Country-Wave Fixed Effects</td>
<td>No</td>
</tr>
<tr>
<td>Individual Controls</td>
<td>No</td>
</tr>
<tr>
<td>Observations</td>
<td>26328</td>
</tr>
<tr>
<td>(R^2)</td>
<td>0.04</td>
</tr>
</tbody>
</table>

Robust standard errors clustered at the language level. The dependent variable is an indicator for being in the labor force (either working for a wage, self-employed, or actively seeking employment). Data is from Afrobarometer Rounds 2 through 5. The analysis includes data from Kenya, Niger, Nigeria, and Uganda; Niger was only added to the Afrobarometer in Round 5, while the other countries appear in all four rounds. Individual controls are age and age-squared and indicators for being identifying as Muslim, Catholic, Protestant, or another religion, plus interactions between these controls and the female dummy.
Table A7: OLS Regressions of African Women’s Educational Attainment

<table>
<thead>
<tr>
<th>Specification:</th>
<th>Primary School</th>
<th>Secondary School</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>OLS (1)</td>
<td>OLS (2)</td>
</tr>
<tr>
<td>Native language is a gender language</td>
<td>-0.31 (0.04)</td>
<td>-0.30 (0.06)</td>
</tr>
<tr>
<td>Country-Wave Fixed Effects</td>
<td>No</td>
<td>Yes</td>
</tr>
<tr>
<td>Individual Controls</td>
<td>No</td>
<td>No</td>
</tr>
<tr>
<td>Observations</td>
<td>13142</td>
<td>13142</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.06</td>
<td>0.12</td>
</tr>
</tbody>
</table>

Robust standard errors clustered at the language level. The dependent variable is an indicator for being in the labor force (either working for a wage, self-employed, or actively seeking employment). Data is from Afrobarometer Rounds 2 through 5. The analysis includes data from Kenya, Niger, Nigeria, and Uganda; Niger was only added to the Afrobarometer in Round 5, while the other countries appear in all four rounds. Individual controls are age and age-squared and indicators for being identifying as Muslim, Catholic, Protestant, or another religion.
Table A8: OLS Regressions of Gender Differences in Educational Attainment in Africa

| Dependent variable: | PRIMARY SCHOOL | | | | | | SECONDARY SCHOOL |
|---------------------|----------------|----------------|----------------|----------------|----------------|----------------|
|                     | OLS            | OLS            | OLS            | OLS            | OLS            | OLS            |
| Specification:      | (1)            | (2)            | (3)            | (4)            | (5)            | (6)            |
| Female x gender language | -0.12          | -0.11          | -0.11          | -0.06          | -0.06          | -0.06          |
|                     | (0.01)         | (0.01)         | (0.02)         | (0.01)         | (0.01)         | (0.02)         |
| [p < 0.001]         | [p < 0.001]    | [p < 0.001]    |                 | [p < 0.001]    | [p < 0.001]    | [0.009]        |
| Native language is a gender language | -0.19          | -0.17          | -0.10          | -0.13          | -0.17          | -0.10          |
|                     | (0.04)         | (0.05)         | (0.04)         | (0.04)         | (0.05)         | (0.03)         |
| [p < 0.001]         | [p < 0.001]    |                 | [0.008]        | [p < 0.001]    |                 | [0.003]        |
| Female              | -0.19          | -0.17          | -0.10          | -0.13          | -0.17          | -0.10          |
|                     | (0.04)         | (0.05)         | (0.04)         | (0.04)         | (0.05)         | (0.03)         |
| [p < 0.001]         | [p < 0.001]    | [0.008]        | [p < 0.001]    | [p < 0.001]    | [p < 0.001]    | [0.003]        |
| Country-Wave Fixed Effects | No            | Yes            | Yes            | No             | Yes            | Yes            |
| Individual Controls | No             | No             | Yes            | No             | No             | Yes            |
| Observations        | 26294          | 26294          | 26294          | 26294          | 26294          | 26294          |
| $R^2$               | 0.06           | 0.12           | 0.21           | 0.03           | 0.11           | 0.15           |

Robust standard errors clustered at the language level. The dependent variable is an indicator for being in the labor force (either working for a wage, self-employed, or actively seeking employment). Data is from Afrobarometer Rounds 2 through 5. The analysis includes data from Kenya, Niger, Nigeria, and Uganda; Niger was only added to the Afrobarometer in Round 5, while the other countries appear in all four rounds. Individual controls are age and age-squared and indicators for being identifying as Muslim, Catholic, Protestant, or another religion, plus interactions between these controls and the female dummy.
Table A9: OLS Regressions of Indian Women’s Labor Force Participation

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>IN LABOR FORCE</th>
</tr>
</thead>
<tbody>
<tr>
<td>Specification:</td>
<td>OLS</td>
</tr>
<tr>
<td></td>
<td>(1)</td>
</tr>
<tr>
<td>Native language is a gender language</td>
<td>-0.08</td>
</tr>
<tr>
<td></td>
<td>(0.07)</td>
</tr>
<tr>
<td></td>
<td>[0.308]</td>
</tr>
<tr>
<td>Individual Controls</td>
<td>No</td>
</tr>
<tr>
<td>Observations</td>
<td>39895</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.01</td>
</tr>
</tbody>
</table>

Robust standard errors clustered at the language level. The dependent variable is an indicator for being in the labor force (reporting one’s primary activity as agriculture, wage labor, self-employment, or salaried/professional work). Data is from India Human Development Survey-II (Desai, Dubey, and Vanneman 2015). Individual controls are age and age-squared and indicators for being identifying as Muslim, Christian, Sikh, or another religion.
Table A10: OLS Regressions of Gender Differences in Labor Force Participation in India

<table>
<thead>
<tr>
<th>Dependent variable</th>
<th>IN LABOR FORCE</th>
</tr>
</thead>
<tbody>
<tr>
<td>Specification:</td>
<td>OLS (1)</td>
</tr>
<tr>
<td>Female × gender language</td>
<td>-0.10 (0.07)</td>
</tr>
<tr>
<td></td>
<td>[0.17]</td>
</tr>
<tr>
<td>Native language is a gender language</td>
<td>0.02 (0.01)</td>
</tr>
<tr>
<td></td>
<td>[0.13]</td>
</tr>
<tr>
<td>Female</td>
<td>-0.56 (0.05)</td>
</tr>
<tr>
<td></td>
<td>[p &lt; 0.001]</td>
</tr>
<tr>
<td>Individual Controls</td>
<td>No</td>
</tr>
<tr>
<td>Observations</td>
<td>75966</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.40</td>
</tr>
</tbody>
</table>

Robust standard errors clustered at the language level. The dependent variable is an indicator for being in the labor force (reporting one’s primary activity as agriculture, wage labor, self-employment, or salaried/professional work). Data is from India Human Development Survey-II (Desai, Dubey, and Vanneman 2015). Individual controls are age and age-squared and indicators for being identifying as Muslim, Christian, Sikh, or another religion.
Table A11: OLS Regressions of Indian Women’s Educational Attainment

<table>
<thead>
<tr>
<th></th>
<th>PRIMARY SCHOOL</th>
<th></th>
<th>SECONDARY SCHOOL</th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>OLS (1)</td>
<td>OLS (2)</td>
<td>OLS (3)</td>
<td>OLS (4)</td>
</tr>
<tr>
<td>Native language is a gender language</td>
<td>-0.14 (0.06)</td>
<td>-0.13 (0.06)</td>
<td>-0.03 (0.02)</td>
<td>-0.02 (0.02)</td>
</tr>
<tr>
<td></td>
<td>[0.033]</td>
<td>[0.043]</td>
<td>[0.103]</td>
<td>[0.158]</td>
</tr>
<tr>
<td>Individual Controls</td>
<td>No</td>
<td>Yes</td>
<td>No</td>
<td>Yes</td>
</tr>
<tr>
<td>Observations</td>
<td>39895</td>
<td>39895</td>
<td>39895</td>
<td>39895</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.02</td>
<td>0.08</td>
<td>0.00</td>
<td>0.02</td>
</tr>
</tbody>
</table>

Robust standard errors clustered at the language level. The dependent variable is an indicator for being in the labor force (reporting one’s primary activity as agriculture, wage labor, self-employment, or salaried/professional work). Data is from India Human Development Survey-II (Desai, Dubey, and Van- neman 2015). Individual controls are age and age-squared and indicators for being identifying as Muslim, Christian, Sikh, or another religion, plus interactions between these controls and the female dummy.
Table A12: OLS Regressions of Gender Differences in Educational Attainment in India

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>PRIMARY SCHOOL</th>
<th>SECONDARY SCHOOL</th>
</tr>
</thead>
<tbody>
<tr>
<td>Specifcation:</td>
<td>OLS (1)</td>
<td>OLS (2)</td>
</tr>
<tr>
<td></td>
<td></td>
<td>OLS (3)</td>
</tr>
<tr>
<td></td>
<td></td>
<td>OLS (4)</td>
</tr>
<tr>
<td>Female × gender language</td>
<td>-0.13 (0.03) [p &lt; 0.001]</td>
<td>-0.12 (0.03) [p &lt; 0.001]</td>
</tr>
<tr>
<td>Native language is a gender language</td>
<td>-0.01 (0.04) [0.767]</td>
<td>-0.01 (0.04) [0.842]</td>
</tr>
<tr>
<td>Female</td>
<td>-0.11 (0.02) [p &lt; 0.001]</td>
<td>0.27 (0.05) [p &lt; 0.001]</td>
</tr>
<tr>
<td>Individual Controls</td>
<td>No</td>
<td>Yes</td>
</tr>
<tr>
<td>Observations</td>
<td>75966</td>
<td>75966</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.05</td>
<td>0.09</td>
</tr>
</tbody>
</table>

Robust standard errors clustered at the language level. The dependent variable is an indicator for being in the labor force (reporting one’s primary activity as agriculture, wage labor, self-employment, or salaried/professional work). Data is from India Human Development Survey-II (Desai, Dubey, and Vanneman 2015). Individual controls are age and age-squared and indicators for being identifying as Muslim, Christian, Sikh, or another religion, plus interactions between these controls and the female dummy.
B Conceptual Framework: Further Analysis and Proofs

B.1 Educational Attainment

There are three possible equilibria: the M-equilibrium (school is a masculine domain), the N-equilibrium (school is not a gendered domain), and the F-equilibrium (school is a feminine domain). Each equilibrium is associated with a cost structure. For example, because school is a masculine domain in the M-equilibrium, girls face a psychic cost when they choose to go to school (but boys do not). The M-equilibrium only exists if the proportion of girls (as a share of all students) that results from that cost structure is below $\lambda$.

If it exists, the N-equilibrium is identical to the case discussed above (in the absence of grammatical gender). Since school is not seen as a gendered domain, boys and girls attend whenever the net return is greater than zero — i.e. whenever $\gamma_i \geq \gamma^N = \gamma^*$ (where $\gamma^*$ is the solution to $R_g(\gamma^*) = 0$, as discussed in Section 3.1) or $\beta_i \geq \beta^N = \beta^*$. Hence, the proportion of students who are female is equal to

$$P_{\text{girls}}^N = \frac{1 - F_\gamma(\gamma^*)}{2 - F_\beta(\beta^*) - F_\gamma(\gamma^*)}. \tag{1}$$

The N-equilibrium exists if and only if $\lambda \leq P_{\text{girls}}^N \leq 1 - \lambda$. In the N-equilibrium, the proportion of children attending school is $1 - \left[F_\beta(\beta^*) + F_\gamma(\gamma^*)\right]/2$.

School is considered a masculine domain if the proportion of students who are female is below $\lambda$. In this case, girls will only attend school if $R_g(\gamma_i) > \phi$. Define $\gamma^M$ as the solution to $R_g(\gamma_i) - \phi = 0$. Since the return to education is increasing in ability, $\gamma^M > \gamma^N$. Girls with $\gamma_i \geq \gamma^M$ will attend school whether or not school is perceived as a masculine domain, but those with $\gamma_i \in (\gamma^N, \gamma^M)$ will only attend school when it is perceived as a feminine or neutral domain. The fact that school is perceived as masculine does not impact the net return to education for boys, so boys will (still) attend school whenever $\beta_i \geq \beta^M = \beta^*$. Hence in the M-equilibrium, the proportion of students who are female is:

$$P_{\text{girls}}^M = \frac{1 - F_\gamma(\gamma^M)}{2 - F_\beta(\beta^*) - F_\gamma(\gamma^M)}. \tag{2}$$

The M-equilibrium exists if and only if $P_{\text{girls}}^M \leq \lambda$. The proportion of children who attend school in the M-equilibrium is $1 - \left[F_\beta(\beta^*) + F_\gamma(\gamma^M)\right]/2$. Since $F_\gamma(\gamma^M) > F_\gamma(\gamma^*)$, fewer children attend school in the M-equilibrium than in the N-equilibrium. The F-equilibrium — in which school is a feminine domain — is defined symmetrically.

It is apparent that $P_{\text{girls}}^M \leq P_{\text{girls}}^N \leq P_{\text{girls}}^F$. Hence, multiple equilibria are possible whenever $P_{\text{girls}}^M \leq \lambda \leq P_{\text{girls}}^N$ or $P_{\text{girls}}^N \leq 1 - \lambda \leq P_{\text{girls}}^F$. To see that at least one equilibrium always exists, first note that the N-equilibrium always exists if $\lambda \leq P_{\text{girls}}^N \leq 1 - \lambda$. If the N-equilibrium does not exist because $\lambda > P_{\text{girls}}^N$, then $\lambda$ must also be greater than $P_{\text{girls}}^M$ — so the M-equilibrium exists. Similarly, if the N-equilibrium does not exist because $P_{\text{girls}}^N > 1 - \lambda$, then $1 - \lambda$ must also be less than $P_{\text{girls}}^F$ — so the F-equilibrium exists.

B.2 Labor Force Participation and the Division of Household Tasks

There are nine candidate equilibria to consider. Home and work can each be either masculine, neutral (non-gendered), or feminine. Each candidate equilibrium is a pair $HW$ where...
$H \in \{M, N, F\}$ characterizes the “home” environment and $W \in \{M, N, F\}$ characterizes the “work” environment. So, the NN equilibrium would be one in which neither home nor work is perceived as a gendered domain, whereas the FM equilibrium would be one in which the home environment is perceived as feminine and the work environment is perceived as masculine.

Figure B1: Candidate Equilibria When Domains Can Be Gendered

<table>
<thead>
<tr>
<th>Work</th>
<th>Feminine</th>
<th>Neutral</th>
<th>Masculine</th>
</tr>
</thead>
<tbody>
<tr>
<td>Feminine</td>
<td>FF</td>
<td>FN</td>
<td>FM</td>
</tr>
<tr>
<td>Home</td>
<td>Neutral</td>
<td>NF</td>
<td>NN</td>
</tr>
<tr>
<td>Masculine</td>
<td>MF</td>
<td>MN</td>
<td>MM</td>
</tr>
</tbody>
</table>

For each candidate equilibrium $HW$, we define $P_{\text{mom}}^{HW}$, $P_{\text{dad}}^{HW}$, and $P_{\text{nanny}}^{HW}$ as the proportion of households where (respectively) the mother, the father, or a nanny stays home with the children. $P_{\text{mom}}^{HW} + P_{\text{dad}}^{HW} + P_{\text{nanny}}^{HW} = 1$. We then define $Q_{\text{home}}^{HW} = P_{\text{mom}}^{HW} + P_{\text{nanny}}^{HW}$ as the proportion of households where the person (at home) taking care of the children is female. In equilibrium, home is perceived as a masculine domain if $Q_{\text{home}}^{HW} < \lambda$, and home is perceived as a feminine domain if $Q_{\text{home}}^{HW} > 1 - \lambda$. We define

$$Q_{\text{work}}^{HW} = \frac{P_{\text{dad}}^{HW} + P_{\text{nanny}}^{HW}}{1 + P_{\text{nanny}}^{HW}}$$

as the proportion of the out-of-home workforce that is female (note that households that hire a nanny send both a man and a woman into the out-of-home labor force, while other households send either a man or a woman). Work is perceived as a masculine domain whenever $Q_{\text{work}}^{HW} < \lambda$; it is perceived as a feminine domain whenever $Q_{\text{work}}^{HW} > 1 - \lambda$.

We begin by documenting a useful algebraic inequality that we will invoke repeatedly in our subsequent exposition. Trivial as it is, we label this inequality to avoid unnecessary repetition in our proofs and discussion.

**Inequality 1.** For any $a \in (0, 1)$ and $b > 0$, $a < (a + b)/(1 + b)$.

A direct consequence of Inequality 1 is that $P_{\text{dad}}^{HW} < Q_{\text{work}}^{HW}$; the proportion of households
where the father is responsible for childcare is lower than the proportion of women in the out-of-home workforce. This follows from the definition of $Q_{work}^{HW}$ above.

In our first set of results, we show that three of the candidate equilibria — NM, MN, and MF — cannot exist. As the lemmas below demonstrate, this does not depend on the strength of the inclination to partition the world into masculine and feminine domains ($\lambda$) or the magnitude of the cost of entering a domain dominated by the opposite sex ($\phi$).

**Lemma 1.** If home is masculine, then work must be feminine.

Proof. Home is perceived as masculine if and only if $Q_{home}^{HW} < \lambda$. This places bounds on the feasible range of values of $Q_{work}^{HW}$:

$$Q_{home}^{HW} < \lambda \iff P_{mom}^{HW} + P_{nanny}^{HW} < \lambda$$

$$\iff 1 - P_{dad}^{HW} < \lambda$$

$$\iff P_{dad}^{HW} > 1 - \lambda$$

$$\Rightarrow Q_{work}^{HW} > 1 - \lambda$$

with the last step following immediately from Inequality 1 and the definition of $Q_{work}^{HW}$.

**Lemma 2.** If work is masculine, then home must be feminine.

Proof. Work is masculine if and only if $Q_{work}^{HW} < \lambda$. Inequality 1 tells us that $Q_{work}^{HW} > P_{dad}^{HW}$, so $P_{dad}^{HW} < \lambda$ or, equivalently, $1 - P_{dad}^{HW} = P_{mom}^{HW} + P_{nanny}^{HW} = Q_{home}^{HW} > 1 - \lambda$.

Lemma 1 and Lemma 2 prove that the MN, MM, and NM equilibria cannot exist. In our next piece of analysis, we will demonstrate that an equilibrium always exists and the multiple equilibria are possible. Before doing so, we introduce some additional notation.

Partition the $\gamma \times \beta$ space into 13 regions a, b, c, d, e, f, g, h, i, j, k, l, and m, as in Figure B2. For capital letters $Z = A, B, C, \ldots M$, define $Z$ as the integral of $f_{\gamma, \beta}(\gamma, \beta)$ over the region $z$. So,

$$A = \int_{\beta = w_n + \phi}^{\beta = w_n - \phi} \int_{\gamma = 0}^{\gamma = \beta - 2\phi} f_{\beta, \gamma}(\beta, \gamma) + \int_{\beta = w_n + \phi}^{\beta = w_n - \phi} \int_{\gamma = 0}^{\gamma = w_n - \phi} f_{\beta, \gamma}(\beta, \gamma)$$

(3)

is the integral of $f_{\gamma, \beta}(\gamma, \beta)$ over the region $a$, and

$$B = \int_{\beta = w_n + \phi}^{\beta = \beta_{max}} \int_{\gamma = w_n}^{\gamma = w_n - \phi} f_{\beta, \gamma}(\beta, \gamma)$$

(4)

is the integral of $f_{\gamma, \beta}(\gamma, \beta)$ over $b$.

Note that for any $f_{\gamma, \beta}(\gamma, \beta)$,


(5)

Moreover, for every candidate equilibrium $HW$, the proportion of households where a woman does the childcare, $Q_{home}^{HW}$, and the proportion of the workforce that is female, $Q_{work}^{HW}$, can be expressed in terms of $A, B, C, \ldots$ etc. Before presenting our existence proof, we characterize $P_{mom}^{HW}, P_{dad}^{HW}, P_{nanny}^{HW}, Q_{home}^{HW}$, and $Q_{work}^{HW}$ for each of the six possible equilibria.
Figure B2: Labor Force Participation when Domains Are Not Gendered
1. The FF equilibrium: home is feminine and work is feminine

In this equilibrium, the payoffs are: $\beta_i - \phi$ if mom stays home, $\gamma_i - \phi$ if dad stays home, $\beta_i + \gamma_i - \phi - w_n$ if the household hires a nanny. So, the household hires a nanny whenever $\beta_i > w_n$ and $\gamma_i > w_n$. The mom stays at home whenever $\beta_i < \gamma_i$ and $\beta_i > w_n$. The dad stays at home whenever $\beta_i < \gamma_i$ and $\beta_i < w_n$. As shown in Figure B3:

- $P^{FF}_{mom} = A + B + E + F$
- $P^{FF}_{nanny} = C + D + G + H$
- $P^{FF}_{dad} = I + J + K + L + M$
- $Q^{FF}_{home} = A + B + C + D + E + F + G + H$
- $Q^{FF}_{work} = \frac{C + D + G + H + I + J + K + L + M}{1 + C + D + G + H}$

The FF equilibrium exists if and only if $Q^{FF}_{home} > 1 - \lambda$ and $Q^{FF}_{work} > 1 - \lambda$.

---

1^Note that these are identical to the regions we observe in the “no gendered domains” equilibrium — it is just a question of where the implied proportions female end up relative to $1 - \lambda$. 

B5
2. The FN equilibrium: home is feminine and work is neutral

In this equilibrium, the payoffs are: \( \beta_i \) if mom stays home, \( \gamma_i - \phi \) if dad stays home, \( \beta_i + \gamma_i - w_n \) if the household hires a nanny. So, the household hires a nanny whenever \( \beta_i > w_n - \phi \) and \( \gamma_i > w_n \). The mom stays at home whenever \( \beta_i > \gamma_i - \phi \) and \( \gamma_i < w_n \). The dad stays at home whenever \( \beta_i < \gamma_i - \phi \) and \( \beta_i < w_n - \phi \). As shown in Figure B4:

- \( P^{FN}_{mom} = A + B + E + F + I \)
- \( P^{FN}_{nanny} = C + D + G + H + J + K \)
- \( P^{FN}_{dad} = L + M \)
- \( Q^{FN}_{home} = A + B + C + D + E + F + G + H + I + J + K \)
- \( Q^{FN}_{work} = \frac{C + D + G + H + J + K + L + M}{1 + C + D + G + H + J + K} \)

The FN equilibrium exists if and only if \( Q^{FN}_{home} > 1 - \lambda \) and \( \lambda < Q^{FN}_{work} > 1 - \lambda \).
3. The FM equilibrium: home is feminine and work is masculine

In this equilibrium, the payoffs are: $\beta_i$ if mom stays home, $\gamma_i - 2\phi$ if dad stays home, $\beta_i + \gamma_i - \phi - w_n$ if the household hires a nanny. So, the household hires a nanny whenever $\beta_i > w_n - \phi$ and $\gamma_i > w_n + \phi$. The mom stays at home whenever $\beta_i > \gamma_i - 2\phi$ and $\gamma_i < w_n + \phi$. The dad stays at home whenever $\beta_i < \gamma_i - 2\phi$ and $\beta_i < w_n - \phi$. As shown in Figure B5:

- $P_{mom}^{FM} = A + B + C + E + F + G + I + J + L$
- $P_{nanny}^{FM} = D + H + K$
- $P_{dad}^{FM} = M$
- $Q_{home}^{FM} = A + B + C + D + E + F + G + H + I + J + K + L = 1 - M$
- $Q_{work}^{FM} = \frac{D + H + K + M}{1 + D + H + K}$

The FM equilibrium exists if and only if $Q_{home}^{FM} > 1 - \lambda$ and $Q_{work}^{FM} < \lambda$.
4. The NF equilibrium: home is neutral and work is feminine

In this equilibrium, the payoffs are: $\beta_i - \phi$ if mom stays home, $\gamma_i$ if dad stays home, $\beta_i + \gamma_i - \phi - w_n$ if the household hires a nanny. So, the household hires a nanny whenever $\beta_i > w_n + \phi$ and $\gamma_i > w_n$. The mom stays at home whenever $\beta_i > \gamma_i + \phi$ and $\gamma_i < w_n$. The dad stays at home whenever $\beta_i < \gamma_i + \phi$ and $\beta_i < w_n + \phi$. As shown in Figure B6:

- $P_{mom}^{NF} = A + B + E$
- $P_{nanny}^{NF} = C + D$
- $P_{dad}^{NF} = F + G + H + I + J + K + L + M$
- $Q_{home}^{NF} = A + B + C + D + E$
- $Q_{work}^{NF} = \frac{C + D + F + G + H + I + J + K + L + M}{1 + C + D}$

The NF equilibrium exists if and only if $\lambda < Q_{home}^{NF} < 1 - \lambda$ and $Q_{work}^{NF} > 1 - \lambda$. 
5. The NN equilibrium: home is neutral and work is neutral

In this equilibrium, the payoffs are: \( \beta_i - \phi \) if mom stays home, \( \gamma_i - \phi \) if dad stays home, \( \beta_i + \gamma_i - \phi - w_n \) if the household hires a nanny. So, the household hires a nanny whenever \( \beta_i > w_n \) and \( \gamma_i > w_n \). The mom stays at home whenever \( \beta_i > \gamma_i \) and \( \gamma_i < w_n \). The dad stays at home whenever \( \beta_i < \gamma_i \) and \( \beta_i < w_n \). As shown in Figure B7:

- \( P_{\text{mom}}^{NN} = A + B + E + F \)
- \( P_{\text{nanny}}^{NN} = C + D + G + H \)
- \( P_{\text{dad}}^{NN} = I + J + K + L + M \)
- \( Q_{\text{home}}^{NN} = A + B + C + D + E + F + G + H \)
- \( Q_{\text{work}}^{NN} = \frac{C + D + G + H + I + J + K + L + M}{1 + C + D + G + H} \)

The NN equilibrium exists if and only if \( \lambda < Q_{\text{home}}^{NN} < 1 - \lambda \) and \( \lambda < Q_{\text{work}}^{NN} < 1 - \lambda \).
6. The MF equilibrium: home is masculine and work is feminine

In this equilibrium, the payoffs are: $\beta_i - 2\phi$ if mom stays home, $\gamma_i$ if dad stays home, $\beta_i + \gamma_i - \phi - w_n$ if the household hires a nanny. So, the household hires a nanny whenever $\beta_i > w_n + \phi$ and $\gamma_i > w_n - \phi$. The mom stays at home whenever $\beta_i > \gamma_i + 2\phi$ and $\gamma_i < w_n - \phi$. The dad stays at home whenever $\beta_i < \gamma_i + 2\phi$ and $\beta_i < w_n + \phi$. As shown in Figure B8:

- $P_{\text{MF}}^{\text{mom}} = A$
- $P_{\text{MF}}^{\text{nanny}} = B + C + D$
- $P_{\text{MF}}^{\text{dad}} = E + F + G + H + I + J + K + L + M$
- $Q_{\text{MF}}^{\text{home}} = A + B + C + D$
- $Q_{\text{MF}}^{\text{work}} = \frac{B + C + D + E + F + G + H + I + J + K + L + M}{1 + B + C + D}$

The MF equilibrium exists if and only if $Q_{\text{MF}}^{\text{home}} < \lambda$ and $Q_{\text{MF}}^{\text{work}} > 1 - \lambda$. 

B10
Proposition 1. An equilibrium exists.

Proof of Proposition 1. We structure the proof as follows. First, we show that when \( A + B + C + D < \lambda \), the MF equilibrium always exists. Second, we show that whenever \( A + B + C + D + E + F + G + H > 1 - \lambda \), at least one equilibrium exists in which home is a feminine domain. Finally, we consider the intermediate case where \( A + B + C + D \geq \lambda \) but \( A + B + C + D + E + F + G + H \leq 1 - \lambda \); we show that either the NF or the NN equilibrium will always exist in this intermediate case.

1. Case 1: \( A + B + C + D < \lambda \).
   
   First, note that \( Q_{\text{home}}^{MF} = A + B + C + D \), so \( Q_{\text{home}}^{MF} < \lambda \) (i.e. home is masculine).

   Next, note that
   
   \[
   Q_{\text{home}}^{MF} < \lambda \Rightarrow 1 - Q_{\text{home}}^{MF} > 1 - \lambda \\
   \Rightarrow P_{\text{dad}}^{MF} > 1 - \lambda \\
   \Rightarrow Q_{\text{work}}^{MF} > 1 - \lambda
   \]

   by Inequality 1 and the definition of \( Q_{\text{work}}^{MF} \). So, \( A + B + C + D < \lambda \) implies that \( Q_{\text{home}}^{MF} < \lambda \) and \( Q_{\text{work}}^{MF} > 1 - \lambda \); hence, \( A + B + C + D < \lambda \) implies that the MF equilibrium exists.

2. Case 2: \( A + B + C + D + E + F + G + H > 1 - \lambda \).

   First, note that \( A + B + C + D + E + F + G + H > 1 - \lambda \) implies that \( Q_{\text{home}}^{FF} > 1 - \lambda \), \( Q_{\text{home}}^{FN} > 1 - \lambda \), and \( Q_{\text{home}}^{FM} > 1 - \lambda \) (this follows directly from the definitions of \( Q_{\text{home}}^{FF} \), \( Q_{\text{home}}^{FN} \), and \( Q_{\text{home}}^{FM} \) which are listed above). Next, note that if \( Q_{\text{work}}^{FM} < \lambda \), then the FM equilibrium exists. If this condition doesn’t hold,

   \[
   Q_{\text{work}}^{FM} > \lambda \iff \frac{M + (D + H + K)}{1 + (D + H + K)} > \lambda \\\n   \Rightarrow \frac{M + (D + H + K) + (C + G + J)}{1 + (D + H + K) + (C + G + J)} > \lambda \\\n   \Rightarrow \frac{L + M + (D + H + K) + (C + G + J)}{1 + (D + H + K) + (C + G + J)} > \lambda
   \]

   and this final fraction is equal to \( Q_{\text{work}}^{FN} \). If \( Q_{\text{work}}^{FM} \) is also less than \( 1 - \lambda \), then the FN equilibrium exists. Finally, if \( Q_{\text{work}}^{FM} > \lambda \) (so the FN equilibrium does not exist), we know that the FF equilibrium must exist because:

   \[
   \frac{C + D + G + H + J + K + L + M}{1 + C + D + G + H + J + K} > \lambda \Rightarrow \frac{C + D + G + H + I + J + K + L + M}{1 + C + D + G + H + J + K} > \lambda \\
   \Rightarrow \frac{C + D + G + H + I + J + K + L + M}{1 + C + D + G + H} > \lambda
   \]

   which guarantees that the FF equilibrium exists because the last fraction is equal to \( Q_{\text{work}}^{FF} \). So, either the FF, the FN, or the FM equilibrium must exist whenever \( A + B + C + D + E + F + G + H > 1 - \lambda \).
Case 3: $A + B + C + D \geq \lambda$ and $A + B + C + D + E + F + G + H \leq 1 - \lambda$.

First, note that $A + B + C + D \geq \lambda$ guarantees that $Q_{\text{home}}^{NF} \geq \lambda$ and $Q_{\text{home}}^{NN} \geq \lambda$ (and the equalities are strict when all regions $a$, $b$, $c$, etc. are non-empty). So, the requirement that home is a neutral domain in the NF and NN equilibria is satisfied. If it is also the case that

$$\frac{C + D + G + H + I + J + K + L + M}{1 + C + D + G + H} \leq 1 - \lambda$$

then $Q_{\text{work}}^{NN} \leq 1 - \lambda$ (by the definition of $Q_{\text{work}}^{NN}$), so the NN equilibrium exists. However, if

$$\frac{C + D + G + H + I + J + K + L + M}{1 + C + D + G + H} > 1 - \lambda$$

then $Q_{\text{work}}^{NF} > 1 - \lambda$ because

$$\frac{C + D + G + H + I + J + K + L + M}{1 + C + D + G + H} > 1 - \lambda$$

$$\Rightarrow \frac{(C + D + G + H + I + J + K + L + M) + F}{1 + C + D + G + H} > 1 - \lambda$$

$$\Rightarrow \frac{(C + D + G + H + I + J + K + L + M) + F}{1 + C + D} > 1 - \lambda.$$  

So, when $A + B + C + D \geq \lambda$ and $A + B + C + D + E + F + G + H \leq 1 - \lambda$, then either the NN equilibrium or the NF equilibrium must exist. \qed

B12
C Data Construction Appendix

C.1 Using the Ethnologue to Define the Set of Living Languages

For every known language, the Ethnologue provides an estimate of the number of native speakers (if any) in every country. Data are drawn from a range of sources including national censuses and surveys compiled by linguists.

There are 7,457 languages in the Ethnologue database. We drop those that are extinct or have no native speakers, sign languages, and dying languages that had fewer than 100 native speakers when last assessed by Ethnologue researchers. This leaves 6,190 living, oral languages which together account for an estimated 6.50 billion native speakers.

C.2 The World Atlas of Language Structures

One of the best known sources of systematic data on language structure is the World Atlas of Language Structures (WALS). We use the WALS (Corbett 2013a, Corbett 2013b, Corbett 2013c) to classify the gender structure of 525 Ethnologue languages. The WALS contains data on the grammatical gender structure on 257 languages. However, data from the WALS must be used with caution because the linguist who compiled the gender information, Gre-ville Corbett, advocates the use of a somewhat non-standard definition of grammatical gender that includes systems of anaphoric pronominal agreement (Corbett 1991). This is problematic when one wishes to combine the WALS data with information from other sources that do not classify systems of pronominal agreement based on the gender of the referent as examples of grammatical gender. We address this by excluding WALS data on languages that are classified as “strictly semantic” (i.e. agreement class can always be inferred from the meaning of the noun) since Corbett considers pronominal agreement an example of such a system. Languages that are classified in the WALS as either lacking a grammatical gender system or having a system that is “semantic and formal” are unam-biguous. We use WALS data on 188 languages; these languages map to 525 languages in the Ethnologue. In many cases, a single language in the WALS database (e.g. Arabic or Kanuri) will map to a macrolanguage or a higher point in the Ethnologue language tree (accounting for the disparity in the number of languages).

C.3 George Abraham Grierson’s Linguistic Survey of India

A rich source of information about the languages of South Asia is George Abraham Grier-son’s eleven-volume Linguistic Survey of India (Grierson 1903a, 1903b, 1904, 1905, 1907, 1908, 1909, 1916, 1919, 1921), which was compiled between 1891 and 1921. It covers more than 300 South Asian languages and dialects. We use Greierson’s data to classify 52 South Asian languages, and as a second source to confirm the grammatical structure of an additional 469 languages.

C.4 George L. Campbell’s Compendium of the World’s Languages

We use data from George L. Campbell’s Compendium of the World’s Languages (Campbell 1991) to characterize the grammatical gender structure of 949 languages, and as a second source to confirm the grammatical structure of an additional 253 languages.
C.5 First-Person Accounts

We are particularly grateful to Premila Chand, Sameer Chand, and Satish Chand for first-person accounts of the grammatical agreement structure in Fiji Hindi, and to A. K. Rakim and Dr. Kyaw Hla for their detailed descriptions of grammatical agreement in Rohingya.

C.6 Additional Sources

Additional data on the grammatical gender structures of languages comes from academic articles and teaching materials focused on individual languages or language families. The complete list of sources used to classify languages is provided in the references section.

C.7 Identifying Gender Languages

We classify languages as gender or non-gender in several different ways. First, some languages are explicitly identified as gender or non-gender languages in linguistic or pedagogical materials. For example, the UCLA Language Materials Project characterizes Serbian by stating: “Three grammatical genders (masculine, feminine, and neuter) and two numbers (singular and plural) are also distinguished” (UCLA Language Materials Project 2014). Some sources are equally explicit about the absence of grammatical gender. For example, *A Reference Grammar of Maithili* states: “Modern Maithili, however, has no grammatical gender. In other words, in modern Maithili distinctions of gender are determined solely by the sex of the animate noun” (Yadav 1996). In other cases, grammatical materials characterize the different noun classes present within a language (or the absence of a noun classification system), and provide examples of words that fall into each class.

For each language, we record specific quotes characterizing the gender structure. Whenever possible, we use two independent sources to confirm the structure of each language. We identify two independent sources characterizing the grammatical structure of 2,561 languages, or 58.9 percent of languages successfully classified.

To illustrate our classification process, we list the sources and relevant quotes for the 100 most widely spoken languages that are not included in the WALS database (below).

C.8 Source Information for 100 Widely Spoken Languages

*Information on the sources used to classify the 100 most widely spoken languages that are not included in the WALS database.*

**Aceh**

- **Classification:** No grammatical gender
- **Source 1:** Campbell (1991)
- **Reference 1:** No grammatical gender, no marking for case or number (p.12).
- **Source 2:** Corbett (2006)
- **Reference 2:** In almost all the languages of the Austronesian family, [grammatical gender] is simply missing.
Afrikaans

**Classification:** No grammatical gender  
**Source 1:** Campbell (1991)  
**Reference 1:** The division into common and neuter nouns, retained in Dutch, has been lost in Afrikaans (p. 19).  
**Source 2:** UCLA Language Materials Project (2014)  
**Reference 2:** The Dutch distinction of common vs. neuter genders for nouns is lost in Afrikaans. A single definite article die is used for all nouns.

Akan

**Classification:** No grammatical gender  
**Source 1:** Creissels (2000)  
**Reference 1:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).  
**Source 2:** Creissels, Dimmendaal, Frajzyngier, and Konig (2008)  
**Reference 2:** Another type of gender system, in which the sex distinction plays no role, is encountered in all major branches of the Niger-Congo phylum... In addition to the irrelevance of the masculine vs. feminine distinction, Niger-Congo gender systems, usually referred to as noun-class systems, share the following characteristics... A few Niger-Congo languages (e.g. Ijo, the Ubangian language Zande, the Mande language Jo) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest at the level of the relation between the noun and its modifiers (pp. 115-117).

Assamese

**Classification:** No grammatical gender  
**Source 1:** Aikhenvald (2003)  
**Reference 1:** In Indic and Iranian languages the masculine and feminine declensional paradigms merged. This resulted in a complete loss of gender oppositions in Assamese, Bengali, Nepali, Oriya, Persian, Beluzhi, and Ossete (p.379).  
**Source 2:** Kilarski (2013)  
**Reference 2:** Gender has been lost in several languages of the Indo-Iranian branch of the Indo-European, e.g. Assamese and Bengali (Indo-Aryan) and Persian and Ossetic (Iranian).

Bali

**Classification:** No grammatical gender  
**Source:** Corbett (2006)  
**Quote:** In almost all the languages of the Austronesian family, [grammatical gender] is simply missing.
Bamanankan

**Classification:** No grammatical gender

**Source 1:** Creissels (2000)

**Reference 1:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

**Source 2:** Creissels, Dimmendaal, Frajzyngier, and Konig (2008)

**Reference 2:** Another type of gender system, in which the sex distinction plays no role, is encountered in all major branches of the Niger-Congo phylum... In addition to the irrelevance of the masculine vs. feminine distinction, Niger-Congo gender systems, usually referred to as noun-class systems, share the following characteristics... A few Niger-Congo languages (e.g. Ijo, the Ubangian language Zande, the Mande language Jo) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest at the level of the relation between the noun and its modifiers (pp. 115-117).

Bemba

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).

**Source 2:** Creissels (2000)

**Reference 2:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Bengali

**Classification:** No grammatical gender

**Source 1:** Kilarski (2013)

**Reference 1:** Gender has been lost in several languages of the Indo-Iranian branch of the Indo-European, e.g. Assamese and Bengali (Indo-Aryan) and Persian and Ossetic (Iranian).

**Source 2:** Campbell (1991)

**Reference 2:** Bengali has lost the grammatical gender system of Indo-Aryan, and has replaced it with a natural taxonomy of animate versus non-animate (p. 229).

Betawi

**Classification:** No grammatical gender

**Source:** Aikhenvald (2003)

**Quote:** The reduction and loss of class/gender distinctions is a universal feature of the pidginization and creolization of languages(p. 388).
Bhili

**Classification:** Grammatical gender with more than two noun classes

**Source 1:** Masica (1991)

**Reference 1:** In NIA, three genders are preserved only in parts of the west: on the chart, in Gujarati, Marathi, and Konkani; also in dialects in their vicinity not on the chart (Bhili, Khandesi - not, however, in Halbi); finally, in the Bhadarwahi-Bhalesi-Khashali group of extreme northwestern West Pahari (p. 220).

**Source 2:** Grierson (1907)

**Reference 2:** The neuter gender is often used to denote feminine beings... All adjectives which do not end in U are uninflected. Those ending in U are inflected for gender, number, and partly for case (p. 12).

Bhojpuri

**Classification:** No grammatical gender

**Source 1:** Masica (1991)

**Reference 1:** To the east, [gender] begins to attenuate, already in the band of languages represented on the chart by Awadhi (the least attenuated), Nepali, Maithili, Bhojpuri, and Chhattisgarhi, to which may be added Bagheli, Magahi, and Angika. There gender accord is typically restricted to female animales... optional or loose even then (e.g. in Chhattisgarhi), and greatly reduced in syntactic scope... Although Awadhi (and Western Bhojpuri, e.g. of Banaras) often show gender agreement similar to that of (Western) Hindi, in the southern extensions of Eastern Hindi, namely Bagheli and Chhattisgarhi, this is greatly attenuated, in the latter almost to the point of disappearance - completely so from the verb. The influence of Standard Hindi as the official language of this region is now a factor confusing the situation, and it may often be a question of the revival or even of the introduction of gender agreement rather than of its preservation (p. 221).

**Source 2:** Grierson (1903b)

**Reference 2:** In Standard Bhojpuri, adjectives do not change for gender (p. 50).

Bugis

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** The personal articles i-la (masc.) and i-we (fem.) are applied to proper nouns, names of boats, weapons, etc (p. 269).

**Source 2:** Corbett (2006)

**Reference 2:** In almost all the languages of the Austronesian family, [grammatical gender] is simply missing.
Bulgarian

Classification: Grammatical gender with more than two noun classes
Source 1: Leafgren (2011)
Reference 1: Nouns in Bulgarian can be divided into three broad categories based on their grammatical gender: masculine, feminine or neuter. Grammatical gender is important because it determines formal agreement and correspondence features, at least in the singular, of adjectives, many pronouns and adjectival pronouns, and parts of compound verbal constructions which agree with the subject of the clause.
Source 2: UCLA Language Materials Project (2014)
Reference 2: The Bulgarian noun is inflected for number and gender (masculine, feminine, and neuter) while the case inflection, contrary to other Slavic languages, has been lost. The adjective agrees with the noun it modifies in number and gender. There is a small number of adjectives of Turkish origin that show no inflection.

Catalan

Classification: Grammatical gender with only two classes (masculine, feminine)
Source 1: Campbell (1991)
Reference 1: Nouns in Catalan are masculine or feminine (p. 311).
Source 2: UCLA Language Materials Project (2014)
Reference 2: As in Spanish, nouns inflect for gender and number, pronouns inflect for gender, number, person, and case, and verbs agree with their subjects in both gender and number and inflect for tense and mood. Additionally, adjectives agree with the nouns they modify in gender and number.

Cebuano

Classification: No grammatical gender
Source 1: Campbell (1991)
Reference 1: [Detailed description of nouns (p. 313-319) in Cebuano makes no mention of noun classes.]
Source 2: Corbett (2006)
Reference 2: In almost all the languages of the Austronesian family, [grammatical gender] is simply missing.

Central Kurdish

Classification: No grammatical gender
Source: Campbell (1991)
Quote: Gender is not distinguished grammatically (p. 926).
Chhattisgarhi

Classification:  No grammatical gender
Source 1:  Masica (1991)
Reference 1:  To the east, [gender] begins to attenuate, already in the band of languages represented on the chart by Awadhi (the least attenuated), Nepali, Maithili, Bhojpuri, and Chhattisgarhi, to which may be added Bagheli, Magahi, and Angika. There gender accord is typically restricted to female animales... optional or loose even then (e.g. in Chhattisgarhi), and greatly reduced in syntactic scope... Although Awadhi (and Western Bhojpuri, e.g. of Banaras) often show gender agreement similar to that of (Western) Hindi, in the southern extensions of Eastern Hindi, namely Bagheli and Chhattisgarhi, this is greatly attenuated, in the latter almost to the point of disappearance - completely so from the verb. The influence of Standard Hindi as the official language of this region is now a factor confusing the situation, and it may often be a question of the revival or even of the introduction of gender agreement rather than of its preservation (p. 221).
Source 2:  Grierson (1904)
Reference 2:  Tadbhava adjectives in A form the feminine in I... This rule is, however, very arbitrarily followed. Other adjectives do not change for gender (p. 28).

Chittagonian

Classification:  No grammatical gender
Source 1:  Grierson (1903a)
Reference 1:  Adjectives do not change for gender.
Reference 2:  Chittagonian is a non-standard dialect of Bengali.

Croatian

Classification:  Grammatical gender with more than two noun classes
Reference 1:  Every noun in [Bosnian, Serbian, or Croatian] belongs to one of three genders - masculine, feminine, or neuter.
Source 2:  UCLA Language Materials Project (2014)
Reference 2:  Three grammatical genders (masculine, feminine, and neuter) and two numbers (singular and plural) are also distinguished. Case, grammatical gender and number are represented by inflectional morphemes. Adjectives agree with their noun in grammatical gender, number and case. Main verbs and participles agree with the subject only in person, number and gender.
Czech

**Classification:** Grammatical gender with more than two noun classes  
**Source 1:** Danaher (2016)

**Reference 1:** Czech nouns are marked for grammatical gender. Czech has three grammatical genders: Masculine (M), Feminine (F), and Neuter (N). M and F partly overlap with the natural gender of human beings, but grammatical gender is a feature of all nouns (names of inanimate things, places, abstractions...), and it plays an important role in how Czech grammar works.

**Source 2:** UCLA Language Materials Project (2014)

**Reference 2:** Nouns which are feminine, masculine, and neuter are declined in six declensions. Adjectives agree with nouns in number, gender, and case. Number (singular and plural) is distinguished as is gender by inflectional endings on stems.

Danish

**Classification:** No grammatical gender  
**Source 1:** Campbell (1991)

**Reference 1:** Danish has two genders: common and neuter (p.459).

**Source 2:** UCLA Language Materials Project (2014)

**Reference 2:** Nouns in Danish belong to one of two classes, so-called common gender (en is the indefinite singular article) or neuter gender (et is the indefinite singular article).

Deccan

**Classification:** Grammatical gender with more than two noun classes  
**Source:** Grierson (1905)

**Quote:** There are three genders: masculine, feminine, and neuter. The neuter is used to denote animate beings, and also animate beings in the plural where both genders are included (p. 23). [This makes it sound like natural gender, but the subsequent page lists liquor and tongue as feminine nouns.]

Dholuo

**Classification:** No grammatical gender  
**Source 1:** Hurskainen (1999)

**Reference 1:** Eastern Nilotic languages have a fully developed gender system, while Western and Southern branches of Nilotic languages do not (p. 681).

**Source 2:** Dimmendaal (2000)

**Reference 2:** Eastern Nilotic languages distinguish between masculine and feminine gender (as well as neuter gender in the Teso-Turkana group within this branch of the Nilotic) as an obligatory inflectional category of the noun. Southern and Western Nilotic languages on the other hand have gender marking as a derivational category for certain nouns (in particular those referring to names of animals) as well as for personal names (p. 173).
Dutch

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** Nouns are divided into those of the common gender, with singular definite article in DE, and neuters with singular definite article HET (p. 480).

**Source 2:** UCLA Language Materials Project (2014)

**Reference 2:** Dutch nouns are divided into two genders: common and neuter. Nouns of the common gender take the singular definite article de, while neuter nouns take het. Both genders take the plural definite article de, and the indefinite article een. The Dutch attributive adjective precedes the noun. When the adjective modifies a common gender noun, the adjective takes an ending E: een goede man a good man, but een goed boek a good book.

Eastern Maninkakan

**Classification:** No grammatical gender

**Source 1:** Creissels (2000)

**Reference 1:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

**Source 2:** Creissels, Dimmendaal, Frajzyngier, and Konig (2008)

**Reference 2:** Another type of gender system, in which the sex distinction plays no role, is encountered in all major branches of the Niger-Congo phylum... In addition to the irrelevance of the masculine vs. feminine distinction, Niger-Congo gender systems, usually referred to as noun-class systems, share the following characteristics... A few Niger-Congo languages (e.g. Ijo, the Ubangian language Zande, the Mande language Jo) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest at the level of the relation between the noun and its modifiers (pp. 115-117).
English

**Classification:** No grammatical gender

**Source 1:** Aikhenvald (2003)

**Reference 1:** Some systems based on animacy and sex (and traditionally called gender systems) do not, in fact, satisfy the criteria set out here. English distinguishes three genders just in 3rd person pronouns, HE/SHE/IT. They involve the opposition: male/female/inanimate. There are a few conventionalized metaphorical extensions, e.g. ships are commonly referred to with the feminine pronoun. There is no gender agreement within a noun phrase or or with a verb in a clause. Gender markers in English simply have an anaphoric function, as they also do in Japanese where masculine and feminine forms are distinguished only in 3rd person pronouns with a human referent. Strictly speaking, these are not noun classes (p. 21).

**Source 2:** Hellinger (2003)

**Reference 2:** English has no grammatical gender at all. While Old English had three gender classes, feminine, masculine, and neuter, the category of grammatical gender was lost by the end of the 14th century due to the decay of inflectional endings and the disintegration of declensional classes (cf. Strang 1970, Kastovsky 2000). Unlike German, which has a number of elements inside and outside the noun phrase (determiners, adjectives, pronouns) which vary according to the noun’s grammatical gender, Modern English shows no such morphological agreement. English is no longer a (grammatical) gender language (p. 107).

Flemish

**Classification:** Grammatical gender with more than two noun classes

**Source 1:** De Vogelaer (2009)

**Reference 1:** The results of the 2006 questionnaire do not show a radical breakdown of the gender system: in all dialects the three genders are still used.

**Source 2:** Donaldson (2008)

**Reference 2:** Dutch nouns belong to one of two genders, common gender and neuter. The former is an amalgamation of what were formerly masculine and feminine. In the south of Holland and in Belgium the difference between masculine and feminine is still heeded in the use of pronouns (p. 34). In Belgium the old distinction between masculine and feminine is still very much alive (p. 73). Those speakers for whom certain non-personal nouns are still regarded as feminine may use ZE as an object pronoun instead of HEM (but never HAAR/D’R which can only be used with reference to people) (p. 76).
Ganda

**Classification:** No grammatical gender  
**Source 1:** Campbell (1991)  
**Reference 1:** The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).  
**Source 2:** Creissels (2000)  
**Reference 2:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but is concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Ghanaian Pidgin English

**Classification:** No grammatical gender  
**Source:** Aikhenvald (2003)  
**Quote:** The reduction and loss of class/gender distinctions is a universal feature of the pidginization and creolization of languages. Indo-European-based creole languages do not have any gender distinctions (p. 388).

Gikuyu

**Classification:** No grammatical gender  
**Source 1:** Campbell (1991)  
**Reference 1:** The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).  
**Source 2:** Creissels (2000)  
**Reference 2:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but is concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Goan Konkani

**Classification:** Grammatical gender with more than two noun classes  
**Source 1:** Masica (1991)  
**Reference 1:** In NIA, three genders are preserved only in parts of the west: on the chart, in Gujarati, Marathi, and Konkani; also in dialects in their vicinity not on the chart (Bhili, Khandesi - not, however, in Halbi); finally, in the Bhadarwahi-Bhalesi-Khashali group of extreme northwestern West Pahari (p. 220).  
**Source 2:** Grierson (1905)  
**Reference 2:** Gender is usually distinguished in the same way as in standard Marathi (p. 169).
**Gujarati**

**Classification:** Grammatical gender with more than two noun classes  
**Source 1:** Masica (1991)  
**Reference 1:** In NIA, three genders are preserved only in parts of the west: on the chart, in Gujarati, Marathi, and Konkani; also in dialects in their vicinity not on the chart (Bhili, Khandesi - not, however, in Halbi); finally, in the Bhadarwahi-Bhalesi-Khashali group of extreme northwestern West Pahari (p. 220).  
**Source 2:** Campbell (1991)  
**Reference 2:** Gujarati distinguishes masculine, feminine, and neuter genders (p. 658).

**Haitian Creole**

**Classification:** No grammatical gender  
**Source 1:** Muhleisen and Walicek (2010)  
**Reference 1:** Creole languages are generally regarded as gender-less regardless of whether the lexifier utilizes it (e.g., Dutch, French, Portuguese) or not (English) (p. 17)... Grammatical gender is generally represented as absent in Caribbean Creole languages (p. 18).  
**Source 2:** McWhorter (2005)  
**Reference 2:** DeGraff's claim that Haitian has gender inflection is mistaken. DeGraff usefully points out that Haitian has feminine allomorphs for various suffixes denoting occupation, role, or quality such as... RADOTE/RADOTEZ... Yet this is not grammatical gender but natural gender (p. 24).

**Haryanvi**

**Classification:** Grammatical gender with only two classes (masculine, feminine)  
**Source:** Grierson (1916)  
**Quote:** Bangaru [Haryanvi] is the dialect of Western Hindi which is spoken in the eastern Punjab (p. 1).

**Hiligaynon**

**Classification:** No grammatical gender  
**Source:** Corbett (2006)  
**Quote:** In almost all the languages of the Austronesian family, [grammatical gender] is simply missing.
Ilocano

Classification: No grammatical gender
Source 1: Aikhenvald (2003)
Reference 1: Ilocano does no have genders; the natural gender of humans may be distinguished lexically (p. 313).
Source 2: Corbett (2006)
Reference 2: In almost all the languages of the Austronesian family, [grammatical gender] is simply missing.

Italian

Classification: Grammatical gender with only two classes (masculine, feminine)
Source 1: Proudfoot and Cardo (2005)
Reference 1: All Italian nouns have either a masculine or a feminine gender. Gender is purely a grammatical term. Nouns referring to human beings or animals sometimes have the same grammatical gender as their natural gender, but not always.
Source 2: Campbell (1991)
Reference 2: Italian has two genders... The definite article is marked for gender and number: masc. IL/I; fem. LA/LE (pp. 778-779).

Japanese

Classification: No grammatical gender
Source 1: Campbell (1991)
Reference 1: There is no grammatical gender; no articles (p. 808).
Source 2: UCLA Language Materials Project (2014)
Reference 2: Japanese nouns and adjectives do not have the categories of gender or number.

Javanese

Classification: No grammatical gender
Source 1: Kuntjara (2001)
Reference 1: Neither Javanese nor Indonesian have grammatical gender, not even pronominal gender distinctions like English he or she (p.202).
Source 2: Campbell (1991)
Reference 2: No gender (p. 816).
Kamba

Classification: No grammatical gender

Source 1: Campbell (1991)

Reference 1: The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).

Source 2: Creissels (2000)

Reference 2: A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but is concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Kanauji

Classification: Grammatical gender with only two classes (masculine, feminine)

Source 1: Masica (1991)

Reference 1: Two-gender systems: Braj M F (p. 220, in Figure 8.3).

Source 2: Grierson (1916)

Reference 2: Adjectives as in ordinary Hindi except that strong masculine forms end in O instead of A (p. 85).

Kannada

Classification: No grammatical gender

Source 1: Krishnamurti (2001)

Reference 1: [Figure 8.4 on p. 137 shows that there are three singular classes (male humans, female humans, everything else) and two plural classes (humans, non-humans).]


Reference 2: The common rule is that nouns denoting males belong to the masculine gender, nouns denoting females belong to the feminine gender, and all other nouns... belong to the neuter gender (p. 106). [Table 5 on p. 106: shows noun class system is masc. humans/fem. humans/all other nouns in singular and human/non-human in plural.]

Kazakh

Classification: No grammatical gender

Source 1: Campbell (1991)

Reference 1: There is no grammatical gender (p. 65).

Source 2: Johanson (2013[2002])

Reference 2: Turkic has no classifiers of grammatical gender. Hence, gender does not play a role in grammatical agreement.
Kinyarwanda

Classification: No grammatical gender
Source 1: Campbell (1991)
Reference 1: The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).
Source 2: Creissels (2000)
Reference 2: A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Kituba

Classification: No grammatical gender
Source: Aikhenvald (2003)
Quote: The reduction and loss of class/gender distinctions is a universal feature of the pidginization and creolization of languages (p. 388).

Korean

Classification: No grammatical gender
Source: UCLA Language Materials Project (2014)
Quote: Gender and number are not marked, and the language lacks articles, fusional morphology, relative pronouns, conjunctions and agglutination. Nouns are not inflected as such; rather, there is a class of postpositional particles or suffixes which may be used to mark 7 cases (nominative, genitive, accusative, dative, locative, instrumental, and comitative). Grammatical gender and number are not marked.

Lambadi

Classification: Grammatical gender with only two classes (masculine, feminine)
Source: Grierson (1908)
Quote: As regards gender, the rule of Western Hindi is generally followed, that there are only two genders, a masculine and a feminine. In one or two dialects of Western Hindi we have noted sporadic instances of the use of a neuter gender. In Rajasthani these occasional instances become more and more common as we go west and south till we find the neuter gender firmly established in Gujarati (p. 5)... Adjectives follow the genitive postpositions in their inflexions. Thus, ACHCHHO, good; fem., ACHCHHI; masc., ACHCHHHA (p. 7).
Lombard

**Classification:** Grammatical gender with only two classes (masculine, feminine)

**Source 1:** Lewis, Simons, and Fennig, eds., (2016)

**Reference 1:** Similar to French and Italian.

**Source 2:** Campbell (1991)

**Reference 2:** A dialectical division of Italy running roughly along the lines of the Northern Appennines has long been recognized. To the north of this line are Piedmontese, Lombardian, Venetian, etc.; to the south lie Tuscan, Umbrian, Neapolitan, Calabrese, and Sicilian... Most of the dialects are still very much alive... Italian has two genders and two numbers... The adjective agrees in gender and number with the noun (pp. 777-779).

Luba-Kasai

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).

**Source 2:** Creissels (2000)

**Reference 2:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Madura

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** Gender may be indicated by such words as MALE and FEMALE (p. 1035).

**Source 2:** Corbett (2006)

**Reference 2:** In almost all the languages of the Austronesian family, [grammatical gender] is simply missing.
Magahi

**Classification:** No grammatical gender

**Source 1:** Masica (1991)

**Reference 1:** To the east, [gender] begins to attenuate, already in the band of languages represented on the chart by Awadhi (the least attenuated), Nepali, Maithili, Bhojpuri, and Chhattisgarhi, to which may be added Bagheli, Magahi, and Angika. There gender accord is typically restricted to female animales... optional or loose even then (e.g. in Chhattisgarhi), and greatly reduced in syntactic scope... Although Awadhi (and Western Bhojpuri, e.g. of Banaras) often show gender agreement similar to that of (Western) Hindi, in the southern extensions of Eastern Hindi, namely Bagheli and Chhattisgarhi, this is greatly attenuated, in the latter almost to the point of disappearance - completely so from the verb. The influence of Standard Hindi as the official language of this region is now a factor confusing the situation, and it may often be a question of the revival or even of the introduction of gender agreement rather than of its preservation (p. 221).

**Source 2:** Grierson (1903b)

**Reference 2:** Adjectives do not change for gender (p. 50).

Maithili

**Classification:** No grammatical gender

**Source 1:** Yadav (1996)

**Reference 1:** Modern Maithili, however, has no grammatical gender. In other words, in modern Maithili, distinctions of gender are determined soley by the sex of the animate noun.

**Source 2:** Masica (1991)

**Reference 2:** To the east, [gender] begins to attenuate, already in the band of languages represented on the chart by Awadhi (the least attenuated), Nepali, Maithili, Bhojpuri, and Chhattisgarhi, to which may be added Bagheli, Magahi, and Angika. There gender accord is typically restricted to female animales... optional or loose even then (e.g. in Chhattisgarhi), and greatly reduced in syntactic scope... Although Awadhi (and Western Bhojpuri, e.g. of Banaras) often show gender agreement similar to that of (Western) Hindi, in the southern extensions of Eastern Hindi, namely Bagheli and Chhattisgarhi, this is greatly attenuated, in the latter almost to the point of disappearance - completely so from the verb. The influence of Standard Hindi as the official language of this region is now a factor confusing the situation, and it may often be a question of the revival or even of the introduction of gender agreement rather than of its preservation (p. 221).
Malayalam

**Classification:** No grammatical gender

**Source 1:** Andronov (2003)

**Reference 1:** The common rule is that nouns denoting males belong to the masculine gender, nouns denoting females belong to the feminine gender, and all other nouns... belong to the neuter gender (p. 106). [Table 5 on p. 106: shows that there are no longer grammatical gender agreement classes.]

**Source 2:** Krishnamurti (2001)

**Reference 2:** [Figure 8.4 on p. 137 shows that there are three singular classes (male humans, female humans, everything else) and two plural classes (humans, non-humans).]

Malvi

**Classification:** Grammatical gender with only two classes (masculine, feminine)

**Source:** Grierson (1908)

**Quote:** As regards gender, the rule of Western Hindi is generally followed, that there are only two genders, a masculine and a feminine. In one or two dialects of Western Hindi we have noted sporadic instances of the use of a neuter gender. In Rajasthani these occasional instances become more and more common as we go west and south till we find the neuter gender firmly established in Gujarati (p. 5)... Adjectives follow the genitive postpositions in their inflexions. Thus, ACHCHHO, good; fem., ACHCHHI; masc., ACHCHHA (p. 7).

Marwari

**Classification:** Grammatical gender with only two classes (masculine, feminine)

**Source 1:** Masica (1991)

**Reference 1:** Two-gender systems: Marwari M F (p. 220, in Figure 8.3).

**Source 2:** Grierson (1908)

**Reference 2:** As regards gender, the rule of Western Hindi is generally followed, that there are only two genders, a masculine and a feminine. In one or two dialects of Western Hindi we have noted sporadic instances of the use of a neuter gender. In Rajasthani these occasional instances become more and more common as we go west and south till we find the neuter gender firmly established in Gujarati (p. 5)... Adjectives follow the genitive postpositions in their inflexions. Thus, ACHCHHO, good; fem., ACHCHHI; masc., ACHCHHA (p. 7).
Merwari

**Classification:** Grammatical gender with only two classes (masculine, feminine)

**Source 1:** Masica (1991)

**Reference 1:** Two-gender systems: Marwari M F (p. 220, in Figure 8.3).

**Source 2:** Grierson (1908)

**Reference 2:** As regards gender, the rule of Western Hindi is generally followed, that there are only two genders, a masculine and a feminine. In one or two dialects of Western Hindi we have noted sporadic instances of the use of a neuter gender. In Rajasthani these occasional instances become more and more common as we go west and south till we find the neuter gender firmly established in Gujarati (p. 5)... Adjectives follow the genitive postpositions in their inflexions. Thus, ACHCHHO, good; fem., ACHCHHI; masc., ACHCHHA (p. 7).

Mewari

**Classification:** Grammatical gender with only two classes (masculine, feminine)

**Source 1:** Masica (1991)

**Reference 1:** Two-gender systems: Marwari M F (p. 220, in Figure 8.3).

**Source 2:** Grierson (1908)

**Reference 2:** As regards gender, the rule of Western Hindi is generally followed, that there are only two genders, a masculine and a feminine. In one or two dialects of Western Hindi we have noted sporadic instances of the use of a neuter gender. In Rajasthani these occasional instances become more and more common as we go west and south till we find the neuter gender firmly established in Gujarati (p. 5)... Adjectives follow the genitive postpositions in their inflexions. Thus, ACHCHHO, good; fem., ACHCHHI; masc., ACHCHHA (p. 7).

Moore

**Classification:** No grammatical gender

**Source 1:** Creissels (2000)

**Reference 1:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but is concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

**Source 2:** Creissels, Dimmendaal, Frajzyngier, and Konig (2008)

**Reference 2:** Another type of gender system, in which the sex distinction plays no role, is encountered in all major branches of the Niger-Congo phylum... In addition to the irrelevance of the masculine vs. feminine distinction, Niger-Congo gender systems, usually referred to as noun-class systems, share the following characteristics... A few Niger-Congo languages (e.g. Ijo, the Ubangian language Zande, the Mande language Jo) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest at the level of the relation between the noun and its modifiers (pp. 115-117).
Napoletano-Calabrese

Classification: Grammatical gender with only two classes (masculine, feminine)
Reference 1: Similar to standard Italian.
Source 2: Campbell (1991)
Reference 2: A dialectical division of Italy running roughly along the lines of the Northern Appennines has long been recognized. To the north of this line are Piedmontese, Lombardian, Venetian, etc.; to the south lie Tuscan, Umbrian, Neapolitan, Calabrese, and Sicilian... Most of the dialects are still very much alive... Italian has two genders and two numbers... The adjective agrees in gender and number with the noun (pp. 777-779).

Nepali

Classification: No grammatical gender
Source 1: Grierson (1916)
Reference 1: The distinction of gender is purely sexual. The so-called grammatical gender does not occur, and hence many nouns which are feminine in Hindi are masculine in [Nepali] (p. 22).
Source 2: Campbell (1991)
Reference 2: Gender plays little part in Nepali structure, though the markers of its historical presence could hardly be entirely absent form a new Indo-Aryan language. In nouns denoting human beings, O and A are masculine endings and I is feminine (p. 1213).

North Azerbaijani

Classification: No grammatical gender
Source 1: Campbell (1991)
Reference 1: There is no grammatical gender (p. 65).
Source 2: Johanson (2013[2002])
Reference 2: Turkic has no classifiers of grammatical gender. Hence, gender does not play a role in grammatical agreement.

Northern Kurdish

Classification: No grammatical gender
Source: Campbell (1991)
Quote: Gender is not distinguished grammatically (p. 926).
Northern Sotho

Classification: No grammatical gender
Source 1: Campbell (1991)
Reference 1: The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).
Source 2: Creissels (2000)
Reference 2: A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Norwegian

Classification: Grammatical gender with more than two noun classes
Source 1: Hellinger and Bußman (2003)
Reference 1: [Frisian, Icelandic, and Norwegian] have retained the Indo-European three gender system with masc, fem, and neuter (p. 143).
Source 2: UCLA Language Materials Project (2014)
Reference 2: Nouns in Norwegian fall into one of three genders (masculine, feminine and neuter), however in most dialects (and especially bokmal) masculine and feminine have merged to COMMON. The dialect in Bergen on the West Coast has only 2 genders, no feminine forms. Nynorsk maintains the 3 gender system more rigidly.

Odia

Classification: No grammatical gender
Source 1: Aikhenvald (2003)
Reference 1: In Indic and Iranian languages the masculine and feminine declensional paradigms merged. This resulted in a complete loss of gender oppositions in Assamese, Bengali, Nepali, Oriya, Persian, Beludzhi, and Ossete (p.379).
Source 2: Campbell (1991)
Reference 2: Natural gender is distinguished for animates, either lexically or by Sanskrit ending (p. 1281).
Polish

**Classification:** Grammatical gender with more than two noun classes

**Source 1:** Feldstein (2001)

**Reference 1:** Each Polish noun has a specific gender, called masculine, neuter, or feminine. In order to use any given Polish noun, one must know its gender, as well as the behavior of its particular declensional subclass, in case it does not follow the basic type for that gender. Based on agreement with adjectives, nouns can readily be identified as masculine, neuter, or feminine in the singular.

**Source 2:** Campbell (1991)

**Reference 2:** Polish has three genders, two numbers, and seven cases... As attribute, adjective normally precedes the noun, and is in concord for gender, number, and case (p. 1357).

Portuguese

**Classification:** Grammatical gender with only two classes (masculine, feminine)

**Source 1:** Hutchinson and Lloyd (1996)

**Reference 1:** There are two genders: masculine and feminine.

**Source 2:** Campbell (1991)

**Reference 2:** Nouns are masculine or feminine... Article: Definite: O (masc.), a (fem.)... Adjective: A basic opposition is [NULL] for masculine, -A for feminine, but many adjectives have identical forms for both genders (pp. 1369-1370).

Rangpuri

**Classification:** No grammatical gender

**Source 1:** Wilde (2008)

**Reference 1:** Noun classes bear no semantic correlation with sex (p. 63).

**Source 2:** Toulmin (2006)

**Reference 2:** As in in the other e.Mg. lects (Oriya, Bangla, Assamiya, etc.), gender is not an inflectional category (p. 47).

Romanian

**Classification:** Grammatical gender with more than two noun classes

**Source 1:** Cojocaru (2003)

**Reference 1:** Romanian, as any inflective language, is governed by nominal agreement: the forms of different modifiers (adjectives, pronominal adjectives, ordinal numerals) depend on the gender and number of the noun. Romanian is the only Romance language that preserved three genders from the Latin: masculine, feminine, and neuter.

**Source 2:** UCLA Language Materials Project (2014)

**Reference 2:** Three grammatical genders, masculine, feminine, and irregular (masculine in the singular and feminine in the plural) are distinguished.
Rundi

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).

**Source 2:** Creissels (2000)

**Reference 2:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Santhali

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** Gender is not distinguished, though natural gender may be lexically marked (p. 1454).

**Source 2:** Grierson (1904)

**Reference 2:** Nouns do not differ for gender. The natural gender is distinguished by using different words or by adding words denoting male, female, respectively... Nouns, on the other hand, can be divided into two classes, viz. those that denote animate being, and those that denote inanimate objects respectively (p. 23).

Serbian

**Classification:** Grammatical gender with more than two noun classes

**Source 1:** Alexander (2006)

**Reference 1:** Every noun in [Bosnian, Serbian, or Croatian] belongs to one of three genders - masculine, feminine, or neuter.

**Source 2:** UCLA Language Materials Project (2014)

**Reference 2:** Three grammatical genders (masculine, feminine, and neuter) and two numbers (singular and plural) are also distinguished.

Shan

**Classification:** No grammatical gender

**Source 1:** Cushing (1871)

**Reference 1:** The masculine and feminine genders are distinguished - 1st, by common words, as father, mother, husband, wife, grandfather, grandmother, lord, lady. 2nd. By affixes [SYMBOL] and [SYMBOL] are used to distinguish gender of the human species, while [SYMBOL] and [SYMBOL] are used to distinguish that of the brutes, as... a cock, a hen (p. 14).

**Source 2:** Strecker (1987)

**Reference 2:** The Tai languages are uninflected (p. 754).
Sicilian

**Classification:** Grammatical gender with only two classes (masculine, feminine)

**Source:** Campbell (1991)

**Quote:** A dialectical division of Italy running roughly along the lines of the Northern Appennines has long been recognized. To the north of this line are Piedmontese, Lombardian, Venetian, etc.; to the south lie Tuscan, Umbrian, Neapolitan, Calabrese, and Sicilian... Most of the dialects are still very much alive... Italian has two genders and two numbers... The adjective agrees in gender and number with the noun (pp. 777-779).

Sindhi

**Classification:** Grammatical gender with only two classes (masculine, feminine)

**Source 1:** Campbell (1991)

**Reference 1:** Masculine and feminine nouns share all five vocalic endings, but -O and -U are typically masculine, -I and -E are typically feminine... As attribute, adjective precedes noun. Many adjectives are indeclinable, but those with vocalic endings are usually inflected for gender (p. 1500).

**Source 2:** Masica (1991)

**Reference 2:** Two-gender systems: Sindhi M F (p. 220, in Figure 8.3).

Sinhala

**Classification:** No grammatical gender

**Source 1:** Masica (1991)

**Reference 1:** It is sometimes held that Sinhalese also preserves the three genders of OIA, but this is misleading. The system has been completely restructured into one based on natural gender(as in neighboring Dravidian), with the basic distinction between inanimates on the one hand (the so-called Neuter, whose membership therefore includes the inanimate Masculines and Feminines of OIA as well as most Neuters of OIA) and animates on the other hand, secondarily divided into Masculine and Feminine... Moreover, gender is primarily a declensional phenomenon in Sinhalese; syntactic agreement is very marginal, and confined to the literary language (pp. 220-221).

**Source 2:** Campbell (1991)

**Reference 2:** The basic contrast is between animates and inanimates; animates are further sub-divided into masculine and feminine... Neuter, i.e. inanimate, nouns usually end in -YA, -VA, -A (p. 1505).
Slovak

**Classification:** Grammatical gender with more than two noun classes

**Source 1:** Campbell (1991)

**Reference 1:** Slovak has three genders (p. 1517).

**Source 2:** UCLA Language Materials Project (2014)

**Reference 2:** Nouns which are feminine, masculine, and neuter are declined in six declensions, and adjectives agree in number, gender, and case. Number (singular and plural) is distinguished as is gender by inflectional endings on stems. The six case endings are nominative, accusative, genitive, dative, instrumental, and locative. Slovak also marks an animate/inanimate distinction for masculine nouns. The language has no definite article.

Somali

**Classification:** Grammatical gender with only two classes (masculine, feminine)

**Source 1:** Campbell (1991)

**Reference 1:** Somali has two grammatical genders, masculine and feminine (p. 1527).

**Source 2:** Appleyard (2011)

**Reference 2:** The typical Afroasiatic grammatical gender system comprising masculine and feminine runs throughout Cushitic morphosyntax (p. 44).

South Azerbaijani

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** There is no grammatical gender (p. 65).

**Source 2:** Johanson (2013[2002])

**Reference 2:** Turkic has no classifiers of grammatical gender. Hence, gender does not play a role in grammatical agreement.

Southern Balochi

**Classification:** No grammatical gender

**Source 1:** Windfuhr (2009)

**Reference 1:** There is no grammatical gender in any dialect of Balochi (p. 651).

**Source 2:** Grierson (1921)

**Reference 2:** There is no distinction of grammatical gender in Balochi. Male and female are distinguished either by the use of different words - as in GURAND, a ram, GAD, a ewe - or by the addition of words such as NAR, male, or MADAG, female (p. 340).
Southern Sotho

Classification: No grammatical gender
Source 1: Campbell (1991)
Reference 1: The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).
Source 2: Creissels (2000)
Reference 2: A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Sukuma

Classification: No grammatical gender
Source 1: Campbell (1991)
Reference 1: The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).
Source 2: Creissels (2000)
Reference 2: A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Sunda

Classification: No grammatical gender
Source 1: Campbell (1991)
Reference 1: [Detailed description of nouns (p. 1557) in Sundanese makes no mention of noun classes.]
Source 2: Corbett (2006)
Reference 2: In almost all the languages of the Austronesian family, [grammatical gender] is simply missing.

Swedish

Classification: No grammatical gender
Source 1: Campbell (1991)
Reference 1: There are two genders: common and neuter (p. 1569).
Source 2: UCLA Language Materials Project (2014)
Reference 2: Nouns in Swedish are either common gender [sometimes called utrum] (approximately 75 percent of all nouns) or neuter.
Sylheti
**Classification:** No grammatical gender
**Source:** Grierson (1903a)
**Quote:** Adjectives do not change for gender.

Tagalog
**Classification:** No grammatical gender
**Source 1:** Stolz (2013)
**Reference 1:** The class of nouns and adjectives that reflect gender distinctions is very small in Tagalog. It is difficult therefore to determine exactly to what extent grammatical gender is integrated into the grammatical system of Tagalog. What can be said with certainty is that Tagalog grammatical gender is marginal and contact induced (p. 100).
**Source 2:** Campbell (1991)
**Reference 2:** No grammatical gender (p. 1587).

Tajiki
**Classification:** No grammatical gender
**Source 1:** Aikhenvald (2003)
**Reference 1:** In Indic and Iranian languages the masculine and feminine declensional paradigms merged. This resulted in a complete loss of gender oppositions in Assamese, Bengali, Nepali, Oriya, Persian, Beludzhi, and Ossete (p.379).
**Source 2:** Campbell (1991)
**Reference 2:** Persian has no grammatical gender and no articles (p. 1342).

Tamil
**Classification:** No grammatical gender
**Source 1:** Andronov (2003)
**Reference 1:** The common rule is that nouns denoting males belong to the masculine gender, nouns denoting females belong to the feminine gender, and all other nouns... belong to the neuter gender. Such an expanded system is attested in one case only... Harijan Tamil (p. 106).
**Source 2:** Campbell (1991)
**Reference 2:** Tamil divides nominals into two classes: rational/human as opposed to non-human (neuter). Within the rational class, a distinction is made between masculine and feminine (p. 1605).
Tatar

**Classification:** No grammatical gender  
**Source 1:** Campbell (1991)  
**Reference 1:** There is no grammatical gender (p. 65).  
**Source 2:** Johanson (2013[2002])  
**Reference 2:** Turkic has no classifiers of grammatical gender. Hence, gender does not play a role in grammatical agreement.

Telugu

**Classification:** Grammatical gender with only two classes (masculine, feminine)  
**Source 1:** Andronov (2003)  
**Reference 1:** The common rule is that nouns denoting males belong to the masculine gender, nouns denoting females belong to the feminine gender, and all other nouns...belong to the neuter gender (p. 106). [Table 5 on p. 106: shows noun class system is male humans/female humans and all other nouns in singular and human/non-human in plural.]  
**Source 2:** Campbell (1991)  
**Reference 2:** In the singular, the opposition is between masculine and non-masculine, the latter category including all nouns denoting females; in the plural, the dichotomy changes to rational versus non-rational, with females promoted to rational status. Grammatical gender is not formally marked, but it is made explicit, e.g. by concord of noun with verb (p. 1625).

Tigrigna

**Classification:** Grammatical gender with only two classes (masculine, feminine)  
**Source 1:** Campbell (1991)  
**Reference 1:** Two genders, masculine and feminine, are found throughout the family (p. 1470).  
**Source 2:** UCLA Language Materials Project (2014)  
**Reference 2:** Tigrinya nouns are either masculine or feminine and are inflected for number. Gender is not marked on the noun, but on nominal dependents like articles and adjectives. Verbs agree with their subjects and objects in person, number, and gender.

Tsonga

**Classification:** No grammatical gender  
**Source 1:** Campbell (1991)  
**Reference 1:** The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).  
**Source 2:** Creissels (2000)  
**Reference 2:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but is concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).
Tswana

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).

**Source 2:** Creissels (2000)

**Reference 2:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Turkmen

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** There is no grammatical gender (p. 65).

**Source 2:** Johanson (2013[2002])

**Reference 2:** Turkic has no classifiers of grammatical gender. Hence, gender does not play a role in grammatical agreement.

Umbundu

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).

**Source 2:** Creissels (2000)

**Reference 2:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Uyghur

**Classification:** No grammatical gender

**Source 1:** Campbell (1991)

**Reference 1:** There is no grammatical gender (p. 65).

**Source 2:** Johanson (2013[2002])

**Reference 2:** Turkic has no classifiers of grammatical gender. Hence, gender does not play a role in grammatical agreement.
Varhadi-Nagpuri

**Classification:** Grammatical gender with more than two noun classes  
**Source:** Grierson (1905)  
**Quote:** The Marathi of Berar is usually spoken of as Varhadi or Berari, and Nagpuri is the traditional name of the dialect spoken in the Central Provinces... essentially identical with the form of Marathi spoken in Berar and the Central Provinces (pp. 217-219).

Venetian

**Classification:** Grammatical gender with only two classes (masculine, feminine)  
**Source:** Campbell (1991)  
**Quote:** A dialectical division of Italy running roughly along the lines of the Northern Appennines has long been recognized. To the north of this line are Piedmontese, Lombardian, Venetian, etc.; to the south lie Tuscan, Umbrian, Neapolitan, Calabrese, and Sicilian... Most of the dialects are still very much alive... Italian has two genders and two numbers... The adjective agrees in gender and number with the noun (pp. 777-779).

Wolof

**Classification:** No grammatical gender  
**Source 1:** Campbell (1991)  
**Reference 1:** There is no grammatical gender (p. 1766).  
**Source 2:** Creissels (2000)  
**Reference 2:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).

Xhosa

**Classification:** No grammatical gender  
**Source 1:** Campbell (1991)  
**Reference 1:** The class system has nothing to do with gender; nor is it, at least in origin, connected with an animate/inanimate dichotomy (p. 196).  
**Source 2:** Creissels (2000)  
**Reference 2:** A few Niger-Congo languages (Ijo, Zande) are reported to have a masculine vs. feminine distinction, but it concerns only pronominal gender and does not manifest itself at the level of the relation between the noun and the modifiers (p. 242).
References


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**HLA, K. (2018): personal communication.**


